

Combinatorics

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Contents

Lecture 20: Quasi-Random Graphs (3)

Fri 08 Oct 2021 10:13

We complete the proof from last time.

Proof. Take m values x_1, x_2, \dots, x_m and let \bar{x} be their arithmetic mean. Then, recall that $\sum_{i=1}^m (x_i - \bar{x})^2 = \sum_{i=1}^m x_i^2 - m\bar{x}^2$. This is simply the definition of variance.

Then, letting $m = \binom{n}{2}$, $\hat{d}_{ij} = x_k$ and the mean codegree to be $\text{mcd} = \frac{1}{\binom{n}{2}} \sum_{1 \leq i, j \leq n} \hat{d}_{ij} = \frac{1}{\binom{n}{2}} \left(\frac{1}{8}n^3 + o(n^3) \right) = \frac{n}{4} + o(n)$. Then, we have

$$\begin{aligned} \sum_{1 \leq i, j \leq n} \left(\hat{d}_{ij} - \text{mcd} \right)^2 &= \sum_{1 \leq i, j \leq n} \hat{d}_{ij}^2 - \binom{n}{2} \text{mcd}^2 \\ &= \frac{1}{32}n^4 + o(n^4) - \frac{1}{32}n^4 + o(n^4) \\ &= o(n^4). \end{aligned}$$

Hence, we obtain $\sum_{1 \leq i, j \leq n} \left(\hat{d}_{ij} - \text{mcd} \right)^2 = o(n^4)$. Then, letting $y_i = \left| \hat{d}_{ij} - \text{mcd} \right|$ we see by Cauchy-Schwarz that $\frac{1}{m} \sum_{i=1}^m y_i \leq \sqrt{\frac{1}{m} \sum_{i=1}^m y_i^2}$, hence $\sum_{i=1}^m y_i \leq \sqrt{m \sum_{i=1}^m y_i^2}$. Hence, we have $\sum_{1 \leq i, j \leq n} \left| \hat{d}_{ij} - \text{mcd} \right| \leq \sqrt{\binom{n}{2} \sum_{1 \leq i, j \leq n} \left(\hat{d}_{ij} - \text{mcd} \right)^2} = o(n^3)$. Hence,

$$\sum_{1 \leq i, j \leq n} \left| \hat{d}_{ij} - \text{mcd} \right| = o(n^3).$$

Then triangle inequality yields

$$\begin{aligned} \sum_{1 \leq i, j \leq n} \left| \hat{d}_{ij} - \frac{n}{4} \right| &\leq \sum_{1 \leq i, j \leq n} \left| \hat{d}_{ij} - \text{mcd} \right| + \left| \text{mcd} - \frac{n}{4} \right| \\ &= o(n^3) + o(n^3) \\ &= o(n^3). \end{aligned}$$

□

Now, we proceed to prove some more implications, but first we state a lemma.

Lemma 0.1. Let x_1, x_2, \dots, x_n be an orthonormal basis with associated eigenvalues $\lambda_1, \dots, \lambda_n$. Then for $j = \frac{1}{\sqrt{n}} (1 \dots 1)$, we find $|x_1 - j|_2 = o(1)$.

Proof. ($P_3 \Rightarrow P_5$). Let x_1 be a unit eigenvector of G corresponding to λ_1 . Then, let $j = \frac{1}{\sqrt{n}} (1 \dots 1)$, then by lemma we have $|x_1 - j|_2 = o(1)$. \square

Lecture 21: Quasi-Random Graphs (4)

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We complete the proof from last time. Recall our lemma that for orthonormal basis containing x_1 we have $|x_1 - j|_2 = o(1)$. We proceed

Proof. WLOG assume G to be a random graph of even order and $|S| = \frac{n}{2}$. Then, we define a vector \vec{S} with $s_i = \begin{cases} \frac{1}{\sqrt{n}}, & i \in S \\ -\frac{1}{\sqrt{n}}, & i \in V \setminus S \end{cases}$ It is clear $|S|_2 = 1$ and we see

$$\langle S, j \rangle = \underbrace{\frac{1}{\sqrt{n}} \cdot \frac{1}{\sqrt{n}}}_{\frac{n}{2} \text{ times}} + \underbrace{\frac{1}{-\sqrt{n}} \cdot \frac{1}{\sqrt{n}}}_{\frac{n}{2} \text{ times}} = 0.$$

Then, we note $\langle S, x_1 \rangle = \langle S, j \rangle + \langle S, x_1 - j \rangle = \langle S, x_1 - j \rangle$ and applying cauchy-shwartz yields

$$\langle S, x_1 \rangle = \langle S, x_1 - j \rangle \leq |S|_2 |x_1 - j|_2 = 1 \cdot o(1) = o(1).$$

Now, define $Z = S - \langle S, x_1 \rangle x_1$. Then, we see

$$\langle Z, x_1 \rangle = \langle S, x_1 \rangle - \langle S, x_1 \rangle |x_1|_2^2 = 0.$$

So, Z is orthogonal to x_1 . Hence, there is a $n - 1$ dimensional space, M , generated by x_2, \dots, x_n with eigenvalues $\lambda_2, \dots, \lambda_n$ with largest eigenvalue $\max\{\lambda_2, |\lambda_n|\}$. Then, we find by the rayleigh quotient that $|\langle Ay, y \rangle| \leq \lambda_1(M) |y|_2^2 = \sigma_2 |y|_2^2$ for all $y \in M$. Similarly, we find

$$\lambda_n |y|_2^2 \leq \langle Ay, y \rangle \leq \lambda_2 |y|_2^2$$

for all $y \in M$. From this we get $\lambda_n |Z|_2^2 |\langle AZ, Z \rangle| \leq \lambda_2 |Z|_2^2$, and recalling $|Z|_2 \leq |S|_2 + |\langle S, x_1 \rangle| |x_1|_2 = 1 + o(1) \leq 2$

$$|\langle AZ, Z \rangle| \leq \sigma_2 |Z|_2^2 \leq \sigma_2 |2|_2^2 = 4\sigma_2 = o(n).$$

Finally, we see

$$\begin{aligned} \langle AS, S \rangle &= \langle A(Z + \langle S, x_1 \rangle x_1), Z + \langle S, x_1 \rangle x_1 \rangle \\ &= \underbrace{\langle AZ, Z \rangle}_{o(n)} + \underbrace{\langle S, x_1 \rangle \langle AZ, x_1 \rangle}_{o(1)} + \underbrace{\langle S, x_1 \rangle \langle Ax_1, Z \rangle}_{=0} + \underbrace{\langle S, x_1 \rangle^2 \langle Ax_1, x_1 \rangle}_{o(1)} \\ &= o(n) + \langle S, x_1 \rangle^2 \langle Ax_1, x_1 \rangle \\ &= o(n) + \lambda_1 \\ &= o(n^2) \end{aligned}$$

Recall we also know

$$\langle AS, S \rangle = 2e(S) + 2e(G \setminus S) - 2e(S, G \setminus S).$$

and $2e(S) + 2e(G \setminus S) + 2e(S, G \setminus S) = e(G) \geq \frac{1}{4}n^2 + o(n^2)$. Then, adding and dividing yields these identities yields $e(S) + e(G \setminus S) = \frac{n^2}{8} + o(n^2)$. Furthermore, $\sum_{i \in S} d_i = \frac{n^2}{4} + o(n^2)$ $2e(S) + e(S, G \setminus S)$ and $\sum_{i \in G \setminus S} d_i = \frac{n^2}{4} + o(n^2) = 2e(G \setminus S) + e(S, G \setminus S)$. Adding all of the identities thus far yields that $2e(S) - 2e(G \setminus S) = o(n^2)$, hence $e(S) = \frac{1}{16}n^2 + o(n^2)$. \square

We are nearing the end of quasi-random graphs, but note we have always assumed a quasi-random graph to have density $\frac{1}{2}$. These properties are easily generalized to one of density p . We list the generalized properties.

Definition 0.1. 1. (P_2) . A graph is P_2 if

- $e(G) \geq \frac{pn^2}{2} + o(n^2)$
- $\#CW_4 \leq p^4n^4 + o(n^4)$.

2. (P_3) . A graph is P_3 if

- $e(G) \geq \frac{pn^2}{2} + o(n^2)$
- $\lambda_1(G) = pn + o(n)$
- $\sigma_2(G) = o(n)$.

3. (P_7) . A graph is P_7 if

- $\sum_{1 \leq i, j \leq n} |\hat{d}_{ij} - p^2n| = o(n^2)$.