# Mixed / Multilevel Models

Francisco Rodríguez-Sánchez

https://frodriguezsanchez.net

## Example dataset: trees

· Data on 1000 trees from 10 sites.

### head(trees)

dead	sex	height	dbh	site	
0	male	36.1	29.68	4	1
0	male	42.3	33.29	5	2
0	female	41.9	28.03	2	3
0	female	46.5	39.86	5	4
0	female	43.9	47.94	1	5
0	male	26.2	10.82	1	6

## Example dataset: trees

- · Data on 1000 trees from 10 sites.
- Trees per site: 4 392.

### head(trees)

	site	dbh	height	sex	dead
1	4	29.68	36.1	male	0
2	5	33.29	42.3	male	0
3	2	28.03	41.9	female	0
4	5	39.86	46.5	female	0
5	1	47.94	43.9	female	0
6	1	10.82	26.2	male	0

Q: What's the relationship between tree diameter and height?

### A simple linear model

```
lm.simple <- lm(height ~ dbh, data = trees)</pre>
Call:
lm(formula = height ~ dbh, data = trees)
Residuals:
    Min 10 Median 30
                                   Max
-13.3270 -2.8978 0.1057 2.7924 12.9511
Coefficients:
          Estimate Std. Error t value Pr(>|t|)
(Intercept) 19.33920 0.31064 62.26 <2e-16 ***
dbh 0.61570 0.01013 60.79 <2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

ŀ

#### Remember our model structure

$$y_i \sim N(\mu_i, \sigma^2)$$
  
 $\mu_i = \alpha + \beta x_i$ 

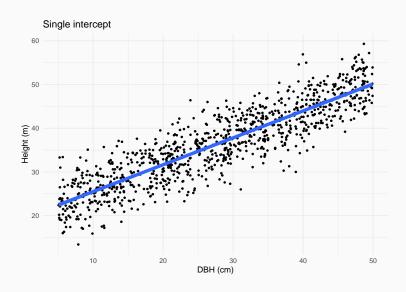
In this case:

Height
$$_i \sim N(\mu_i, \sigma^2)$$
  
 $\mu_i = \alpha + \beta DBH_i$ 

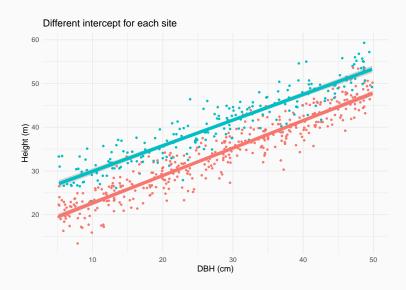
lpha: expected height when DBH = 0

eta: how much height increases with every unit increase of DBH

## There is only one intercept



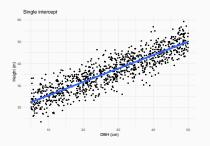
# What if allometry varies among sites?

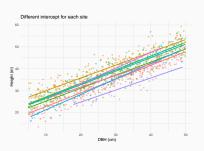


## Fitting a varying intercepts model with lm

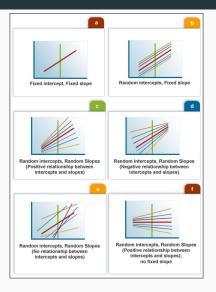
```
Call:
lm(formula = height ~ factor(site) + dbh, data = trees)
Residuals:
    Min
              10
                   Median
                               30
                                       Max
         -1.9885
-10.1130
                   0.0582
                           2.0314 11.3320
Coefficients:
               Estimate Std. Error t value Pr(>|t|)
(Intercept)
              16.699037
                         0.260565 64.088 < 2e-16 ***
factor(site)2
             6.504303
                         0.256730 25.335 < 2e-16 ***
factor(site)3 4.357457
                         0.354181 12.303 < 2e-16 ***
factor(site)4 1.934650
                         0.356102 5.433 6.98e-08 ***
factor(site)5 3.637432
                         0.339688 10.708 < 2e-16 ***
factor(site)6
             4,204511
                         0.421906 9.966 < 2e-16 ***
factor(site)7
                                           0.7916
              -0.176193
                         0.666772 -0.264
factor(site)8
              -5.312648
                         0.893603 -5.945 3.82e-09 ***
factor(site)9 5.437049
                         1.087766 4.998 6.84e-07 ***
factor(site)10 2.263338
                         1.369986 1.652
                                            0.0988 .
dbh
               0.617075
                         0.007574 81.473 < 2e-16 ***
```

# Single vs varying intercept





### Mixed models enable us to account for variability



www.esourceresearch.org/

### Mixed model with varying intercepts

$$y_{i} = a + \alpha_{j} + b \cdot x_{i} + \varepsilon_{i}$$
$$\alpha_{j} \sim N(0, \tau^{2})$$
$$\varepsilon_{i} \sim N(0, \sigma^{2})$$

In our example:

$$\begin{aligned} \textit{Height}_i &= a + \textit{site}_j + b \cdot \textit{DBH}_i + \varepsilon_i \\ &\textit{site}_j \sim \textit{N}\left(0, \tau^2\right) \\ &\varepsilon_i \sim \textit{N}\left(0, \sigma^2\right) \end{aligned}$$

Mixed models estimate varying parameters

(intercepts and/or slopes)

with pooling among levels

(rather than considering them fully independent)

• complete pooling: Single overall intercept.

· complete pooling: Single overall intercept.

```
· lm (height ~ dbh)
```

- · complete pooling: Single overall intercept.
  - · lm (height ~ dbh)
- · no pooling: One independent intercept for each site.

- · complete pooling: Single overall intercept.
  - · lm (height ~ dbh)
- · no pooling: One independent intercept for each site.
  - · lm (height ~ dbh + site)

- · complete pooling: Single overall intercept.
  - · lm (height ~ dbh)
- no pooling: One independent intercept for each site.
  - · lm (height ~ dbh + site)
- partial pooling: Inter-related intercepts.

- · complete pooling: Single overall intercept.
  - · lm (height ~ dbh)
- no pooling: One independent intercept for each site.
  - · lm (height ~ dbh + site)
- · partial pooling: Inter-related intercepts.
  - · lmer(height ~ dbh + (1 | site))

1. Fixed effects constant across individuals, random effects vary.

http://andrewgelman.com/2005/01/25/why\_i\_dont\_use/

- 1. Fixed effects constant across individuals, random effects vary.
- 2. Effects are fixed if they are interesting in themselves; random if interest in the underlying population.

http://andrewgelman.com/2005/01/25/why\_i\_dont\_use/

- 1. Fixed effects constant across individuals, random effects vary.
- 2. Effects are fixed if they are interesting in themselves; random if interest in the underlying population.
- 3. Fixed when sample exhausts the population; random when the sample is small part of the population.

http://andrewgelman.com/2005/01/25/why i dont use/

- 1. Fixed effects constant across individuals, random effects vary.
- 2. Effects are fixed if they are interesting in themselves; random if interest in the underlying population.
- 3. Fixed when sample exhausts the population; random when the sample is small part of the population.
- 4. Random effect if it's assumed to be a realized value of random variable.

http://andrewgelman.com/2005/01/25/why i dont use/

- 1. Fixed effects constant across individuals, random effects vary.
- 2. Effects are fixed if they are interesting in themselves; random if interest in the underlying population.
- 3. Fixed when sample exhausts the population; random when the sample is small part of the population.
- 4. Random effect if it's assumed to be a realized value of random variable.
- 5. Fixed effects estimated using least squares or maximum likelihood; random effects estimated with shrinkage.

http://andrewgelman.com/2005/01/25/why\_i\_dont\_use/

## What is a random effect, really?

Varies by group

Random effects are estimated with *partial pooling*, while fixed effects are not (infinite variance).

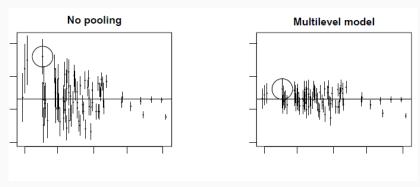
### What is a random effect, really?

- · Varies by group
- Variation estimated with probability model

Random effects are estimated with *partial pooling*, while fixed effects are not (infinite variance).

# Shrinkage improves parameter estimation

Especially for groups with low sample size



From Gelman & Hill p. 253

## Fitting mixed/multilevel models

Correlation of Fixed Effects:

(Intr)

```
library("lme4")
mixed <- lmer(height ~ dbh + (1|site), data = trees)
Linear mixed model fit by REML ['lmerMod']
Formula: height ~ dbh + (1 | site)
  Data: trees
REML criterion at convergence: 5108.3
Scaled residuals:
                                Max
   Min
           10 Median 30
-3.3199 -0.6607 0.0227 0.6716 3.7328
Random effects:
Groups Name Variance Std.Dev.
site (Intercept) 11.195 3.346
Residual
                9.261 3.043
Number of obs: 1000, groups: site, 10
Fixed effects:
           Estimate Std. Error t value
(Intercept) 19.011468 1.100444 17.28
dhh
   0.616927 0.007572 81.47
```

17

#### Retrieve model coefficients

### coef(mixed)

```
$site
   (Intercept)
                     dbh
1
      16.70800 0.6169271
      23.19162 0.6169271
3
      21.04229 0.6169271
4
      18.64086 0.6169271
5
      20.32995 0.6169271
6
      20.88200 0.6169271
      16.61686 0.6169271
8
      11.88302 0.6169271
9
      21.84779 0.6169271
10
      18,97228 0,6169271
```

### Broom: model estimates in tidy form

```
library(broom.mixed)
tidy(mixed)
```

```
# A tibble: 4 x 6
 effect group
                             estimate std.error statistic
                term
 <chr> <chr>
                <chr>>
                                <dbl>
                                       <dbl>
                                                 <dbl>
1 fixed <NA> (Intercept)
                          19.0 1.10
                                               17.3
2 fixed <NA>
                dbh
                                0.617 0.00757 81.5
3 ran pars site
                sd (Intercept)
                                3.35
                                      NΑ
                                                 NΑ
4 ran_pars Residual sd__Observation
                                3.04
                                      NA
                                                 NA
```

See also broom.mixed

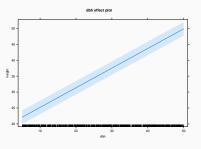
## Visualising model: allEffects

#### allEffects(mixed)

model: height ~ dbh

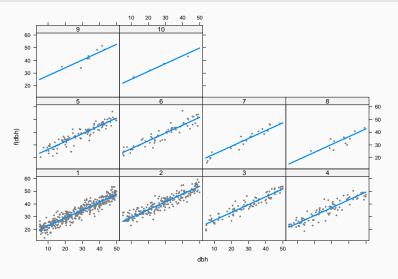
dbh effect

dbh 5 20 30 40 50 22.09610 31.35001 37.51928 43.68855 49.85782



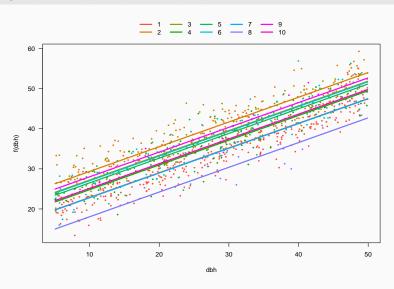
# Visualising model: visreg

visreg(mixed, xvar = "dbh", by = "site", re.form = NULL)

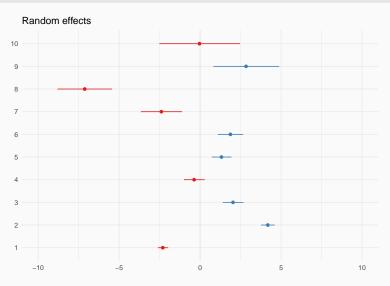


# Visualising model

visreg(mixed, xvar = "dbh", by = "site", re.form = NULL, overlay = TRUE)



# Visualising model: sjPlot

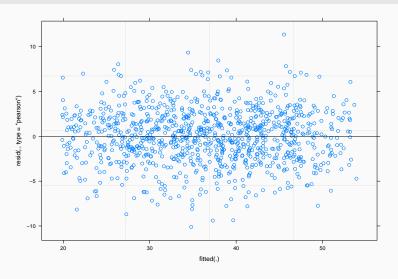


## Using merTools to understand fitted model

```
library("merTools")
shinyMer(mixed)
```

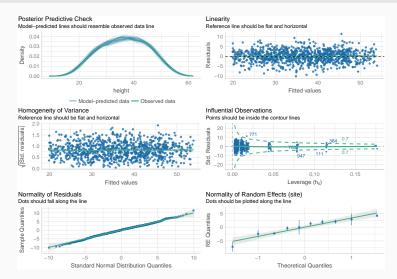
# Checking residuals

## plot(mixed)



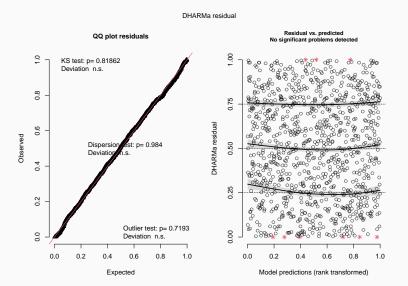
# Checking residuals

# library("performance") check\_model(mixed)



# Checking residuals (DHARMa)

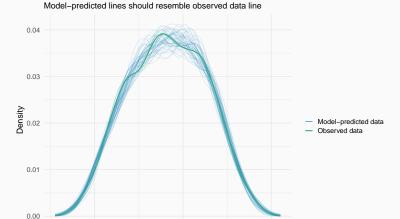
DHARMa::simulateResiduals(mixed, plot = TRUE, re.form = NULL)



# Model checking with simulated data

### check\_predictions(mixed)

#### Posterior Predictive Check



heiaht

### R-squared for GLMMs

Many approaches! Somewhat polemic (e.g. see this).

Nakagawa & Schielzeth propose marginal (considering fixed effects only) and conditional  $R^2$  (including random effects too):

#### r2(mixed)

# R2 for Mixed Models

Conditional R2: 0.888 Marginal R2: 0.753

# Growing the hierarchy: adding site-level predictors

# Model with group-level predictors

We had:

$$y_i = a + \alpha_j + b \cdot x_i + \varepsilon_i$$
$$\alpha_j \sim N(0, \tau^2)$$
$$\varepsilon_i \sim N(0, \sigma^2)$$

Now

$$egin{aligned} y_i &= a + lpha_j + b \cdot x_i + arepsilon_i \ & lpha_j \sim \mathit{N}(\mu_j, au^2) \ & \mu_j &= \delta \cdot \mathit{Predictor}_j \ & arepsilon_i \sim \mathit{N}(0, \sigma^2) \end{aligned}$$

# Are height differences among sites related to temperature?

$$\begin{aligned} \textit{Height}_i &= \textit{site}_j + b \cdot \textit{DBH}_i + \varepsilon_i \\ &\quad \textit{site}_j \sim \textit{N}(\mu_j, \tau^2) \\ \mu_j &= a + \delta \cdot \textit{Temperature}_j \\ &\quad \varepsilon_i \sim \textit{N}(0, \sigma^2) \end{aligned}$$

# Are height differences among sites related to temperature?

```
sitedata <- read.csv("data/sitedata.csv")
sitedata</pre>
```

```
site temp
      1 15.1
2
      2 22.0
3
      3 20.1
4
      4 20.4
5
      5 20.0
6
      6 20.1
      7 17.5
8
      8 14.6
9
      9 19.2
10
     10 16.0
```

# Merging trees and site data

```
trees.full <- merge(trees, sitedata, by = "site")
head(trees.full)</pre>
```

```
      site
      dbh height
      sex dead temp

      1
      1 21.05
      32.2
      male
      0 15.1

      2
      1 46.63
      45.9 female
      0 15.1

      3
      1 43.86
      45.5
      male
      0 15.1

      4
      1 29.03
      35.5
      male
      0 15.1

      5
      1 6.02
      21.1
      male
      0 15.1

      6
      1 40.82
      38.7
      male
      0 15.1
```

#### Fit multilevel model

```
group.pred <- lmer(height ~ dbh + (1 | site) + temp, data = trees.full)</pre>
Linear mixed model fit by REML ['lmerMod']
Formula: height ~ dbh + (1 | site) + temp
  Data: trees.full
REML criterion at convergence: 5098.2
Scaled residuals:
   Min 10 Median 30
                                Max
-3.3247 -0.6517 0.0192 0.6663 3.7268
Random effects:
Groups Name Variance Std.Dev.
site (Intercept) 3.158 1.777
Residual
           9.266 3.044
Number of obs: 1000, groups: site, 10
Fixed effects:
           Estimate Std. Error t value
(Intercept) -1.730910 4.671330 -0.371
dbh
       0.616894 0.007571 81.484
temp 1.115104 0.248000 4.496
```

Correlation of Fixed Effects:
(Intr) dbh

### Centre (and scale) continuous variables

```
mean(sitedata$temp)
```

[1] 18.5

```
trees.full$temp.c <- trees.full$temp - 18</pre>
```

Temperatures now referred as deviations from 18 °C (close to average)

#### Fit multilevel model

```
group.pred <- lmer(height ~ dbh + (1 | site) + temp.c, data = trees.full)</pre>
Linear mixed model fit by REML ['lmerMod']
Formula: height ~ dbh + (1 | site) + temp.c
  Data: trees.full
REML criterion at convergence: 5098.2
Scaled residuals:
   Min 10 Median 30
                                Max
-3.3247 -0.6517 0.0192 0.6663 3.7268
Random effects:
Groups Name Variance Std.Dev.
site (Intercept) 3.158 1.777
Residual
           9.266 3.044
Number of obs: 1000, groups: site, 10
Fixed effects:
           Estimate Std. Error t value
(Intercept) 18.340954 0.655054 27.999
dbh
        0.616894 0.007571 81.484
temp.c 1.115104 0.248000 4.496
```

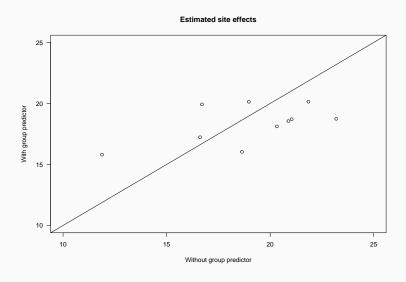
Correlation of Fixed Effects:
(Intr) dbh

37

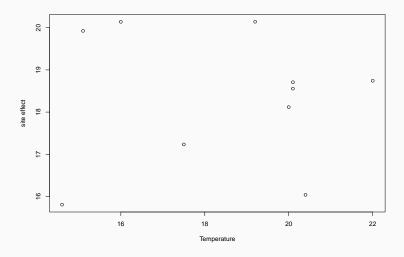
### Examine model with merTools

shinyMer(group.pred)

# Comparing site effects with and without group predictor



# Are site effects related to temperature?



There is overall difference in height among sites (different intercepts)

AND

Relationship between DBH and Height varies among sites (different slopes)

```
mixed.slopes <- lmer(height ~ dbh + (1 + dbh | site), data=trees)</pre>
```

```
Linear mixed model fit by REML ['lmerMod']
Formula: height ~ dbh + (1 + dbh | site)
  Data: trees
REML criterion at convergence: 5105.1
Scaled residuals:
   Min 10 Median 30 Max
-3.3342 -0.6599 0.0375 0.6916 3.7756
Random effects:
Groups Name Variance Std.Dev. Corr
 site (Intercept) 1.566e+01 3.95671
        dbh
            3.087e-04 0.01757 -1.00
Residual 9.226e+00 3.03744
Number of obs: 1000, groups: site, 10
Fixed effects:
          Estimate Std. Error t value
```

(Intercept) 18.95272 1.29190 14.67

0.61837 0.00946 65.37

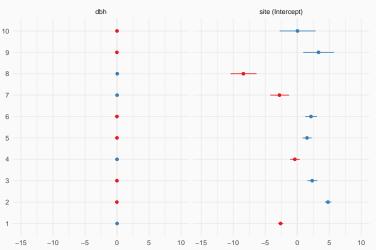
dbh

```
$site
   (Intercept)
                      dbh
1
      16.34655 0.6299443
2
      23.74733 0.5970814
3
      21.28802 0.6080019
4
      18,57844 0,6200337
5
      20,47961 0,6115916
6
      21.09608 0.6088542
7
      16.17675 0.6306983
8
      10.54681 0.6556978
9
      22.27301 0.6036281
10
      18,99463 0,6181856
attr(,"class")
```

[1] "coef.mer"

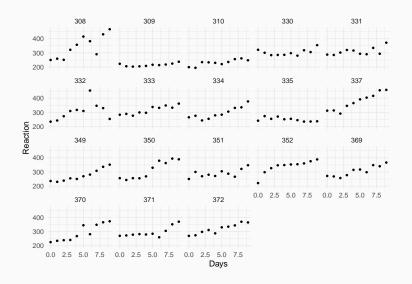
# Visualising model: sjPlot

#### Random effects



# More examples

# sleepstudy (repeated measures)



# Varying intercepts and slopes (lme4)

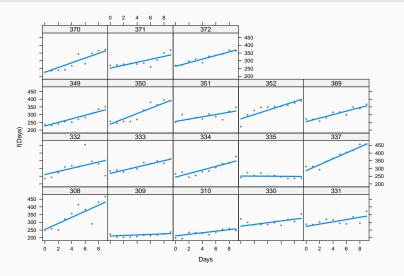
(Intr)

```
sleep <- lmer(Reaction ~ Days + (1+Days|Subject), data = sleepstudy)</pre>
Linear mixed model fit by REML ['lmerMod']
Formula: Reaction ~ Days + (1 + Days | Subject)
  Data: sleepstudy
REML criterion at convergence: 1743.6
Scaled residuals:
   Min
       10 Median 30
                                Max
-3.9536 -0.4634 0.0231 0.4634 5.1793
Random effects:
Groups Name
               Variance Std.Dev. Corr
Subject (Intercept) 612.10 24.741
         Days 35.07 5.922 0.07
Residual
           654.94 25.592
Number of obs: 180. groups: Subject. 18
Fixed effects:
          Estimate Std. Error t value
(Intercept) 251.405 6.825 36.838
Davs 10.467 1.546 6.771
Correlation of Fixed Effects:
```

48

# Varying intercepts and slopes (lme4)

visreg(sleep, xvar = "Days", by = "Subject", re.form = NULL)



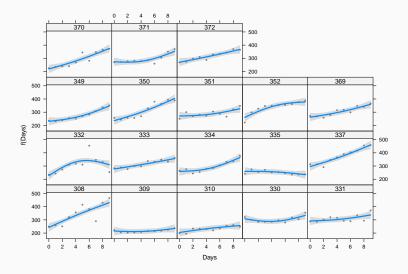
# Fitting multilevel models (GAMM) with mgcv

```
sgamm <- mgcv::gam(Reaction ~ s(Days, Subject, k = 3, bs = "fs"),</pre>
                  data = sleepstudy, method = "REML")
Family: gaussian
Link function: identity
Formula:
Reaction \sim s(Days, Subject, k = 3, bs = "fs")
Parametric coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 298.51 9.05 32.98 <2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Approximate significance of smooth terms:
                  edf Ref.df F p-value
```

s(Davs.Subject) 45.67 53 17.11 <2e-16 \*\*\*

# Fitting multilevel models (GAMM) with mgcv

visreg(sgamm, xvar = "Days", by = "Subject")



### Fitting multilevel models (GAMM) with mgcv

# Hierarchical generalized additive models: an introduction with mgcv

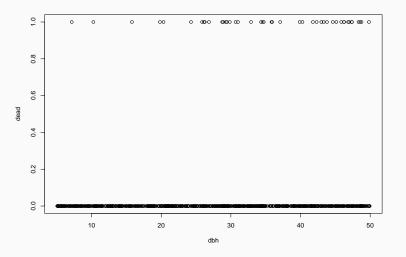
```
Eric J Pedersen Corresp., 1,2, David L. Miller 3,4, Gavin L. Simpson 5, Noam Ross 6
```

https://doi.org/10.7287/peerj.preprints.27320v1

# Multilevel logistic regression

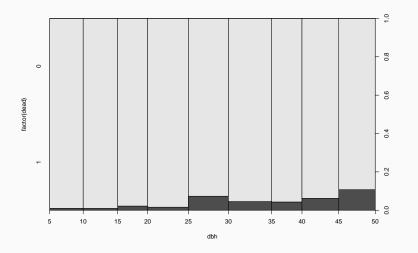
# Q: Relationship between tree size and mortality





# Q: Relationship between tree size and mortality





# Fit simple logistic regression

```
simple.logis <- glm(dead ~ dbh, data = trees, family=binomial)</pre>
Call:
glm(formula = dead ~ dbh, family = binomial, data = trees)
Deviance Residuals:
   Min 1Q Median 3Q
                                      Max
-0.4805 -0.3520 -0.2647 -0.1928 2.9690
Coefficients:
           Estimate Std. Error z value Pr(>|z|)
(Intercept) -4.77874   0.50902   -9.388   < 2e-16 ***
dhh
      0.05365 0.01377 3.895 9.82e-05 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
   Null deviance: 360.91 on 999 degrees of freedom
Residual deviance: 343.69 on 998 degrees of freedom
AIC: 347.69
Number of Fisher Scoring iterations: 6
```

# Logistic regression with independent site effects

```
logis2 <- glm(dead ~ dbh + factor(site), data = trees, family=binomial)</pre>
Call:
glm(formula = dead ~ dbh + factor(site), family = binomial, data = trees)
Deviance Residuals:
   Min
            1Q Median
                            3Q
                                    Max
-0.6359 -0.3449 -0.2561 -0.1852
                                2.9763
Coefficients:
               Estimate Std. Error z value Pr(>|z|)
(Intercept)
               -4.80123 0.54985 -8.732 <2e-16 ***
dhh
               0.05371 0.01381 3.889 0.0001 ***
factor(site)2 -0.29692 0.46073 -0.644
                                          0.5193
factor(site)3 0.21275 0.52799 0.403
                                          0.6870
factor(site)4 0.39841 0.53025 0.751
                                          0.4524
factor(site)5 -0.42557 0.64018 -0.665
                                          0.5062
factor(site)6 0.66861 0.53656 1.246
                                          0.2127
factor(site)7 0.11862 1.06211 0.112
                                          0.9111
factor(site)8 0.43899 1.08058 0.406
                                          0.6846
factor(site)9 -13.63389 840.90382 -0.016
                                          0.9871
factor(site)10 -13.17148 1042.21823 -0.013
                                          0.9899
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

# Fit multilevel logistic regression

```
mixed.logis <- glmer(dead ~ dbh + (1|site), data=trees, family = binomial)
Generalized linear mixed model fit by maximum likelihood (Laplace
 Approximation) [glmerMod]
Family: binomial (logit)
Formula: dead ~ dbh + (1 | site)
  Data: trees
   AIC
             BIC logLik deviance df.resid
  349.7 364.4 -171.8 343.7
                                     997
Scaled residuals:
   Min
       1Q Median 3Q
                                 Max
-0.3498 -0.2528 -0.1888 -0.1370 9.0031
Random effects:
Groups Name Variance Std.Dev.
site (Intercept) 0 0
Number of obs: 1000, groups: site, 10
Fixed effects:
           Estimate Std. Error z value Pr(>|z|)
(Intercept) -4.77874    0.50904   -9.388    < 2e-16 ***
dhh
        0.05365 0.01377 3.895 9.83e-05 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

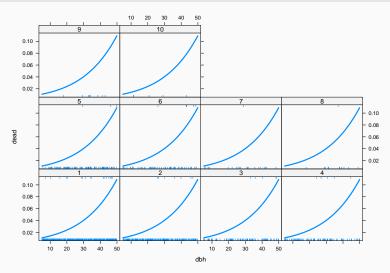
#### Retrieve model coefficients

#### coef(mixed.logis)

```
$site
   (Intercept) dbh
1
    -4.778744 0.05364989
2
  -4.778744 0.05364989
3
    -4.778744 0.05364989
4
    -4.778744 0.05364989
5
    -4.778744 0.05364989
6
    -4.778744 0.05364989
    -4.778744 0.05364989
8
   -4.778744 0.05364989
9
    -4.778744 0.05364989
10
   -4.778744 0.05364989
```

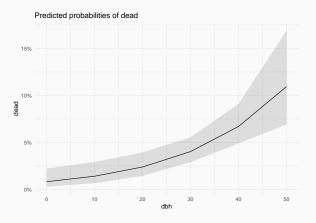
# Visualising model: visreg

visreg(mixed.logis, xvar = "dbh", by = "site", scale = "response")



# Visualising model: sjPlot

#### \$dbh



- Perfect for  $structured\ data\ (space-time)$ 

- Perfect for structured data (space-time)
- $\cdot$  Predictors enter at the appropriate level

- · Perfect for structured data (space-time)
- Predictors enter at the appropriate level
- · Accommodate variation in treatment effects

- Perfect for structured data (space-time)
- Predictors enter at the appropriate level
- Accommodate variation in treatment effects
- · More efficient inference of regression parameters

- Perfect for structured data (space-time)
- Predictors enter at the appropriate level
- Accommodate variation in treatment effects
- · More efficient inference of regression parameters
- $\boldsymbol{\cdot}$  Using all the data to perform inferences for groups with  $\boldsymbol{small}$   $\boldsymbol{sample}$   $\boldsymbol{size}$

Varying intercepts

Varying intercepts

```
\cdot y ~ x + (1 | group)
```

Varying intercepts

$$\cdot$$
 y ~ x + (1 | group)

Varying intercepts and slopes

Varying intercepts

$$\cdot$$
 y ~ x + (1 | group)

Varying intercepts and slopes

$$\cdot$$
 y ~ x + (1 + x | group)

Varying intercepts

$$\cdot$$
 y ~ x + (1 | group)

· Varying intercepts and slopes

$$\cdot$$
 y ~ x + (1 + x | group)

· Varying intercepts, 2 groups (crossed)

Varying intercepts

$$\cdot$$
 y ~ x + (1 | group)

· Varying intercepts and slopes

$$\cdot$$
 y ~ x + (1 + x | group)

· Varying intercepts, 2 groups (crossed)

$$\cdot$$
 y ~ x + (1 | group1) + (1 | group2)

Varying intercepts

$$\cdot$$
 y ~ x + (1 | group)

· Varying intercepts and slopes

$$\cdot$$
 y ~ x + (1 + x | group)

· Varying intercepts, 2 groups (crossed)

$$\cdot$$
 y ~ x + (1 | group1) + (1 | group2)

Varying intercepts, 2 groups (nested)

Varying intercepts

$$\cdot$$
 y ~ x + (1 | group)

· Varying intercepts and slopes

$$\cdot$$
 y ~ x + (1 + x | group)

Varying intercepts, 2 groups (crossed)

$$\cdot$$
 y ~ x + (1 | group1) + (1 | group2)

· Varying intercepts, 2 groups (nested)

$$\cdot$$
 y ~ x + (1 | group/subgroup)

Varying intercepts

$$\cdot$$
 y ~ x + (1 | group)

· Varying intercepts and slopes

$$\cdot$$
 y ~ x + (1 + x | group)

Varying intercepts, 2 groups (crossed)

$$\cdot$$
 y ~ x + (1 | group1) + (1 | group2)

- · Varying intercepts, 2 groups (nested)
  - $\cdot$  y ~ x + (1 | group/subgroup)
  - This is equivalent to  $y \sim x + (1 \mid group1) + (1 \mid group2)$  with distinct labelling of group levels.

· Varying intercepts

$$\cdot$$
 y ~ x + (1 | group)

· Varying intercepts and slopes

$$\cdot$$
 y ~ x + (1 + x | group)

Varying intercepts, 2 groups (crossed)

$$\cdot$$
 y ~ x + (1 | group1) + (1 | group2)

- Varying intercepts, 2 groups (nested)
  - $\cdot$  y ~ x + (1 | group/subgroup)
  - This is equivalent to  $y \sim x + (1 \mid group1) + (1 \mid group2)$  with distinct labelling of group levels.
- · Varying intercepts and slopes, 2 groups (crossed)

· Varying intercepts

$$\cdot$$
 y ~ x + (1 | group)

· Varying intercepts and slopes

$$\cdot$$
 y ~ x + (1 + x | group)

· Varying intercepts, 2 groups (crossed)

$$\cdot$$
 y ~ x + (1 | group1) + (1 | group2)

- Varying intercepts, 2 groups (nested)
  - $\cdot$  y ~ x + (1 | group/subgroup)
  - This is equivalent to  $y \sim x + (1 \mid group1) + (1 \mid group2)$  with distinct labelling of group levels.
- · Varying intercepts and slopes, 2 groups (crossed)

$$\cdot$$
 y ~ x + (1 + x | group1) + (1 + x | group2)

#### **GLMM FAQ**

https://bbolker.github.io/mixedmodels-misc/glmmFAQ.html