

Supercritical GMC

Today: guess critical value of param. in model + some intuition.
Based on work with F. Denzato.

Conv f_1 : $\mathbb{E} X(x) X(y) \sim -\log|x-y|, x, y \in \mathbb{R}^d$
 \Rightarrow can't be reduced as rand. f_1 , but can as rand. d^2 !

Test f_2 : $\mathbb{E} X(\varphi) X(\psi) \sim -\int \int \varphi(x) \psi(y) \log|x-y| dx dy$
 log not quite pos. definite. (*)

Interested in rand. meas.:
 (I) $\mu_\gamma(dx) = c e^{\gamma X(x)} dx$ and "exponentiate".
 \hookrightarrow renorm const!

Q: what's multiplicative about GMC?

Q: what could one mean by (*) \rightarrow regularize and "renormalise" taking limit, show indep. of approx. modification.

Take $\bar{K}: \mathbb{R}^d \rightarrow \mathbb{R}^V, \bar{K} \in C_c^\infty(\mathbb{R}^d), \int \bar{K} = 1$
 \bar{K} a space-time volume noise, i.e. Grand. dist. in \mathbb{R}^{d+1}
 $\mathbb{E} \bar{K}(\varphi) \bar{K}(\psi) = \langle \varphi, \psi \rangle, \varphi, \psi$ white noise.
 "Disjoint regions in space time \Rightarrow indep."

and field $X_t(x) = \int_{\mathbb{R}^d} \int_0^t \bar{K}(e^{i(x-y)}) e^{\frac{dr}{2}} \bar{K}(y) dr$

$$:= \bar{K} \left(\int_0^t \int_{\mathbb{R}^d} \bar{K}(e^{i(x-y)}) e^{\frac{dr}{2}} \bar{K}(y) dy dr \right)$$

well-def. for any test $f \in L^2(\mathbb{R}^{d+1})$.

Observe, X_t (smooth in x , for fixed t)
 and for fixed $x, X_t(x) \leq BM$.

$$\text{Now, } \mathbb{E} X_t(x) X_s(y) = \int_{\mathbb{R}^d} \int_0^{\min(t,s)} \mathbb{E}(\bar{K}(e^{i(x-z)}) \bar{K}(e^{i(y-z)})) e^{\frac{dr}{2}} dr$$

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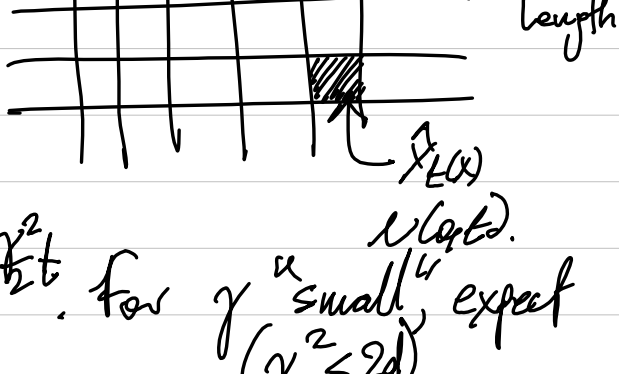
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Gaussian fields $X_t(x) \xrightarrow{\text{law}} X(x)$, $x \in \mathbb{R}^d$
 $\mathbb{E} X(x)X(y) \approx -\log|x-y| + O(1)$

Tried to give meaning to $e^{\gamma X_t - \frac{\gamma^2}{2}t} \rightarrow e^{\gamma X}$ "Wick product".
 "Worked" for $\gamma^2 \leq 2d$.

$\gamma^2 \leq 2d$: Super-toy model.



Look at $e^{\gamma X_t - \frac{\gamma^2}{2}t}$. For γ "small", expect some LLN. \downarrow Leb.

$$P(X_t(x) > \alpha t) = P(\sqrt{t}N(0,1) > \alpha t) = P(N(0,1) > \alpha \sqrt{t}) \sim \frac{e^{-\frac{\alpha^2 t}{2}}}{\sqrt{t}}$$

Contribution from "exceptional" boxes s.t. $X_t(x) \approx \alpha t$. \rightarrow $N(1)/N(\alpha t) \approx \alpha t$
 $e^{\alpha t - \frac{\alpha^2}{2}t + \alpha t - \frac{\gamma^2}{2}t}$
 $\# \text{ boxes} \sim \frac{1}{t^d}$ Fraction form \odot
 $= e^{-\frac{(\alpha-\gamma)^2}{2}t}$. If $\alpha \neq \gamma \Rightarrow$ "small cont."
 When $\alpha = \gamma$ have $\sim O(1)$ cont. (needs to be neglected poly. terms)

Limiting meas \perp w.r.t. Leb. because on C_f thick-pts.

$$A_\gamma := \{x \in \mathbb{R}^d : \lim_{t \rightarrow \infty} \frac{X_t(x)}{t} = \gamma\} \quad (\text{Hardy-Littlewood dim} = 1, \text{ sparse})$$

Q: largest value of $\hat{X}_t(x)$?

$$P(\hat{X}_t(x) > K) = P(N > \frac{K}{\sqrt{t}}) \sim \frac{e^{-\frac{K^2}{2t}}}{\sqrt{t}} \sim e^{-\frac{K^2}{2t}}$$

$$\Rightarrow -\frac{K^2}{2t} + \frac{1}{2} \log t - \log K = -dt$$

$$\Rightarrow -K \log t + 2t \log K + K^2 = 2dt^2$$

$$\text{Ansatz: } K = \sqrt{2d} \cdot t - \delta$$

$$\Rightarrow \log t - 2\delta \sqrt{2d} = 0$$

$$\Rightarrow \delta^2 = \frac{\log t}{2\sqrt{2d}} + O(1)$$

$$\Rightarrow (\#) K = \sqrt{2d} \cdot t - \frac{\log t}{2\sqrt{2d}} + O(1)$$

So for γ large, should do different normalisation, get localisation, limit collection of Dirac meas.

$$\gamma^2 > 2d: e^{\gamma X_t(x) + (d - \gamma \sqrt{2d})t}$$

$O(1)$ boxes contribute $(\#)$ Gaussian dist. \rightarrow $e^{-\frac{(\sqrt{2d}t + \delta)^2}{2t}} = e^{-\frac{2d}{2}t - \sqrt{2d}\delta + \frac{\delta^2}{2t}}$

"Out in the tail, dist $\sim \exp(-t)$ ". So, tells us contributions of each box and view as point process on \mathbb{R}^d .

$$(+) \sum_{x \in \mathbb{Q}^d} \delta_{x, X_t(x) - \sqrt{2d}t} \rightarrow \text{small contributions} \rightarrow -\infty \text{ (don't see them)}$$

Only exceptional values $X_t(x) = \sqrt{2d}t + \delta$. So expect $(+) \Rightarrow$ P.P.P. $(dx \otimes e^{-\sqrt{2d}\delta} d\delta)$.

$$\text{So } \int_{\text{test f}} e^{\gamma X_t(x) + (d - \gamma \sqrt{2d})t} dx \Rightarrow \sum_{(x,s) \in \text{P.P.P.}(\dots)} f(x) e^{\gamma s} \quad (*)$$

$$\text{let } m = e^{\gamma s} \quad dm = \gamma e^{\gamma s} ds$$

$$\text{So } e^{-\sqrt{2d}s - \gamma s} dm = \gamma e^{-\sqrt{2d}s} ds$$

$$m^{1-\beta} dm, \quad \beta = \frac{\sqrt{2d}}{\gamma} \in (0,1)$$

$$\Rightarrow \odot \sum_{(x,m) \in \text{P.P.P.}(dx \otimes m^{1-\beta} dm)} m \cdot f(x)$$

not integrable at zero. P.P.P. has inf. many points near zero (going to accumulate). But quantity above is well-defined due to m .

$\gamma = \sqrt{2d} \Rightarrow$ not integrable but have a lot of mass from bulk \rightarrow still conv. to Leb. so just change normalisation to $t^{\frac{1}{2}} e^{\gamma X_t - \frac{\gamma^2}{2}t}$.

Q: recover spatial structure of fields $X_t(x)$ in limit, other than Leb.

$$\text{recall } X_t = X_s + X_{s,t} \xrightarrow{\text{law}} X_s + X_{s,t} \xrightarrow{\text{law}} X_s + X_{s,t} \xrightarrow{\text{law}} X_s + X_{s,t}$$

But X_s term \rightarrow contradiction, so will hope to conv. to $(\#)$

$$\text{Call. } \left[\begin{array}{l} \text{P.P.P. } \beta(dx) \\ v \sim \text{P.P.P. } \beta(\mu) \\ \int f(x) v(dx) = \sum_{(x,m) \in \text{P.P.P.}(\mu \otimes m^{1-\beta} dm)} m f(x) \end{array} \right]$$

$$\text{Claim: } X_{t-s}(e^{\gamma \cdot}) \xrightarrow{\text{law}} \text{PPP}_\beta(dx) \text{ and put correct st-dep. prefactors to } \mu_\beta \dots$$

Now, $e^{\gamma X_t} \sim e^{\gamma X_s} \text{PPP}_\beta(dx)$ keep track of exp. normalisations log terms more tricky.

$$\text{Recall: } v \sim \text{PPP}(\mu)$$

$$\mathbb{E} e^{-\int f(x) v(dx)} = \prod_x \frac{\mathbb{E} e^{-f(x) v(dx)}}{e^{-\mu(dx)(1-e^{-f(x)})}} = \int (1-e^{-f(x)}) \mu(dx)$$

$$\text{So for } v \sim \text{PPP}_\beta(\mu)$$

$$\mathbb{E} e^{-\int f(x) v(dx)} \odot \exp\left(-\int (1-e^{-m f(x)}) m^{1-\beta} \mu(dx)\right)$$

$$\text{Let } q = m f(x) \quad dq = dm f(x)$$

$$\odot \exp\left(-\int_0^\infty (1-e^{-q}) q^{-1-\beta} dq \int f(x) \mu(dx)\right)$$

$$= \exp\left(-c_\beta \int f(x) \mu(dx)\right)$$

$$\text{So } e^{\gamma X_s} \text{PPP}_\beta(dx) = \text{PPP}_\beta(c_\beta \gamma X_s(dx)) \odot \text{PPP}_\beta(c_\beta \gamma X_s(dx))$$

Suggests normalise $X_t \Rightarrow$ conv. to non-deg. limit \Rightarrow rand. purely atomic measure with rand. intensity given in terms of centric GMC. Two layers of randomness become indep.

$$\odot \rightarrow \text{PPP}_\beta(\text{GMC}_{\sqrt{2d}})$$

Conj. Duplantier, Sheffield & Vargas γ -1/2 power law for masses "known" by physicists before.

$$\text{With } \mu_t^{(\gamma)}(dx) := \mathbb{E} e^{\gamma X_t(x) - c(\dots)t} dx, \text{ look at map } s \mapsto \mu_{t,s,\gamma} \text{ (measure-valued)}$$

\hookrightarrow have to make sense of limiting measure changes (expect locations to be the same)

$$\text{So expect: } \left[\begin{array}{l} \text{GMC}_{\sqrt{2d}} \\ \text{at small scales, } \alpha e^{-s}, \text{ field is "smooth"} \end{array} \right]$$

Expect:

$$e^{\gamma X_t + (d - \sqrt{2d}\gamma)t}$$

$$X_t = X_s + \underbrace{\tilde{X}_{t-s}(e^{\gamma \cdot})}_{\substack{\text{max in box} \\ \sim \sqrt{2d(t-s)} \text{ in} \\ \text{each box}}} \sim \text{super-ty model}$$

So global max dx unit scales down
 ~ (of order) $\sqrt{2d(t-s)} + \sqrt{\frac{d}{2}} \cdot s$
 (Since $\text{Exp}(-\sqrt{2d}k) \sim e^{-\frac{d}{2}k^2} \sim \frac{s \cdot d}{\sqrt{2d}}$)

"Width" of the peak around the global max not
 $(X_{t-s} \text{ has peaks of width } \sim e^{-(t-s)} \text{ & rescale})$
 by e^{γ}

So concentration of max. peaks \sim
 $\text{Exp}(-td + \gamma\sqrt{2d}(t-s) + \gamma\sqrt{\frac{d}{2}}s)$

$$A \leftarrow \left(e^{\gamma X_{t-s}(t-s)} \right)$$

Now,
$$e^{\gamma X_t + (d - \sqrt{2d}\gamma)t} = e^{(d - \sqrt{2d}\gamma)t + \gamma\sqrt{\frac{d}{2}}s} e^{\gamma X_{t-s}} (e^{\gamma\sqrt{2d}X_{t-s} - sd})^{1/2}$$

"Correct normalization" $\Rightarrow \text{PPP}(\text{Lab})$
 "correct norm" $\Rightarrow \text{GMC}_{\text{Lab}}$

Now can convolve with mollifier at scale $\varepsilon = e^{-t}$

$$X(\varepsilon) = X * \rho_\varepsilon, \quad \rho_\varepsilon(x) = \varepsilon^{-d} \rho(x/\varepsilon)$$

and recover group limit. For large class of mollifiers:

$$X(\varepsilon) = X_t + W(e^{\frac{t}{2}}) + o(\varepsilon) \quad (\varepsilon = e^{-t})$$

where W is a stationary, smooth Gaussian field with decaying covariance.

Intuition about local maxima & glob. max \gg loc.

At $O(1)$ scales, properties of max don't change due to decay in cov.

Let $v \sim \text{PPP}(m^{-1}d\mu), (Z_i)_i \text{ i.i.d.}$

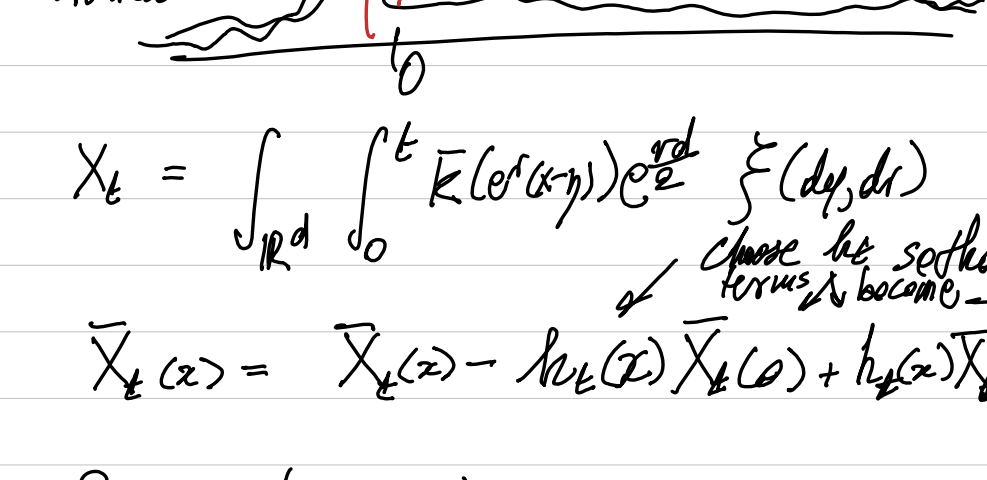
$$v = \sum \delta_{m_i}, \quad \tilde{v} := \sum \delta_{m_i Z_i}$$

$\Rightarrow \tilde{v} \sim \text{PPP}(cm^{1-\beta}d\mu)$ (computer code trans.)

$\Rightarrow W$ might change weights in limit mildly.

Q: How does glob. max of X_t behave if one conditions on width of max. rescaled from e^{-t} to t ?

Let $\tilde{X}_t = X_t(e^{-t}) \rightarrow O(1)$ correlations
 Look at $\tilde{X}_t - \tilde{X}_t(0)$, conditioned on $\tilde{X}_t(0) = \text{stat}$



$$X_t = \int_{\mathbb{R}^d} \int_0^t K(e^{r-x}) e^{\frac{rd}{2}} \xi(dy, dr)$$

$$\tilde{X}_t(x) = \tilde{X}_t(x) - h_t(x) \tilde{X}_t(0) + h_t(x) \tilde{X}_t(0)$$

So computing covariance:

$$\mathbb{E}[\tilde{X}_t(0) \tilde{X}_t(x) h_t(x) - t h_t^2(x)] = 0$$

So can take $h_t(x) = \frac{1}{t} \int_0^t K(e^{-s}x) ds$

Have decomposition of \tilde{X}_t

$$\tilde{X}_t(x) = \int_0^t K(e^{-r}x) d\tilde{X}_r(0) + Z_t(x)$$

(at the level of processes) \nearrow BM t indep.

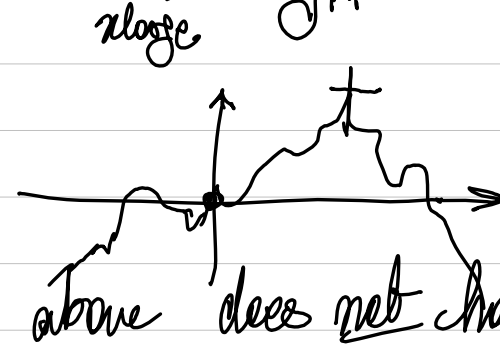
$$\mathbb{E} Z_t(x) Z_t(y) = \int_0^{\text{stat}} (K(e^{-x}y) - K(e^{-x})K(e^{-y})) dt$$

Now $\tilde{X}_t(x)$ cond. on $\tilde{X}_t(0) \sim \sqrt{2d}t$ (after cond. $\tilde{X}_t(0) \sim \text{BM}$ width)

$$\Rightarrow \sqrt{2d} \int_0^t (K(e^{-r}x) - 1) dr + \int_0^t (K(e^{-r}x) - 1) dB + Z_t(x)$$

$$\xrightarrow{\text{as } t \rightarrow \infty} \sqrt{2d} \int_0^\infty (K(e^{-r}x) - 1) dr + \int_0^\infty (K(e^{-r}x) - 1) dB + Z_\infty(x)$$

Intuition:



turns out field above does not have a global max.

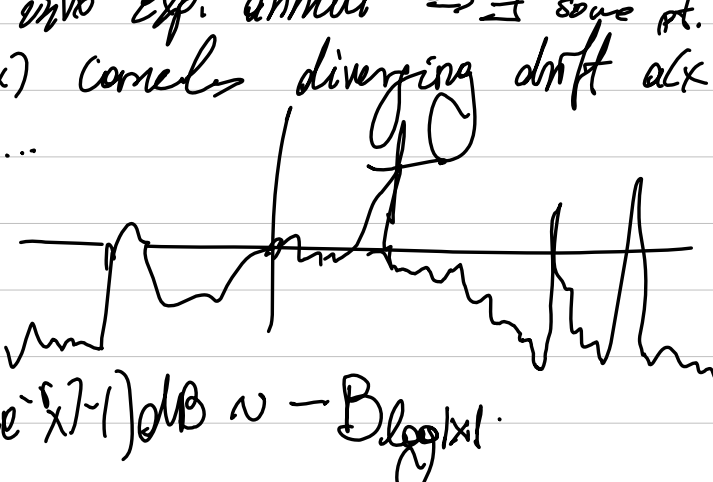
Why? $Z_t(x) = \int_{\mathbb{R}^d} \int_0^t (K(e^{-x}y) - K(e^{-x})K(e^{-y})) e^{-\frac{rd}{2}} \xi(dy, dr)$

For $x, y \in \mathbb{R}^d \Rightarrow Z_\infty(x) \sim Z_\infty(y) \sim X_t(x)$

Expect: $\sup_{t \in \mathbb{R}^d} Z_\infty(x) \sim \sqrt{2d} \cdot \ln \sqrt{2d} \log |x|$

"Chop space into exp. annuli" $\Rightarrow \exists$ some pt. where $Z_\infty(x)$ comes diverging drift $a(x)$.

So chance ...

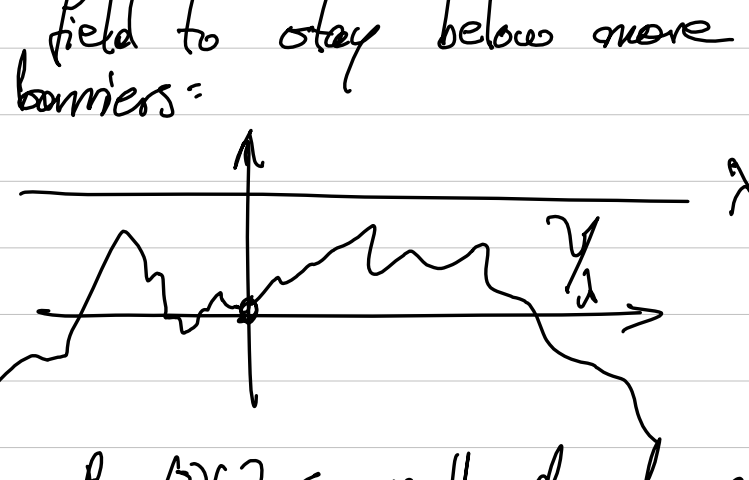


and $\int_0^\infty (K(e^{-r}x) - 1) dB \sim -B \log |x|$

Need to further condition on origin being close to origin.

Condition process to be < 1 on growing balls \Rightarrow get random field. (see paper)

Condition field to stay below more general barriers:



(and BM $\sim \text{Bes}(\frac{3}{2})(0)$ so really do have (localised) Global max. and then change coordinates.

Want/hope: law of limiting field indep. of barrier λ .

However, it does depend on λ .

Call Ψ := field near local maxima

"location" \sim unit
 "max." \sim $\text{PPP}(e^{-\frac{F(x)}{2}} d\psi \otimes \mathcal{H}(d\Psi))$
 "shape": ?

write $\mathcal{P} := \text{law } \Psi$

Have $\gamma_\lambda(0) = 0, \gamma_\lambda(y) = \gamma(y+z) + s, 0 \leq s \leq \lambda$

$$\Rightarrow s = -\gamma(x) \Rightarrow -\lambda \leq \gamma(x) \leq 0$$

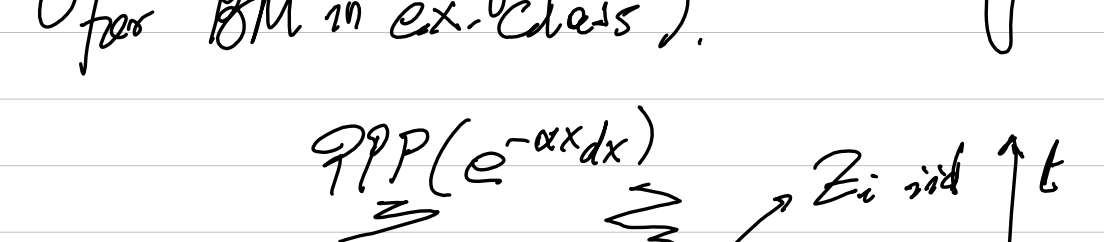
So for $\gamma_\lambda(x) \xrightarrow{\text{loc. max}} \text{PPP}(e^{-\frac{F(x)}{2}} \mathbb{1}_{\{\gamma(x) \geq -\lambda\}} dx \otimes \mathcal{P}(d\Psi))$

$$\text{So } \mathbb{E} F(\gamma_\lambda) \propto \mathbb{E} F(\Psi) \cdot \int e^{\frac{\sqrt{2d} \gamma_\lambda}{2}} \mathbb{1}_{\{\gamma(x) \geq -\lambda\}} dx$$

Can insert formula to obtain law of Ψ in terms of λ and show that this "agrees" w/ indep. of λ (see analogue for BM in ex. class)

$$\text{PPP}(e^{-\alpha x} dx) \xrightarrow{\text{Zi iid}} \sum_{i=1}^{\infty} Z_i(t)$$

look at evolution in t:



what guarantees stationarity in time?

stat. increments? \rightarrow not true.

need stat. increments modulo some tilt.

$$(Z_s, \tilde{Z})(t) = Z(t+s) - Z(s)$$

$$Z_0 = 0, \tilde{Z} \in \mathcal{C}(\mathbb{R}_+; \mathbb{R})$$

Take $F_\lambda: \mathbb{R} \times \mathbb{C}_0 \rightarrow \mathbb{R}$

$$F_\lambda(x_j + Z_j(s_i), Z_s; Z_j)$$

$$\mathbb{E} e^{-\int_0^\infty (1 - \mathbb{1}_{\{e^{-F_\lambda(x_j + Z_j(s_i), Z_s; Z_j)}\}}) ds} = e^{-\alpha x} \mathbb{E} d\mathcal{P}(B)$$

Now take $0 \leq t \leq \inf S_i$

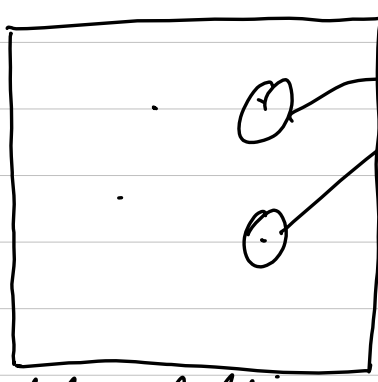
$$\mathbb{E} e^{-\int_0^\infty (1 - \mathbb{1}_{\{e^{-F_\lambda(x_j + Z_j(s_i), Z_s; Z_j)}\}}) ds} = e^{-\alpha x} \mathbb{E} d\mathcal{P}(B)$$

$$= \exp(-(\dots F_\lambda(x + (Z_\infty B)_{x-t}, Z_\infty + (Z_\infty B)_{x-t})) \times e^{-\alpha x} \mathbb{E} d\mathcal{P}(B))$$

Want: (Stationarity)

$$\int \tilde{F}(x+B) e^{-\alpha B} \mathcal{P}(dB) = \int \tilde{F}(B) \mathcal{P}(dB)$$

Gaussian field



maximal points.

Goal: study field around extremal points.

More specifically, talk about BM.

Brownian motion around extreme points

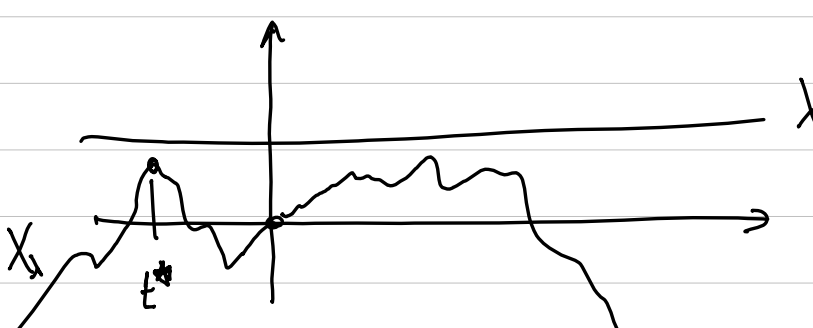
Let B be a two-sided B.M. $B_0 = 0$.

Strategy:

1) Take $\lambda > 0, R > 0$.

Let $X_{\lambda,R} : \mathbb{R} \rightarrow \mathbb{R}$ conditioned on event $\left\{ \sup_{[-R,R]} B(t) < \lambda \right\}$

Prop 1: $X_{\lambda,R} \xrightarrow{\text{law}} X_\lambda$ ($R \rightarrow \infty$)
and $\lambda - X_\lambda$ is Bessel(3) process.



Want to recenter around t^* .

Define Ψ_λ s.t. $\forall F : \mathbb{C} \rightarrow \mathbb{R}$ $\tau_{t^*} X(s) = X(t+s) - X(t^*)$
 $E[F(\Psi_\lambda)] \propto E \left[\frac{F(\tau_{t^*} X_\lambda)}{|\{u : X_\lambda(u) \geq X_\lambda(t^*) - \lambda\}|} \right]$ (shift operator)

Prop 2: Ψ_λ is indep. of λ .

Proof (Prop 1):

Let \tilde{B} be another B.M. $\tilde{B}(0) = 0$
 \tilde{X}_i conditioned to be positive at $[-R,R]$. Change $0 < t_1 < t_2$

Compute $P(\tilde{X}_{\lambda,R}(t_2) = x_2 \mid \tilde{X}_{\lambda,R}(t_1) = x_1)$

$$= \frac{P(\tilde{X}(t_2) = x_2, \tilde{X}(t_1) = x_1)}{P(\tilde{X}(t_1) = x_1)}$$

$$= \frac{P(\tilde{B}(t_2) = x_2, \tilde{B}(t_1) = x_1, \inf_{[-R,R]} \tilde{B} > 0)}{P(\tilde{B}(t_1) = x_1, \inf_{[-R,R]} \tilde{B} > 0)}$$

$$P(\tilde{B}(t_2) = x_2, \tilde{B}(t_1) = x_1, \inf_{[-R,R]} \tilde{B} > 0) \cdot \frac{P(\inf_{[x_2, t_2]} \tilde{B} > 0)}{P(\inf_{[x_1, t_1]} \tilde{B} > 0)}$$

Let ϕ^0 be the transition prob. of BM killed at 0,
 $\phi^0 = \phi^0(0, t_1, t_1, x_1) \phi^0(t_1, x_1, t_2, x_2)$

$$\times (1 - 2P_0(B(R-t_2) > x_2))$$

$$= P(|B(R-t_2)| < x_2)$$

$$\sim \frac{x_2}{\sqrt{R-t_2}}$$

As $R \rightarrow +\infty$, we get

$$\lim_{R \rightarrow \infty} P(\tilde{X}_{\lambda,R}(t_2) = x_2 \mid \tilde{X}_{\lambda,R}(t_1) = x_1) = \phi^0(t_1, x_1, t_2, x_2) \cdot \frac{x_2}{x_1}$$

and obtain X_λ has same finite dim. distr. of Bessel(3) process.

Prop 3: (sampling property of X_λ).

For any G , have

$$E G(X_\lambda) = E \left[\frac{\int G(\tau_{t^*} X_\lambda) \mathbb{1}_{X_\lambda(t^*) \geq X_\lambda(t) - \lambda} dt}{|\{u : X_\lambda(u) \geq X_\lambda(t^*) - \lambda\}|} \right]$$

"Sample t^* unif. and recenter and condition."

Proof (Prop 2): let $\lambda_2 < \lambda_1$.

$$E F(\Psi_{\lambda_2}) \stackrel{?}{=} E \left(\frac{F(\tau_{t^*} X_{\lambda_2})}{|\{u : X_{\lambda_2}(u) \geq X_{\lambda_2}(t^*) - \lambda_2\}|} \right)$$

$X_{\lambda_2} = X_{\lambda_1}$ conditioned to be $\leq \lambda_2$.

$$\stackrel{?}{=} E \left[\frac{F(\tau_{t^*} X_{\lambda_1}) \mathbb{1}_{X_{\lambda_1}(t^*) \leq \lambda_2}}{|\{u : X_{\lambda_1}(u) \geq X_{\lambda_1}(t^*) - \lambda_2\}|} \right] = F(\Psi_{\lambda_2}).$$

$\hookrightarrow =: G$

Shift X by u uniform on $\{u : X_{\lambda_1}(u) \geq X_{\lambda_1}(t^*) - \lambda_2\}$ indep. of X_{λ_1}

$$\mathbb{1}_{(X_{\lambda_1}(t^*) - X_{\lambda_1}(t) \leq \lambda_2)}.$$