Statistical Analysis of EEG-experiment data

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Contents

1 Introduction

```r
knitr::opts_chunk$set(echo = TRUE)
```

1.1 The data

(Fictitious data)

ERP experiment

- 20 Subjects,
- 6 Channels: O1, O2, PO7, PO8, P7, P8
- Stimuli: pictures. Conditions:
  - 1 (f): fear (face)
  - 2 (h): happiness (face)
  - 3 (d): disgust (face)
  - 4 (n): neutral (face)
  - 5 (o): object (face)
- Measure: Area around the component P170

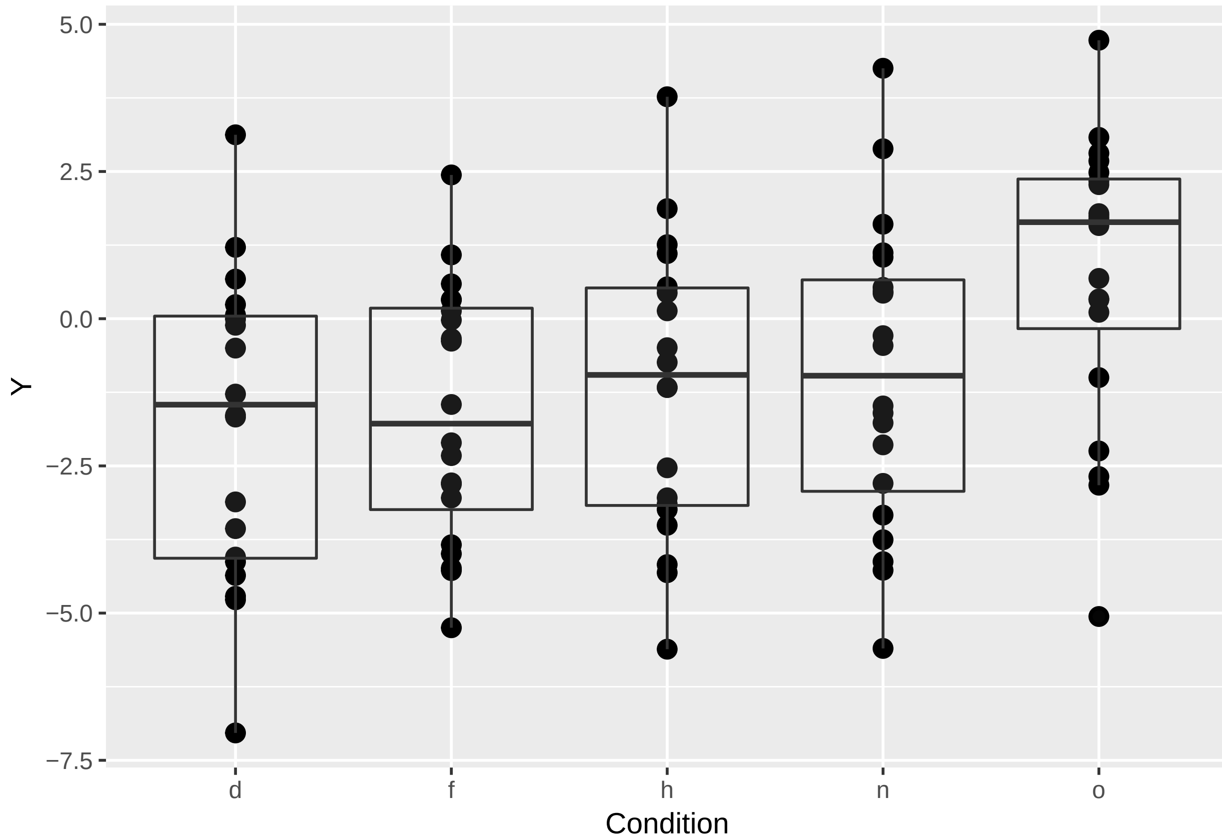
Setting parameters, importing the data:

```r
rm(list=ls())

load("./dataset/datiEEG.Rdata")

# dati2 is the same as dati, only selecting:
# ( Chan = "O1" ) & ( Condition = "f" or "n" )
load("./dataset/dati2EEG.Rdata")

# VERY IMPORTANT:
contrasts(dati$Chan) <- contr.sum(6)
contrasts(dati$Condition) <- contr.sum(5)
contrasts(dati$Subj) <- contr.sum(nlevels(dati$Subj))

contrasts(dati2$Condition) <- contr.sum(2)
contrasts(dati2$Subj) <- contr.sum(nlevels(dati2$Subj))
```

1.2 Motivation (EDA)

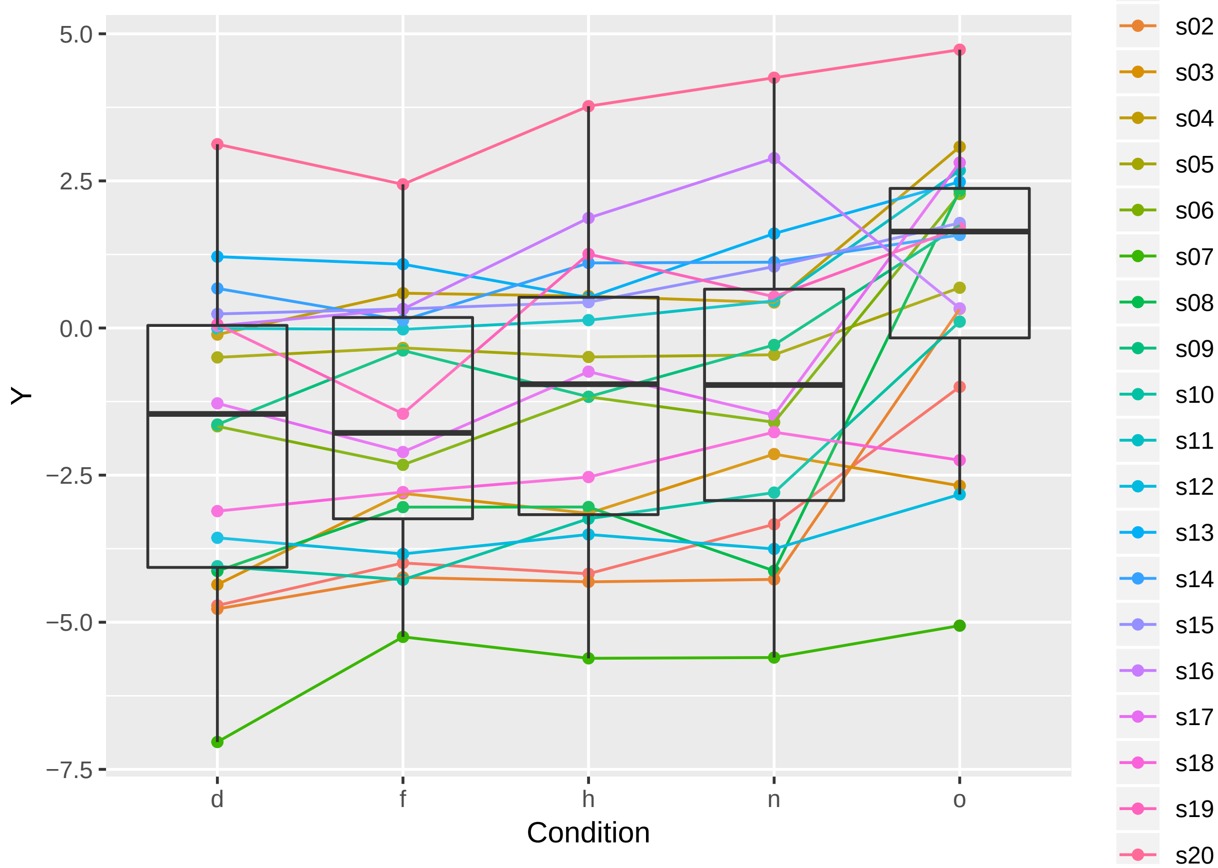
For Channel `O1`:

```r
library(ggplot2)
p <- ggplot(subset(dati,Chan=="O1"),aes(Condition,Y))
p+geom_point(size = 3) +geom_boxplot(alpha=.1)
```



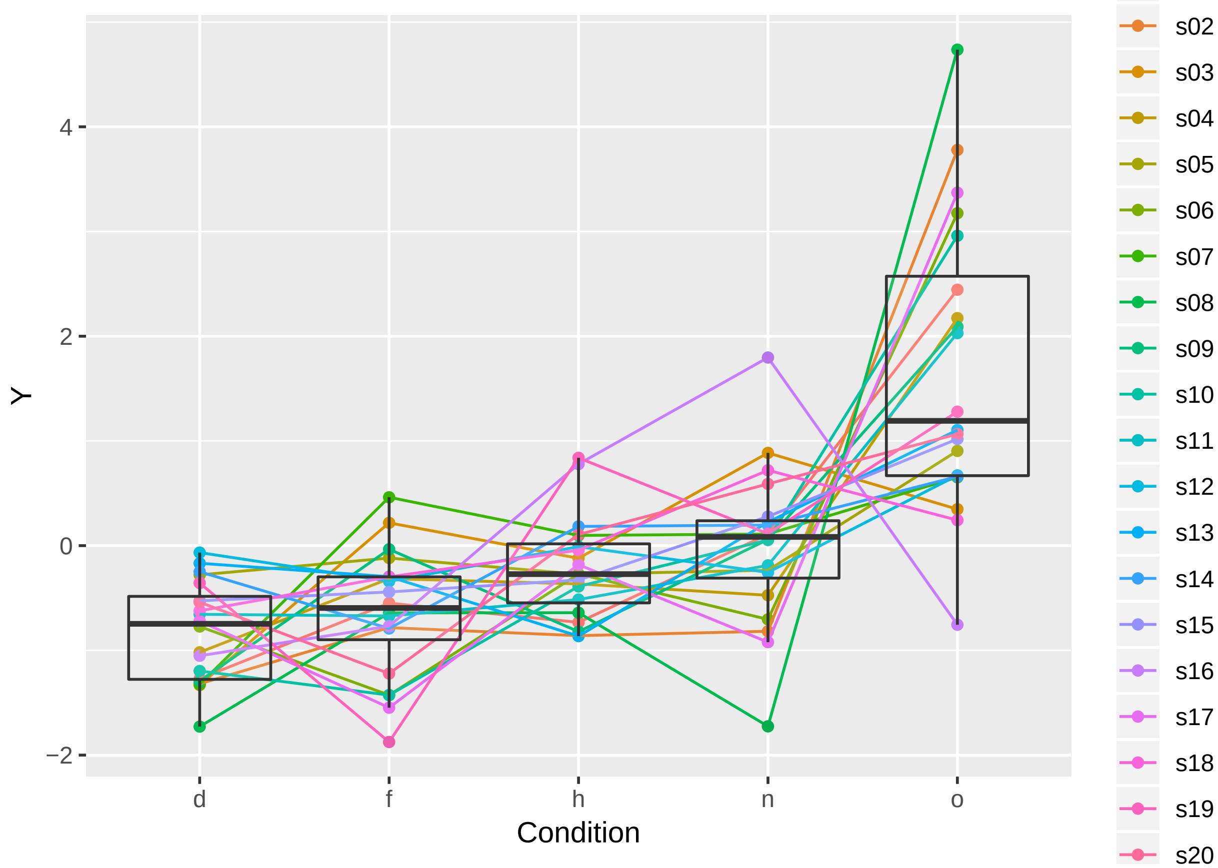
Is there a specificity of the subject?

```
dati01=subset(dati,Chan=="01")
library(ggplot2)
p <- ggplot(dati01,aes(Condition,Y))
p+geom_point(aes(group = Subj, colour = Subj))+
  geom_line(aes(group = Subj, colour = Subj))+
   geom_boxplot(alpha=.1)
```



We subtract the subject-specific effect (i.e. subject's mean) to each observation.

```
dati01=subset(dati,Chan=="O1")
Y=scale(matrix(dati01$Y,5),scale=FALSE)
dati01$Y=as.vector(Y)

library(ggplot2)
p <- ggplot(dati01,aes(Condition,Y))
p+geom_point(aes(group = Subj, colour = Subj))+
  geom_line(aes(group = Subj, colour = Subj))+
   geom_boxplot(alpha=.1)
```



The dispersion of the data has been largely reduced. This effect is the one taken in account by the models for repeated measures.

2 Repeated Mesures ANOVA

2.1 Introduction

wiki reference: https://en.wikipedia.org/wiki/Repeated_measures_design

A nice explanation can be found (in particular see 7.9 and 7.10):
Jonathan Baron (2011) Notes on the use of R for psychology experiments and questionnaires https://www.sas.upenn.edu/~baron/from_cattell/rpsych/rpsych.html

and in the Course materal of
ST 732, Applied Longitudinal Data Analysis, NC State University by Marie Davidian https://www.stat.ncsu.edu/people/davidian/courses/st732/notes/chap5.pdf from https://www.stat.ncsu.edu/people/davidian/courses/st732/

2.2 2 conditions, matched observations

Let consider the reduced problem: channel `Chan=="O1` and `Condition=="n"` or `Condition=="f"`.

How to compare the two conditions? First try:

```r
t.test(dati2$Y[dati2$Condition=="n"],
       dati2$Y[dati2$Condition=="f"])
```

```
##
##  Welch Two Sample t-test
##
## data:  dati2$Y[dati2$Condition == "n"] and dati2$Y[dati2$Condition == "f"]
## t = 0.8449, df = 36.861, p-value = 0.4036
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
##  -0.8868791  2.1552491
## sample estimates:
## mean of x mean of y
## -0.964530 -1.598715
```

Is it ok?

NO! We don't take in account the fact that measures are taken on the same subject!

```r
t.test(dati2$Y[dati2$Condition=="n"],
       dati2$Y[dati2$Condition=="f"],paired=TRUE)
```

```
##
##  Paired t-test
##
## data:  dati2$Y[dati2$Condition == "n"] and dati2$Y[dati2$Condition == "f"]
## t = 3.287, df = 19, p-value = 0.003877
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
##  0.2303616 1.0380084
## sample estimates:
## mean of the differences
##                0.634185
```

```r
## equivalent to
t.test(dati2$Y[dati2$Condition=="n"]-
         dati2$Y[dati2$Condition=="f"])
```

```
##
##  One Sample t-test
##
## data:  dati2$Y[dati2$Condition == "n"] - dati2$Y[dati2$Condition ==     "f"]
## t = 3.287, df = 19, p-value = 0.003877
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
##  0.2303616 1.0380084
## sample estimates:
## mean of x
##  0.634185
```

Can you write it as a linear model?

```r
mod2=lm(Y~ Condition+Subj,data=dati2)
anova(mod2)
```

```
## Analysis of Variance Table
##
## Response: Y
##           Df  Sum Sq Mean Sq F value    Pr(>F)
## Condition  1   4.022  4.0219  10.804  0.003877 **
## Subj      19 207.022 10.8959  29.270 3.118e-10 ***
## Residuals 19   7.073  0.3722
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Compare the results. (Different or the same?)

2.3 Linear models with repeated measures

Let's consider (and fit) a linear model with `Chan*Condition`:

```r
modlmf=lm(Y~ Chan*Condition,data=dati)
anova(modlmf)
```

```
## Analysis of Variance Table
##
## Response: Y
##                 Df Sum Sq Mean Sq F value Pr(>F)
## Chan             5  871.9 174.376 25.4499 <2e-16 ***
## Condition        4 1022.9 255.714 37.3209 <2e-16 ***
## Chan:Condition  20   66.6   3.328  0.4857 0.9719
## Residuals      570 3905.5   6.852
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

We don't take in account the fact that measures are taken on the same subject!

Can we just add the `Subj` term?

```r
modlmf=lm(Y~ Chan*Condition+Subj,data=dati)
anova(modlmf)
```

```
## Analysis of Variance Table
##
## Response: Y
##                 Df  Sum Sq Mean Sq F value Pr(>F)
## Chan             5  871.88 174.376 68.0714 <2e-16 ***
## Condition        4 1022.86 255.714 99.8233 <2e-16 ***
## Subj            19 2494.02 131.264 51.2418 <2e-16 ***
## Chan:Condition  20   66.56   3.328  1.2992 0.1724
## Residuals      551 1411.48   2.562
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Answer: yes and no.
The estimates are ok, but we need to take care of the residuals SS in the testing phase.

All the SS that we need can be found in the saturated linear model. We compute them now and we use them later.

```
modlmf=lm(Y~ Chan*Condition*Subj,data=dati)
anova(modlmf)
```

```
## Warning in anova.lm(modlmf): ANOVA F-tests on an essentially perfect fit
## are unreliable
```

```
## Analysis of Variance Table
##
## Response: Y
##                      Df  Sum Sq Mean Sq F value Pr(>F)
## Chan                  5  871.88 174.376
## Condition             4 1022.86 255.714
## Subj                 19 2494.02 131.264
## Chan:Condition       20   66.56   3.328
## Chan:Subj            95 1017.54  10.711
## Condition:Subj       76  246.95   3.249
## Chan:Condition:Subj 380  146.99   0.387
## Residuals             0    0.00
```

2.4 Repeated measures

```
# The standard way
mod=aov(Y~ Chan*Condition+Subj + Error(Subj/(Chan*Condition)),data=dati)
summary(mod)
```

```
##
## Error: Subj
##      Df Sum Sq Mean Sq
## Subj 19   2494   131.3
##
## Error: Subj:Chan
##           Df Sum Sq Mean Sq F value   Pr(>F)
## Chan       5  871.9  174.38   16.28 1.42e-11 ***
## Residuals 95 1017.5   10.71
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Error: Subj:Condition
##           Df Sum Sq Mean Sq F value Pr(>F)
## Condition  4 1022.9  255.71    78.7 <2e-16 ***
## Residuals 76  246.9    3.25
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Error: Subj:Chan:Condition
```

```
##                  Df Sum Sq Mean Sq F value Pr(>F)
## Chan:Condition   20  66.56   3.328   8.604 <2e-16 ***
## Residuals       380 146.99   0.387
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

A better output and slightly more compleate analysis (Sphericity Corrections):

```r
library(ez)
mod=ezANOVA(dv=Y, wid=Subj, within=.(Chan,Condition),data=dati,type=3)
```

```
## Warning in log(det(U)): Si è prodotto un NaN
```

```r
print(mod)
```

```
## $ANOVA
##           Effect DFn DFd         F            p p<.05        ges
## 2           Chan   5  95 16.280163 1.422895e-11     * 0.18250183
## 3      Condition   4  76 78.697466 2.998429e-26     * 0.20754506
## 4 Chan:Condition  20 380  8.604227 5.232560e-21     * 0.01675807
##
## $`Mauchly's Test for Sphericity`
##      Effect          W            p p<.05
## 2      Chan 0.03433646 3.910057e-07     *
## 3 Condition 0.06754172 4.802965e-07     *
##
## $`Sphericity Corrections`
##           Effect       GGe        p[GG] p[GG]<.05       HFe        p[HF]
## 2           Chan 0.4368229 3.441213e-06         * 0.4957490 9.287363e-07
## 3      Condition 0.4114825 4.226482e-12         * 0.4454085 6.399344e-13
## 4 Chan:Condition 0.1134660 4.452748e-04         * 0.1296611 2.121437e-04
##   p[HF]<.05
## 2         *
## 3         *
## 4         *
```

To see the relation between repeated measures and linear model, again, the Baron material is a good start. Specially see section "7.9.3 The Appropriate Error Terms"

2.5 Spend your DF in a different way!

Same number of DF, but spent in a different way

```r
dati$Lateral=dati$Chan
levels(dati$Lateral)
```

```
## [1] "O1"  "O2"  "P7"  "P8"  "PO7" "PO8"
```

```r
levels(dati$Lateral)[c(1,3,5)]="Left"
levels(dati$Lateral)[-1]="Right"
levels(dati$Lateral)
```

```
## [1] "Left"  "Right"
```

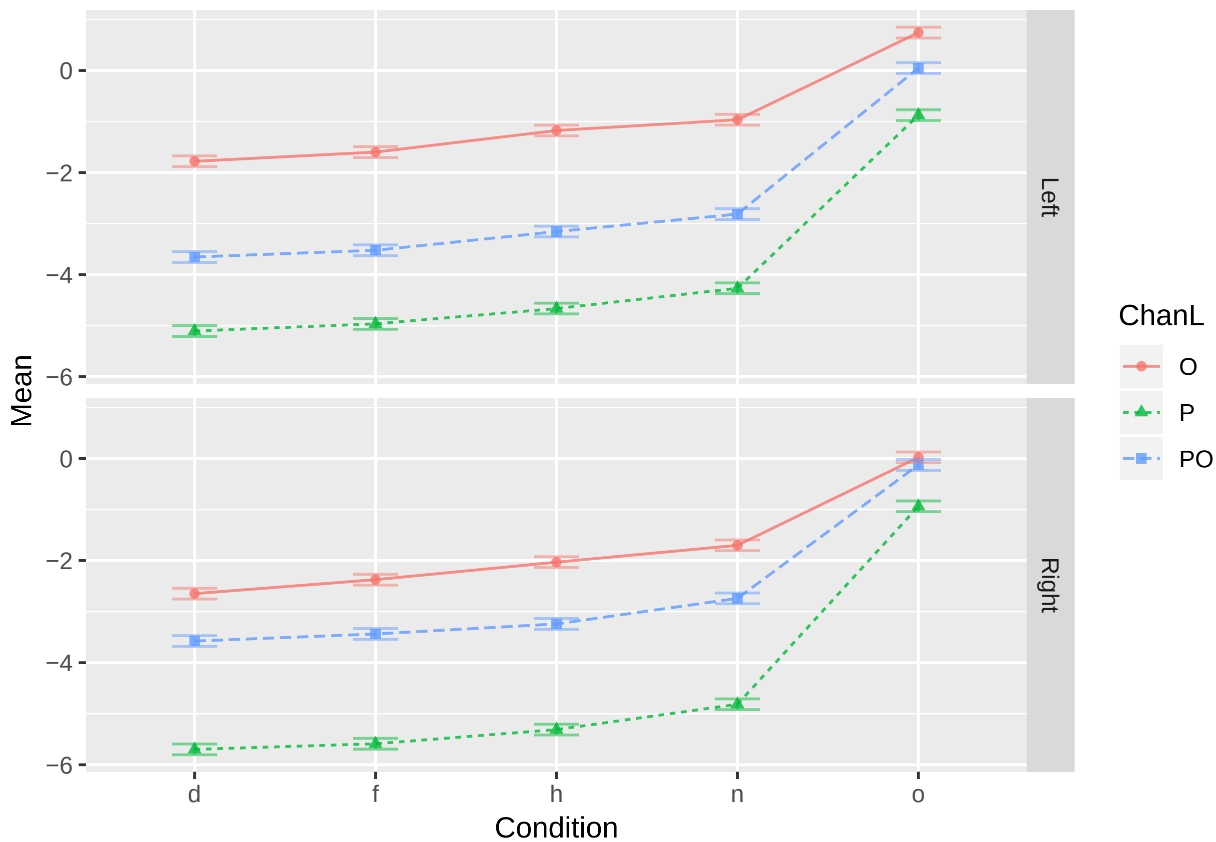
```r
contrasts(dati$Lateral) <- contr.sum(2)

dati$ChanL=dati$Chan
# https://en.wikipedia.org/wiki/Regular_expression
# Digits: \d
levels(dati$ChanL)=gsub("\\d","",levels(dati$ChanL))

contrasts(dati$ChanL) <- contr.sum(3)
```

```r
# The standard way
# mod=aov(Y~ ChanL*Lateral*Condition+Subj + Error(Subj/(ChanL*Lateral*Condition)),data=dati)
# summary(mod)
#

library(ez)
mod=ezANOVA(dv=Y, wid=Subj, within=.(Condition,Lateral,ChanL),data=dati,type=3)
print(mod)
```

```
## $ANOVA
##                    Effect DFn DFd          F            p p<.05
## 2               Condition   4  76 78.6974657 2.998429e-26     *
## 3                 Lateral   1  19  0.8340652 3.725436e-01
## 4                   ChanL   2  38 56.1729241 4.478690e-12     *
## 5       Condition:Lateral   4  76  0.1443064 9.649821e-01
## 6         Condition:ChanL   8 152 35.1954651 5.609013e-31     *
## 7           Lateral:ChanL   2  38  2.8073524 7.292266e-02
## 8 Condition:Lateral:ChanL   8 152  2.6449974 9.620701e-03     *
##            ges
## 2 0.2075450604
## 3 0.0070380089
## 4 0.1749978790
## 5 0.0001853800
## 6 0.0159782333
## 7 0.0040220980
## 8 0.0006202045
##
## $`Mauchly's Test for Sphericity`
##                    Effect            W            p p<.05
## 2               Condition 6.754172e-02 4.802965e-07     *
## 4                   ChanL 6.299144e-01 1.561471e-02     *
## 5       Condition:Lateral 9.051612e-03 1.045354e-13     *
## 6         Condition:ChanL 1.397323e-05 7.887294e-21     *
## 7           Lateral:ChanL 8.087935e-01 1.480945e-01
## 8 Condition:Lateral:ChanL 3.462436e-05 2.626994e-18     *
##
## $`Sphericity Corrections`
##                    Effect       GGe        p[GG] p[GG]<.05       HFe
## 2               Condition 0.4114825 4.226482e-12         * 0.4454085
## 4                   ChanL 0.7298814 2.181290e-09         * 0.7752249
## 5       Condition:Lateral 0.3235149 7.714124e-01           0.3371918
## 6         Condition:ChanL 0.2488273 2.427683e-09         * 0.2778788
```

```
## 7           Lateral:ChanL 0.8394850 8.359370e-02           0.9115913
## 8 Condition:Lateral:ChanL 0.2410174 8.634269e-02           0.2677123
##          p[HF] p[HF]<.05
## 2 6.399344e-13         *
## 4 7.702179e-10         *
## 5 7.811776e-01
## 6 3.463383e-10         *
## 7 7.863060e-02
## 8 7.969075e-02
```

```
ezPlot(dv=Y, wid=Subj, within=.(ChanL,Lateral,Condition),data=dati,
       x=Condition,split=ChanL,row=Lateral)
```



2.6 Sphericity

The sphericity assumption is an assumption about the structure of the covariance matrix in a repeated measures design. Before we describe it lets consider a simpler (but stricter) condition.

Compound symmetry

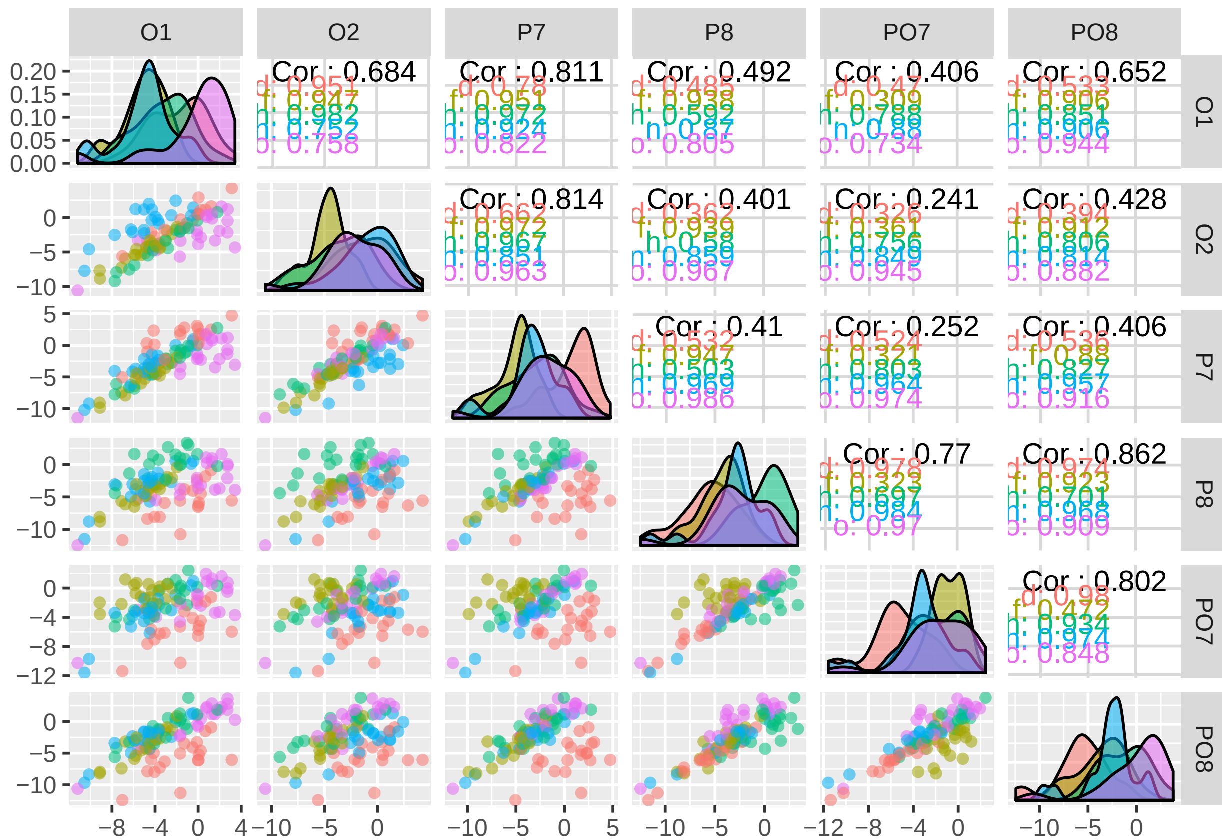
Compound symmetry is met if all the covariances (the off-diagonal elements of the covariance matrix) are equal and all the variances are equal in the populations being sampled. Provided the observed covariances are roughly equal in our samples (and the variances are OK too) we can be pretty confident that compound symmetry is not violated.

If compound symmetry is met, then sphericity is also met.

compound symmetry is met when the correlation between `Condition f` and `Condition h` is equal to the correlation between `Condition f` and `Condition o` or `Condition h` and `Condition n`, etc (same for any other factor within subject, such as `Chan`). But a more direct way to think about compound symmetry is to say that it requires that all subjects in each group change in the same way over trials. In other words the slopes of the lines regressing the dependent variable on time are the same for all subjects. Put that way it is easy to see that compound symmetry can really be an unrealistic assumption.

Is compound symmetry met in our data?

```r
# install.packages("GGally")
library(GGally)
Y=matrix(dati$Y,byrow = TRUE,nrow = 20*5)
Y=data.frame(Y)
names(Y)=levels(dati$Chan)
ggpairs(Y,aes(colour = dati$Condition[1:100], alpha = 0.4))
```



Not really! (correlations often differ)

Sphericity
The sphericity assumption is that the all the variances of the differences are equal (in the population sampled). In practice, we'd expect the observed sample variances of the differences to be similar if the sphericity assumption was met.

We can check sphericity assumption using the covariance matrix, but it turns out to be fairly laborious. Remember that variance of differences can be computed as:

$$S_x^2 - S_y^2 = S_x^2 + S_y^2 - 2S_{xy}$$

Further reading: http://homepages.gold.ac.uk/aphome/spheric.html

This is often an unrealistic assumption in EEG data (spatial location of channel relates to correlation between measures)

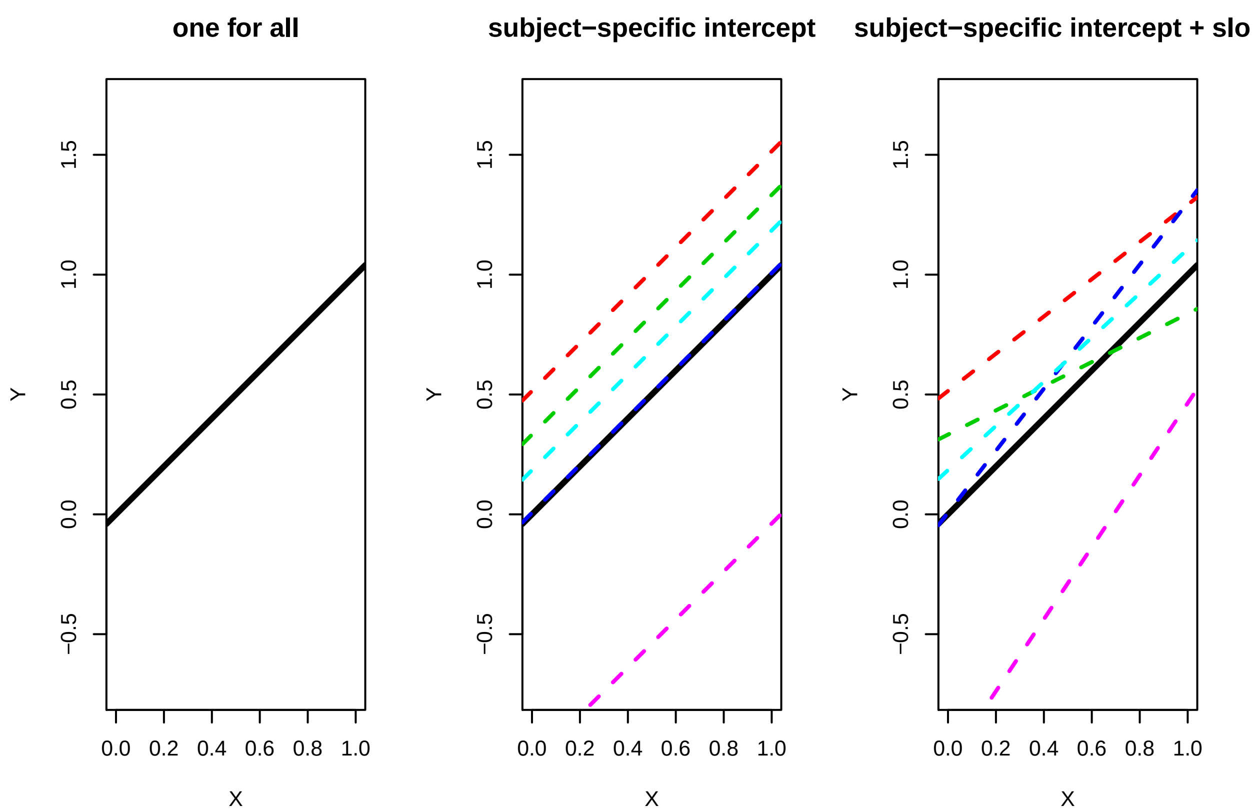
2.7 (Further) Limitations

- (Design and) Data must be balanced
- Repeated Measures Anova doesn't allow for missing data (e.g. subjects/condiction/channel cells)
- It only handle factors, no quantitative variables

Mixed model is a more flexible approach.

3 Mixed models

3.1 Motivation/Introduction



Mixed models allow for more flexible modelization.

I assume you are expert on mixed models, if not https://en.wikipedia.org/wiki/Mixed_model
and much more on: http://webcom.upmf-grenoble.fr/LIP/Perso/DMuller/M2R/R_et_Mixed/documents/Bates-book.pdf
and
https://cran.r-project.org/web/packages/lme4/vignettes/lmer.pdf

3.2 The model

Models with random effects can be defined as:

$$Y_{n\times 1} = X_{n\times p}B_{p\times 1} + Z_{n\times q}b_{q\times 1} + \varepsilon_{n\times 1}$$

where

$$\varepsilon \sim \mathcal{N}(0, \sigma^2 I_n)$$

In the models we will consider, the random effects are modeled as a multivariate normal random variable:

$$b \sim \mathcal{N}(0, \Sigma_{q\times q}),$$

In a *linear mixed model* the conditional distribution $(Y|\mathcal{B} = b)$ is a *spherical* multivariate Gaussian.

In our case $n = \#Subjects \times \#Conditions \times \#Channels = 20 \times 5 \times 6 = 600$. X is the matrix of (dummified) predictors. Z can take many dimensions and values. Examples follow.

Random effect for Subject (Random Intercept)

Z is the matrix of dummy variables of the column `dati$Subj`.

```
library(lmerTest)
# library(lme4)
# contrasts(dati$Lateral)=contr.sum
# contrasts(dati$ChanL)=contr.sum
# contrasts(dati$Condition)=contr.sum
mod=lmer(Y~ Condition*Lateral*ChanL +(1|Subj),data=dati)

car::Anova(mod)
```

```
## Analysis of Deviance Table (Type II Wald chisquare tests)
##
## Response: Y
##                            Chisq Df Pr(>Chisq)
## Condition               399.2932  4  < 2.2e-16 ***
## Lateral                  10.8062  1   0.001012 **
## ChanL                   323.3939  2  < 2.2e-16 ***
## Condition:Lateral         0.2827  4   0.990905
## Condition:ChanL          24.7559  8   0.001710 **
## Lateral:ChanL             6.1568  2   0.046032 *
## Condition:Lateral:ChanL   0.9461  8   0.998566
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Random effect for Subject and Channel

Actually, instead of `Channel`, we use the combination of `ChanL*Lateral`. Same prediction ability (6 channels in Channel and 3X2 combination of ChanL and Lateral), just a different point of view.

Z is the matrix of dummy variables of the column `dati$Subj`.

```
mod2=lmer(Y~ 0+Lateral*ChanL*Condition +(0+Lateral*ChanL|Subj),data=dati)
summary(mod2)
```

```
## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula: Y ~ 0 + Lateral * ChanL * Condition + (0 + Lateral * ChanL |
##     Subj)
##    Data: dati
##
## REML criterion at convergence: 1935.8
##
## Scaled residuals:
##     Min      1Q  Median      3Q     Max
## -4.1899 -0.4579 -0.0040  0.5130  3.5464
##
## Random effects:
##  Groups   Name            Variance Std.Dev. Corr
##  Subj     LateralLeft     3.8560   1.9637
##           LateralRight    6.9945   2.6447    0.63
##           ChanL1          0.6787   0.8238    0.20  0.24
##           ChanL2          0.4654   0.6822   -0.47 -0.45 -0.87
##           Lateral1:ChanL1 0.1853   0.4305   -0.07  0.18 -0.18  0.02
##           Lateral1:ChanL2 0.1642   0.4052   -0.28 -0.14  0.29 -0.16 -0.82
##  Residual                 0.8518   0.9229
## Number of obs: 600, groups:  Subj, 20
##
## Fixed effects:
##                             Estimate Std. Error         df t value
## LateralLeft                -2.516722   0.442312  19.000071  -5.690
## LateralRight               -2.946310   0.593772  18.999986  -4.962
## ChanL1                      1.381018   0.191768  19.000039   7.201
## ChanL2                     -1.490923   0.161576  19.025001  -9.227
## Condition1                 -1.011749   0.075355 475.000008 -13.426
## Condition2                 -0.849248   0.075355 475.000008 -11.270
## Condition3                 -0.531396   0.075355 475.000008  -7.052
## Condition4                 -0.151912   0.075355 475.000008  -2.016
## Lateral1:ChanL1             0.180370   0.110021  19.150923   1.639
## Lateral1:ChanL2             0.032409   0.105109  19.239316   0.308
## Lateral1:Condition1         0.015606   0.075355 475.000008   0.207
## Lateral1:Condition2         0.004035   0.075355 475.000008   0.054
## Lateral1:Condition3         0.049614   0.075355 475.000008   0.658
## Lateral1:Condition4        -0.013239   0.075355 475.000008  -0.176
## ChanL1:Condition1           0.149200   0.106568 475.000008   1.400
## ChanL2:Condition1          -0.167232   0.106568 475.000008  -1.569
## ChanL1:Condition2           0.213801   0.106568 475.000008   2.006
## ChanL2:Condition2          -0.204413   0.106568 475.000008  -1.918
## ChanL1:Condition3           0.278675   0.106568 475.000008   2.615
## ChanL2:Condition3          -0.233478   0.106568 475.000008  -2.191
## ChanL1:Condition4           0.169969   0.106568 475.000008   1.595
## ChanL2:Condition4          -0.166870   0.106568 475.000008  -1.566
## Lateral1:ChanL1:Condition1   0.022442   0.106568 475.000008   0.211
## Lateral1:ChanL2:Condition1   0.034156   0.106568 475.000008   0.321
## Lateral1:ChanL1:Condition2  -0.011970   0.106568 475.000008  -0.112
## Lateral1:ChanL2:Condition2   0.060872   0.106568 475.000008   0.571
## Lateral1:ChanL1:Condition3  -0.017273   0.106568 475.000008  -0.162
## Lateral1:ChanL2:Condition3   0.026281   0.106568 475.000008   0.247
## Lateral1:ChanL1:Condition4  -0.014015   0.106568 475.000008  -0.132
```

```
## Lateral1:ChanL2:Condition4   0.039652   0.106568 475.000008   0.372
##                              Pr(>|t|)
## LateralLeft                  1.74e-05 ***
## LateralRight                 8.65e-05 ***
## ChanL1                       7.70e-07 ***
## ChanL2                       1.87e-08 ***
## Condition1                    < 2e-16 ***
## Condition2                    < 2e-16 ***
## Condition3                   6.26e-12 ***
## Condition4                    0.04437 *
## Lateral1:ChanL1               0.11745
## Lateral1:ChanL2               0.76114
## Lateral1:Condition1           0.83602
## Lateral1:Condition2           0.95731
## Lateral1:Condition3           0.51060
## Lateral1:Condition4           0.86061
## ChanL1:Condition1             0.16216
## ChanL2:Condition1             0.11725
## ChanL1:Condition2             0.04540 *
## ChanL2:Condition2             0.05569 .
## ChanL1:Condition3             0.00921 **
## ChanL2:Condition3             0.02895 *
## ChanL1:Condition4             0.11139
## ChanL2:Condition4             0.11805
## Lateral1:ChanL1:Condition1    0.83330
## Lateral1:ChanL2:Condition1    0.74872
## Lateral1:ChanL1:Condition2    0.91062
## Lateral1:ChanL2:Condition2    0.56813
## Lateral1:ChanL1:Condition3    0.87131
## Lateral1:ChanL2:Condition3    0.80531
## Lateral1:ChanL1:Condition4    0.89543
## Lateral1:ChanL2:Condition4    0.71000
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
##
## Correlation matrix not shown by default, as p = 30 > 12.
## Use print(x, correlation=TRUE)  or
##     vcov(x)        if you need it
```

```
#alternative:
# mod=lmer(Y~ Condition*Lateral*ChanL +(1+ChanL|Subj),data=dati)
#   +(1+Chan|Subj) (equivalent to +(Chan|Subj)) is a linear combination of +(0+Chan|Subj). I better lik

car::Anova(mod2)
```

```
## Analysis of Deviance Table (Type II Wald chisquare tests)
##
## Response: Y
##                              Chisq Df Pr(>Chisq)
## Lateral                    31.7515  2  1.274e-07 ***
## ChanL                      81.2666  2  < 2.2e-16 ***
## Condition                1200.8753  4  < 2.2e-16 ***
```

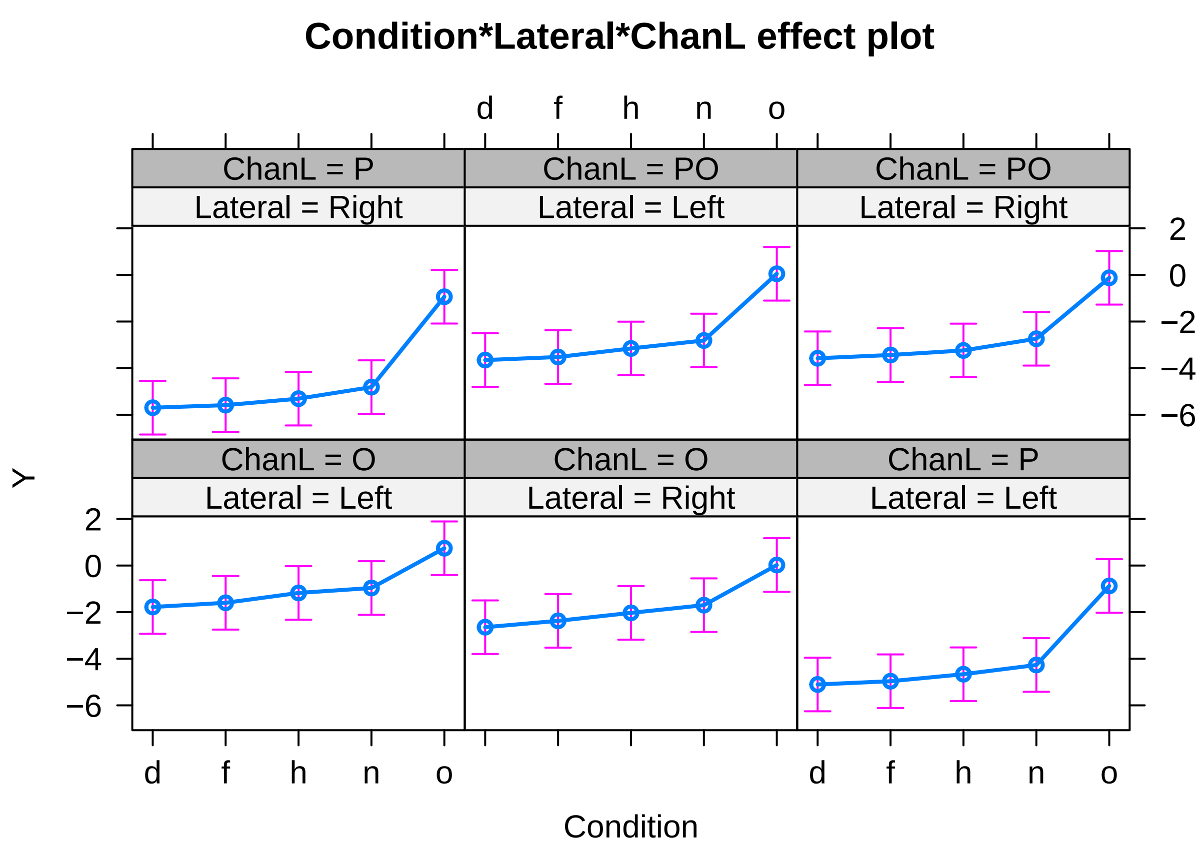
```
## Lateral:ChanL                7.7561  2    0.02069 *
## Lateral:Condition            0.8502  4    0.93160
## ChanL:Condition             74.4533  8  6.346e-13 ***
## Lateral:ChanL:Condition      2.8455  8    0.94367
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
# More flexible, but hard to fit:
# mod3=lmer(Y~ Condition*ChanL+Lateral +(0+Lateral|Subj)+(0+Condition|Subj)+(0+ChanL|Subj),data=dati)
#
```

3.3 Plotting tools

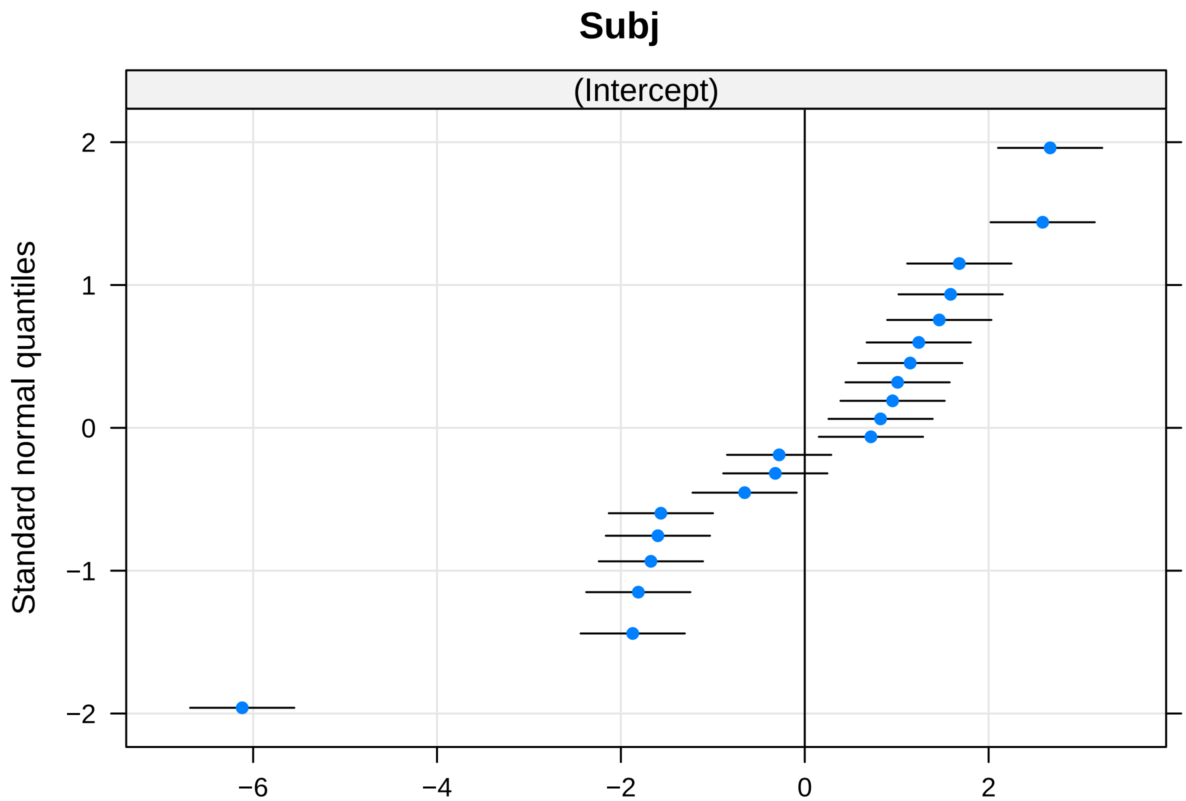
for the first model:

```
library(effects)
plot(allEffects(mod))
```



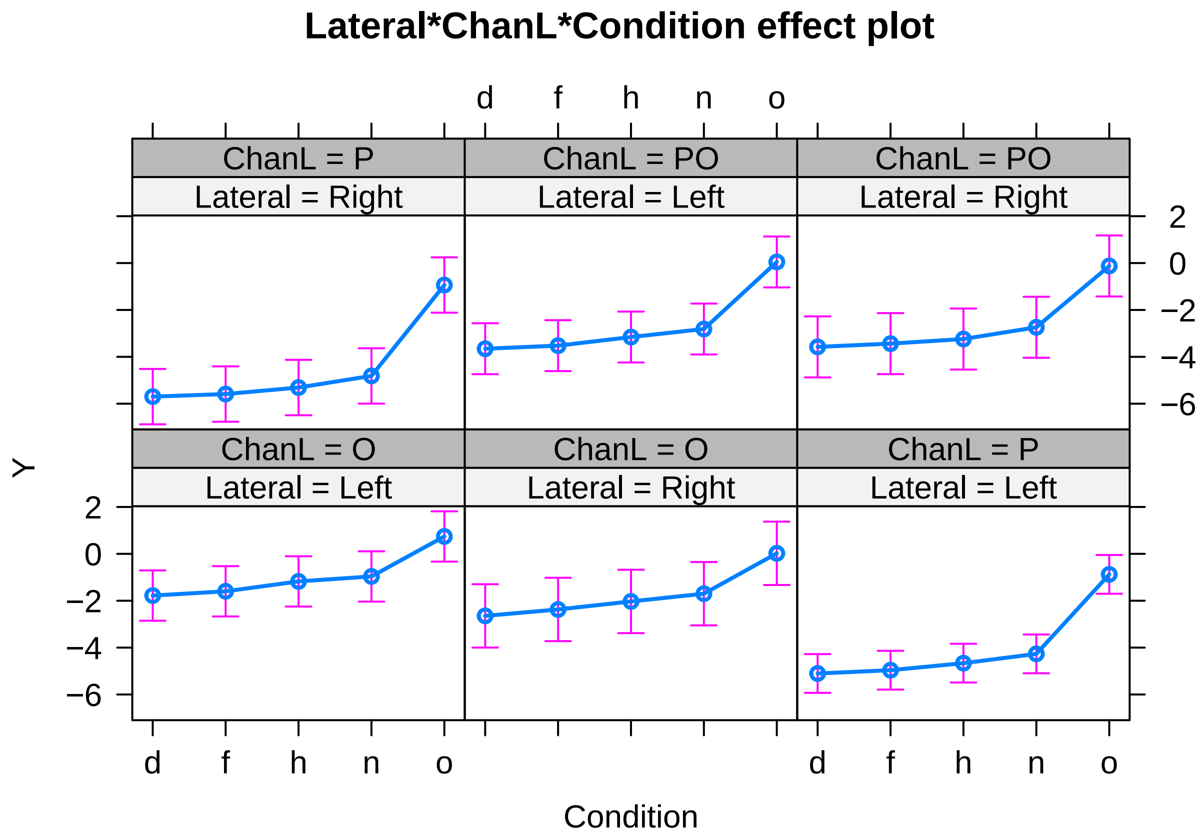

```
#plot random effects:
require(lattice)
qqmath(ranef(mod, condVar=TRUE))
```

```
## $Subj
```



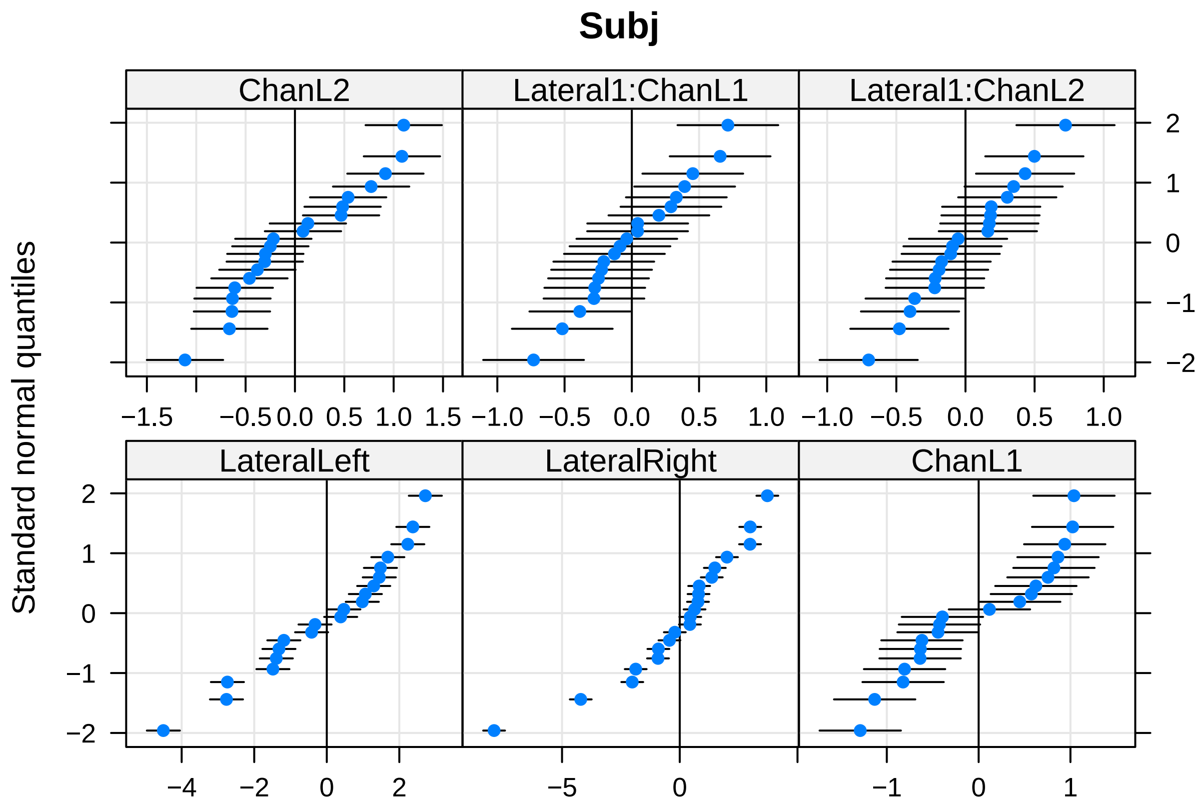
The second model:

```
library(effects)
plot(allEffects(mod2))
```

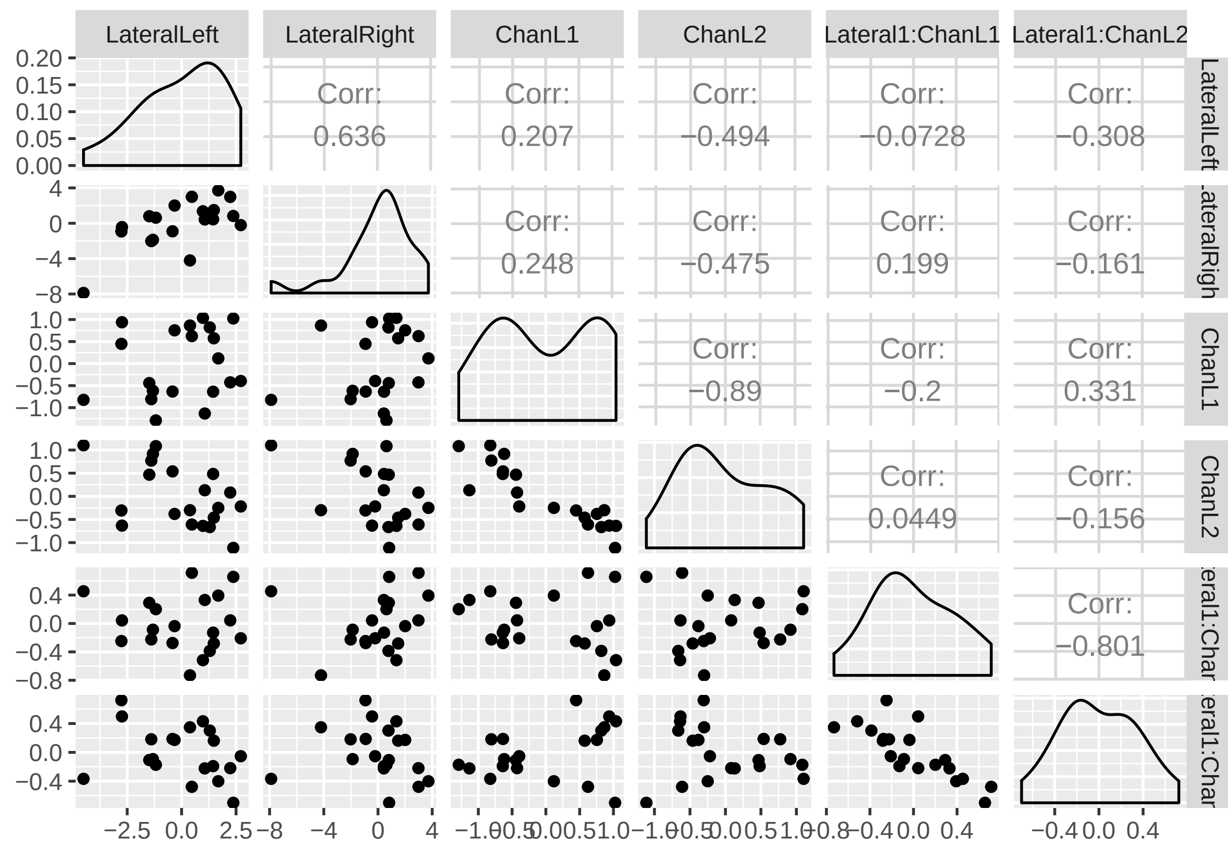



```r
#plot random effects:
require(lattice)
qqmath(ranef(mod2, condVar=TRUE))
```

```
## $Subj
```

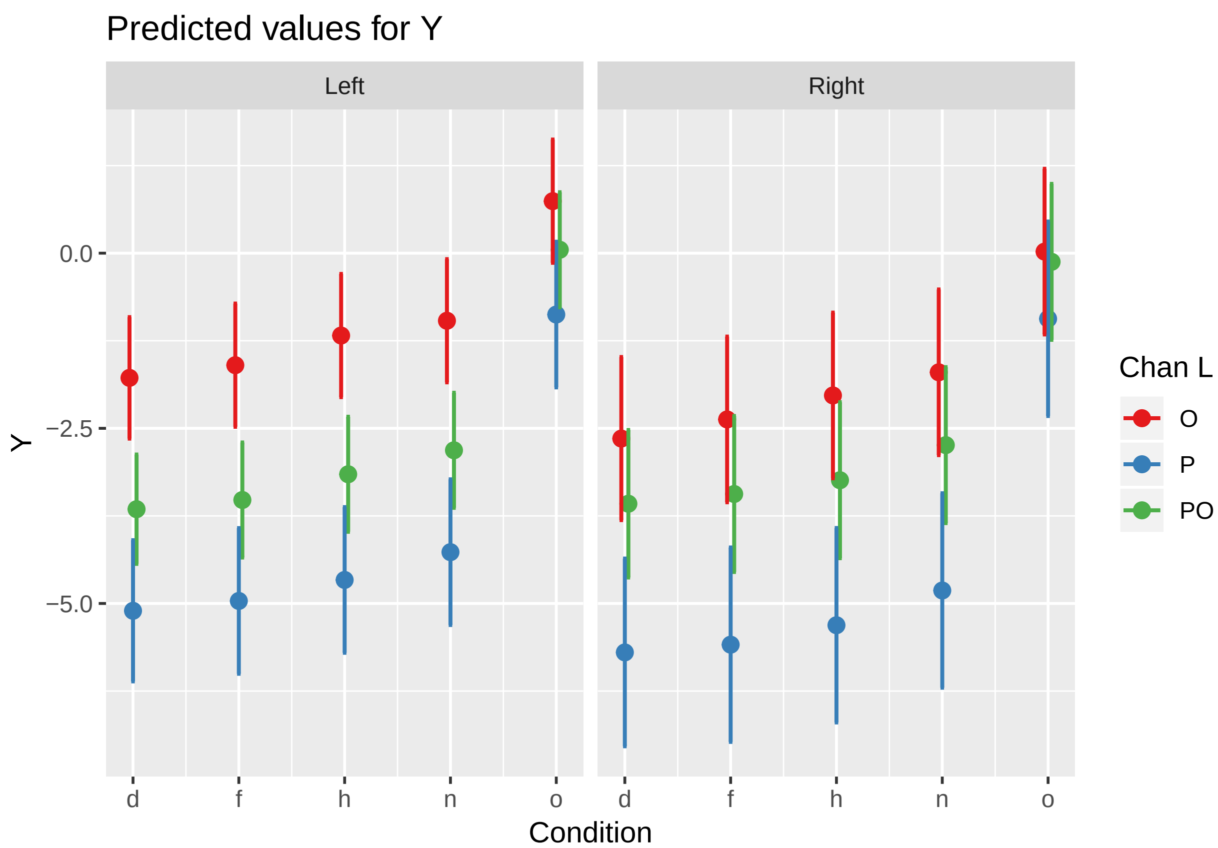



```r
# scatter plot
ggpairs(ranef(mod2, condVar=TRUE)[[1]])
```



An alternative plotting tool:

```
library(sjPlot)
library(ggplot2)
plot_model(mod2, type = "pred", terms = c("Condition", "ChanL","Lateral"))
```



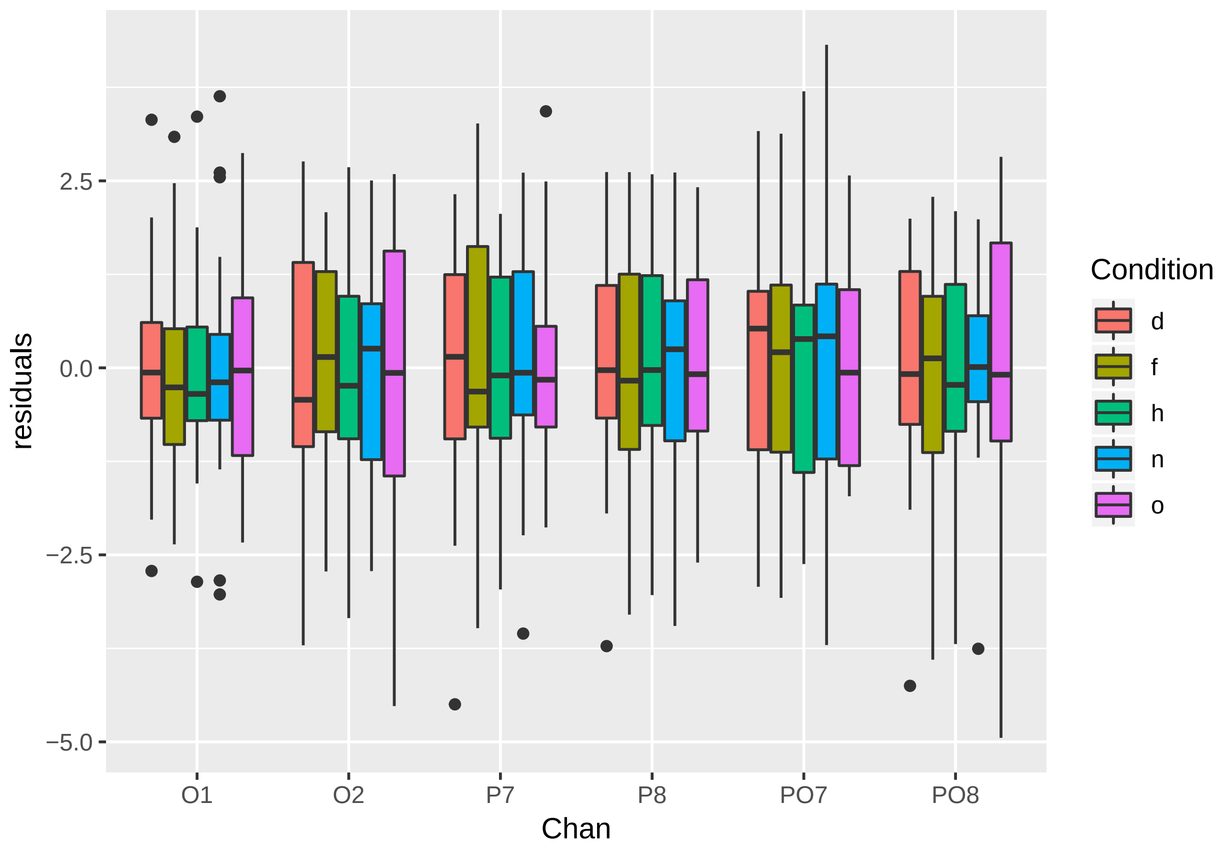
3.4 Validity of the assumptions

- Independence of the residuals?
- Normality of the residuals?
- Homoscedasticity of the residuals (i.e. same variance between subject/channel/condition?
- outliers?
- Leaverage? (influential observations)

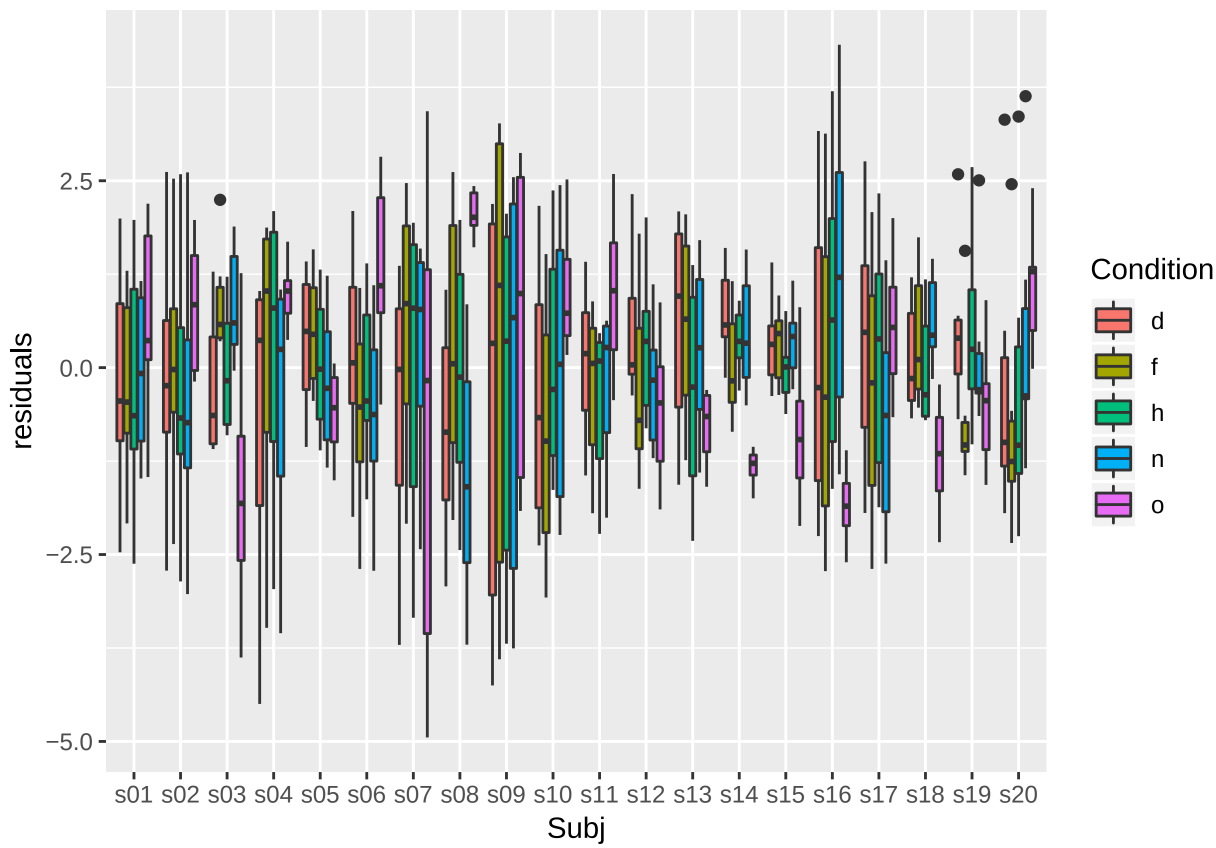
Please, do not test for normality, for homoscedasticity, sphericity etc.

Use Exploratory data Analysis, instead!

```
dati$residuals=residuals(mod)
p <- ggplot(dati, aes(x=Chan, y=residuals,fill=Condition)) +   geom_boxplot()
p
```




```
p <- ggplot(dati, aes(x=Subj, y=residuals,fill=Condition)) +   geom_boxplot()
p
```



3.5 Contrasts and post-hoc

3.5.1 Post-hoc

```
library(multcomp)
```

```
## Loading required package: mvtnorm

## Loading required package: survival

## Loading required package: TH.data

## Loading required package: MASS

##
## Attaching package: 'TH.data'

## The following object is masked from 'package:MASS':
##
##     geyser
```

```r
summary(glht(mod2, linfct = mcp(Condition = "Tukey")))
```

```
## Warning in mcp2matrix(model, linfct = linfct): covariate interactions found
## -- default contrast might be inappropriate
```

```
##
##   Simultaneous Tests for General Linear Hypotheses
##
## Multiple Comparisons of Means: Tukey Contrasts
##
##
## Fit: lmer(formula = Y ~ 0 + Lateral * ChanL * Condition + (0 + Lateral *
##     ChanL | Subj), data = dati)
##
## Linear Hypotheses:
##            Estimate Std. Error z value Pr(>|z|)
## f - d == 0   0.1625     0.1191   1.364   0.6509
## h - d == 0   0.4804     0.1191   4.032   <0.001 ***
## n - d == 0   0.8598     0.1191   7.217   <0.001 ***
## o - d == 0   3.5561     0.1191  29.846   <0.001 ***
## h - f == 0   0.3179     0.1191   2.668   0.0589 .
## n - f == 0   0.6973     0.1191   5.853   <0.001 ***
## o - f == 0   3.3936     0.1191  28.482   <0.001 ***
## n - h == 0   0.3795     0.1191   3.185   0.0125 *
## o - h == 0   3.0757     0.1191  25.814   <0.001 ***
## o - n == 0   2.6962     0.1191  22.629   <0.001 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## (Adjusted p values reported -- single-step method)
```

3.5.2 Custom contrasts

An example:

- neutral vs object in O1 (left)
- disgust vs neutral in O1 (left)
- fear vs neutral in O1 (left)
- happy vs neutral in O1 (left)

```r
library(multcomp)
ncoeff=length(coefficients(mod2)[[1]])
contr <- rbind("n - o" = c(0,0,0,0,0,0,1,-1, rep(0,ncoeff-8)),
               "d - n" = c(0,0,0,0,0,0,-1, rep(0,ncoeff-7)),
                 "f - n" = c(0,0,0,0,1,0,-1, rep(0,ncoeff-7)),
                 "h - n" = c(0,0,0,0,0,1,-1, rep(0,ncoeff-7)))
compa= glht(mod2, linfct = contr)
summary(compa, test = adjusted("none"))
```

```
##
##   Simultaneous Tests for General Linear Hypotheses
##
```

```
## Fit: lmer(formula = Y ~ 0 + Lateral * ChanL * Condition + (0 + Lateral *
##     ChanL | Subj), data = dati)
##
## Linear Hypotheses:
##            Estimate Std. Error z value Pr(>|z|)
## n - o == 0 -0.37948    0.11915  -3.185  0.00145 **
## d - n == 0  0.53140    0.07536   7.052 1.77e-12 ***
## f - n == 0 -0.48035    0.11915  -4.032 5.54e-05 ***
## h - n == 0 -0.31785    0.11915  -2.668  0.00764 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## (Adjusted p values reported -- none method)
```

```
# with multiple comparisons
summary(compa)
```

```
##
##   Simultaneous Tests for General Linear Hypotheses
##
## Fit: lmer(formula = Y ~ 0 + Lateral * ChanL * Condition + (0 + Lateral *
##     ChanL | Subj), data = dati)
##
## Linear Hypotheses:
##            Estimate Std. Error z value Pr(>|z|)
## n - o == 0 -0.37948    0.11915  -3.185  0.00468 **
## d - n == 0  0.53140    0.07536   7.052  < 0.001 ***
## f - n == 0 -0.48035    0.11915  -4.032  < 0.001 ***
## h - n == 0 -0.31785    0.11915  -2.668  0.02342 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## (Adjusted p values reported -- single-step method)
```

4 Multivariate ANOVA (MANOVA)

4.1 Motivation

ehi, wait a moment… the trials for `object` condition are much more than any other condition, the variance of its estimated component must be (much?) lower, homoschedastcity doesn't hold!!

Let's use a different approach Reshape the data from `long` to `wide` format.

to simplify the example, let's consider the comparison between conditions `neutral` vs `object`.

4.2 Reshaping the data

Let' now compute the vectors of contrasts (one vector of reach channel, length equal to number of subjects): `Fear vs Neutral`

```
Y=matrix(dati$Y,byrow = TRUE,nrow = 20)
colnames(Y)=paste(dati$Condition,dati$ChanL,dati$Lateral,sep = "_")[1:30]
```

```r
colnames(Y)
```

```
##  [1] "f_P_Left"   "h_P_Left"   "d_P_Left"   "n_P_Left"   "o_P_Left"
##  [6] "f_PO_Left"  "h_PO_Left"  "d_PO_Left"  "n_PO_Left"  "o_PO_Left"
## [11] "f_O_Left"   "h_O_Left"   "d_O_Left"   "n_O_Left"   "o_O_Left"
## [16] "f_P_Right"  "h_P_Right"  "d_P_Right"  "n_P_Right"  "o_P_Right"
## [21] "f_PO_Right" "h_PO_Right" "d_PO_Right" "n_PO_Right" "o_PO_Right"
## [26] "f_O_Right"  "h_O_Right"  "d_O_Right"  "n_O_Right"  "o_O_Right"
```

```r
contr=matrix(0,30,6)
contr[c(1,4),1]=c(1,-1)
contr[c(1,4)+5,2]=c(1,-1)
contr[c(1,4)+10,3]=c(1,-1)
contr[c(1,4)+15,4]=c(1,-1)
contr[c(1,4)+20,5]=c(1,-1)
contr[c(1,4)+25,6]=c(1,-1)

dim(contr)
```

```
## [1] 30  6
```

```r
head(contr)
```

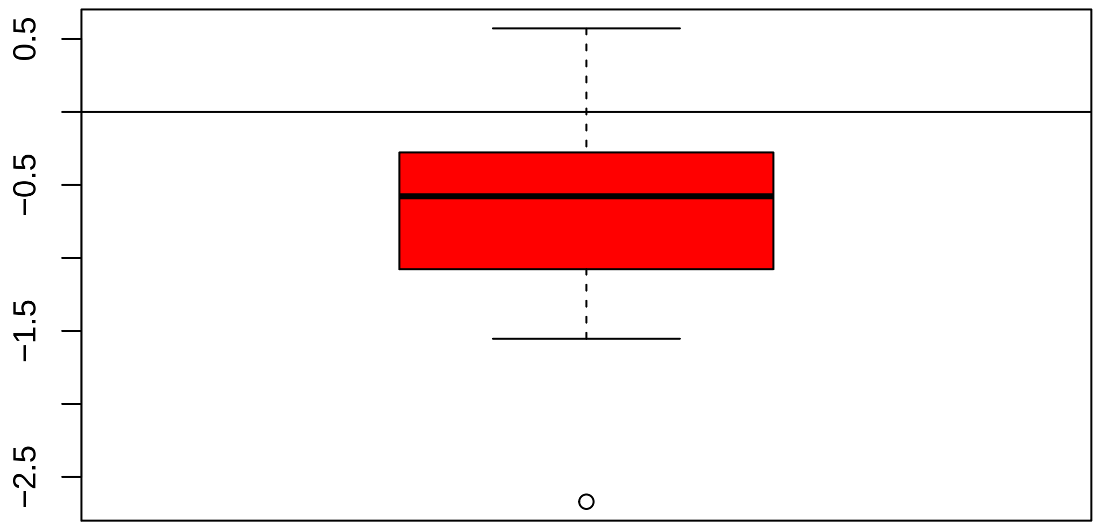
```
##      [,1] [,2] [,3] [,4] [,5] [,6]
## [1,]    1    0    0    0    0    0
## [2,]    0    0    0    0    0    0
## [3,]    0    0    0    0    0    0
## [4,]   -1    0    0    0    0    0
## [5,]    0    0    0    0    0    0
## [6,]    0    1    0    0    0    0
```

```r
Yfn=Y%*%contr
colnames(Yfn)= levels(dati$Chan)
dim(Yfn)
```

```
## [1] 20  6
```

What we see in O1?

```r
boxplot(Yfn[,1],col=2)
abline(0,0)
```



Same test as above, but under a different model

```r
t.test(Yfn[,1])
```

```
##
##  One Sample t-test
##
## data:  Yfn[, 1]
## t = -4.3327, df = 19, p-value = 0.0003587
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
##  -1.0327953 -0.3599747
## sample estimates:
## mean of x
## -0.696385
```

We can run the analysis over all channels

```r
(uni_t=apply(Yfn,2,t.test))
```

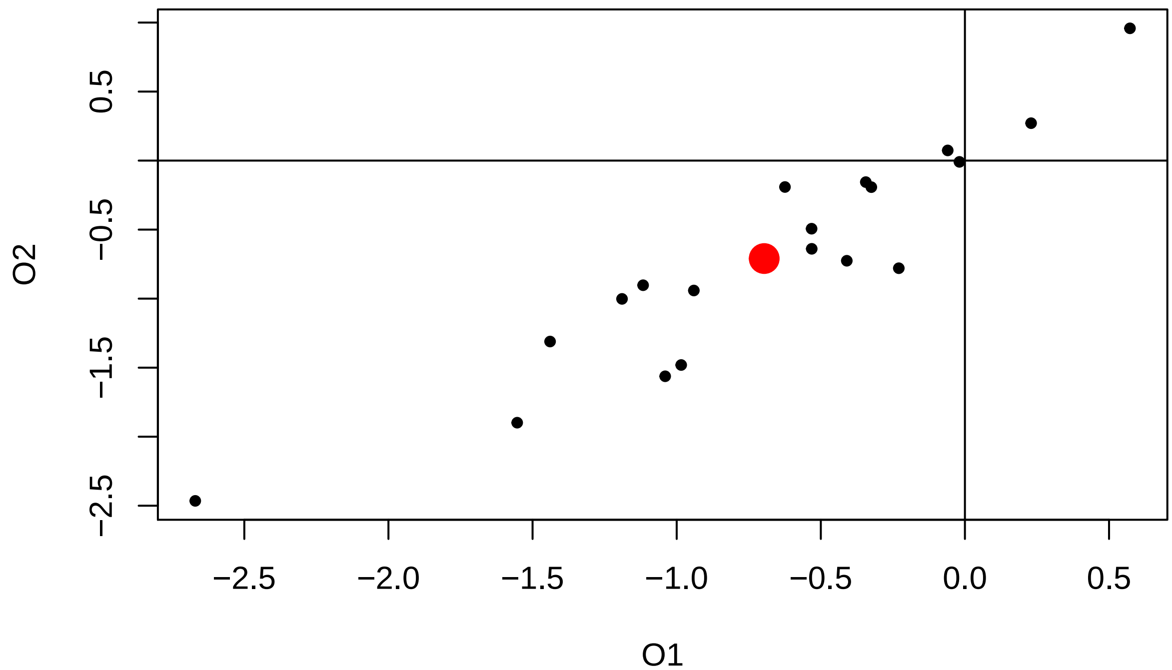
```
## $O1
##
##  One Sample t-test
##
## data:  newX[, i]
```

```
## t = -4.3327, df = 19, p-value = 0.0003587
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
##  -1.0327953 -0.3599747
## sample estimates:
## mean of x
## -0.696385
##
##
## $O2
##
##  One Sample t-test
##
## data:  newX[, i]
## t = -3.9761, df = 19, p-value = 0.0008092
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
##  -1.0831578 -0.3360722
## sample estimates:
## mean of x
## -0.709615
##
##
## $P7
##
##  One Sample t-test
##
## data:  newX[, i]
## t = -3.287, df = 19, p-value = 0.003877
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
##  -1.0380084 -0.2303616
## sample estimates:
## mean of x
## -0.634185
##
##
## $P8
##
##  One Sample t-test
##
## data:  newX[, i]
## t = -5.1301, df = 19, p-value = 5.95e-05
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
##  -1.0889013 -0.4578487
## sample estimates:
## mean of x
## -0.773375
##
##
## $PO7
##
##  One Sample t-test
```

```
##
## data:  newX[, i]
## t = -3.8837, df = 19, p-value = 0.0009992
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
##  -1.0736034 -0.3216666
## sample estimates:
## mean of x
## -0.697635
##
##
## $P08
##
##  One Sample t-test
##
## data:  newX[, i]
## t = -3.1978, df = 19, p-value = 0.004737
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
##  -1.1131969 -0.2324531
## sample estimates:
## mean of x
## -0.672825
```

4.3 Manova

```r
plot(Yfn[,1:2],pch=20)
abline(v=0)
abline(h=0)
points(mean(Yfn[,1]),mean(Yfn[,2]),cex=3,col=2,pch=20)
```



Manova test, overall among all channels:
H_0 neutral=object in ANY of the channels. https://en.wikipedia.org/wiki/Multivariate_analysis_of_variance https://en.wikipedia.org/wiki/Hotelling%27s_T-squared_distribution

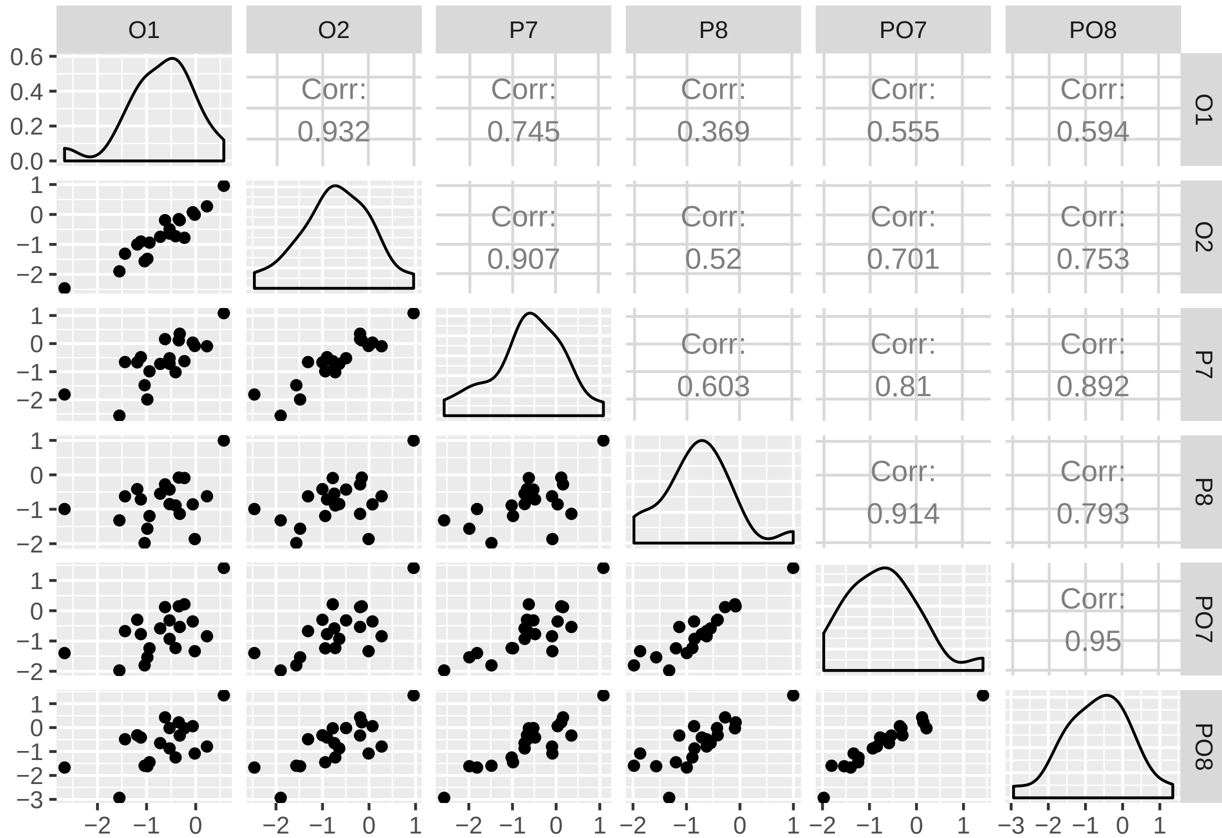
```
modman <- manova(Yfn ~ 1)
anova(modman)
```

```
## Analysis of Variance Table
##
##             Df  Pillai approx F num Df den Df    Pr(>F)
## (Intercept)  1 0.76928     7.78      6     14 0.0007989 ***
## Residuals   19
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
# equivalent to anova(modman,manova(Yfn ~ 0))
```

Assumptions: multivariate normality

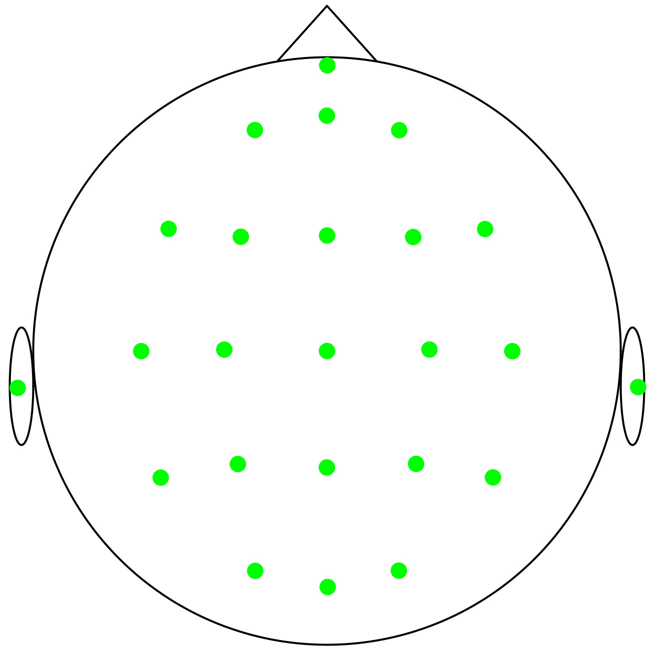
```
subj=dati$Subj[(1:20)*30]
ggpairs(data.frame(Yfn))
```



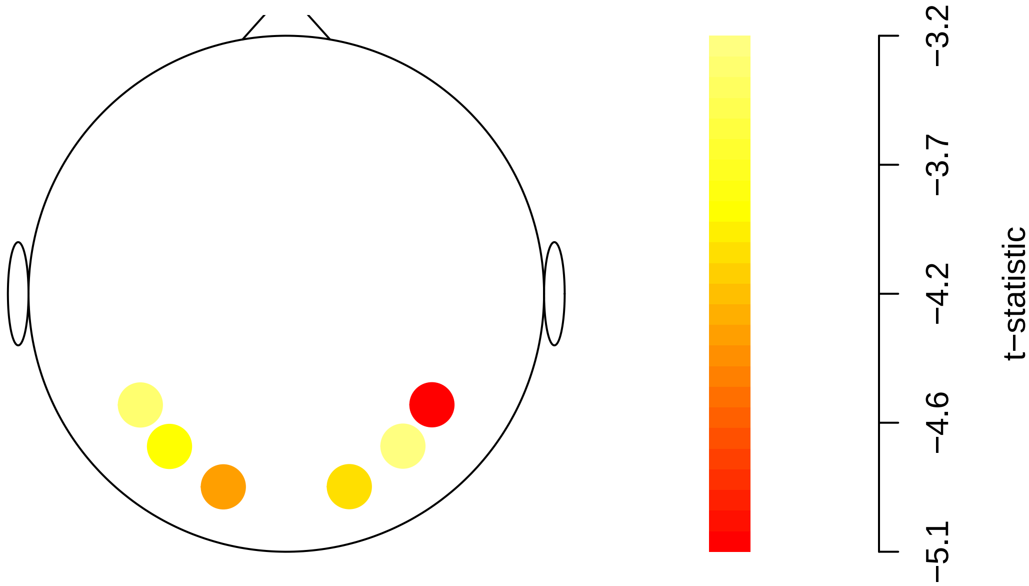
Not so bad, actually.

5 Mapping results on a scalp

```r
# install.packages("eegkit")
library(eegkit)

# plot 2d cap without labels
eegcap("10-20", plotlabels = FALSE)
```



```r
# get the t-statistic for each channel:
t_chan=sapply(uni_t,function(chan)chan$statistic)
names(t_chan)=gsub("\\.t","",names(t_chan))

# match to eeg coordinates
data(eegcoord)
cidx <- match(names(t_chan),rownames(eegcoord))

# # plot t-stat in 3d
#  open3d()
# eegspace(eegcoord[cidx,1:3],t_chan)

# plot t-stat in 2d
eegspace(eegcoord[cidx,4:5],t_chan,cex.point = 3,colorlab="t-statistic",mycolors=heat.colors(4))
```



If you like to play with library `ggplot2` this may help:
http://www.mattcraddock.com/blog/2017/02/25/erp-visualization-creating-topographical-scalp-maps-part-1/

6 (minimal) Bibliography

Jonathan Baron (2011) Notes on the use of R for psychology experiments and questionnaires https://www.sas.upenn.edu/~baron/from_cattell/rpsych/rpsych.html

and Course materal of
ST 732, Applied Longitudinal Data Analysis, NC State University by Marie Davidian https://www.stat.ncsu.edu/people/davidian/courses/st732/notes/chap5.pdf from https://www.stat.ncsu.edu/people/davidian/courses/st732/

About Type I, II, III SS: https://mcfromnz.wordpress.com/2011/03/02/anova-type-iiiiii-ss-explained/

About Mixed models:
http://webcom.upmf-grenoble.fr/LIP/Perso/DMuller/M2R/R_et_Mixed/documents/Bates-book.pdf
and
https://cran.r-project.org/web/packages/lme4/vignettes/lmer.pdf