Generalised Linear Models: Logistic regression

Q: Survival of passengers on the Titanic ~ Class

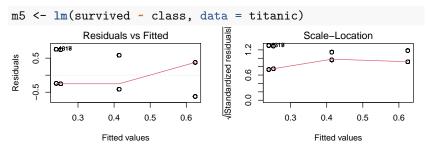
Read titanic_long.csv dataset and fit linear model (survival \sim class).

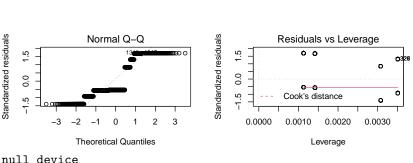
```
class age sex survived
1 first adult male 1
2 first adult male 1
3 first adult male 1
4 first adult male 1
5 first adult male 1
6 first adult male 1
```

Quiz

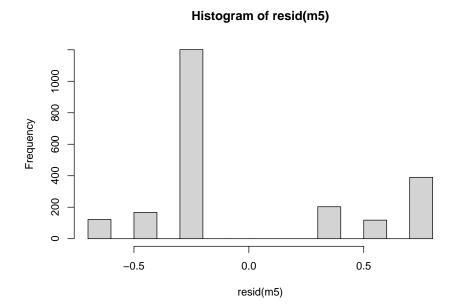
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Let's check linear model:





Weird residuals!



What if your residuals are clearly non-normal, or variance not constant (heteroscedasticity)?

Binary variables (0/1)

Counts (0, 1, 2, 3, ...)

Generalised Linear Models to the rescue!

 $1. \ \, \textbf{Response variable} \, \textbf{-} \, \, \textbf{distribution family}$

- 1. Response variable distribution family
 - ▶ Bernouilli Binomial

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 - Poisson

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 - Gamma

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- 2. Predictors (continuous or categorical)

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- 2. Predictors (continuous or categorical)
- 3. Link function

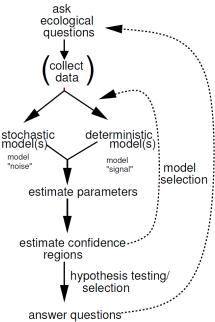
- 1. Response variable distribution family
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 - ▶ Binomial: logit, probit

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 - Poisson: log...

- 1. Response variable distribution family
 - ▶ Bernouilli Binomial
 - Poisson
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 - etc
- 2. Predictors (continuous or categorical)
- 3. Link function
 - ► Gaussian: identity
 - ▶ Binomial: logit, probit
 - Poisson: log...
 - ► See family.

The modelling process



Bernouilli - Binomial distribution (Logistic regression)

Response variable: **Yes/No** (e.g. survival, sex, presence/absence) Link function: logit (others possible, see family)

$$logit(p) = \ln\left(\frac{p}{1-p}\right)$$

Then

$$Pr(alive) = a + bx$$

$$logit(Pr(alive)) = a + bx$$

$$Pr(alive) = invlogit(a + bx) = \frac{e^{a+bx}}{1 + e^{a+bx}}$$

Back to survival of Titanic passengers

How many survived in each class?

```
table(titanic$class, titanic$survived)
```

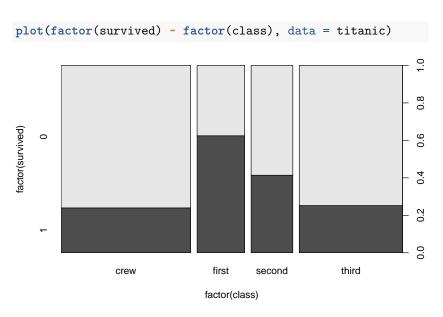
```
crew 673 212
first 122 203
second 167 118
third 528 178
```

Back to survival of Titanic passengers (dplyr)

Passenger survival according to class

```
titanic %>%
  group_by(class, survived) %>%
  summarise(count = n())
# A tibble: 8 \times 3
# Groups: class [4]
  class survived count
  <chr> <int> <int>
                    673
1 crew
                0
                    212
2 crew
3 first
                0 122
4 first
                  203
                0 167
5 second
                  118
6 second
7 third
                    528
8 third
                    178
```

Or graphically...



Mosaic plots (ggplot2)

```
ggplot(titanic) +
 geom_mosaic(aes(x = product(survived, class))) +
 labs(x = "", y = "Survived")
```

Fitting GLMs in R: glm

```
tit.glm <- glm(survived ~ class, data = titanic, family = binomial)</pre>
```

which corresponds to

$$\begin{split} logit(Pr(survival)_i) = a + b \cdot class_i \\ logit(Pr(survival)_i) = a + b_{first} + c_{second} + d_{third} \end{split}$$

Fitting GLMs in R: glm

```
tit.glm <- glm(survived ~ class, data = titanic, family = binomial)
Call:
glm(formula = survived ~ class, family = binomial, data = titanic)
Deviance Residuals:
   Min
            10 Median
                             3Q
                                    Max
-1.3999 -0.7623 -0.7401 0.9702 1.6906
Coefficients:
          Estimate Std. Error z value Pr(>|z|)
classfirst 1.66434 0.13902 11.972 < 2e-16 ***
classsecond 0.80785 0.14375 5.620 1.91e-08 ***
classthird 0.06785 0.11711 0.579 0.562
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
   Null deviance: 2769.5 on 2200 degrees of freedom
Residual deviance: 2588.6 on 2197 degrees of freedom
AIC: 2596.6
Number of Fisher Scoring iterations: 4
These estimates are in logit scale!
```

Interpreting logistic regression output

```
Parameter estimates (logit-scale)
(Intercept) classfirst classsecond classthird
-1.15515905 1.66434399 0.80784987 0.06784632
```

We need to back-transform: apply inverse logit

Crew probability of survival:

```
plogis(coef(tit.glm)[1])
(Intercept)
    0.239548
```

Looking at the data, the proportion of crew who survived is [1] 0.239548

Q: Probability of survival for 1st class passengers?

Must add intercept (baseline) to the parameter estimate:

```
plogis(coef(tit.glm)[1] + coef(tit.glm)[2])
(Intercept)
  0.6246154
```

Again this value matches the data:

```
sum(titanic$survived[titanic$class == "first"]) /
nrow(titanic[titanic$class == "first", ])
```

[1] 0.6246154

Model interpretation using effects package

0.2395480 0.6246154 0.4140351 0.2521246

```
library(effects)
allEffects(tit.glm)
model: survived ~ class

class effect
class
    crew first second third
```

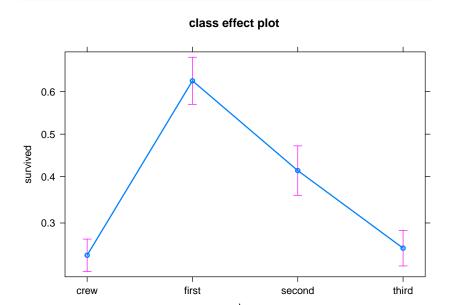
Presenting model results

kable(xtable::xtable(tit.glm), digits = 2)

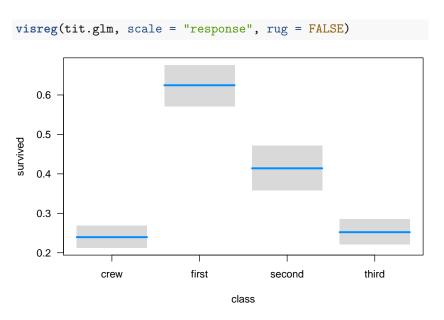
	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-1.16	0.08	-14.67	0.00
classfirst	1.66	0.14	11.97	0.00
classsecond	0.81	0.14	5.62	0.00
classthird	0.07	0.12	0.58	0.56

Visualising model: effects package

plot(allEffects(tit.glm))



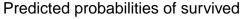
Visualising model: visreg package

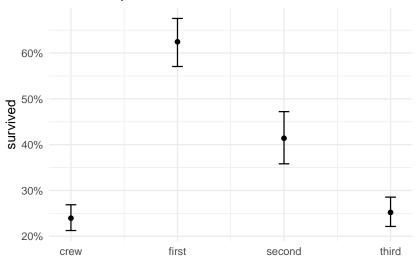


Visualising model: sjPlot package

```
sjPlot::plot_model(tit.glm, type = "eff")
```

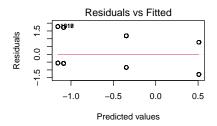
\$class

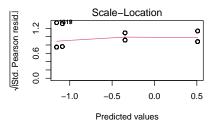


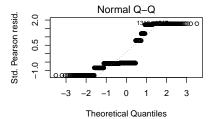


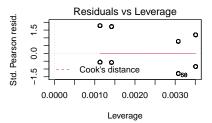
Logistic regression: model checking

Not very useful







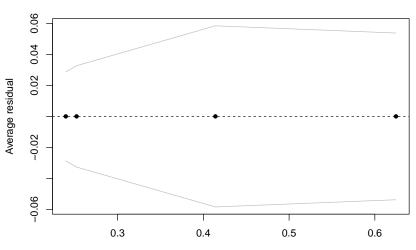


null device

Binned residual plots for logistic regression

```
predvals <- predict(tit.glm, type="response")
arm::binnedplot(predvals, titanic$survived - predvals)</pre>
```

Binned residual plot



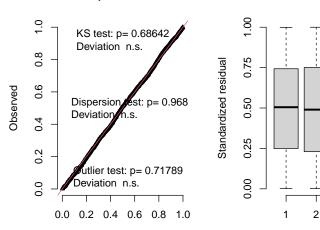
Expected Values

Residual diagnostics with DHARMa

```
library(DHARMa)
simulateResiduals(tit.glm, plot = TRUE)
```

DHARMa residual diagnostics

QQ plot residuals



Model checking with simulated data

```
library(bayesplot)
sims <- simulate(tit.glm, nsim = 100)</pre>
ppc_bars(titanic$survived, yrep = t(as.matrix(sims)))
        1200
      Count
         800
         400
                      0.0
                                 0.5
                                           1.0
                                                     1.5
```

Pseudo R-squared for GLMs

```
library(performance)
r2(tit.glm)
```

\$R2_Tjur Tjur's R2 0.08650663

But many caveats apply! (e.g. see here and here)

1. Visualise data

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- 2. **Fit model**: glm. Don't forget to specify family!

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- 2. Fit model: glm. Don't forget to specify family!
- 3. Examine model: summary
- Back-transform parameters from logit into probability scale (e.g. allEffects)
- 5. **Plot model**: plot(allEffects(model)), visreg, plot_model...

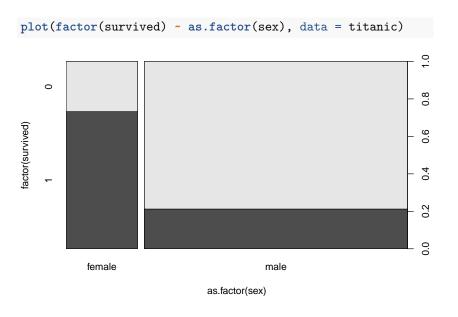
- 1. Visualise data
- 2. Fit model: glm. Don't forget to specify family!
- 3. Examine model: summary
- Back-transform parameters from logit into probability scale (e.g. allEffects)
- Plot model: plot(allEffects(model)), visreg, plot_model...
- 6. Examine residuals: DHARMa::simulateResiduals.

Q: Did men have higher survival than women?

Quiz

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Plot first



Fit model

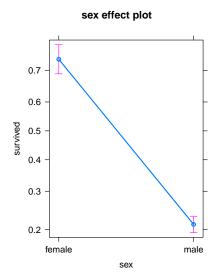
Call:

ATC: 2339

```
glm(formula = survived ~ sex, family = binomial, data = titanic)
Deviance Residuals:
   Min 10 Median 30 Max
-1.6226 -0.6903 -0.6903 0.7901 1.7613
Coefficients:
           Estimate Std. Error z value Pr(>|z|)
(Intercept) 1.0044 0.1041 9.645 <2e-16 ***
sexmale -2.3172 0.1196 -19.376 <2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
   Null deviance: 2769.5 on 2200 degrees of freedom
Residual deviance: 2335.0 on 2199 degrees of freedom
```

Effects

model: survived ~ sex
sex effect
sex
female male
0.7319149 0.2120162



Q: Did women have higher survival because they travelled more in first class?

Let's look at the data

```
table(titanic$class, titanic$survived, titanic$sex)
    = female
          3 20
  crew
  first
          4 141
  second 13
            93
  third 106
             90
   = male
        670 192
  crew
        118
  first
             62
```

third 422 Mmmm...

second 154

25

88

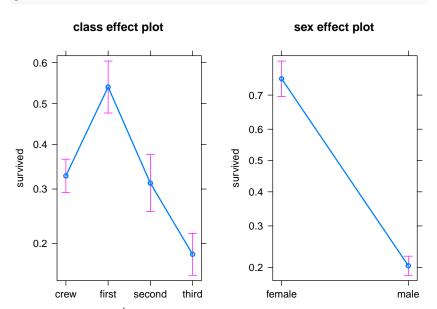
Quiz

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Fit additive model with both factors

Plot additive model

plot(allEffects(tit.sex.class))

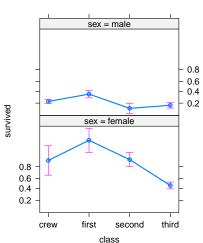


Fit model with both factors (interactions)

```
tit.sex.class <- glm(survived ~ class * sex, family = binomial,
glm(formula = survived ~ class * sex, family = binomial, data =
                 coef.est coef.se
(Intercept)
                  1.90 0.62
classfirst
                1.67 0.80
classsecond
               0.07 0.69
classthird -2.06 0.64
           -3.15 0.62
sexmale
classfirst:sexmale -1.06 0.82
classsecond:sexmale -0.64 0.72
classthird:sexmale 1.74 0.65
 n = 2201, k = 8
 residual deviance = 2163.7, null deviance = 2769.5 (difference
```

Effects

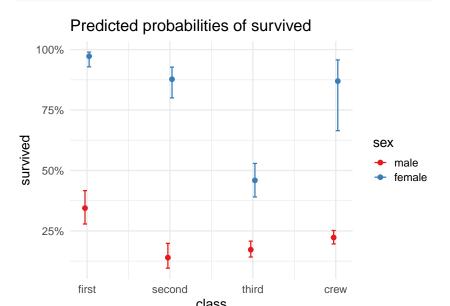
class*sex effect plot



So, women had higher probability of survival than men, even within the same class.

Effects (sjPlot)

```
plot_model(tit.sex.class, type = "int")
```



Logistic regression for proportion data

Read Titanic data in different format

Read titanic_prop.csv data.

```
X Class Sex Age No Yes
1 1 1st Female Adult 4 140
2 2 1st Female Child 0 1
3 3 1st Male Adult 118 57
4 4 1st Male Child 0 5
5 5 2nd Female Adult 13 80
6 6 2nd Female Child 0 13
```

These are the same data, but summarized (see Freq variable).

Use cbind(n.success, n.failures) as response

```
prop.glm <- glm(cbind(Yes, No) ~ Class, data = tit.prop, family</pre>
Call:
glm(formula = cbind(Yes, No) ~ Class, family = binomial, data =
Deviance Residuals:
   Min 10 Median 30 Max
-9.6404 -0.2915 1.5698 5.0366 10.1516
Coefficients:
          Estimate Std. Error z value Pr(>|z|)
(Intercept) 0.5092 0.1146 4.445 8.79e-06 ***
Class2nd -0.8565 0.1661 -5.157 2.51e-07 ***
Class3rd -1.5965 0.1436 -11.114 < 2e-16 ***
ClassCrew -1.6643 0.1390 -11.972 < 2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
```

Effects

```
model: cbind(Yes, No) ~ Class

Class effect
Class

1st 2nd 3rd Crew
0.6246154 0.4140351 0.2521246 0.2395480

Compare with former model based on raw data:
```

0.2395480 0.6246154 0.4140351 0.2521246

```
model: survived ~ class

class effect
class

crew first second third
```

Same results!

Logistic regression with continuous predictors

Example dataset: GDP and infant mortality

Read UN_GDP_infantmortality.csv.

country	mortality	gdp
Length: 207	Min. : 2.00	Min. : 36
Class :character	1st Qu.: 12.00	1st Qu.: 442
Mode :character	Median : 30.00	Median : 1779
	Mean : 43.48	Mean : 6262
	3rd Qu.: 66.00	3rd Qu.: 7272
	Max. :169.00	Max. :42416
	NA's :6	NA's :10

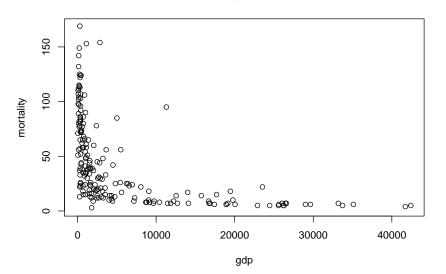
Q: Is infant mortality related to GDP?

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EDA

plot(mortality ~ gdp, data = gdp, main = "Infant mortality (per

Infant mortality (per 1000 births)



Fit model

```
gdp.glm <- glm(cbind(mortality, 1000 - mortality) ~ gdp,</pre>
              data = gdp, family = binomial)
Call:
glm(formula = cbind(mortality, 1000 - mortality) ~ gdp, family =
   data = gdp)
Deviance Residuals:
   Min 10 Median 30 Max
-9.2230 -3.5163 -0.5697 2.4284 13.5849
Coefficients:
             Estimate Std. Error z value Pr(>|z|)
(Intercept) -2.657e+00 1.311e-02 -202.76 <2e-16 ***
gdp -1.279e-04 3.458e-06 -36.98 <2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
```

Effects

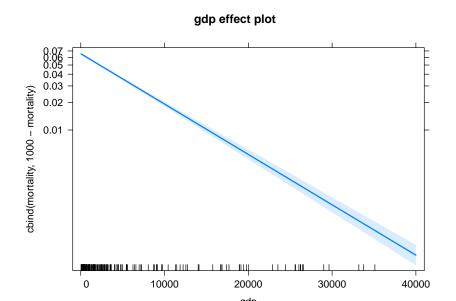
```
allEffects(gdp.glm)
model: cbind(mortality, 1000 - mortality) ~ gdp

gdp effect
gdp
40 10000 20000 30000 40000
```

0.0652177296 0.0191438829 0.0054028095 0.0015096074 0.0004206154

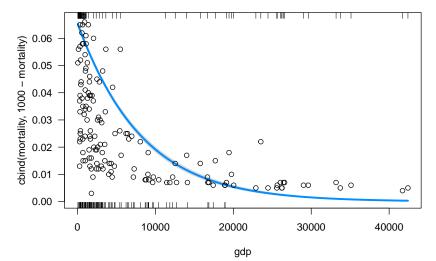
Effects plot

plot(allEffects(gdp.glm))



Plot model using visreg:

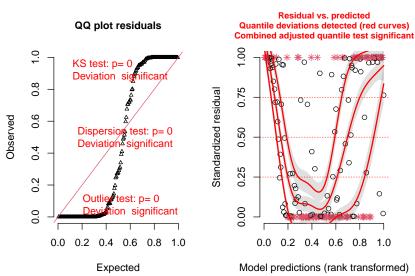
```
visreg(gdp.glm, scale = "response")
points(mortality/1000 ~ gdp, data = gdp)
```



Residuals diagnostics with DHARMa

simulateResiduals(gdp.glm, plot = TRUE)

DHARMa residual diagnostics





Testing for overdispersion (DHARMa)

```
simres <- simulateResiduals(gdp.glm, refit = TRUE)
testDispersion(simres, plot = FALSE)</pre>
```

DHARMa nonparametric dispersion test via mean deviance residus. simulated-refitted

data: simres
dispersion = 21, p-value < 2.2e-16
alternative hypothesis: two.sided</pre>

Overdispersion in logistic regression with proportion data

```
Call:
glm(formula = cbind(mortality, 1000 - mortality) ~ gdp, family =
   data = gdp)
```

Deviance Residuals:

```
Min 1Q Median 3Q Max
-9.2230 -3.5163 -0.5697 2.4284 13.5849
```

Coefficients:

```
Estimate Std. Error t value Pr(>|t|)
(Intercept) -2.657e+00 5.977e-02 -44.465 < 2e-16 ***
gdp -1.279e-04 1.577e-05 -8.111 5.96e-14 ***
```

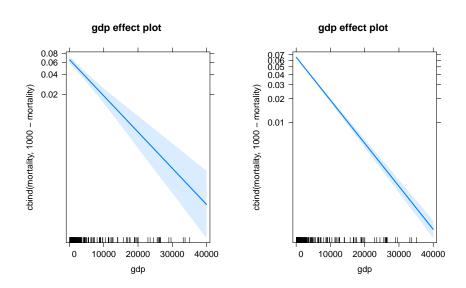
```
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

(Dispersion parameter for quasibinomial family taken to be 20.79

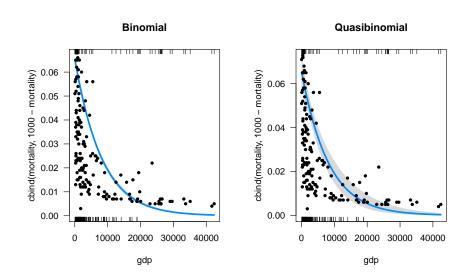
Mean estimates do not change after accounting for overdispersion

```
model: cbind(mortality, 1000 - mortality) ~ gdp
gdp effect
gdp
                    10000
                                  20000
          40
                                               30000
                                                            40000
0.0652177296 0.0191438829 0.0054028095 0.0015096074 0.0004206154
 model: cbind(mortality, 1000 - mortality) ~ gdp
gdp effect
gdp
                    10000
                                  20000
                                               30000
          40
                                                            40000
0.0652177296 0.0191438829 0.0054028095 0.0015096074 0.0004206154
```

But standard errors (uncertainty) do!



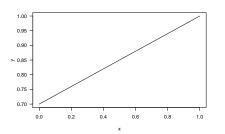
Plot model and data

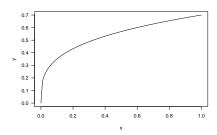




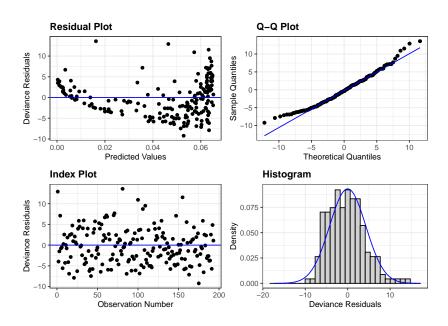
Whenever you fit logistic regression to **proportion** data, check family quasibinomial.

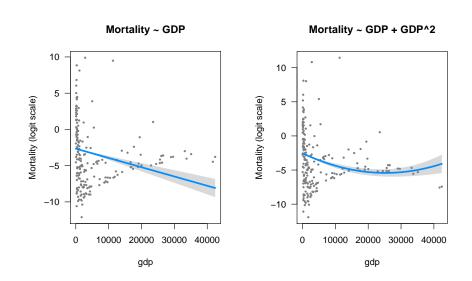
 $\label{eq:continuous} \begin{array}{l} y\sim x+z\\ \text{Really? Not everything has to be linear! Actually, it often is not.} \end{array}$ Think about shape of relationship. See chapter 3 in Bolker's book.

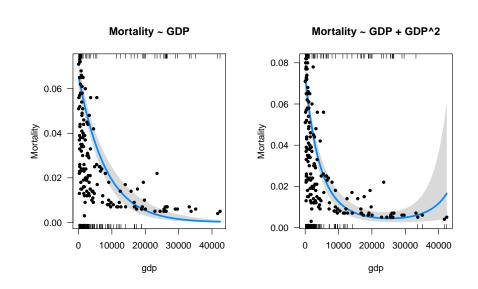


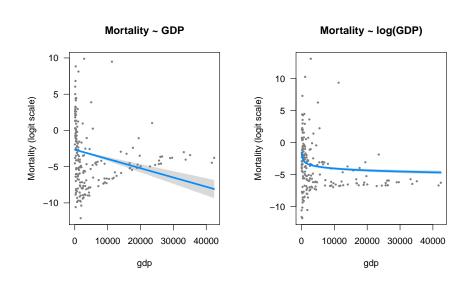


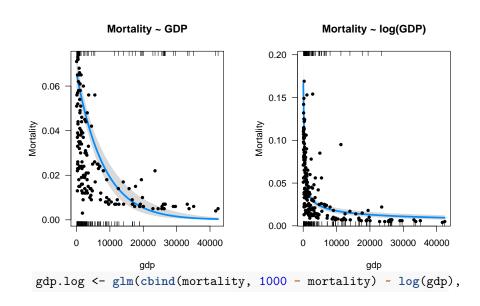
```
visreg(gdp.glm, ylab = "Mortality (logit scale)")
     10
      5
Mortality (logit scale)
     -5
    -10
                         10000
                                         20000
                                                        30000
                                                                        40000
                                            gdp
```











More examples

 seedset.csv: Comparing seed set among plants (Data from Harder et al. 2011)

Seed set among plants

```
seed <- readr::read csv("data/seedset.csv")</pre>
head(seed)
# A tibble: 6 x 6
 species plant pcmass fertilized seeds ovulecnt
 <chr>
           <dbl> <dbl>
                            <dbl> <dbl>
                                         <dbl>
1 ferruginea
               2 0
                              70
                                    52
                                           330
               2 0.2
2 ferruginea
                             321
                                   188
                                           461
3 ferruginea 2 0.485
                                   278
                             351
                                           435
4 ferruginea 2 0.737
                             386
                                   301
                                           430
5 ferruginea
               2 1
                             367
                                   342
                                           419
                             185
                                    39
                                           470
6 ferruginea
seed$plant <- as.factor(seed$plant)</pre>
```

Questions:

https://pollev.com/franciscorod726

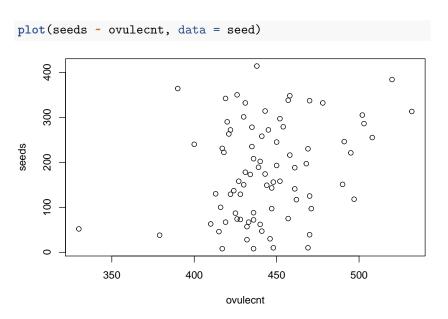
ls seed set related to proportion of outcross pollen (pcmass)?

Questions:

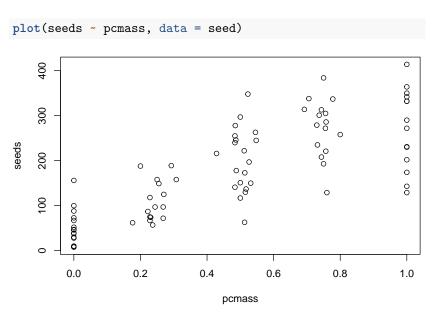
https://pollev.com/franciscorod726

- ▶ Is seed set related to proportion of outcross pollen (pcmass)?
- ▶ Which plant had lower seed set?

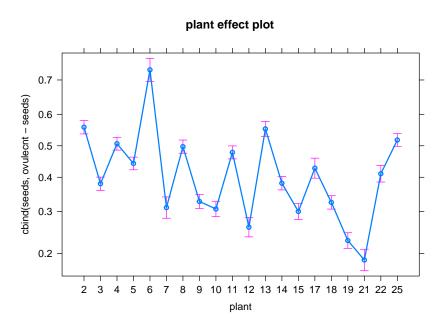
Number of seeds vs Number of ovules



Number of seeds vs Proportion outcross pollen



Seed set across plants



Seed set ∼ outcross pollen

