

LAWLER 1980

- He implicitly shows that the probability of the LERW is the same as that of the Laplacian RW.
- In particular, he shows that it is the same as the conditioned probability of a ~~path~~ RW with "taboo region" equal to the already-visited points \equiv Laplacian RW:
- Indeed the Laplacian RW is a RW where the probability is given by solving the (discrete) Laplace equation with zero boundary condition on the already-present path and 1 at some other point.

Q??: In his 1980 paper, Lawler shows this correspondence in $d \geq 3$.

Does it hold also for $d \leq 2$?? I think that in a later paper he says yes, with some slight adjustments.

• Let's follow its derivation.

- First, define the RW and some often useful quantities.
- On a square lattice \mathbb{Z}^d , with $\{e_1, \dots, e_d\}$ the unit vectors, we consider the random variables Y_1, Y_2, \dots defined on the probability space (Ω, \mathcal{B}, P) with uniform distribution:

$$\text{DEF } P\{Y_i = e_j\} = \frac{1}{2d}, \quad j=1, \dots, 2d$$

↖ of Links at each point

SIMPLE RANDOM WALK S defined as:

$$S(n, \omega) := \sum_{i=1}^n Y_i(\omega) \quad \text{with } S(0, \omega) = 0 \in \mathbb{Z}^d \quad \forall \omega \in \Omega$$

↑ the extracted "random number"

Now, let's define the transition probability function for S :

$$P_k(x, y) := P_x \{ S(n) = y \} \equiv P \{ S(k+n) = y \mid S(n) = x \} \quad (\text{P}\{S(n)=x\}>0)$$

i.e. the probability that, given the RW at point x at the time n ,

we get a RW passing through y at time $k+n$.

P_x means: probability with S starting from x

N.B. This process is independent from n (i.e. No Memory, no interest in the past)

We also set

$$P_k := P_k(0, 0) = P_k(x, x) \quad \forall x \in \mathbb{Z}^d$$

i.e. the probability to come back to an already-visited point and therefore form a loop

Now, for any set $A \subset \mathbb{Z}^d$ and $x \in \mathbb{Z}^d$, let

$$E^A(x) := P_x \{ S(n) \notin A \ \forall n > 0 \} = P \{ S(n) \notin A \ \forall n > 0 \mid S(0) = x \} = \frac{P \{ S(n) \notin A \ \forall n > 0, S(0) = x \}}{P \{ S(0) = x \}}$$

i.e. the probability that a RW starting from x avoids the set A at any time after zero (It could start in A : $S(0) = x \in A$, but then leave it for good)

Now assume $A \subset \mathbb{Z}^d$ s.t. $E^A(0) > 0$, i.e. there is a finite probability that

a RW starting from the origin never visits A .
THIS IS A COLLECTION OF ω 's

Now call $\Omega^A \subset \Omega$ the set of those RW that avoid A ,

and let P^A be the conditioned probability:

$$P^A(B) := \frac{P(B)}{P(\Omega^A)} \quad \forall B \subset \Omega^A$$

i.e. the Conditioned probability to get a subset B of Random Walks that avoid A ($B \subset \Omega^A$), provided that we get a RW avoiding A

↳ Then

DEF

Laplacian RW

RW WITH TABOO SET A : S^A

= It is a Markov Process on Ω^A with stationary transition probabilities.

$P^A(x, x+e_i) := P^A \{ S^A(n+1) = x + e_i \mid S^A(n) = x \}$ INDEPENDENT FROM n
i.e. the probability to go from x to $x+e_i$ respecting the Taboo set A

N.B.

Even if he doesn't say it, This is essentially the Laplacian RW

Let's Manipulate $P^A(x, x+e_i)$, using the definition of P^A :

$$\begin{aligned} P^A(x, x+e_i) &= P \{ S(n+1) = x + e_i \mid S(n) = x, S(j) \notin A \ \forall j \geq 1 \} \\ &= \frac{P \{ S(n+1) = x + e_i, S(n) = x, S(j) \notin A \ \forall j \geq 1 \}}{P \{ S(n) = x, S(j) \notin A \ \forall j \geq 1 \}} / P \{ S(n) = x \} \end{aligned}$$

$$\text{Def of } E^A(x) \underset{\text{With } n \rightarrow 0}{=} \frac{P \{ S(n+1) = x + e_i, S(j) \notin A \ \forall j \geq 1 \}}{P \{ S(n) = x \}} \cdot \frac{P \{ S(n+1) = x + e_i \}}{P \{ S(n) = x \}}$$

$$\begin{aligned} \text{Def of } E^A(x) &\underset{\text{With } n \rightarrow 0}{=} \frac{P \{ S(n+1) = x + e_i, S(j) \notin A \ \forall j \geq 1 \}}{P \{ S(n) = x \}} \cdot \frac{P \{ S(n+1) = x + e_i \}}{P \{ S(n) = x \}} \\ &\stackrel{(1)}{=} \frac{P \{ S(n+1) = x + e_i \}}{P \{ S(n) = x \}} \cdot E^A(x) \\ &\stackrel{(2)}{=} \frac{P \{ S(n+1) = x + e_i \}}{P \{ S(n) = x \}} \cdot \frac{P \{ S(n+1) = x + e_i \}}{P \{ S(n) = x \}} \cdot E^A(x) \\ &\stackrel{(3)}{=} \frac{1}{2^d} \cdot E^A(x) \cdot E^A(x+e_i) \end{aligned}$$

Then we define by Induction:

$$P_k^A(x, y) := \sum_{z \in \mathbb{Z}^d} P^A(x, z) P_{k-1}^A(z, y)$$

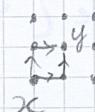
$$\text{I THINK w/ } P_k^A = P^A$$

c.g.

$$P_2^A(x, y) = \sum_{z \in \mathbb{Z}^d} P^A(x, z) P^A(z, y)$$

$$\text{Since } P^A(x, z) \propto \delta(|x-z|=1) \Rightarrow P_2^A(x, y) \propto \delta(|x-y|=3)$$

$$P_3^A(x, y) = \sum_{z \in \mathbb{Z}^d} P^A(x, z) P_2^A(z, y) \propto \delta(|x-y|=3)$$



Sum over all possible trajectories
Path integral with Boundary conditions:
(i) start at x
(ii) end at y
(iii) Avoid A

↳ Then, again by Induction I can show:

$$P_k^A(x,y) = \frac{E^A(y)}{E^A(x)} P\{S(n+k)=y, S(j) \notin A \forall j=h+1, \dots, h+k \mid S(n)=x\}$$

PROOF: $K=2$) $P_2^A(x,y) = \sum_{z \in \mathbb{Z}^d} P^A(x,z) P^A(z,y) = \sum_{z \in \mathbb{Z}^d} \frac{S^A(|x-z|=1)}{2d} \frac{E^A(y)}{E^A(z)} = \frac{E^A(y)}{E^A(x)} \frac{S^A(|x-y|=2)}{(2d)^2}$

Prob. that after $K=2$ steps from x we reach y , Avoiding A .

and $\frac{S^A(|x-y|=1)}{2d}$ is the probability that among the $2d$ neighboring vertices, the path chooses y (Always avoiding A)

$$\Rightarrow \frac{S^A(|x-y|=1)}{2d} = P\{S(n+1)=y, S(j) \notin A \forall j=h+1, \dots, h+k \mid S(n)=x\}$$

$$\Rightarrow \text{Valid for } K=1 \checkmark$$

- Now: assuming it's valid for K , let's show that it holds for $K+1$:

$$\hookrightarrow P_{K+1}^A(x,y) = \sum_{z \in \mathbb{Z}^d} P^A(x,z) P_k^A(z,y)$$

$$= \sum_{z \in \mathbb{Z}^d} P^A(x,z) \frac{E^A(y)}{E^A(z)} P\{S(n+k)=y, S(j) \notin A \forall j=h+1, \dots, h+k \mid S(n)=z\}$$

$$= \sum_{z \in \mathbb{Z}^d} \frac{E^A(z=x+e_i)}{2d} S^A(|x-z|=1) \frac{E^A(y)}{E^A(z)} P\{S(n+k)=y, S(j) \notin A \forall j=h+1, \dots, h+k \mid S(n)=z\}$$

$$\stackrel{\text{"def." of}}{=} \frac{E^A(y)}{E^A(x)} \sum_{z \in \mathbb{Z}^d} P_x\{S(1)=z, S(j) \notin A \forall j=1, \dots, K\} P_z\{S(k)=y, S(j) \notin A \forall j=1, \dots, K\}$$

$$= \frac{E^A(y)}{E^A(x)} P_x\{S(K+1)=y, S(j) \notin A \forall j=1, \dots, K+1\} \quad \text{X}$$

We basically add an extra step $x \rightarrow z$ then $z \rightarrow y$, always avoiding A

NB

NB As a consequence, if $A \subset B$, i.e. B is a larger region we want to avoid

$$\Rightarrow P_n^A(x,x) \geq P_n^B(x,x)$$

Simply because the probability to avoid a larger region is smaller (and $\frac{E^B(x)}{E^A(x)} = 1$)

= And in Particular since the Simple RW is Just a Random Walk with the empty set as the Taboo set:

$$P_k^A(x,x) \leq P_k^{\emptyset}(0,0) \quad \forall x \in \mathbb{Z}^d$$

i.e. any RW with any Taboo Set is less likely to return to its origin

then the unrestricted, Simple RW \Rightarrow if $P_k(0,0) = 0$ i.e. TRANSIENT, then $P_k^A(0,0) = 0$ TRANSIENT

Now, we define the loop-crossing procedure, to create the LERW.

There are several ways in which one can define this procedure. Let's follow the one proposed by Lawler:

- Assume S a RW on \mathbb{Z}^d with $d \geq 3$ WHAT ABOUT $d \leq 2$
in $d \leq 2$ the RW is RECURSIVE, i.e. it will come back to the origin with probability 1
(WE NEED INFINITE TIME TO DO THIS)

= Let

$$T_1(w) := \min \{j \text{ s.t. } \exists i < j \text{ with } S(i,w) = S(j,w)\}$$

i.e. consider the smallest time at which the RW visits a point it has already visited (\equiv forms a loop)

= and let

$$\varphi_1(w) := \text{that time } i < T_1(w) \text{ for which } S(i,w) = S(T_1(w),w), \quad \text{i.e. the time at which the RW first visited that point}$$

= Then Set

$$\hat{S}_1(n,w) := \begin{cases} S(n,w) & , n \leq \varphi_1(w) \\ S(n + (T_1(w) - \varphi_1(w)), w) & , n > \varphi_1(w) \end{cases}$$

i.e. \hat{S}_1 is the same RW as S up to time $\varphi_1(w)$ (when it visited the "knot" of a loop for the first time), then it skips the Loop (between times $\varphi_1(w)$ and $T_1(w)$), and finally it is S again

- Now we iterate this process: until that is no loop left.

THERE IS NO STOPPING CONDITION! it takes $K \rightarrow \infty$

Namely, for $K > 1$:

$$T_K(w) := \min \{j \text{ s.t. } \exists i < j \text{ with } \hat{S}_{K-1}(i, w) = \hat{S}_{K-1}(j, w)\}$$

$$\varphi_K(w) := \text{that } i < T_K(w) \text{ for which } \hat{S}_{K-1}(i, w) = \hat{S}_{K-1}(T_K(w), w)$$

and

$$\hat{S}_K(n, w) := \begin{cases} \hat{S}_{K-1}(n, w) & n \leq \varphi_K(w) \\ \hat{S}_{K-1}(n + (T_K(w) - \varphi_K(w)), w) & n > \varphi_K(w) \end{cases}$$

Finally:

$$\hat{S}(n, w) := \lim_{K \rightarrow \infty} \hat{S}_K(n, w)$$

$\boxed{\hat{S} \in \text{LERW}}$

Well defined because for $d \geq 3$,
 $S(n)$ approaches infinity with Prob. = 1
 (it is Transient) (almost surely)
 and thus \hat{S} is not empty. In $d=2$ there would be nothing left

B

\hat{S} is a Non-Markov Process: the distribution for each step is

dependent on the entire past history.

Indeed, we cannot know if a piece of the RW will be kept until we have finished the process, because it could come back to itself, thus forming a loop that needs to be erased

However, what is remarkable is that this has the same statistics as the Laplacian RW, which is instead Markovian

? Is it TRUE? NO, Laplacian RWs are Non-Markovian too
 (Farey loops (e.g.))

B LERWs are self-avoiding by construction

Now, we want to compute the transition probability for a LERW.

Let $\Gamma = [x_0, x_1, \dots, x_n]$ be a n -step self-avoiding path, $\neq \text{Walk}$ (a walk has a specific statistic)

i.e. an ordered set of points in \mathbb{Z}^d s.t.:

$$\begin{cases} x_0 = 0 \\ \|x_i - x_{i-1}\| = 1 \quad \forall i \in \mathbb{N} \\ x_i \neq x_j \quad \text{for } i \neq j \end{cases}$$

We want to Compute:

$$P\{\hat{S}(n+1) = x_{n+1} \mid [\hat{S}(0), \dots, \hat{S}(n)] = [x_0, \dots, x_n]\}$$

i.e. the transition prob. that at the next step ($n+1$), \hat{S} will be at x_{n+1} , provided that at each previous step it was a self-avoiding path $[x_0, \dots, x_n]$

Let's start by setting

(2D: we would need to place a cut-off, otherwise $\exists \tilde{n} \text{ s.t. } S(\tilde{n}) = S(n)$ with Prob. = 1)

$$B_K^\Gamma := \{w \text{ s.t. } [\hat{S}(0, w), \dots, \hat{S}(n, w)] = [x_0, \dots, x_n], S(K, w) = \hat{S}(n, w)$$

and $S(\tilde{n}, w) \neq S(n, w)$ for $\tilde{n} > n\}$

i.e. the set of all random events $w \in \Omega$ s.t. we have a LERW up to time n , which is the same as $\Gamma = [x_0, \dots, x_n]$; and s.t. at time K the

RW S ~~never~~ overlaps $\hat{S}(n)$, but ~~at any~~ later time $\tilde{n} > n$ ~~never again~~.
 K is the last time that S visits x_n .

Now, let's define a RW T on B_K^Γ :

$$[T(j) := S(j+K) - S(K)]$$

i.e. a RW normalized in such a way that $T(0) = 0$ (starts from the origin)

and that later it never comes back there (since in B_K^Γ : $S(\tilde{n}, w) \neq S(n, w)$ for $\tilde{n} > n$)

nor to any previous points of \hat{S}
 as it is shown later on

NB

NB In order for w to be in B_K^{π} , two conditions must be met:

(1) If one Loop-erases S for the first K steps, one must get the

n -step LERW $\hat{S} = [x_0, \dots, x_n]$;

$\boxed{\hat{S} \in LE(S)}$

(2) $\forall m > K$ $S(m, w) \notin \{x_0, \dots, x_n\}$, i.e. S will never come back on its first n -steps for any later time $m > K$.

This is so because:

= $S(m, w) = x_n$ for $m > K$ is impossible by definition of B_K^{π} ;

= And if $S(m, w) = x_j$ for $m > K$ and $j < n$, then the loop-erasure

would eliminate the loop containing $S(K, w) = x_n$. BECAUSE THE loop ERASURE IS TAKEN TO INFINITY

This would imply $\hat{S}(n, w) \neq x_n$ because $S(m, w) \neq x_n \forall m > K$.

Therefore, this $w \notin B_K^{\pi}$.

Now, let's see what (1) and (2) imply on T :

the random walk

(1): Since $\{S(i) \text{ with } i=0, \dots, K\}$ and $\{S(j+k) - S(k) \text{ with } j=1, 2, \dots\}$

BECAUSE RWs are Markov Processes and so $S(j+k)$ knows nothing of $S(j)$

are independent sets of random variables, this condition has no implication on the RW T ;

(2): This condition is exactly the definition of a RW starting at x_n with taboo set $\{x_0, \dots, x_n\}$, which is "its previous trail".

$\Rightarrow T$ acts on B_K^{π} as a RW with taboo set $\{x_0, \dots, x_n\}$,

$\Rightarrow \boxed{T = S^{\{x_0, \dots, x_n\}}}$

LRW $\nabla \nabla$ & THIS MEANS THAT IT WON'T ERASE ITS NEXT STEP $\nabla \nabla$

NB

NB All of this is actually independent on the K one chooses.

\hookrightarrow We can make slight changes:

$B^{\pi} := \{w \text{ s.t. } [\hat{S}(0, w), \dots, \hat{S}(n, w)] = [x_0, \dots, x_n]\}$

i.e. the set of w s.t. we get a LERW with trail $[x_0, \dots, x_n]$

and on B^{π} define

$$\boxed{T(j, w) := S(j+k_n(w), w) - S(k_n(w), w)}$$

$$w/ \boxed{k_n(w) := \max \{K \text{ s.t. } S(K, w) = \hat{S}(n, w)\}}$$

i.e. the last time S ~~was~~ visited $\hat{S}(n, w)$

\Rightarrow Again, T is a RW with Taboo set $\{x_0, \dots, x_n\}$!

Now, we use the transition prob. for RWs with a taboo set to compute the transition prob. for \hat{S} , which is what we were looking for:

$$\text{recall: } P^A(x, x+e_i) = \frac{E^A(x+e_i)}{(2d) E^A(x)}$$

This is a little sketchy (*)

$$= P\{T(1) = x+e_i\} = \boxed{P}$$

$$= \text{If we set } A_n = \{x_0, \dots, x_n\}: \quad \boxed{P\{T(1) = x+e_i\} = P^{A_n}(x, x+e_i)}$$

$$\boxed{P\{\hat{S}(n+1) = x+e_i \mid [\hat{S}(0), \dots, \hat{S}(n)] = [x_0, \dots, x_n]\} = P^{A_n}(x, x+e_i)}$$

NB

NB This completes the proof/computation of P for the LERW \hat{S} , but at the same time it shows that

$$\boxed{LERW \equiv LRW}$$

(*) This is TRUE because: given a w , the LERW is completely FIXED.

By construction, T is a RW. At each step, one should define a new T , i.e. "update" its taboo set

The fact $\nabla \nabla$ could form a loop later on, but never in its taboo set

However $\hat{S}(w)$ continues after time K etc. Never intersecting itself.

- On B_K^{π} i.e. for these w 's, we take: $T(j) := S(j+k) - S(k) =$ which gives the LERW \hat{S} without its first n points after loop erasure.
- The thing is: I just consider its next step: $\boxed{T(1)}$. By construction it will not form a loop AND it will not be cancelled by any future loop because it would mean that it came back on its TABOO SET!

2D: we would take the max. up to the cut-off

The $S(\tilde{n}, w) \neq S(n, w)$ $\forall \tilde{n} > n$
Condition in B_K^{π} has been replaced by the def. of $k_n(w)$