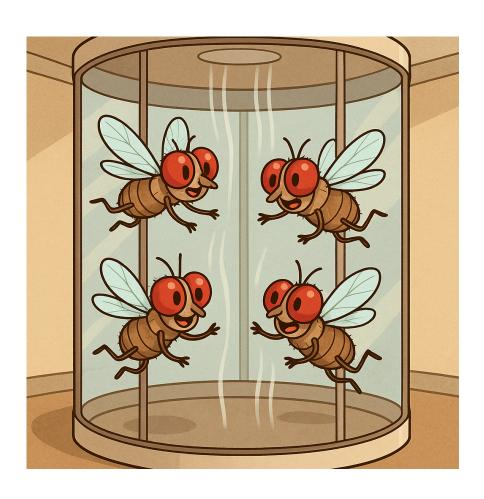
Statistical Analysis of Flight Assay Data in R Jun Ishigohoka



Environment

We recommend you use Rstudio.

- Download flight_assay_analysis.Rmd from https://github.com/PallaresLab/EMB L_Drosophila_course_2025/blob/main/scripts/flight_assay_analysis.Rmd, place it in the same directory as your flight assay data.
- Open flight_assay_analysis.Rmd in Rstudio.
- Follow the instructions

If you prefer other IDEs or base R console, you can copy-paste commands from the code blocks in flight_assay_analysis.pdf from https://github.com/PallaresLab/EMB L_Drosophila_course_2025/blob/main/protocols/flight_assay_analysis.pdf or flight_assay_analysis.html from https://github.com/PallaresLab/EMBL_Drosophila_course_2025/blob/main/protocols/flight_assay_analysis.html.

Preparation

If you don't have the packages, install them.

```
packages <- c("ggplot2", "car")
for (pkg in packages) {
   if (!requireNamespace(pkg, quietly = TRUE)) {
     install.packages(pkg)
   }
}</pre>
```

Load libraries.

```
library(ggplot2)
library(car) # Anova() and leveneTest()
```

Read data in two data frames. For this protocol, test data will be used.

```
d_1 <- read.csv("data/test_data/R4_female.csv")
d_2 <- read.csv("data/test_data/R4_male.csv")
head(d_1)</pre>
```

```
X.1
                      X
                             Y
##
           Area
## 1
          5.456 407.768 9.192
## 2
       2 6.666 343.943 10.074
       3 11.087 306.180 11.690
## 3
## 4
       4 4.932 376.921 16.866
       5 4.569 148.680 25.018
## 5
## 6
       6 5.389 330.855 26.006
```

Add a column for the factor you are comparing.

```
d_1$sex <- "female"
d_2$sex <- "male"
```

Concatenate the two data frames. Then set it as a factor.

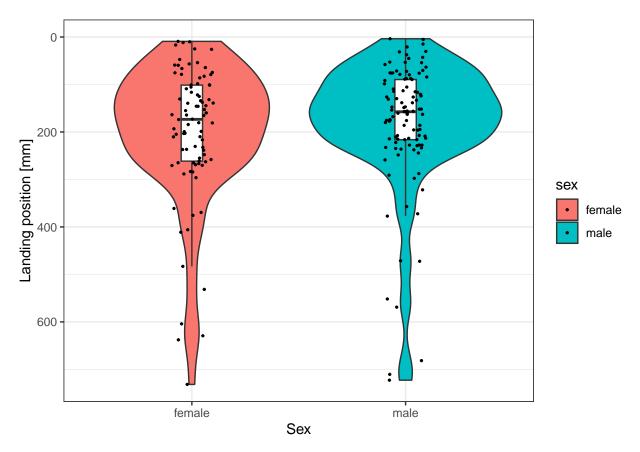
```
d <- rbind(d_1, d_2)
d$sex <- factor(d$sex, levels = c("female", "male"))</pre>
```

Visualisation

Let's visualise the distribution of landing height. Once you have visualised the distributions, try to describe them.

If you have ggplot2...

```
ggplot(data = d, # data to be plotted
          mapping = aes(x=sex, y=Y, fill = sex)) + # columns in the data frame to be used
           geom_violin() +
                              # violin plot
3
           geom_boxplot(
                              # box plot
4
                              # width of the box plot
             width = 0.1,
5
             fill = "white", # colour inside the box plot
                              # whether to have points for outliers.
             outliers = F
                              # FALSE because of geom_fitter() below
             ) +
9
                              # each data point with "jitter" along x axis
           geom_jitter(
10
                              # size of points
             size = 0.5,
11
             width = 0.1
                              # the amount of jitter
12
             ) +
13
           scale_y_reverse( # to make the plot same direction as the tapes
14
             limits = c(NA, 0) # so that the lowest Y is 0
15
16
                              # Add labels
     labs(
17
       x = "Sex",
18
       y = "Landing position [mm]") +
19
     theme_bw()
20
```



If you prefer the base plot or do not have ggplot2 installed.

```
boxplot(d$Y ~ as.factor(d$sex),

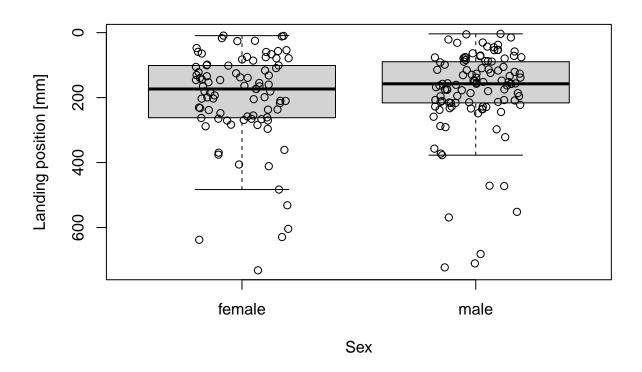
pch = NA,

xlab = "Sex",

ylab = "Landing position [mm]",

ylim = rev(range(d$Y))

points(jitter(as.numeric(as.factor(d$sex))), d$Y)
```



Statistical analysis (hypothesis testing)

Are your data normally distributed?

Many parametric tests assume normality. Let's run Shapiro-Wilk test for each group to test normality.

```
shapiro.test(subset(d, sex == "female")$Y)
##
##
    Shapiro-Wilk normality test
##
          subset(d, sex == "female")$Y
## data:
## W = 0.87191, p-value = 3.497e-07
shapiro.test(subset(d, sex == "male")$Y)
##
##
    Shapiro-Wilk normality test
##
          subset(d, sex == "male")$Y
## data:
## W = 0.80816, p-value = 1.326e-10
```

Does the result fit your description of the distributions?

If they are not normally distributed, think why. Do you expect normally distributed data from this experiment? (Hints: Poisson process; waiting time)

Does your data have equal variance?

Many parametric tests assume equal variance between groups.

Let's run two tests for homogeneity: F-test and Brown-Forsythe test.

Run F-test.

```
var.test(Y ~ sex, d)
##
## F test to compare two variances
##
## data: Y by sex
## F = 1.1212, num df = 87, denom df = 108, p-value = 0.5699
## alternative hypothesis: true ratio of variances is not equal to 1
## 95 percent confidence interval:
## 0.7535657 1.6828055
## sample estimates:
## ratio of variances
##
              1.121162
Run Brown-Forsythe test (car needs to be installed).
leveneTest(Y ~ sex, d, center = median)
## Levene's Test for Homogeneity of Variance (center = median)
          Df F value Pr(>F)
           1 0.8082 0.3698
## group
##
         195
If you do not have car installed, you can run Brown-Forsythe test manually. Compute
absolute deviance from median for each sex manually, and run ANOVA.
d\$z \leftarrow NA
for (sex in levels(d$sex)) {
```

```
d$z <- NA
for (sex in levels(d$sex)) {
   idx <- which(d$sex == sex)
   d[idx,]$z <- abs(d[idx, ]$Y - median(d[idx, ]$Y))
}
summary(aov(d$z ~ d$sex))</pre>
```

```
## Df Sum Sq Mean Sq F value Pr(>F)
## d$sex 1 9114 9114 0.808 0.37
## Residuals 195 2199050 11277
```

Is the mean landing distance different between groups?

If the data are normally distributed and have equal variance, student t-test

```
t.test(d$Y ~ d$sex,
         var.equal = T,
         paired = F
  ##
  ##
      Two Sample t-test
  ##
  ## data: d$Y by d$sex
  ## t = 0.92975, df = 195, p-value = 0.3537
  ## alternative hypothesis: true difference in means between group female and group male
  ## 95 percent confidence interval:
  ## -20.94187 58.29702
  ## sample estimates:
  ## mean in group female
                             mean in group male
  ##
                  200.7005
                                        182.0229
  If data are normally distributed but not with equal variance, Welch's t-test
  t.test(d$Y ~ d$sex,
         var.equal = F,
2
         paired = F
          )
  ##
  ##
      Welch Two Sample t-test
  ##
  ## data: d$Y by d$sex
  ## t = 0.92407, df = 181.57, p-value = 0.3567
  ## alternative hypothesis: true difference in means between group female and group male
  ## 95 percent confidence interval:
  ## -21.20353 58.55868
  ## sample estimates:
  ## mean in group female
                             mean in group male
                  200.7005
  ##
                                        182.0229
  If data are not normally distributed, you can run non-parametric tests or permutation test.
  Option 1: Non-parametric test
  Let's run Mann-Whitney U test.
  wilcox.test(d$Y ~ d$sex,
               paired = F
  ##
```

Wilcoxon rank sum test with continuity correction

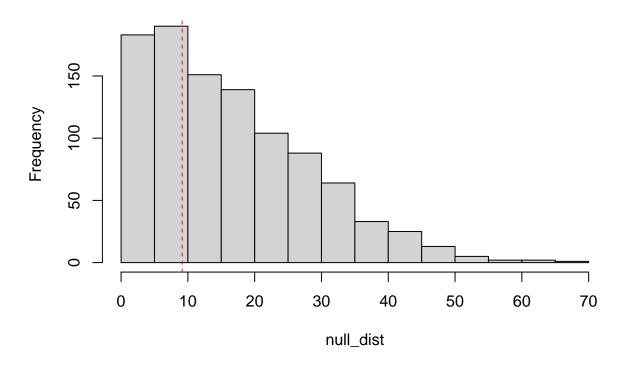
```
##
## data: d$Y by d$sex
## W = 5247, p-value = 0.2575
## alternative hypothesis: true location shift is not equal to 0
```

Option 2: Permutation test

Let's write a function to run a permutation test. Briefly, we shuffle the label, compute difference between groups. We repeat this process many times to get a null distribution. We compute p-value based on the position of the observed difference between groups in the null distribution.

```
# Null distribution of abs(mean(Y | female) - mean(Y | male))
   null dist <- sapply(1:1000, # repeat 1000 times
          function(x){
3
            d tmp <- d
4
            d tmp$sex <- sample(d$sex) # shuffle label</pre>
5
            return(abs(mean(d tmp[d tmp$sex == "female","Y"])
                      - mean(d tmp[d tmp$sex == "male","Y"]))
              )
          }
9
          )
10
11
   # Observed value of mean(female) - mean(male)
12
   obs <- abs(mean(d[d$sex == "female", "Y"] - d[d$sex == "male", "Y"]))
13
   # p value as the rank of observation in the null distribution
15
   p_val <- sum(null_dist > obs )/ length(null_dist)
16
17
   hist(null dist,
18
        main = paste("p_val: ", p_val))
19
   abline(v = obs, col = "red", lty = 2)
20
```





Optional: Statistical modeling

Here, we conduct statistical modeling of landing distance using a generalised linear model (GLM) approach with the exponential distribution, considering landing events as a Poisson process with a constant "landing rate". Because the waiting time of a Poisson process is an exponentially distributed random variable, we can use a GLM with exponential distribution. We compare the result with a linear model (LM), which assumes a normal distribution as the underlying distribution.

GLM

Fit the data to an exponential GLM.

```
glm_1 <- glm(Y ~ sex,
data = d,
family = Gamma(link = "log") # This is how to specify exponential
)</pre>
```

According to the fitted GLM, do the two groups have different landing rate?

If you have car installed

```
Anova(glm_1)
```

Analysis of Deviance Table (Type II tests)

```
##
  ## Response: Y
  ##
         LR Chisq Df Pr(>Chisq)
  ## sex 0.85803 1
                         0.3543
  summary(glm 1, dispersion = 1) # exponential is a Gamma with dispersion = 1
  ##
  ## Call:
  ## glm(formula = Y ~ sex, family = Gamma(link = "log"), data = d)
  ##
  ## Coefficients:
                 Estimate Std. Error z value Pr(>|z|)
  ## (Intercept) 5.30181
                             0.10660 49.735
                                               <2e-16 ***
  ## sexmale
                -0.09768
                             0.14331 -0.682
                                                0.495
  ## ---
  ## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
  ## (Dispersion parameter for Gamma family taken to be 1)
  ##
         Null deviance: 111.53 on 196 degrees of freedom
  ## Residual deviance: 111.07 on 195 degrees of freedom
  ## AIC: 2427.1
  ## Number of Fisher Scoring iterations: 5
  If you do not have car installed
drop1(glm 1, test = "Chi")
  ## Single term deletions
  ##
  ## Model:
  ## Y ~ sex
            Df Deviance AIC scaled dev. Pr(>Chi)
                 111.07 2427.1
  ## sex
                 111.53 2425.9 0.85803
             1
                                             0.3543
  LM
  Fit the data to a linear model.
 lm 1 \leftarrow lm(Y \sim sex,
               data = d
```

If you do not have car installed

```
drop1(lm 1, test = "Chi")
## Single term deletions
##
## Model:
## Y ~ sex
                                  AIC Pr(>Chi)
                           RSS
##
          Df Sum of Sq
## <none>
                        3831693 1949.5
## sex
                 16986 3848679 1948.4
           1
                                         0.3506
summary(lm 1)
##
## Call:
## lm(formula = Y ~ sex, data = d)
##
## Residuals:
##
       Min
                1Q Median
                                 3Q
                                        Max
## -191.51 -94.30 -25.02
                             45.70 540.72
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                 200.70
                             14.94
                                      13.43
                                              <2e-16 ***
## sexmale
                 -18.68
                             20.09
                                      -0.93
                                               0.354
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 140.2 on 195 degrees of freedom
## Multiple R-squared: 0.004413,
                                    Adjusted R-squared:
## F-statistic: 0.8644 on 1 and 195 DF, p-value: 0.3537
```

GLM vs LM

Compare the diagnostic plots¹ between glm_1 and lm_1. Is using LM for landing distance data worse than exponential GLM?

- Is the linearity assumption met?
- Is the data homoscedastic?
- Are residuals distributed normally?
- Are there outliers?

```
par(mfrow = c(2,2), oma = c(0, 0, 2, 0)
```

¹What do the diagnostic plots mean? https://www.sthda.com/english/articles/39-regression-model-diagnostics/161-linear-regression-assumptions-and-diagnostics-in-r-essentials/

```
plot(lm_1)
                                                           Im(Y ~ sex)
                        Residuals vs Fitted
                                                                                                Q-Q Residuals
       009
              8196
                                                          880
                                                                      Standardized residuals
       200 400
                                                           0000000
                                                                             က
              8
Residuals
                                                                             2
       0
                                                                             0
       -200
                                                                             ī
                    185
                                190
                                            195
                                                        200
                                                                                  -3
                                                                                          -2
                                                                                                           0
                                                                                                                   1
                                                                                                                           2
                                                                                                                                   3
                              Fitted values
                                                                                              Theoretical Quantiles
                          Scale-Location
                                                                                          Residuals vs Leverage
       2.0
              8196
                                                          880
                                                                                                                     1988
                                                                                                                               880
/|Standardized residuals
                                                            8
                                                                      Standardized residuals
              8
       1.5
                                                                                                                                 8
                                                            0
                                                                             က
                                                                                                                                 0
                                                                                                                         0
                                                                             ^{\circ}
       1.0
       0.5
                                                                             0
                                                                             ī
       0.0
```

0.000

0.004

Leverage

200

185

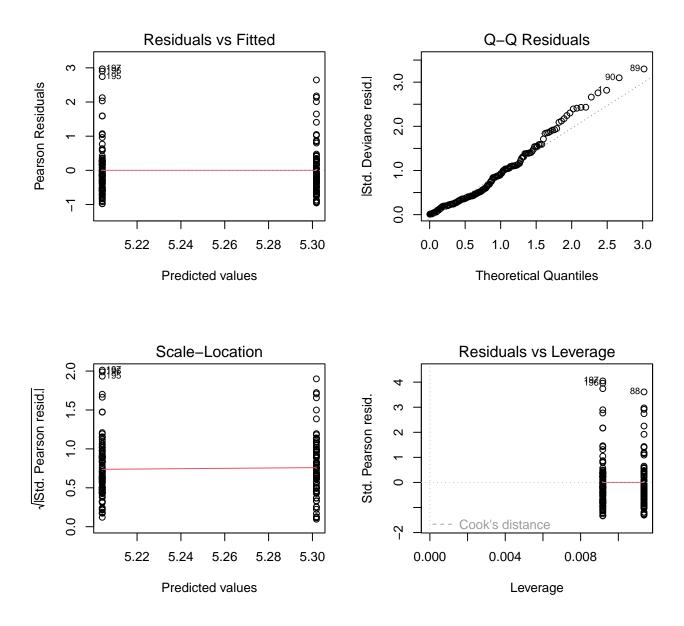
190

Fitted values

195

0.008

glm(Y ~ sex)



What's next?

In real life, we have multiple biological and technical replicates. How do we account for such random effects?

Further reading

- Whitlock & Schluter. The Analysis of Biological Data 3rd edition. Macmillan Learning. 2020
- Zuur et al. Mixed Effects Models and Extensions in Ecology with R. Springer. 2009.