

Mixed / Multilevel models

Example dataset: trees

- Data on 1000 trees from 10 sites.

```
head(trees)
```

	site	dbh	height	sex	dead
1	4	29.68	36.1	male	0
2	5	33.29	42.3	male	0
3	2	28.03	41.9	female	0
4	5	39.86	46.5	female	0
5	1	47.94	43.9	female	0
6	1	10.82	26.2	male	0

Example dataset: trees

- ▶ Data on 1000 trees from 10 sites.
- ▶ Trees per site: 4 - 392.

```
head(trees)
```

	site	dbh	height	sex	dead
1	4	29.68	36.1	male	0
2	5	33.29	42.3	male	0
3	2	28.03	41.9	female	0
4	5	39.86	46.5	female	0
5	1	47.94	43.9	female	0
6	1	10.82	26.2	male	0

Q: What's the relationship between tree diameter and height?

A simple linear model

```
lm.simple <- lm(height ~ dbh, data = trees)
```

Call:

```
lm(formula = height ~ dbh, data = trees)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-13.3270	-2.8978	0.1057	2.7924	12.9511

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	19.33920	0.31064	62.26	<2e-16 ***
dbh	0.61570	0.01013	60.79	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 4.093 on 998 degrees of freedom

Multiple R-squared: 0.7874, Adjusted R-squared: 0.7871

F-statistic: 3695 on 1 and 998 DF, p-value: < 2.2e-16

Remember our model structure

$$y_i \sim N(\mu_i, \sigma^2)$$

$$\mu_i = \alpha + \beta x_i$$

In this case:

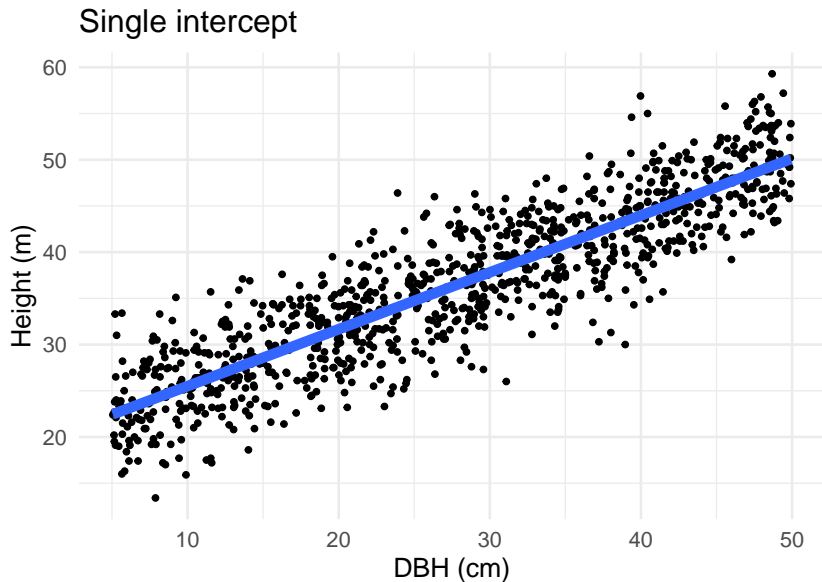
$$Height_i \sim N(\mu_i, \sigma^2)$$

$$\mu_i = \alpha + \beta DBH_i$$

α : expected height when $DBH = 0$

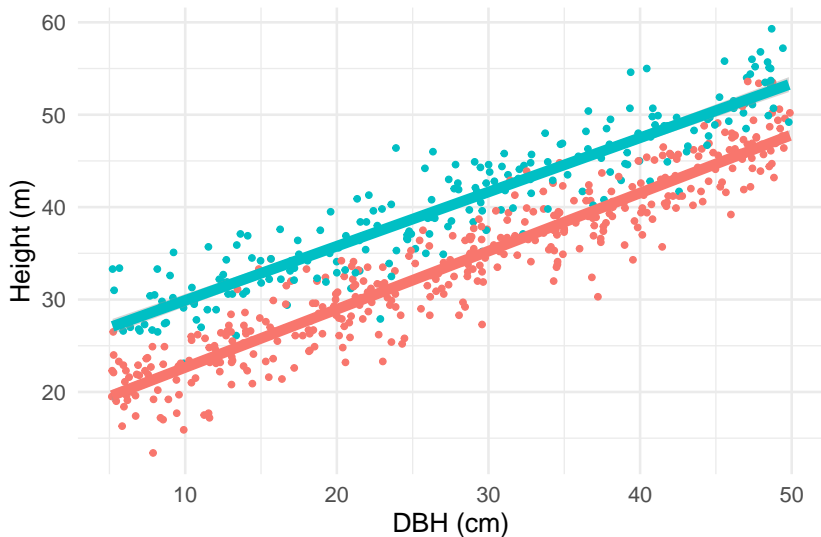
β : how much height increases with every unit increase of DBH

There is only one intercept



What if allometry varies among sites?

Different intercept for each site



Fitting a varying intercepts model with `lm`

Call:

```
lm(formula = height ~ factor(site) + dbh, data = trees)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-10.1130	-1.9885	0.0582	2.0314	11.3320

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	16.699037	0.260565	64.088	< 2e-16	***
factor(site)2	6.504303	0.256730	25.335	< 2e-16	***
factor(site)3	4.357457	0.354181	12.303	< 2e-16	***
factor(site)4	1.934650	0.356102	5.433	6.98e-08	***
factor(site)5	3.637432	0.339688	10.708	< 2e-16	***
factor(site)6	4.204511	0.421906	9.966	< 2e-16	***
factor(site)7	-0.176193	0.666772	-0.264	0.7916	
factor(site)8	-5.312648	0.893603	-5.945	3.82e-09	***
factor(site)9	5.437049	1.087766	4.998	6.84e-07	***
factor(site)10	2.263338	1.369986	1.652	0.0988	.
dbh	0.617075	0.007574	81.473	< 2e-16	***

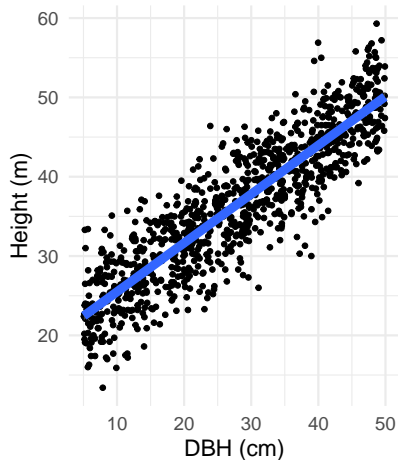
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.043 on 989 degrees of freedom

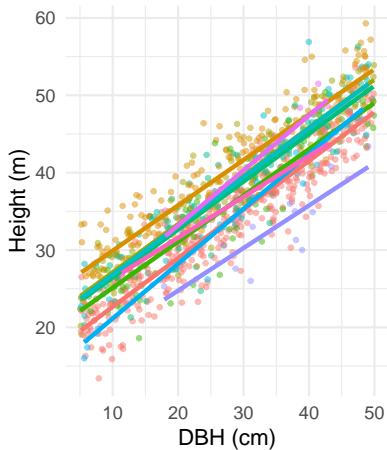
Multiple R-squared: 0.8835, Adjusted R-squared: 0.8823

Single vs varying intercept

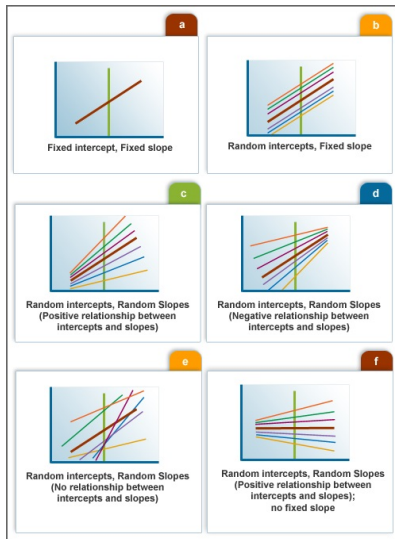
Single intercept



Different intercept for each si

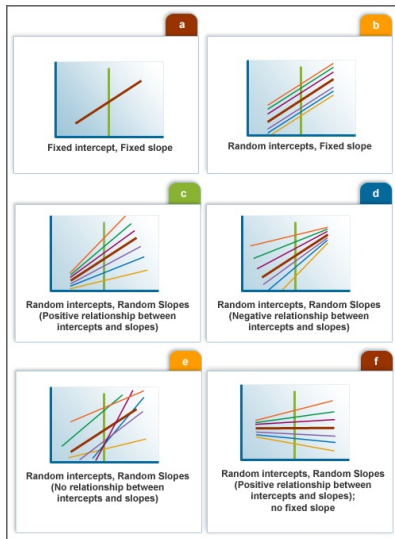


Mixed models enable us to account for variability



► Varying intercepts

Mixed models enable us to account for variability



- Varying intercepts
- Varying slopes

Mixed model with varying intercepts

$$y_i = a + \alpha_j + b \cdot x_i + \varepsilon_i$$

$$\alpha_j \sim N(0, \tau^2)$$

$$\varepsilon_i \sim N(0, \sigma^2)$$

En nuestro ejemplo:

$$Height_i = a + site_j + b \cdot DBH_i + \varepsilon_i$$

$$site_j \sim N(0, \tau^2)$$

$$\varepsilon_i \sim N(0, \sigma^2)$$

Mixed models estimate varying parameters (intercepts and/or slopes) with pooling among levels (rather than considering them fully independent)

Hence there's gradient between

- ▶ **complete pooling**: Single overall intercept.

Hence there's gradient between

- ▶ **complete pooling**: Single overall intercept.
 - ▶ `lm (height ~ dbh)`

Hence there's gradient between

- ▶ **complete pooling**: Single overall intercept.
 - ▶ `lm (height ~ dbh)`
- ▶ **no pooling**: One *independent* intercept for each site.

Hence there's gradient between

- ▶ **complete pooling:** Single overall intercept.
 - ▶ `lm (height ~ dbh)`
- ▶ **no pooling:** One *independent* intercept for each site.
 - ▶ `lm (height ~ dbh + site)`

Hence there's gradient between

- ▶ **complete pooling:** Single overall intercept.
 - ▶ `lm (height ~ dbh)`
- ▶ **no pooling:** One *independent* intercept for each site.
 - ▶ `lm (height ~ dbh + site)`
- ▶ **partial pooling:** Inter-related intercepts.

Hence there's gradient between

- ▶ **complete pooling:** Single overall intercept.
 - ▶ `lm (height ~ dbh)`
- ▶ **no pooling:** One *independent* intercept for each site.
 - ▶ `lm (height ~ dbh + site)`
- ▶ **partial pooling:** Inter-related intercepts.
 - ▶ `lmer(height ~ dbh + (1 | site))`

Random vs Fixed effects?

1. Fixed effects constant across individuals, random effects vary.

http://andrewgelman.com/2005/01/25/why_i_dont_use/

Random vs Fixed effects?

1. Fixed effects constant across individuals, random effects vary.
2. Effects are fixed if they are interesting in themselves; random if interest in the underlying population.

http://andrewgelman.com/2005/01/25/why_i_dont_use/

Random vs Fixed effects?

1. Fixed effects constant across individuals, random effects vary.
2. Effects are fixed if they are interesting in themselves; random if interest in the underlying population.
3. Fixed when sample exhausts the population; random when the sample is small part of the population.

http://andrewgelman.com/2005/01/25/why_i_dont_use/

Random vs Fixed effects?

1. Fixed effects constant across individuals, random effects vary.
2. Effects are fixed if they are interesting in themselves; random if interest in the underlying population.
3. Fixed when sample exhausts the population; random when the sample is small part of the population.
4. Random effect if it's assumed to be a realized value of random variable.

http://andrewgelman.com/2005/01/25/why_i_dont_use/

Random vs Fixed effects?

1. Fixed effects constant across individuals, random effects vary.
2. Effects are fixed if they are interesting in themselves; random if interest in the underlying population.
3. Fixed when sample exhausts the population; random when the sample is small part of the population.
4. Random effect if it's assumed to be a realized value of random variable.
5. Fixed effects estimated using least squares or maximum likelihood; random effects estimated with shrinkage.

http://andrewgelman.com/2005/01/25/why_i_dont_use/

What is a random effect, really?

- ▶ Varies by group

Random effects are estimated with *partial pooling*, while fixed effects are not (infinite variance).

What is a random effect, really?

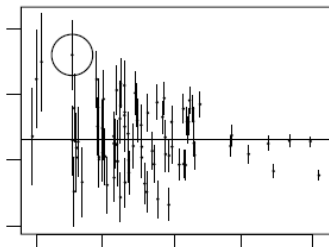
- ▶ Varies by group
- ▶ Variation estimated with **probability model**

Random effects are estimated with *partial pooling*, while fixed effects are not (infinite variance).

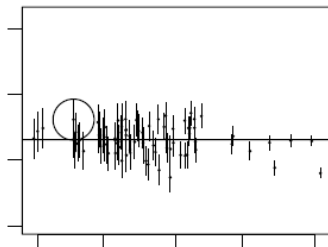
Shrinkage improves parameter estimation

Especially for groups with low sample size

No pooling



Multilevel model



From Gelman & Hill p. 253

Fitting mixed/multilevel models

```
library(lme4)
mixed <- lmer(height ~ dbh + (1|site), data = trees)
```

Linear mixed model fit by REML ['lmerMod']

Formula: height ~ dbh + (1 | site)

Data: trees

REML criterion at convergence: 5108.3

Scaled residuals:

Min	1Q	Median	3Q	Max
-3.3199	-0.6607	0.0227	0.6716	3.7328

Random effects:

Groups	Name	Variance	Std.Dev.
site	(Intercept)	11.195	3.346
Residual		9.261	3.043

Number of obs: 1000, groups: site, 10

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	19.011468	1.100444	17.28
dbh	0.616927	0.007572	81.47

Correlation of Fixed Effects:

(Intr)

dbh -0.167

Retrieve model coefficients

```
coef(mixed)
```

```
$site
```

```
  (Intercept)      dbh
```

1	16.70800	0.6169271
2	23.19162	0.6169271
3	21.04229	0.6169271
4	18.64086	0.6169271
5	20.32995	0.6169271
6	20.88200	0.6169271
7	16.61686	0.6169271
8	11.88302	0.6169271
9	21.84779	0.6169271
10	18.97228	0.6169271

```
attr(,"class")
```

```
[1] "coef.mer"
```

Broom: model estimates in tidy form

```
library(broom.mixed)
tidy(mixed)
```

```
# A tibble: 4 x 6
```

	effect	group	term	estimate	std.error	statistic
	<chr>	<chr>	<chr>	<dbl>	<dbl>	<dbl>
1	fixed	<NA>	(Intercept)	19.0	1.10	17.3
2	fixed	<NA>	dbh	0.617	0.00757	81.5
3	ran_pars	site	sd__(Intercept)	3.35	NA	NA
4	ran_pars	Residual	sd__Observation	3.04	NA	NA

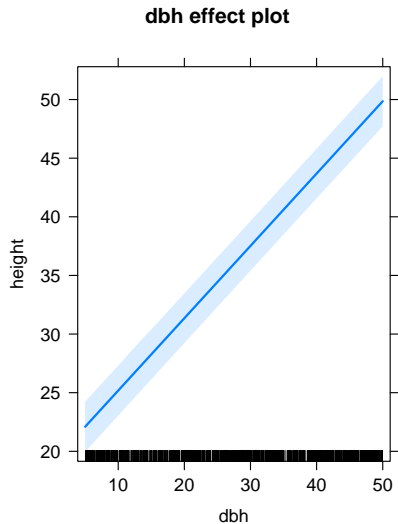
See also [broom.mixed](#)

Visualising model: allEffects

```
model: height ~ dbh
```

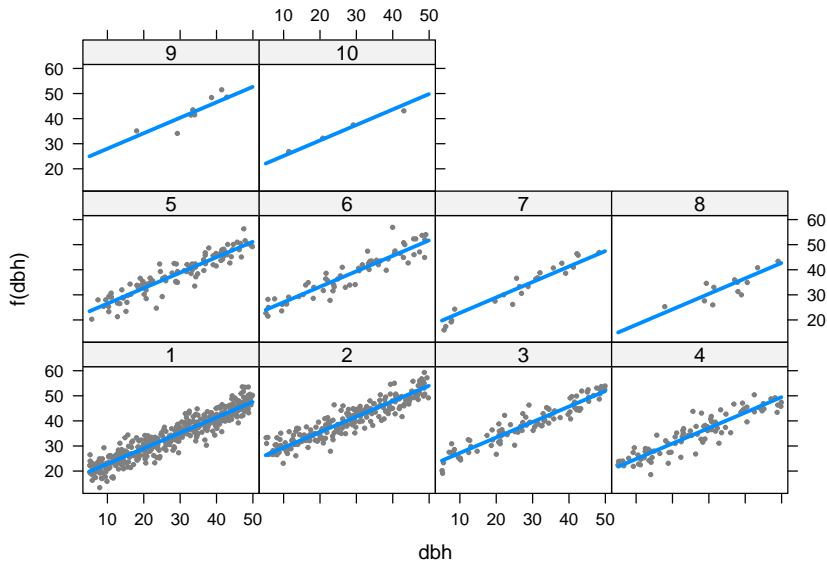
```
dbh effect
```

dbh	5	20	30	40	50
	22.09610	31.35001	37.51928	43.68855	49.85782



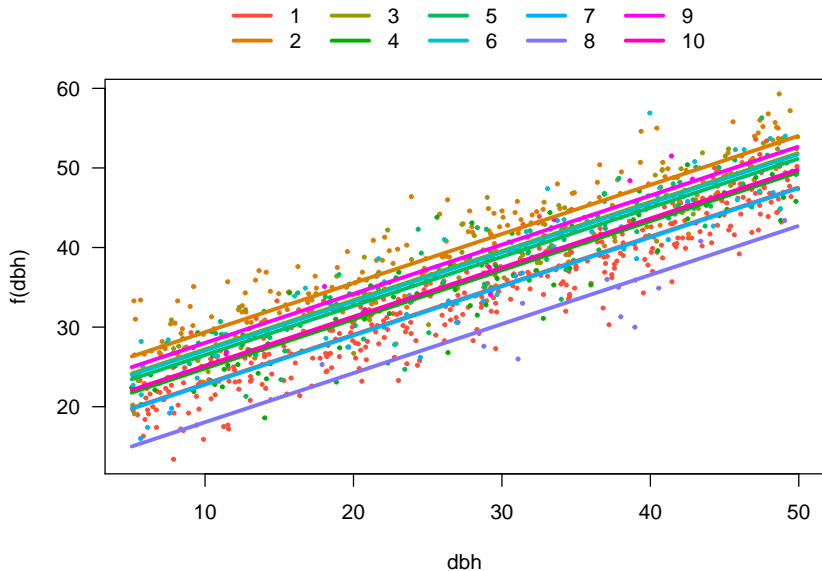
Visualising model: visreg

```
visreg(mixed, xvar = "dbh", by = "site", re.form = NULL)
```



Visualising model

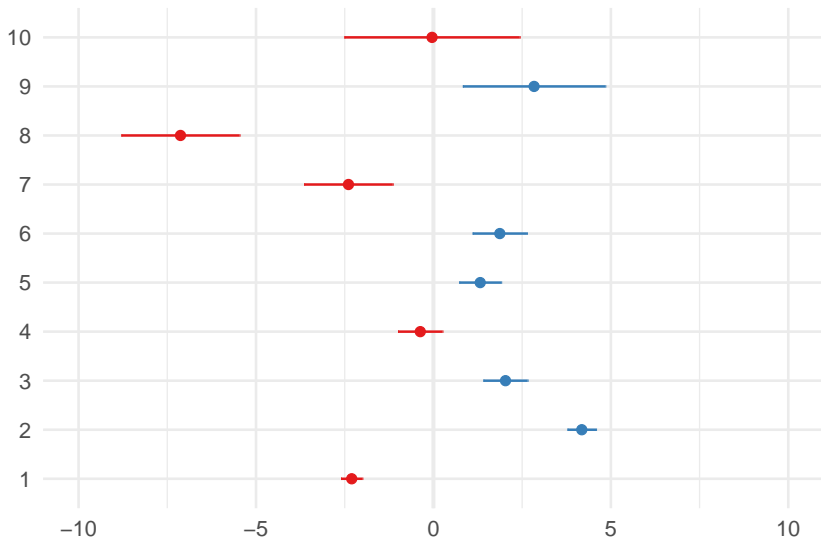
```
visreg(mixed, xvar = "dbh", by = "site", re.form = NULL, overlay
```



Visualising model: sjPlot

```
sjPlot::plot_model(mixed, type = "re")
```

Random effects

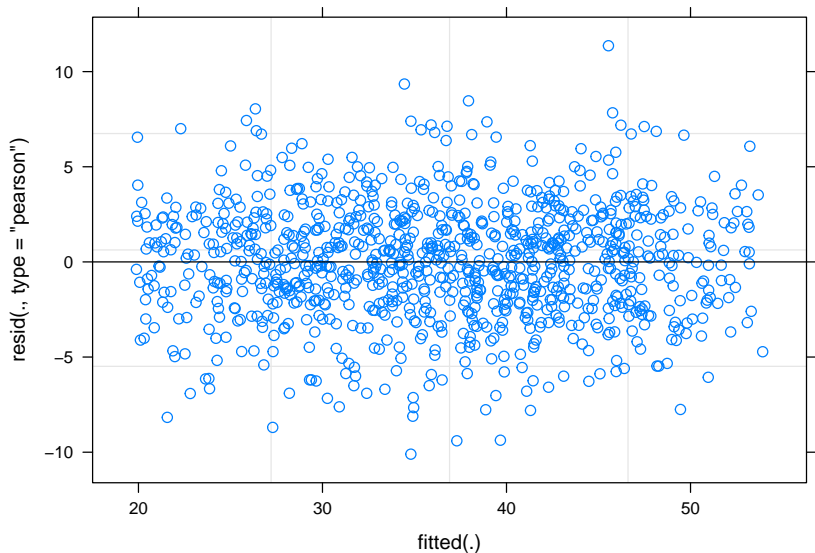


Using merTools to understand fitted model

```
library(merTools)
shinyMer(mixed)
```

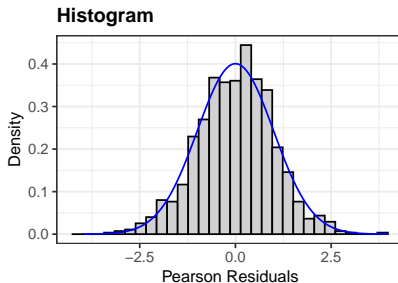
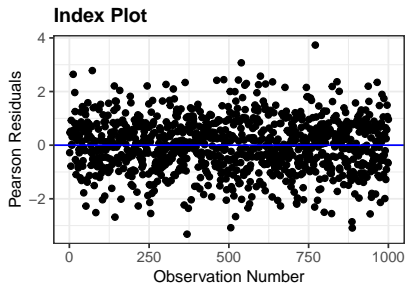
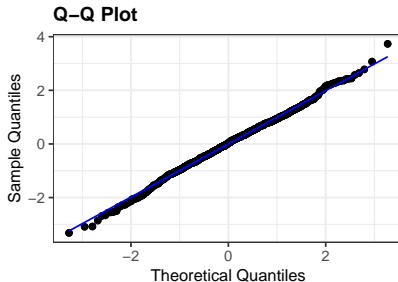
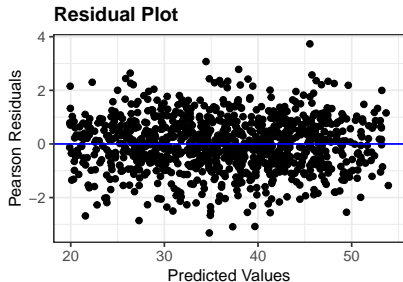
Checking residuals

```
plot(mixed)
```



Checking residuals

```
ggResidpanel::resid_panel(mixed)
```

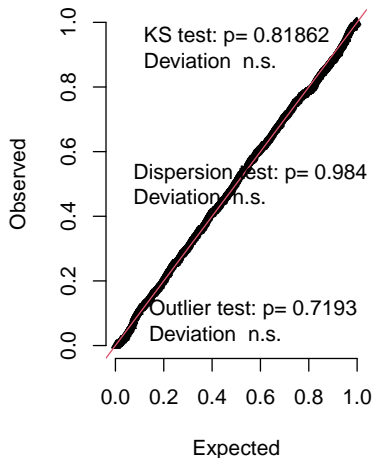


Checking residuals (DHARMA)

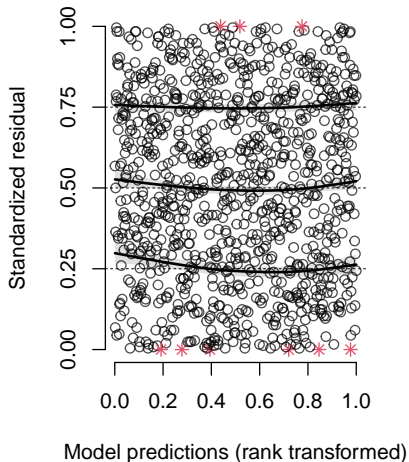
```
DHARMA::simulateResiduals(mixed, plot = TRUE, use.u = TRUE)
```

DHARMA residual diagnostics

QQ plot residuals

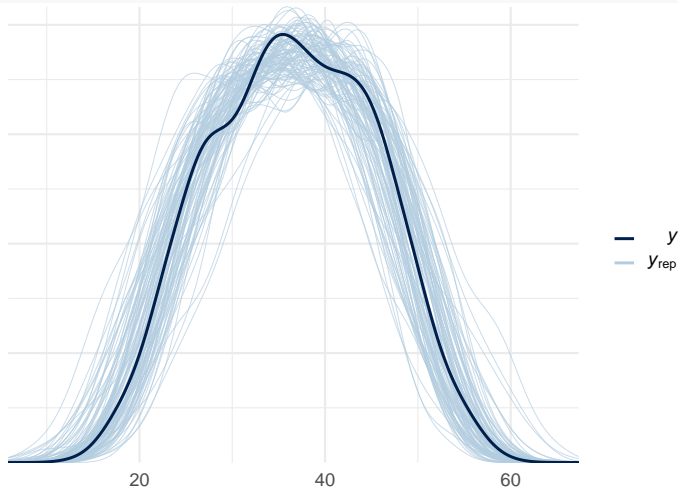


Residual vs. predicted
No significant problems detected



Model checking with simulated data

```
library(bayesplot)
sims <- simulate(mixed, nsim = 100)
ppc_dens_overlay(trees$height, yrep = t(as.matrix(sims)))
```



R-squared for GLMMs

Many approaches! Somewhat polemic (e.g. see [this](#)).
Nakagawa & Schielzeth propose **marginal** (considering fixed effects only)
and **conditional** R^2 (including random effects too):

```
library(MuMIn)  
r.squaredGLMM(mixed)
```

	R2m	R2c
[1,]	0.752535	0.8879656

Growing the hierarchy: adding site-level
predictors

Model with group-level predictors

We had:

$$y_i = a + \alpha_j + b \cdot x_i + \varepsilon_i$$

$$\alpha_j \sim N(0, \tau^2)$$

$$\varepsilon_i \sim N(0, \sigma^2)$$

Now

$$y_i = a + \alpha_j + b \cdot x_i + \varepsilon_i$$

$$\alpha_j \sim N(\mu_j, \tau^2)$$

$$\mu_j = \delta \cdot \text{Predictor}_j$$

$$\varepsilon_i \sim N(0, \sigma^2)$$

Are height differences among sites related to temperature?

$$Height_i = site_j + b \cdot DBH_i + \varepsilon_i$$

$$site_j \sim N(\mu_j, \tau^2)$$

$$\mu_j = a + \delta \cdot Temperature_j$$

$$\varepsilon_i \sim N(0, \sigma^2)$$

Are height differences among sites related to temperature?

```
sitedata <- read.csv("data/sitedata.csv")
```

```
sitedata
```

	site	temp
1	1	15.1
2	2	22.0
3	3	20.1
4	4	20.4
5	5	20.0
6	6	20.1
7	7	17.5
8	8	14.6
9	9	19.2
10	10	16.0

Merging trees and site data

```
trees.full <- merge(trees, sitedata, by = "site")  
head(trees.full)
```

	site	dbh	height	sex	dead	temp
1	1	21.05	32.2	male	0	15.1
2	1	46.63	45.9	female	0	15.1
3	1	43.86	45.5	male	0	15.1
4	1	29.03	35.5	male	0	15.1
5	1	6.02	21.1	male	0	15.1
6	1	40.82	38.7	male	0	15.1

Fit multilevel model

```
group.pred <- lmer(height ~ dbh + (1 | site) + temp, data = trees.full)
```

Linear mixed model fit by REML ['lmerMod']

Formula: height ~ dbh + (1 | site) + temp

Data: trees.full

REML criterion at convergence: 5098.2

Scaled residuals:

	Min	1Q	Median	3Q	Max
	-3.3247	-0.6517	0.0192	0.6663	3.7268

Random effects:

Groups	Name	Variance	Std.Dev.
site	(Intercept)	3.158	1.777
Residual		9.266	3.044

Number of obs: 1000, groups: site, 10

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	-1.730910	4.671330	-0.371
dbh	0.616894	0.007571	81.484
temp	1.115104	0.248000	4.496

Correlation of Fixed Effects:

	(Intr) dbh
dbh	-0.055
temp	-0.991 0.008

Too strong correlation of parameters!

Centre (and scale) continuous variables

```
mean(sitedata$temp)
```

```
[1] 18.5
```

```
trees.full$temp.c <- trees.full$temp - 18
```

Temperatures now referred as deviations from 18 °C (close to average)

Fit multilevel model

```
group.pred <- lmer(height ~ dbh + (1 | site) + temp.c, data = trees.full)
```

```
Linear mixed model fit by REML ['lmerMod']  
Formula: height ~ dbh + (1 | site) + temp.c  
Data: trees.full
```

REML criterion at convergence: 5098.2

Scaled residuals:

	Min	1Q	Median	3Q	Max
	-3.3247	-0.6517	0.0192	0.6663	3.7268

Random effects:

Groups	Name	Variance	Std.Dev.
site	(Intercept)	3.158	1.777
Residual		9.266	3.044

Number of obs: 1000, groups: site, 10

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	18.340954	0.655054	27.999
dbh	0.616894	0.007571	81.484
temp.c	1.115104	0.248000	4.496

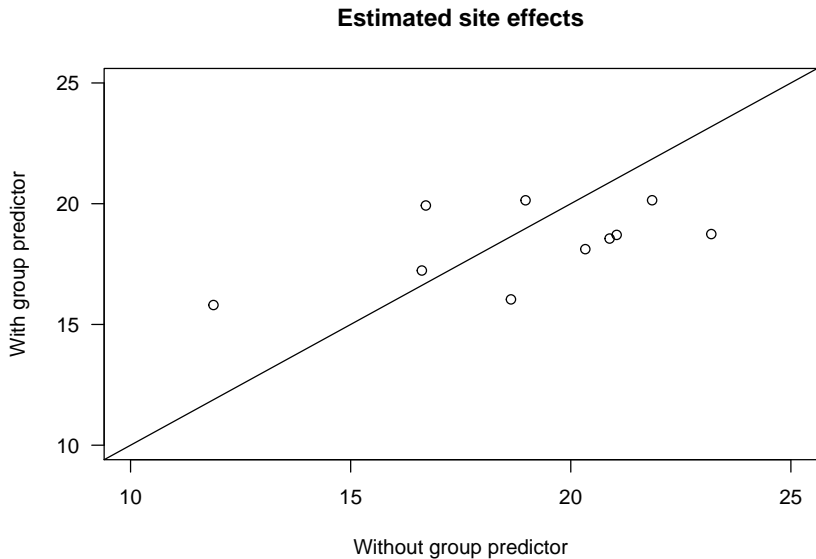
Correlation of Fixed Effects:

	(Intr) dbh
dbh	-0.333
temp.c	-0.250 0.008

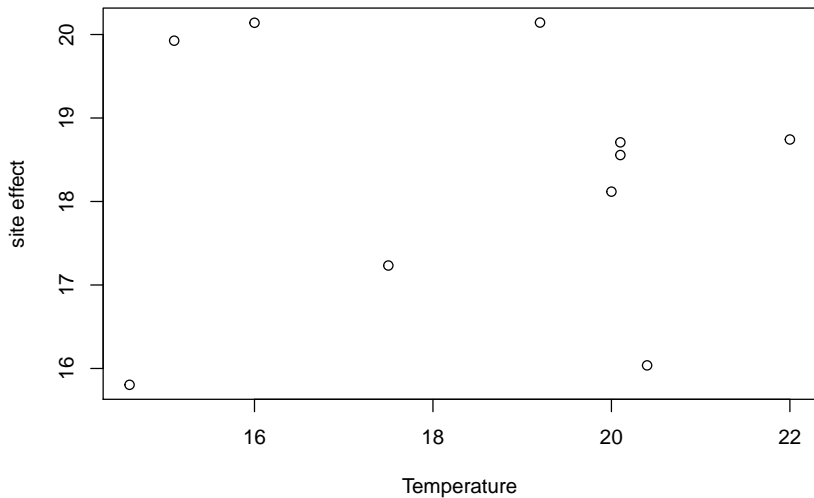
Examine model with merTools

```
shinyMer(group.pred)
```

Comparing site effects with and without group predictor



Are site effects related to temperature?



Varying intercepts and slopes

Varying intercepts and slopes

There is overall difference in height among sites (different intercepts)

AND

Relationship between DBH and Height varies among sites (different slopes)

```
mixed.slopes <- lmer(height ~ dbh + (1 + dbh | site), data=trees)
```

Varying intercepts and slopes

Linear mixed model fit by REML ['lmerMod']

Formula: height ~ dbh + (1 + dbh | site)

Data: trees

REML criterion at convergence: 5105.1

Scaled residuals:

	Min	1Q	Median	3Q	Max
	-3.3342	-0.6599	0.0375	0.6916	3.7756

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
site	(Intercept)	1.566e+01	3.95671	
	dbh	3.087e-04	0.01757	-1.00
Residual		9.226e+00	3.03744	

Number of obs: 1000, groups: site, 10

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	18.95272	1.29190	14.67
dbh	0.61837	0.00946	65.37

Correlation of Fixed Effects:

(Intr)

dbh -0.722

optimizer (nloptwrap) convergence code: 0 (OK)

Varying intercepts and slopes

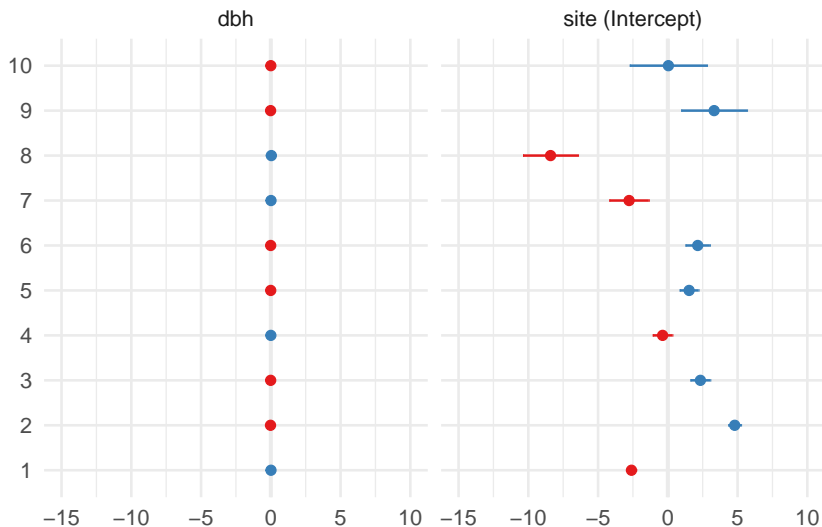
```
$site
  (Intercept)      dbh
1    16.34655 0.6299443
2    23.74733 0.5970814
3    21.28802 0.6080019
4    18.57844 0.6200337
5    20.47961 0.6115916
6    21.09608 0.6088542
7    16.17675 0.6306983
8    10.54681 0.6556978
9    22.27301 0.6036281
10   18.99463 0.6181856
```

```
attr(,"class")
[1] "coef.mer"
```


Visualising model: sjPlot

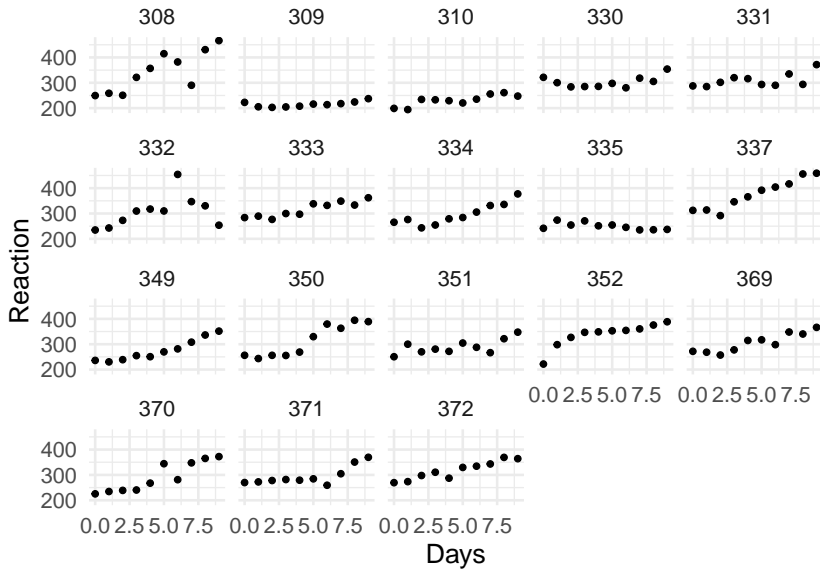
```
plot_model(mixed.slopes, type = "re")
```

Random effects



More examples

sleepstudy (repeated measures)



Varying intercepts and slopes (lme4)

```
sleep <- lmer(Reaction ~ Days + (1+Days|Subject), data = sleepstudy)
```

Linear mixed model fit by REML ['lmerMod']

Formula: Reaction ~ Days + (1 + Days | Subject)

Data: sleepstudy

REML criterion at convergence: 1743.6

Scaled residuals:

	Min	1Q	Median	3Q	Max
	-3.9536	-0.4634	0.0231	0.4634	5.1793

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
Subject	(Intercept)	612.10	24.741	
	Days	35.07	5.922	0.07
Residual		654.94	25.592	

Number of obs: 180, groups: Subject, 18

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	251.405	6.825	36.838
Days	10.467	1.546	6.771

Correlation of Fixed Effects:

(Intr)

Days -0.138

Varying intercepts and slopes (lme4)

```
visreg(sleep, xvar = "Days", by = "Subject", re.form = NULL)
```

Fitting multilevel models (GAMM) with mgcv

```
sgamm <- mgcv::gam(Reaction ~ s(Days, Subject, k = 3, bs = "fs"),  
                   data = sleepstudy, method = "REML")
```

Family: gaussian

Link function: identity

Formula:

Reaction ~ s(Days, Subject, k = 3, bs = "fs")

Parametric coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	295.22	10.49	28.15	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Approximate significance of smooth terms:

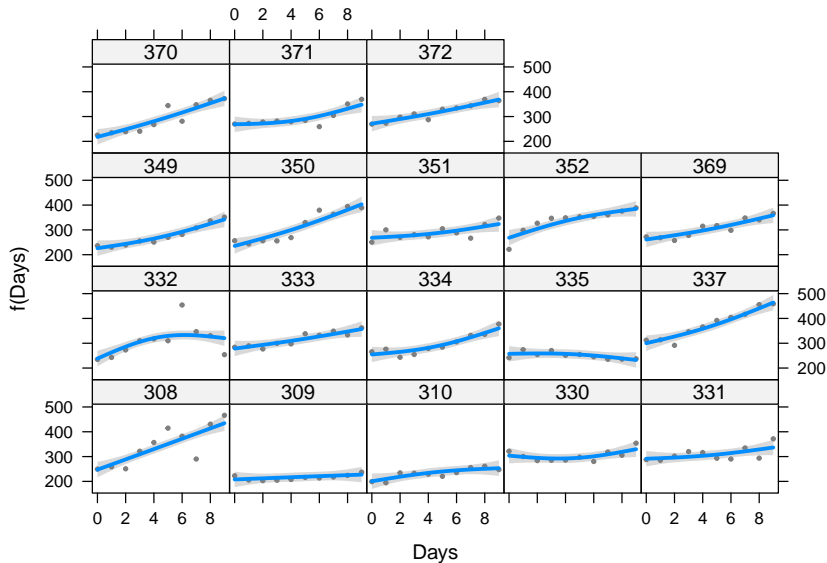
	edf	Ref.df	F	p-value
s(Days,Subject)	42.2	53	16.05	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

R-sq.(adj) = 0.826 Deviance explained = 86.7%

Fitting multilevel models (GAMM) with mgcv

```
visreg(sgammm, xvar = "Days", by = "Subject")
```



Fitting multilevel models (GAMM) with mgcv

Hierarchical generalized additive models: an introduction with mgcv

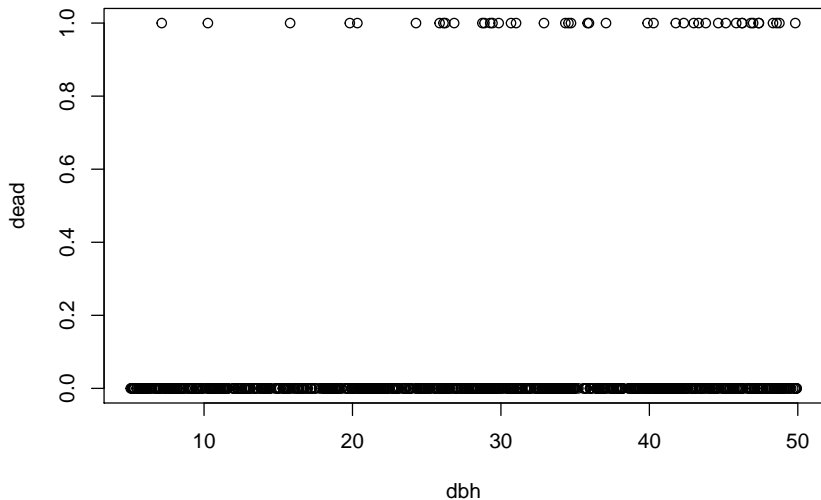
Eric J Pedersen Corresp., 1,2, David L. Miller 3,4, Gavin L. Simpson 5, Noam Ross 6

<https://doi.org/10.7287/peerj.preprints.27320v1>

Multilevel logistic regression

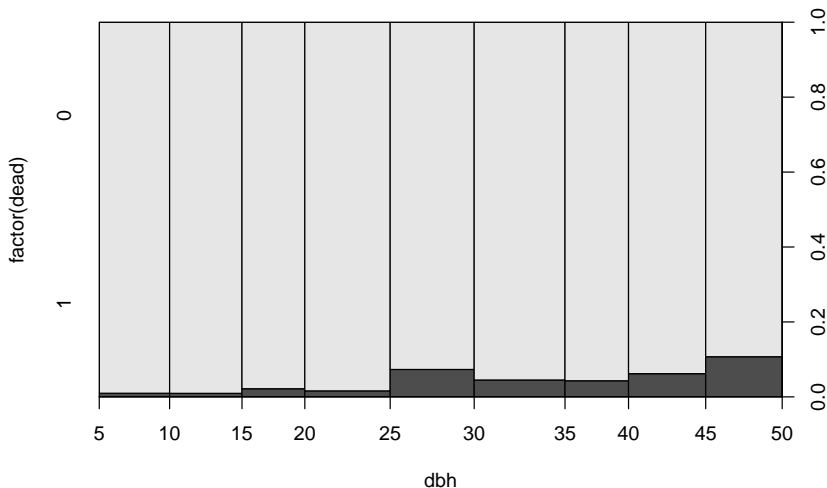
Q: Relationship between tree size and mortality

```
plot(dead ~ dbh, data = trees)
```



Q: Relationship between tree size and mortality

```
plot(factor(dead) ~ dbh, data = trees)
```



Fit simple logistic regression

```
simple.logis <- glm(dead ~ dbh, data = trees, family=binomial)
```

Call:

```
glm(formula = dead ~ dbh, family = binomial, data = trees)
```

Deviance Residuals:

Min	1Q	Median	3Q	Max
-0.4805	-0.3520	-0.2647	-0.1928	2.9690

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-4.77874	0.50902	-9.388	< 2e-16 ***
dbh	0.05365	0.01377	3.895	9.82e-05 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for binomial family taken to be 1)

Null deviance:	360.91	on 999	degrees of freedom
Residual deviance:	343.69	on 998	degrees of freedom

Logistic regression with *independent* site effects

```
logis2 <- glm(dead ~ dbh + factor(site), data = trees, family=binomial)
```

Call:

```
glm(formula = dead ~ dbh + factor(site), family = binomial, data = trees)
```

Deviance Residuals:

Min	1Q	Median	3Q	Max
-0.6359	-0.3449	-0.2561	-0.1852	2.9763

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-4.80123	0.54985	-8.732	<2e-16 ***
dbh	0.05371	0.01381	3.889	0.0001 ***
factor(site)2	-0.29692	0.46073	-0.644	0.5193
factor(site)3	0.21275	0.52799	0.403	0.6870
factor(site)4	0.39841	0.53025	0.751	0.4524
factor(site)5	-0.42557	0.64018	-0.665	0.5062
factor(site)6	0.66861	0.53656	1.246	0.2127
factor(site)7	0.11862	1.06211	0.112	0.9111
factor(site)8	0.43899	1.08058	0.406	0.6846
factor(site)9	-13.63389	840.90382	-0.016	0.9871
factor(site)10	-13.17148	1042.21823	-0.013	0.9899

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

(Diagnostics for the model: binomial family, method = GLS)

Fit multilevel logistic regression

```
mixed.logis <- glmer(dead ~ dbh + (1|site), data=trees, family = binomial)
```

Generalized linear mixed model fit by maximum likelihood (Laplace

Approximation) [glmerMod]

Family: binomial (logit)

Formula: dead ~ dbh + (1 | site)

Data: trees

AIC	BIC	logLik	deviance	df.resid
349.7	364.4	-171.8	343.7	997

Scaled residuals:

Min	1Q	Median	3Q	Max
-0.3498	-0.2528	-0.1888	-0.1370	9.0031

Random effects:

Groups	Name	Variance	Std.Dev.
site	(Intercept)	0	0

Number of obs: 1000, groups: site, 10

Fixed effects:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-4.77874	0.50904	-9.388	< 2e-16 ***
dbh	0.05365	0.01377	3.895	9.83e-05 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Retrieve model coefficients

```
coef(mixed.logis)
```

```
$site
```

```
(Intercept)
```

```
dbh
```

1	-4.778744	0.05364989
2	-4.778744	0.05364989
3	-4.778744	0.05364989
4	-4.778744	0.05364989
5	-4.778744	0.05364989
6	-4.778744	0.05364989
7	-4.778744	0.05364989
8	-4.778744	0.05364989
9	-4.778744	0.05364989
10	-4.778744	0.05364989

```
attr(,"class")
```

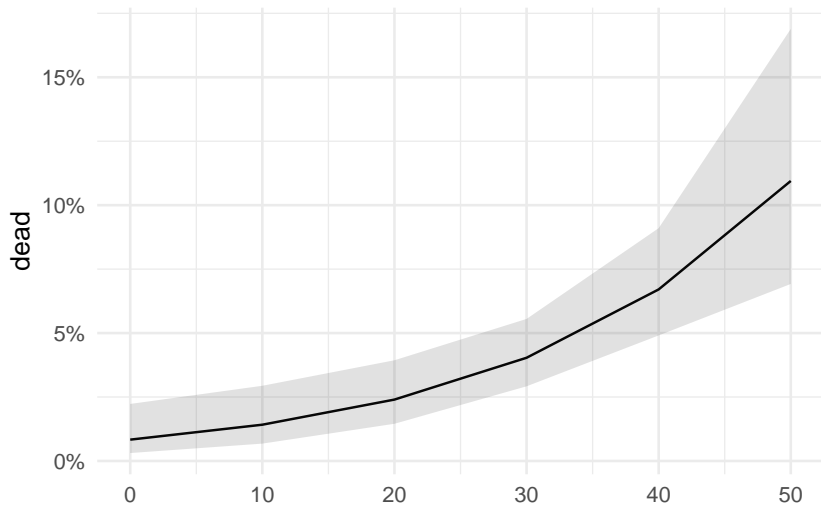
```
[1] "coef.mer"
```

Visualising model: sjPlot

```
plot_model(mixed.logis, type = "eff", show.ci = TRUE)
```

\$dbh

Predicted probabilities of dead



Poisson multilevel regression

Advantages of multilevel models

- ▶ Perfect for **structured data** (space-time)

Advantages of multilevel models

- ▶ Perfect for **structured data** (space-time)
- ▶ Predictors enter at the appropriate level

Advantages of multilevel models

- ▶ Perfect for **structured data** (space-time)
- ▶ Predictors enter at the appropriate level
- ▶ Accommodate **variation** in treatment effects

Advantages of multilevel models

- ▶ Perfect for **structured data** (space-time)
- ▶ Predictors enter at the appropriate level
- ▶ Accommodate **variation** in treatment effects
- ▶ More **efficient inference** of regression parameters

Advantages of multilevel models

- ▶ Perfect for **structured data** (space-time)
- ▶ Predictors enter at the appropriate level
- ▶ Accommodate **variation** in treatment effects
- ▶ More **efficient inference** of regression parameters
- ▶ Using all the data to perform inferences for groups with small sample size

Formula syntax for different models

- ▶ Varying intercepts

Formula syntax for different models

- ▶ Varying intercepts
 - ▶ $y \sim x + (1 \mid \text{group})$

Formula syntax for different models

- ▶ Varying intercepts
 - ▶ $y \sim x + (1 \mid \text{group})$
- ▶ Varying intercepts and slopes

Formula syntax for different models

- ▶ Varying intercepts
 - ▶ $y \sim x + (1 \mid \text{group})$
- ▶ Varying intercepts and slopes
 - ▶ $y \sim x + (1 + x \mid \text{group})$

Formula syntax for different models

- ▶ Varying intercepts
 - ▶ $y \sim x + (1 \mid \text{group})$
- ▶ Varying intercepts and slopes
 - ▶ $y \sim x + (1 + x \mid \text{group})$
- ▶ Varying intercepts, 2 groups (crossed)

Formula syntax for different models

- ▶ Varying intercepts
 - ▶ $y \sim x + (1 \mid \text{group})$
- ▶ Varying intercepts and slopes
 - ▶ $y \sim x + (1 + x \mid \text{group})$
- ▶ Varying intercepts, 2 groups (crossed)
 - ▶ $y \sim x + (1 \mid \text{group1}) + (1 \mid \text{group2})$

Formula syntax for different models

- ▶ Varying intercepts
 - ▶ $y \sim x + (1 \mid \text{group})$
- ▶ Varying intercepts and slopes
 - ▶ $y \sim x + (1 + x \mid \text{group})$
- ▶ Varying intercepts, 2 groups (crossed)
 - ▶ $y \sim x + (1 \mid \text{group1}) + (1 \mid \text{group2})$
- ▶ Varying intercepts, 2 groups (nested)

Formula syntax for different models

- ▶ Varying intercepts
 - ▶ $y \sim x + (1 \mid \text{group})$
- ▶ Varying intercepts and slopes
 - ▶ $y \sim x + (1 + x \mid \text{group})$
- ▶ Varying intercepts, 2 groups (crossed)
 - ▶ $y \sim x + (1 \mid \text{group1}) + (1 \mid \text{group2})$
- ▶ Varying intercepts, 2 groups (nested)
 - ▶ $y \sim x + (1 \mid \text{group/subgroup})$

Formula syntax for different models

- ▶ Varying intercepts
 - ▶ $y \sim x + (1 \mid \text{group})$
- ▶ Varying intercepts and slopes
 - ▶ $y \sim x + (1 + x \mid \text{group})$
- ▶ Varying intercepts, 2 groups (crossed)
 - ▶ $y \sim x + (1 \mid \text{group1}) + (1 \mid \text{group2})$
- ▶ Varying intercepts, 2 groups (nested)
 - ▶ $y \sim x + (1 \mid \text{group/subgroup})$
 - ▶ This is **equivalent** to $y \sim x + (1 \mid \text{group1}) + (1 \mid \text{group2})$ with distinct labelling of group levels.

Formula syntax for different models

- ▶ Varying intercepts
 - ▶ $y \sim x + (1 \mid \text{group})$
- ▶ Varying intercepts and slopes
 - ▶ $y \sim x + (1 + x \mid \text{group})$
- ▶ Varying intercepts, 2 groups (crossed)
 - ▶ $y \sim x + (1 \mid \text{group1}) + (1 \mid \text{group2})$
- ▶ Varying intercepts, 2 groups (nested)
 - ▶ $y \sim x + (1 \mid \text{group/subgroup})$
 - ▶ This is **equivalent** to $y \sim x + (1 \mid \text{group1}) + (1 \mid \text{group2})$ with distinct labelling of group levels.
- ▶ Varying intercepts and slopes, 2 groups (crossed)

Formula syntax for different models

- ▶ Varying intercepts
 - ▶ $y \sim x + (1 \mid \text{group})$
- ▶ Varying intercepts and slopes
 - ▶ $y \sim x + (1 + x \mid \text{group})$
- ▶ Varying intercepts, 2 groups (crossed)
 - ▶ $y \sim x + (1 \mid \text{group1}) + (1 \mid \text{group2})$
- ▶ Varying intercepts, 2 groups (nested)
 - ▶ $y \sim x + (1 \mid \text{group/subgroup})$
 - ▶ This is **equivalent** to $y \sim x + (1 \mid \text{group1}) + (1 \mid \text{group2})$ with distinct labelling of group levels.
- ▶ Varying intercepts and slopes, 2 groups (crossed)
 - ▶ $y \sim x + (1 + x \mid \text{group1}) + (1 + x \mid \text{group2})$

GLMM FAQ

<https://bbolker.github.io/mixedmodels-misc/glmmFAQ.html>