

Statistics of High Frequency Financial Data FINM 33170 and STAT 33910

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Some Highlights of Stochastic Calculus

- Probability Spaces, Conditional Expectation
- Stochastic Integrals, Itô-Processes
- Semimartingales

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Information Sets, σ -fields

- Basic space: (Ω, \mathcal{F})
- Ω is the set of all possible outcomes ω
- \mathcal{F} is the collection of subsets $A \subseteq \Omega$ that will eventually be decidable (it will be observed whether they occurred or not)

A collection \mathcal{A} of subsets of Ω is a σ -field (sets that are decidable by a procedure of observation) if

- $\emptyset, \Omega \in \mathcal{A}$: \emptyset, Ω are decidable (\emptyset didn't occur, Ω did)
- if $A \in \mathcal{A}$, then $A^c = \Omega - A \in \mathcal{A}$: if A occurs, then A^c does not occur, and *vice versa*
- if $A_n, n = 1, 2, \dots$ are all in \mathcal{A} , then $\bigcup_{n=1}^{\infty} A_n \in \mathcal{A}$: if all the A_n are decidable, then so is the event $\bigcup_{n=1}^{\infty} A_n$ (the union occurs if and only if at least one of the A_i occurs)

- All random variables are thought to be a function of the basic outcome $\omega \in \Omega$:
 X is of the form $X(\omega)$
- A random variable X is called \mathcal{A} -measurable if the value of X can be decided on the basis of the information in \mathcal{A} .
 Formally, the requirement is that for all x , the set $\{X \leq x\} = \{\omega \in \Omega : X(\omega) \leq x\}$ be decidable ($\in \mathcal{A}$).
- The σ -field that has the same information as a random variable: $\mathcal{F}(X)$ is the smallest *sigma*-field making X measurable
- $\mathcal{F}(X)$ is the intersection of all σ -fields \mathcal{A} so that X called \mathcal{A} -measurable. Exercise: verify that this is indeed a σ -field
- Ω and ω are usually "anonymous", but tie together the random variables in the same system

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Specifically,

The process $(W_t)_{0 \leq t \leq T}$ is an (\mathcal{F}_t) -Wiener process if it is adapted to (\mathcal{F}_t) and

- $W_0 = 0$;
- $t \rightarrow W_t$ is a continuous function of t ;
- W has independent increments relative to the filtration (\mathcal{F}_t) : if $t > s$, then $W_t - W_s$ is independent of \mathcal{F}_s ;
- for $t > s$, $W_t - W_s$ is normal with mean zero and variance $t - s$ ($N(0, t-s)$).

Brownian motion (W_t) is an (\mathcal{F}_t^W) -Wiener process.

“Predictable”: one can forecast the value over infinitesimal time intervals. Basic example: “simple process”, defined by break points $0 = s_0 = t_0 \leq s_1 < t_1 \leq s_2 < t_2 < \dots \leq s_n < t_n \leq T$, and random variables $H^{(i)}$, observable (measurable) with respect to \mathcal{F}_{s_i} .

$$H_t = \begin{cases} H^{(0)} & \text{if } t = 0 \\ H^{(i)} & \text{if } s_j < t \leq t_j \end{cases} \quad (1)$$

At any time t , the value of H_t is known *before* time t (for $t \neq 0$).

Definition

More generally, a process H_t is predictable if it can be written as a limit of simple functions $H_t^{(n)}$. This means that $H_t^{(n)}(\omega) \rightarrow H_t(\omega)$ as $n \rightarrow \infty$, for all $(t, \omega) \in [0, T] \times \Omega$.

All adapted left continuous processes are predictable (càg, for *continue à gauche*) (JS, Proposition I.2.6 (p. 17))

The expression (2) is defined for simple process integrands as

$$\int_0^T H_t^2 dt < \infty. \quad (4)$$

It will always be the case that the integrator X_t is right continuous with left limits (*càdlàg*, for *continue à droite, limites à gauche*).

$$\int_0^t H_s dX_s = \int_0^T H_s I\{s \leq t\} dX_s \quad (5)$$

can also be taken to be *càdlàg*. If (X_t) is continuous, the integral is then automatically continuous.

X_t is an Itô process relative to filtration (\mathcal{F}_t) provided (X_t) is (\mathcal{F}_t) adapted; and if there is an (\mathcal{F}_t) -Wiener process (W_t) , and (\mathcal{F}_t) -adapted processes (μ_t) and (σ_t) , with

$$\int_0^T |\mu_t| dt < \infty, \text{ and } \int_0^T \sigma_t^2 dt < \infty \quad (6)$$

so that

$$X_t = X_0 + \int_0^t \mu_s ds + \int_0^t \sigma_s dW_s. \quad (7)$$

The process is often written on differential form:

$$dX_t = \mu_t dt + \sigma_t dW_t. \quad (8)$$

We note that the Itô process property is preserved under stochastic integration. If H_t is bounded and predictable, then

$$\int_0^t H_S dX_S = \int_0^t H_S \mu_S dt + \int_0^t H_S \sigma_S dW_S. \quad (9)$$

It is clear from this formula that predictable processes H_t can be used for integration w.r.t. X_t provided

$$\int_0^T |H_t \mu_t| dt < \infty \text{ and } \int_0^T (H_t \sigma_t)^2 dt < \infty. \quad (10)$$

$$\nabla u(i)(x) = u(i)(x) \quad (iii)$$

The Heston Model

A popular model for volatility is due to Heston (1993). In this model, the process X_t is given by

$$dX_t = \left(\mu - \frac{\sigma_t^2}{2} \right) dt + \sigma_t dW_t$$

$$d\sigma_t^2 = \kappa(\alpha - \sigma_t^2)dt + \gamma\sigma_t dZ_t, \text{ with} \quad (12)$$

$$Z_t = \rho W_t + (1 - \rho^2)^{1/2} B_t \quad (13)$$

where (W_t) and (B_t) are two independent Wiener processes, $\kappa > 0$, and $|\rho| \leq 1$. To assure that σ_t^2 does not hit zero, one must also require (Feller (1951)) that $2\kappa\alpha \geq \gamma^2$.

Conditional Expectations: Heuristic Definition

- X, Y random variables, $E|X| < \infty$
- Non-random conditional expectation: $f(y) = E(X|Y = y)$
(from undergraduate courses)
- Random conditional expectation given random variable Y :
 $E(X|Y) = f(Y)$
- Random conditional expectation given σ -field: if \mathcal{A} is
generated by Y , then $E(X|\mathcal{A}) = f(Y)$

Random conditional expectation is very useful in time dependent systems and large dimension. We shall see same phenomenon for Radon-Nikodym derivatives, a.k.a. likelihood ratios, a.k.a. state price densities.

Heuristic definition runs into trouble for measure theoretic reasons, and also practically in time dependent systems and large dimension. But the heuristic is very useful to help you think.

Conditional Expectations: Official Definition

Theorem

Let \mathcal{A} be a σ -field, and let X be a random variable so that $E|X| < \infty$. There is a \mathcal{A} -measurable random variable Z so that for all $A \in \mathcal{A}$,

$$EZI_A = EXI_A, \quad (14)$$

where I_A is the indicator function of A . Z is unique “almost surely”, which is that if Z_1 and Z_2 satisfy the two criteria above, then $P(Z_1 = Z_2) = 1$.

We thus define

$$E(X|\mathcal{A}) = Z \quad (15)$$

where Z is given in the theorem. The conditional expectation is well defined “almost surely”.

For further details and proof of theorem, see Section 34 (p. 445-455) of Billingsley (1995).

For a taste of the Proof

Show that $Z_1 = Z_2$ with probability 1.

[Hist: Consider set $A = \{\omega : Z_1(\omega) \geq Z_2(\omega)\}$, and so on.]

- Conditional constants: if Z is \mathcal{A} -measurable, then

$$E(ZX|\mathcal{A}) = ZE(X|\mathcal{A})$$

- $$E[E(X|\mathcal{A})|\mathcal{A}'] = E(X|\mathcal{A}')$$

- $$E(X|\mathcal{A}) = E(X)$$

- $$E(g(X)|\mathcal{A}) \geq g(E(X|\mathcal{A}))$$

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An (\mathcal{F}_t) adapted process M_t is called a **martingale** if, for all t , $E|M_t| < \infty$, and if, for all $s < t$,

$$E(M_t|\mathcal{F}_S) = M_S. \quad (16)$$

- Central concept in our narrative.
- The concept of martingale applies equally to discrete and continuous time axis.
- A martingale is also known as a *fair game*, for the following reason. In a gambling situation, if M_s is the amount of money the gambler has at time s , then the gambler's expected wealth at time $t > s$ is also M_s .

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$$\begin{aligned} E(W_t | \mathcal{F}_s) &= E(W_t - W_s | \mathcal{F}_s) + W_s \\ &= E(W_t - W_s) + W_s \text{ by independence} \\ &= W_s. \end{aligned} \tag{17}$$

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Stochastic Integrals and Martingales

Example

If H_t is a bounded predictable process, and for any martingale X_t ,

$$M_t = \int_0^t H_s dX_s \quad (19)$$

is a martingale.

Thus, any bounded trading strategy H in an asset M which is a martingale results in a martingale P/L.

To see this, consider first a simple process (1), for which $H_s = H^{(i)}$ when $s_i < s \leq t_j$. For given t , if $s_i > t$, by the properties of conditional expectations,

$$\begin{aligned} E\left(H^{(i)}(X_{t_j} - X_{s_i})|\mathcal{F}_t\right) &= E\left(E(H^{(i)}(X_{t_j} - X_{s_i})|\mathcal{F}_{s_i})|\mathcal{F}_t\right) \\ &= E\left(H^{(i)}E(X_{t_j} - X_{s_i}|\mathcal{F}_{s_i})|\mathcal{F}_t\right) = 0, \end{aligned}$$

and similarly, if $s_i \leq t \leq t_j$, then

$$E\left(H^{(i)}(X_{t_j} - X_{s_i})|\mathcal{F}_t\right) = H^{(i)}(X_t - X_{s_i})$$

so that

$$\begin{aligned} E(M_T|\mathcal{F}_t) &= E\left(\sum_i H^{(i)}(X_{t_j} - X_{s_i})|\mathcal{F}_t\right) \\ &= \sum_{i:t_j < t} H^{(i)}(X_{t_j} - X_{s_i}) + I\{t_i \leq t \leq s_i\} H^{(i)}(X_t - X_{s_i}) = M_t. \end{aligned}$$

The result follows for general bounded predictable integrands by taking limits and using uniform integrability. For definition and results on uniform integrability, see Billingsley (1995).

Example of local martingale (see also Duffie (1996)):

$$X_t = \int_0^t \frac{1}{\sqrt{T-s}} dW_s \quad (20)$$

For $0 \leq t < T$, X_t is a zero mean Gaussian process with independent increments. We shall show below that

$$\text{Var}(X_t) = \int_0^t \frac{1}{T-s} ds = \int_{T-t}^T \frac{1}{u} du = \log \frac{T}{T-t} \rightarrow \infty \text{ as } t \rightarrow T.$$

X_t is not defined at T . However, one can **stop** the process at a convenient time, as follows: Set, for $A > 0$,

$$\tau = \inf\{t \geq 0 : X_t = A\}. \quad (21)$$

One can show that $P(\tau < T) = 1$. Define a modified integral by

$$\begin{aligned} Y_t &= \int_0^t \frac{1}{\sqrt{T-s}} I\{s \leq \tau\} dW_s \\ &= X_{T \wedge t}, \text{ where } s \wedge t = \min(s, t). \end{aligned} \quad (22)$$

0 1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18 19 20 21 22 23 24 25 26 27 28 29 30 31 32 33 34 35 36 37 38 39 40 41 42 43 44 45 46 47 48 49 50 51 52 53 54 55 56 57 58 59 60 61 62 63 64 65 66 67 68 69 70 71 72 73 74 75 76 77 78 79 80 81 82 83 84 85 86 87 88 89 90 91 92 93 94 95 96 97 98 99

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1. **Identify the subject and predicate of the sentence.**

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- Lesson for trading: some condition has to be imposed to make sure that a trading strategy in a martingale cannot result in arbitrage profit.
- The most popular approach to this is to require that the traders wealth at any time cannot go below some fixed amount $-K$. This is the so-called credit constraint. (So strategies are required to satisfy that the integral never goes below $-K$).
- This constraint does not quite guarantee that the integral w.r.t. a martingale is a martingale, but it does prevent arbitrage profit. The technical result is that the integral is a *super-martingale* (see the next section).

References

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Local Martingales

Definition

A process M_t is a local martingale for $0 \leq t \leq T$ provided there is a sequence of stopping times τ_n so that

- (i) $M_{\tau_n \wedge t}$ is a martingale for each n ; and
- (ii) $P(\tau_n \rightarrow T) = 1$ as $n \rightarrow \infty$.

The basic result for stochastic integrals is now that the integral with respect to a local martingale is a local martingale, cf. result I.4.34(b) (p. 47) in JS.

Note: some variations over the concept of local martingale are in use.

Semimartingales

X_t is a semimartingale if it can be written

$$X_t = X_0 + M_t + A_t, 0 \leq t \leq T, \quad (23)$$

where X_0 is \mathcal{F}_0 -measurable, M_t is a local martingale, and A_t is a process of finite variation, *i.e.*,

$$\sup_i \sum |A_{t_{i+1}} - A_{t_i}| < \infty, \quad (24)$$

where the supremum is over all grids $0 = t_0 < t_1 < \dots < t_n = T$, and all n .

In particular, a continuous Itô process is a semimartingale, with

$$M_t = \int_0^t \sigma_t dW_t \text{ and } A_t = \int_0^t \mu_t dt. \quad (25)$$

A *supermartingale* is semimartingale for which A_t is nonincreasing. A *submartingale* is a semimartingale for which A_t is nondecreasing.

Quadratic Variation of a Semimartingale

- Grid of observation times: $\mathcal{G} = \{t_0, t_1, \dots, t_n\}$ with $0 = t_0 < t_1 < \dots < t_n = T$
- Maximum discrepancy: $\Delta(\mathcal{G}) = \max_{1 \leq i \leq n} (t_i - t_{i-1})$.
- Quadratic variation of process X relative to grid \mathcal{G}

$$[X, X]_t^{\mathcal{G}} = \sum (X_{t_{i+1}} - X_{t_i})^2. \quad (26)$$

- Quadratic covariation of X and Y :

$$[X, Y]_t^{\mathcal{G}} = \sum_{t_{i+1} \leq t} (X_{t_{i+1}} - X_{t_i})(Y_{t_{i+1}} - Y_{t_i}). \quad (27)$$

Theorem

For any semimartingale, there is a process $[X, Y]_t$ so that

$$[X, Y]_t^{\mathcal{G}} \xrightarrow{P} [X, Y]_t \text{ for all } t \in [0, T], \text{ as } \Delta(\mathcal{G}) \rightarrow 0. \quad (28)$$

The limit is independent of the sequence of grids \mathcal{G} .

The result follows from Theorem I.4.47 (p. 52) in JS. The t_i can even be stopping times.

- For a continuous Itô process,

$$[X, X]_t = \int_0^t \sigma_s^2 ds. \quad (29)$$

(Cf Thm I.4.52 (p. 55) and I.4.40(d) (p. 48) of JS.) In particular, for a Wiener process W , $[X, X]_t = \int_0^t 1 ds = t$.

- The process $[X, X]_t$ is usually referred to as the quadratic variation of the semimartingale (X_t) . This is an important concept, as seen in the rest of the course. The theorem asserts that this quantity can be estimated consistently from data.
- The result follows from Theorem I.4.47 (p. 52) in JS.
- The t_i can even be stopping times.

Semimartingales with Jumps

- Most commonly studied discontinuous semimartingale:

$$X_t = X_0 + \int_0^t \mu_s ds + \int_0^t \sigma_s dW_s + J_t$$

- J_t is a process whose only quadratic variation comes from jumps: $[X, X]_t = \int_0^t \sigma_s^2 ds + [J, J]_t$ where

$$[J, J]_t = \sum_{0 < s \leq t} \Delta X_s^2 = \sum_{0 < s \leq t} \Delta J_s^2 \text{ and } \Delta X_s = X_s - X_{s-}$$

- Decomposition $J_t = J_t^{(1)} + J_t^{(2)}$
- $J_t^{(1)}$ has finitely many jumps at times that can be predictable
- $J_t^{(2)}$ is a “pure jump” Itô semimartingale
- Under condition (H) (Jacod and Protter (2012), p. 126), the “predictable” quadratic variation $\langle J^{(2)}, J^{(2)} \rangle_t$ is absolutely continuous and locally bounded. This is a generalization of a compound Poisson process.

Properties of Quadratic Variation

- (1) Bilinearity: $[X, Y]_t$ is linear in each of X and Y : so for example, $[aX + bZ, Y]_t = a[X, Y]_t + b[Z, Y]_t$.
- (2) If (W_t) and (B_t) are two independent Wiener processes, then

$$[W, B]_t = 0. \quad (30)$$

Example

Suppose $Z_t = \rho W_t + (1 - \rho^2)^{1/2} B_t$, where W_t and B_t are independent Wiener processes.

One obtains from first principles that

$$\begin{aligned} [W, Z]_t &= \rho[W, W]_t + (1 - \rho^2)^{1/2}[W, B]_t \\ &= \rho t, \end{aligned} \quad (31)$$

since $[W, W]_t = t$ and $[W, B]_t = 0$.

Properties of Quadratic Variation (cont'd)

(3) For stochastic integrals over Itô processes X_t and Y_t ,

$$U_t = \int_0^t H_s dX_s \text{ and } V_t = \int_0^t K_s dY_s, \quad (32)$$

then

$$[U, V]_t = \int_0^t H_s K_s d[X, Y]_s. \quad (33)$$

This is often written on “differential form” as

$$d[U, V]_t = H_t K_t d[X, Y]_t. \quad (34)$$

by invoking the same results that led to (29). For a rigorous statement, see Property I.4.54 (p.55) of Jacod and Shiryaev (2003) (“JS”, from now on)

(4) For any Itô process X , $[X, t] = 0$.

Properties of Quadratic Variation (cont'd)

Example

(Leverage Effect in the Heston model). For the Heston model, recall

$$dX_t = \left(\mu - \frac{\sigma_t^2}{2} \right) dt + \sigma_t dW_t$$

$$d\sigma_t^2 = \kappa(\alpha - \sigma_t^2)dt + \gamma\sigma_t dZ_t, \text{ with} \quad (35)$$

$$Z_t = \rho W_t + (1 - \rho^2)^{1/2} B_t \quad (36)$$

We have seen above that $[W, Z]_t = \rho t$, and so

$$\begin{aligned} d[X, \sigma^2] &= \gamma\sigma_t^2 d[W, Z]_t \\ &= \gamma\sigma^2 \rho dt. \end{aligned} \quad (37)$$

Properties of Quadratic Variation (cont'd)

(5) Invariance under discounting by the short term interest rate. The typical discount rate is the risk free short term interest rate r_t . Recall that $S_t = \exp\{X_t\}$. The discounted stock price is then given by

$$S_t^* = \exp\left\{-\int_0^t r_s ds\right\} S_t. \quad (38)$$

The corresponding process on the log scale is $X_t^* = X_t - \int_0^t r_s ds$, so that if X_t is given by (8), then

$$dX_t^* = (\mu_t - r_t)dt + \sigma_t dW_t. \quad (39)$$

The quadratic variation of X_t^* is therefore the same as for X_t .

Properties of Quadratic Variation (cont'd)

While this result remains true for certain other types of discounting (such as cost-of-carry), it is not true for many other types of discounting. For example, if one discounts by the zero coupon bond Λ_t maturing at time T , the discounted log price becomes $X_t^* = X_t - \log \Lambda_t$. Since the zero coupon bond will itself have volatility, we get

$$[X^*, X^*]_t = [X, X]_t + [\log \Lambda, \log \Lambda]_t - 2[X, \log \Lambda]_t. \quad (40)$$

$$\text{Var}(X|\mathcal{A}) = E((X - E(X|\mathcal{A}))^2|\mathcal{A}) = E(X^2|\mathcal{A}) - E(X|\mathcal{A})^2. \quad (41)$$

and similarly

Theorem

$$\text{Var}(M_t | \mathcal{F}_S) = E((M_t - M_S)^2 | \mathcal{F}_S) = E([M, M]_t - [M, M]_S | \mathcal{F}_S).$$

Spot quantities for Continuous Itô Processes

- $\lim_{h \downarrow 0} \frac{1}{h} \text{Cov}(X_{t+h} - X_t, Y_{t+h} - Y_t | \mathcal{F}_t) = \frac{d}{dt}[X, Y]_t$
- Spot correlation:
 $\text{cor}(X, Y)_t = \lim_{h \downarrow 0} \text{cor}(X_{t+h} - X_t, Y_{t+h} - Y_t | \mathcal{F}_t)$
- Calculating spot correlation:

$$\text{cor}(X, Y)_t = \frac{d[X, Y]_t/dt}{\sqrt{(d[X, X]_t/dt)(d[Y, Y]_t/dt)}}$$
- In the Heston model,

$$\text{cor}(X, \sigma^2)_t = \rho. \quad (42)$$

- In general, if $dX_t = \sigma_t dW_t + dt\text{-term}$, and $dY_t = \gamma_t dB_t + dt\text{-term}$, where W_t and B_t are two Wiener processes, then

$$\text{cor}(X, Y)_t = \text{sgn}(\sigma_t \gamma_t) \text{cor}(W, B)_t. \quad (43)$$

Lévy's Theorem

Theorem

Suppose that M_t is a continuous (\mathcal{F}_t) -local martingale, $M_0 = 0$, so that $[M, M]_t = t$. Then M_t is an (\mathcal{F}_t) -Wiener process.

(Cf. Thm II.4.4 (p. 102) in JS.) More generally, from properties of normal random variables, the same result follows in the vector case: If $M_t = (M_t^{(1)}, \dots, M_t^{(p)})$ is a continuous (\mathcal{F}_t) -martingale, $M_0 = 0$, so that $[M^{(i)}, M^{(j)}]_t = \delta_{ij}t$, then M_t is a vector Wiener process. (δ_{ij} is the Kronecker delta: $\delta_{ij} = 1$ for $i = j$, and $= 0$ otherwise.)

For continuous Itô processes, $\langle X, Y \rangle_t = [X, Y]_t$. For general semimartingales this equality does not hold. Also, except for Itô processes, $\langle X, Y \rangle_t$ cannot generally be estimated consistently from data without further assumptions. For example, If N_t is a Poisson process with intensity λ , then $M_t = N_t - \lambda t$ is a martingale. In this case, $[M, M]_t = N_t$ (observable), while $\langle M, M \rangle_t = \lambda t$ (cannot be estimated in finite time, though see discussion in Stoltenberg, M and Z, 2020).

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Itô's Formula for Continuous Itô processes

Theorem

Suppose that f is a twice continuously differentiable function, and that X_t is an Itô process. Then

$$df(X_t) = f'(X_t)dX_t + \frac{1}{2}f''(X_t)d[X, X]_t. \quad (44)$$

Similarly, in the multivariate case, for $X_t = (X_t^{(1)}, \dots, X_t^{(p)})$,

$$df(X_t) = \sum_{i=1}^p \frac{\partial f}{\partial x^{(i)}}(X_t) dX_t^{(i)} + \frac{1}{2} \sum_{i,j=1}^p \frac{\partial^2 f}{\partial x^{(i)} \partial x^{(j)}}(X_t) d[X^{(i)}, X^{(j)}]_t. \quad (45)$$

(Reference: Theorem I.4.57 in JS.)

Itô's Formula for Continuous Itô processes (cont'd)

We emphasize that (44) is the same as saying that

$$f(X_t) = f(X_0) + \int_0^t f'(X_s) dX_s + \frac{1}{2} \int_0^t f''(X_s) d[X, X]_s. \quad (46)$$

If we write out $dX_t = \mu_t dt + \sigma_t dW_t$ and $d[X, X]_t = \sigma_t^2 dt$, then equation (44) becomes

$$\begin{aligned} df(X_t) &= f'(X_t)(\mu_t dt + \sigma_t dW_t) + \frac{1}{2} f''(X_t) \sigma_t^2 dt \\ &= (f'(X_t) \mu_t + \frac{1}{2} f''(X_t) \sigma_t^2) dt + f'(X_t) \sigma_t dW_t. \end{aligned} \quad (47)$$

We note, in particular, that if X_t is a continuous Itô process, then so is $f(X_t)$.

(Example of Itô's Formula: Stochastic Equation for a Stock Price.) We have so far discussed the model for a stock on the log scale, as $dX_t = \mu_t dt + \sigma_t dW_t$. The price is given as $S_t = \exp(X_t)$. Using Itô's formula, with $f(x) = \exp(x)$, we get

$$dS_t = S_t(\mu_t + \frac{1}{2}\sigma_t^2)dt + S_t\sigma_t dW_t. \quad (48)$$

(Example of Itô's Formula: Proof of Lévy's Theorem.) Take $f(x) = e^{ihx}$, and go on from there. Left for the students. (Hard. See proof in Karatzas and Shreve (1991).)

Example of Itô's Formula: Genesis of the Leverage Effect

A case where quadratic covariation between a process and its volatility can arise from basic economic principles. The following is the origin of the use of the word “leverage effect” to describe such covariation. This type of covariation can arise from many sources, and will later use the term leverage effect to describe the phenomenon broadly.

Suppose that the log value of a firm is Z_t , given as a GBM,

$$dZ_t = \nu dt + \gamma dW_t. \quad (49)$$

For simplicity, suppose that the interest rate is zero, and that the firm has borrowed C dollars (or euros, yuan, ...). If there are M shares in the company, the value of one share is therefore

$$S_t = (\exp(Z_t) - C)/M. \quad (50)$$

From the derivations on p. 132-134 in MZ (2012),

$$\text{cor}(X, \sigma^2)_t = -1. \quad (51)$$

Estimation of volatility. Central limit theorems and stable convergence.