STAT6420 Final Project Cover Page December 5, 2023

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Pledge:
I have neither given nor received any unauthorized aid for this project. I abide by the academic honor code of UGA.

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Summary

In this report, we will do some analysis on the dataset comprising of samples taken in different farms from different animal sources. We would like to answer these five questions using generalized linear regression analysis:

- 1. Are Camplyobacter, Salmonella, and Listeria predictive of each other.
- 2. Model the probability of Camplyobacter in terms of the Type and Month variable
- 3. Model the variable EcoliLog10 in terms of the Type and Month variable
- 4. Are Farm and Flock needed to be controlled to answer the previous two research questions.
- 5. Can both the SampleType and PastureTime be predicted using the proportion of bacteria found? If not, can the SampleType alone be predicted using the proportion of bacteria found?

After analyzing the data, setting up some models and perform some specific test, we arrived at these conclusions for our questions: With our results, we can finally answer our questions presented in the Introduction.

- 1. We do not have enough information to show that Camplyobacter, Salmonella, and Listeria are predictive of each other.
- 2. A model was fitted and we get a dispersion effect of approximately 0.89. So there is lack of overdispersion but there is sign of underdispersion.
- 3. A model was fitted and it was shown that there are no influential points in our dataset. So those problematic points found can be considered to be outliers.
- 4. The Farm and Flock should not be controlled to answer the previous questions because the predictors used in the previous questions are not consistent in every Farm and every Flock.
- 5. We could not predict both the SampleType and PastureTime variables in unison. Rather, we predicted only the SampleType using the proportion of the bacteria samples found. Using the entire dataset, we have a small classification error rate of 6.45%. Using cross-validation, we get an average classification error rate of 7%.

1 Introduction

In this project, we will study the dataset that contains information on the. The goal of these report is to answer these five questions

- 1. Are Camplyobacter, Salmonella, and Listeria predictive of each other.
- 2. Model the probability of Camplyobacter in terms of the Type and Month variable
- 3. Model the variable EcoliLog10 in terms of the Type and Month variable
- 4. Are Farm and Flock needed to be controlled to answer the previous two research questions.
- 5. Can both the SampleType and PastureTime be predicted using the proportion of bacteria found? If not, can the SampleType alone be predicted using using the proportion of bacteria found?

2 Exploratory Data Analyses

2.1 First research question

We want to see the predictive power of Camplyobacter, Salmonella, and Listeria with each other. Hence, we will make three models as follows

```
\begin{split} \text{Camplyobacter} &= \beta_0 + \beta_1 \cdot \text{Salmonella} + \beta_2 \cdot \text{Listeria} \\ &\text{Salmonella} = \beta_0 + \beta_1 \cdot \text{Camplyobacter} + \beta_2 \cdot \text{Listeria} \\ &\text{Listeria} = \beta_0 + \beta_1 \cdot \text{Salmonella} + \beta_2 \cdot \text{Camplyobacter} \end{split}
```

The summary of these models are shown below:

```
> summary(model_Camplyobact)
Call:
lm(formula = "Campylobact~Listeria+Salmonella", data = broilers)
Residuals:
           1Q Median
   Min
                           3Q
                                  Max
-0.4676 -0.4405 -0.4395 0.5595 0.5605
Coefficients:
             Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.4404735 0.0134922 32.647 <2e-16 ***
Listeria -0.0009209 0.0320169 -0.029
                                           0.977
Salmonella 0.0271414 0.0303931
                                 0.893
                                           0.372
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 '' 1
Residual standard error: 0.4973 on 1887 degrees of freedom
  (2 observations deleted due to missingness)
Multiple R-squared: 0.0004237, Adjusted R-squared: -0.0006358
```

```
F-statistic: 0.3999 on 2 and 1887 DF, p-value: 0.6704
> summary(model_Salmonella)
Call.
lm(formula = "Salmonella~Campylobact+Listeria", data = broilers)
Residuals:
           1Q Median 3Q
   Min
-0.1830 -0.1830 -0.1674 -0.1601 0.8555
Coefficients:
          Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.16742 0.01219 13.739 <2e-16 ***
Campylobact 0.01556 0.01743 0.893 0.372
Listeria -0.02291 0.02424 -0.945 0.345
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.3766 on 1887 degrees of freedom
  (2 observations deleted due to missingness)
Multiple R-squared: 0.0008964, Adjusted R-squared: -0.0001626
F-statistic: 0.8465 on 2 and 1887 DF, p-value: 0.4291
> summary(model_Listeria)
lm(formula = "Listeria~Salmonella+Campylobact", data = broilers)
Residuals:
   Min
           1Q Median
                         3Q
                                  Max
-0.1540 -0.1540 -0.1535 -0.1333 0.8671
Coefficients:
             Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.1540067 0.0116058 13.270 <2e-16 ***
Salmonella -0.0206572 0.0218524 -0.945
                                        0.345
Campylobact -0.0004761 0.0165519 -0.029
                                        0.977
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
Residual standard error: 0.3575 on 1887 degrees of freedom
  (2 observations deleted due to missingness)
Multiple R-squared: 0.0004746, Adjusted R-squared: -0.0005848
F-statistic: 0.448 on 2 and 1887 DF, p-value: 0.639
```

We see that the predictors in all the models are not significant which shows signs of colinearity so we will check the correlation matrix.

```
Campylobact 1.000000000 -0.001109826 0.02057247
Listeria -0.001109826 1.000000000 -0.02177444
Salmonella 0.020572470 -0.021774441 1.00000000
```

We see that the correlations between different bacteria samples are very low so signs are slim.

To also consider all possible cases, we will consider a logistic model for each of the

2.2 Second research question

We will study the effect on the sample type and the months on the presence/absence of *Camplyobacter*. We first fit a linear model without interaction terms because there will be too many characters $(4 \times 7 = 28)$. The summary of the model is shown below:

```
Call:
lm (formula = "Campylobact~factor(SampleType) + factor(Month)",
   data = broilers)
Residuals:
           1Q Median 3Q
    Min
                                     Max
-1.01568 -0.21013 -0.01568 0.26482 0.99110
Coefficients:
                     Estimate Std. Error t value Pr(>|t|)
(Intercept)
                      1.13945 0.04212 27.052 < 2e-16 ***
                                0.03053 -7.738 1.64e-14 ***
factor(SampleType)Feces -0.23628
factor(SampleType)Soil -0.70929 0.03053 -23.229 < 2e-16 ***
factor(SampleType)WCR-F -0.91052 0.03704 -24.582 < 2e-16 ***
factor(SampleType)WCR-P -0.85714 0.03695 -23.199 < 2e-16 ***
                     factor (Month) 5
                     -0.19727 0.03705 -5.324 1.13e-07 ***
factor (Month) 6
factor (Month) 7
                     -0.22003 0.03754 -5.862 5.40e-09 ***
factor (Month) 8
                     -0.23105 0.03917 -5.898 4.34e-09 ***
factor (Month) 9
                     -0.24527
                                 0.04029 -6.088 1.38e-09 ***
                     -0.16800
factor (Month) 10
                                0.04007 -4.193 2.88e-05 ***
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 '' 1
Residual standard error: 0.3786 on 1879 degrees of freedom
  (2 observations deleted due to missingness)
Multiple R-squared: 0.423, Adjusted R-squared:
F-statistic: 137.8 on 10 and 1879 DF, p-value: < 2.2e-16
```

We see that all of the predictors are significant but the R^2 is moderately low. We check the residual vs fitted plot in Figure 1.

We see in Figure 1 that there is a lack of a horizontal band but rather a downward pattern shown. So we will also consider logistic regression (not probit because we are considering odds). The summary is shown below:

Residual vs Fitted Plot: Linear Regression

Figure 1: Residual vs fitted plot of regular linear regression

```
Call:
glm(formula = "Campylobact factor(SampleType)+factor(Month)",
    family = binomial(link = "logit"), data = broilers)
Coefficients:
                       Estimate Std. Error z value Pr(>|z|)
                         3.8825 0.3640 10.667 < 2e-16 ***
(Intercept)
factor(SampleType)Feces -1.9172
                                    0.3123 -6.140 8.27e-10 ***
factor(SampleType)Soil
                        -4.0581
                                    0.3151 -12.881 < 2e-16 ***
factor(SampleType)WCR-F -6.5206
                                    0.5428 - 12.012
                                                    < 2e-16 ***
                                    0.3883 - 13.477
                                                    < 2e-16 ***
factor (SampleType) WCR-P -5.2329
factor (Month) 5
                        -0.5767
                                    0.2534 - 2.276 0.02287 *
factor (Month) 6
                        -1.0494
                                    0.2396 -4.380 1.19e-05 ***
factor (Month) 7
                        -1.2041
                                    0.2430 -4.954 7.26e-07 ***
factor (Month) 8
                        -1.3088
                                    0.2603
                                            -5.027 4.97e-07 ***
                        -1.3930
                                    0.2668 -5.221 1.78e-07 ***
factor (Month) 9
factor (Month) 10
                        -0.8052
                                    0.2676 -3.009 0.00262 **
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 '' 1
(Dispersion parameter for binomial family taken to be 1)
    Null deviance: 2597.2 on 1889 degrees of freedom
Residual deviance: 1670.0
                          on 1879 degrees of freedom
  (2 observations deleted due to missingness)
AIC: 1692
```

Number of Fisher Scoring iterations: 6

We will check the deviance plot to see if a transformation is required.

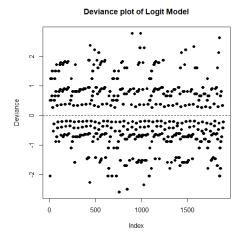


Figure 2: Deviance plot of logistic regression

We see in Figure 2 that the points are distributed evenly so there is no need for a transformation.

2.3 Third research question

We will study the effect on the sample type and the animal source on the logarithmic scale of the *E. coli* concentration in the sample. We first fit a linear model with interaction terms and the <code>EcoliLog10</code> is shifted by 0.1 so the boxcox plot can be shown. The summary is shown below:

```
Call:
lm(formula = "I(EcoliLog10 + 0.1) ~ factor(SampleType) + factor(AnimalSource)
    + factor(SampleType) * factor(AnimalSource) ", data = fecalsoil)
Residuals:
   Min
             10 Median
                              3Q
                                     Max
-5.7246 -0.8384 0.0370 0.9811 4.3188
Coefficients:
                                                   Estimate Std. Error
(Intercept)
                                                     6.63741 0.05691
factor(SampleType)Soil
                                                   -2.11200
                                                                0.08061
factor (AnimalSource) Cattle
                                                   -1.68056
                                                                0.22025
factor (AnimalSource) Layer
                                                    0.59819
                                                                0.16086
                                                                0.20972
factor (AnimalSource) Swine
                                                   -0.81281
factor(SampleType)Soil:factor(AnimalSource)Cattle -0.90452
                                                                0.31151
factor(SampleType)Soil:factor(AnimalSource)Layer
                                                   -1.26201
                                                                0.22753
factor(SampleType)Soil:factor(AnimalSource)Swine
                                                   -2.41290
                                                                0.29662
                                                   t value Pr(>|t|)
(Intercept)
                                                   116.629 < 2e-16 ***
factor(SampleType)Soil
                                                   -26.200 < 2e-16 ***
factor (AnimalSource) Cattle
                                                     -7.630 3.98e-14 ***
factor (AnimalSource) Layer
                                                     3.719 0.000207 ***
factor (AnimalSource) Swine
                                                    -3.876 0.000111 ***
factor(SampleType)Soil:factor(AnimalSource)Cattle -2.904 0.003739 **
```

We see that the interaction terms are very significant. We see that there is not much fanning in or out in the residual plot and the boxcox plot shows that there is no need for a transformation in Figure 3.

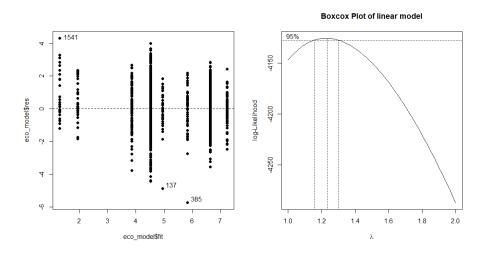


Figure 3: Residual vs Fitted Plot and Boxcox Plot of Linear regression

A Gamma model can be considered but we see from Figure 4, a fitted Gamma model and the linear model have the same fit.

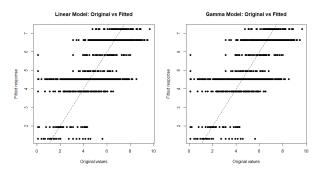


Figure 4: Fitted response vs original response for both linear regression fit and gamma regression fit

So we will stick with the linear model we fitted. Hence, we will check the leverage plot and the cooks distance plot: We see in Figures 3 and 5 that entry 385, 137 and 1541 have high residuals, high leverages

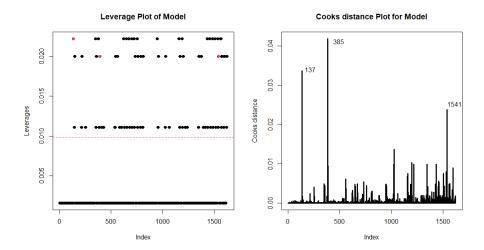


Figure 5: Leverage Plot and Cooks Distance Plot. The red dots are the entries 137, 385 and 1541.

and high cooks distances relative to the other entries. So they are considered to be either outliers or influential points. For more information, we will check the cooks distance if we remove each entry from the model.

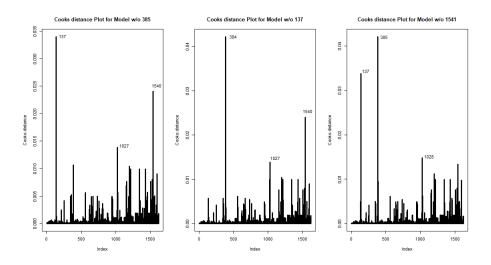


Figure 6: Cooks distances when the exceptional entries are removed

We see that not much changes in the cooks distance, so those entries are not considered to be influential points. So there are no strong influential points in this dataset.

2.4 Fifth research question

The question is if we can predict both the Sample Type and Pasture time by using the proportions of bacteria found (A, B, C, D, E). We cannot predict both the Sample Type and the Pasture time because they are two different groups which could be independent. However, we can predict the Sample Type because there are only two variable in the Sample Type so we can use logistic regression. We only need four predictors

because A + B + C + D + E = 1. If we run all four possible combinations of the four parameters, we see that the deviance for all of them are significant so we will choose the model that has the most significant predictors. The summary of our chosen model is shown below:

```
Call:
qlm(formula = Sample dummy ^ A + B + C + D, family = binomial)
          Estimate Std. Error z value Pr(>|z|)
(Intercept) 50.70 12.00 4.225 2.39e-05 ***
                       12.20 -4.557 5.19e-06 ***
             -55.60
             -50.93
                        12.28 -4.149 3.34e-05 ***
С
             -43.03
                       12.64 -3.404 0.000665 ***
\Box
                        19.27 -2.801 0.005098 **
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
   Null deviance: 300.711 on 216 degrees of freedom
Residual deviance: 69.343 on 212 degrees of freedom
AIC: 79.343
Number of Fisher Scoring iterations: 8
```

3 Statistical Analyses

3.1 First research question

Let $\beta_{Campylobacter}$, $\beta_{Salmonella}$ and $\beta_{Listeria}$ be the coefficients of the responditions. Our hypotheses are as follow:

$$H_0: \beta_{Campylobacter} = \beta_{Salmonella} = \beta_{Listeria} = 0$$

 $H_A:$ at least one of these parameters is 0

We will use is a *t*-test in our process using the models we prepared. As we see in both the linear and logistic models, none of the predictors are significant so we fail to reject the null hypothesis. Hence, we do not have enough to show that Camplyobacter, Salmonella, and Listeria are predictive of each other.

3.2 Second research question

We will check for overdispersion in the binomal model fitted. The calculation of the dispersion factor is as follows:

Deviance Error degree of freedom

Since this is a binomial model, the dispersion factor is expected to be very close to 1. Hence, the dispersion factor of our model is shown below:

> campyl_model_binom_logit\$deviance/campyl_model_binom_logit\$df.residual
[1] 0.8887724

The dispersion factor is smaller than 1 so there is no sign of overdispersion but a small sign of underdispersion.

3.3 Fourth research question

The models used in the second and third response should not be controlled for the Farm and Flock variables because it is possible that the predictors used in the previous research questions might not be in each Farm or Flock. For instance, we will predict the presence of Camplyobacter using the Type and Month but only for the Farm labeled as A.

```
glm(formula = "Campylobact factor(SampleType)+factor(Month)",
      family = binomial(link = "logit"), data = broilers[broilers$Farm ==
           "A", ])
Coefficients:
                                 Estimate Std. Error z value Pr(>|z|)
                                  6.5748 1.1852 5.547 2.90e-08 ***
(Intercept)
factor(SampleType)Feces -1.6478 1.0649 -1.547 0.121765 factor(SampleType)Soil -5.1429 1.0502 -4.897 9.73e-07 ***
factor(SampleType) WCR-F -22.9223 883.4291 -0.026 0.979300
factor(SampleType)WCR-P -8.2900 1.4567 -5.691 1.26e-08 ***
factor(Month) 5 -0.8190 0.7092 -1.155 0.248167
factor(Month) 6 -2.3167 0.6347 -3.650 0.000262 ***
factor(Month) 7 -2.0530 0.6532 -3.143 0.001672 **
factor(Month) 8 -3.2910 0.7047 -4.670 3.01e-06 ***
factor(Month) 9 -3.1195 0.7630 -4.088 4.34e-05 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
(Dispersion parameter for binomial family taken to be 1)
     Null deviance: 619.91 on 449 degrees of freedom
Residual deviance: 256.00 on 440 degrees of freedom
   (2 observations deleted due to missingness)
AIC: 276
Number of Fisher Scoring iterations: 17
```

We see Farm A lacks the 10 Month data so there will be a vast difference in the estimate of the other parameters.

3.4 Fifth research question

We will check the confusion matrix of our results. The matrix is shown in table 1:

Hence, the error rate can be evaluated as $\frac{10+4}{107+10+4+96} = \frac{14}{217} = 6.45\%$.

Now we will perform cross-validation on our dataset: breaking our training set and test set by 70 to 30. The confusion matrix for our cross-validation is shown in Table 2 We see that our error rate is $\frac{1+1}{40+1+1+30} = \frac{2}{72} \approx 2.78\%$.

		Actual values	
		Fecal	Soil
Predicted values	Fecal	107	10
	Soil	4	96

Table 1: Confusion matrix for logistic regression

	Actual values		
		Fecal	Soil
Predicted values	Fecal	40	1
	Soil	1	30

Table 2: Confusion matrix for logistic regression with cross-validation

If we repeat the cross-validation 10 times, we get this list of error rates, the summary of our list is shown below:

We see that the error rate average and median is around 7% and the error rate does not get larger than 10%.

4 Conclusions

With our results, we can finally answer our questions presented in the Introduction.

- 1. We do not have enough information to show that Camplyobacter, Salmonella, and Listeria are predictive of each other.
- 2. A model was fitted and we get a dispersion effect of approximately 0.89. So there is lack of overdispersion but there is sign of underdispersion.
- 3. A model was fitted and it was shown that there are no influential points in our dataset. So those problematic points found can be considered to be outliers.
- 4. The Farm and Flock should not be controlled to answer the previous questions because the predictors used in the previous questions are not consistent in every Farm and every Flock.
- 5. We could not predict both the SampleType and PastureTime variables in unison. Rather, we predicted only the SampleType using the proportion of the bacteria samples found. Using the entire dataset, we have a small classification error rate of 6.45%. Using cross-validation, we get an average classification error rate of 7%.

5 Appendix: R code

```
library(readxl)
 library(faraway)
3 library(MASS)
 library(leaps)
  dev.new()
  broilers <- read_excel("PasturedPoultryFarms.xlsx", sheet = "Broilers")</pre>
  fecalsoil <- read excel("PasturedPoultryFarms.xlsx", sheet =</pre>
10
   → "FecalSoil")
11
  Compositional <- read_excel("PasturedPoultryFarms.xlsx", sheet =</pre>
12
   13
14
  15
    #
16
17
    #
                                  First Ques
18
    #
20
    21
22
  model_Camplyobact <- lm("Campylobact~Listeria+Salmonella", data =</pre>
   → broilers)
  model_Salmonella <- lm("Salmonella~Campylobact+Listeria",data =</pre>
   → broilers)
  model_Listeria <- lm("Listeria~Salmonella+Campylobact", data = broilers)</pre>
25
26
  summary(model_Camplyobact)
27
  summary (model Salmonella)
  summary(model_Listeria)
  model_Camplyobact_binom <- glm("Campylobact~Listeria+Salmonella",data =</pre>
     broilers,
                         family = binomial)
32
  model_Salmonella_binom <- glm("Salmonella~Campylobact+Listeria", data =</pre>
   → broilers,
```

```
family = binomial)
34
  model_Listeria_binom <- qlm("Listeria~Salmonella+Campylobact",data =</pre>
   → broilers,
                             family = binomial)
36
37
  cor(broilers[,c("Campylobact","Listeria","Salmonella")],
38
      use = "complete.obs")
39
40
41
  summary(model_Camplyobact_binom)
42
  summary (model Salmonella binom)
43
  summary(model_Listeria_binom)
45
46
47
48
  49
    #
50
        #
     \hookrightarrow
51
    #
                                   Second Oues
52
    #
53
        #
    #
    55
56
57
  campyl_model <- lm("Campylobact~factor(SampleType) + factor(Month)",</pre>
58
                     data = broilers)
59
  summary(campyl_model)
  plot(campyl_model$fit, campyl_model$res,pch = 19,xlab = "Fitted

    values",

       ylab = "Residual values", main = "Residual vs Fitted Plot: Linear
62

→ Regression")
63
  boxcox(lm("I(Campylobact+0.1)~factor(SampleType)+factor(Month)",
64
            data = broilers))
65
  campyl_model_1 <-</pre>
     lm("log(Campylobact+0.1) factor(SampleType)+factor(Month)",
                       data = broilers)
68
  summary(campyl model 1)
69
```

```
70
  plot(campyl_model_1$fit, campyl_model_1$res,pch = 19)
71
73
  campyl_model_binom_logit <-</pre>
   data = broilers, family = binomial(link =
75
                          → "logit"))
  summary(campyl_model_binom_logit)
76
77
  plot(residuals(campyl_model_binom_logit, type = "deviance"), pch = 19,
79
       ylab="Deviance", main ="Deviance plot of Logit Model")
80
  abline (h = 0, lty = 2)
81
82
  #Dispersion factor
83
  campyl_model_binom_logit$deviance/campyl_model_binom_logit$df.residual
85
  #
87
    #
88
     \hookrightarrow
    #
                                 Third Ques
89
    #
91
    92
93
94
  fecalsoil$EcoliLog10 <- as.numeric(fecalsoil$EcoliLog10)</pre>
96
  eco_model <- lm("I(EcoliLog10 + 0.1) ~ factor(SampleType) +</pre>
98
   → factor(AnimalSource) + factor(SampleType)*factor(AnimalSource)",
                 data = fecalsoil)
99
  summary(eco_model)
100
101
102
  par(mfrow = c(1,2))
  plot(eco model$fit,eco model$res,pch = 19)
  abline (h=0, lty = 2)
105
  identify(eco model$fit,eco model$res,atpen = T, tolerance = 0.5)
```

```
107
   boxcox(eco\_model, lambda = seq(1, 2, by = 0.05))
108
   title ("Boxcox Plot of linear model")
   par(mfrow = c(1,1))
110
111
   step(eco_model)
112
113
   Cpplot(leaps(model.matrix(eco_model)[,-1],na.omit(fecalsoil$EcoliLog10)+0.1))
114
115
   maxadjr(leaps(model.matrix(eco_model)[,-1],na.omit(fecalsoil$EcoliLog10)+0.1,
116
                  method="adjr2"), best = 8)
117
118
   eco model gamma <- qlm("I(EcoliLog10+0.1) ~ factor(SampleType)+
119
      factor(AnimalSource) + factor(SampleType)*factor(AnimalSource)",
                      data = fecalsoil, family = Gamma(link = "inverse"))
120
121
   summary(eco_model_gamma)
122
123
124
   par(mfrow = c(1,2))
125
   plot(na.omit(fecalsoil$EcoliLog10)+0.1,eco_model$fit,pch = 19,
126
         xlab = "Original values",ylab="Fitted response",main = "Linear
127
          → Model: Original vs Fitted")
   abline (a=0, b=1, lty = 2)
128
129
   plot(na.omit(fecalsoil$EcoliLog10)+0.1,eco_model_gamma$fit,pch = 19,
130
         xlab = "Original values", ylab="Fitted response", main = "Gamma
131
           Model: Original vs Fitted")
   abline (a=0, b=1, lty = 2)
132
   par(mfrow = c(1,1))
133
134
   par(mfrow = c(1,2))
   plot(hat(model.matrix(eco_model)),pch = 19,ylab = 'Leverages',
        main = 'Leverage Plot of Model')
137
   abline(h =
    2*ncol(model.matrix(eco_model))/nrow(model.matrix(eco_model)),lty =
    \rightarrow 2, col = 2)
   points (137, hat (model.matrix (eco_model)) [137], col = 2, pch = 19)
139
   points(385, hat(model.matrix(eco_model))[385], col = 2, pch = 19)
   points(1541, hat(model.matrix(eco_model))[1541], col = 2, pch = 19)
142
   plot(cooks.distance(eco_model), type = 'h', lwd = 3, ylab = 'Cooks
       distance',
        main = 'Cooks distance Plot for Model')
145
```

```
identify(1:nrow(model.matrix(eco_model)),cooks.distance(eco_model),
146
            tolerance = 0.5, atpen = TRUE)
147
   par(mfrow = c(1,1))
148
149
   eco_model_385 <- lm("I(EcoliLog10 + 0.1) ~ factor(SampleType)+
150
      factor(AnimalSource) + factor(SampleType)*factor(AnimalSource)",
                   data = na.omit(fecalsoil)[-385,])
151
152
   eco_model_137 <- lm("I(EcoliLog10 + 0.1) ~ factor(SampleType)+
153
    → factor(AnimalSource) + factor(SampleType)*factor(AnimalSource)",
                       data = na.omit(fecalsoil)[-137,])
154
155
   eco model 1541 <- lm("I(EcoliLog10 + 0.1) ~ factor(SampleType)+
156
    → factor(AnimalSource) + factor(SampleType) * factor(AnimalSource) ",
                       data = na.omit(fecalsoil)[-1541,])
157
158
159
   par(mfrow = c(1,3))
   plot(cooks.distance(eco_model_385),type = 'h',lwd = 3,ylab = 'Cooks
161

→ distance',
        main = 'Cooks distance Plot for Model w/o 385')
162
   identify(1:nrow(model.matrix(eco_model_385)),cooks.distance(eco_model_385),
163
            tolerance = 0.5, atpen = TRUE)
164
165
   plot(cooks.distance(eco_model_137),type = 'h',lwd = 3,ylab = 'Cooks

→ distance',
        main = 'Cooks distance Plot for Model w/o 137')
167
   identify(1:nrow(model.matrix(eco model 137)),cooks.distance(eco model 137),
168
            tolerance = 0.5, atpen = TRUE)
169
170
   plot(cooks.distance(eco_model_1541),type = 'h',lwd = 3,ylab = 'Cooks
171

→ distance',
        main = 'Cooks distance Plot for Model w/o 1541')
172
   identify(1:nrow(model.matrix(eco_model_1541)),cooks.distance(eco_model_1541),
173
            tolerance = 0.5, atpen = TRUE)
174
175
   par(mfrow = c(1,1))
176
177
   178
     #
179
180
     #
     #
                                      Fourth Ques
181
```

```
182
    #
183
    184
185
186
   summary(glm("Campylobact factor(SampleType)+factor(Month)",
187
             data = broilers[broilers$Farm == "A",], family =
188

→ binomial(link = "logit")))
189
190
191
   192
    #
193
        #
194
    #
                                 Fifth Oues
195
196
    #
197
    198
199
200
  attach (Compositional)
201
   Sample_dummy <- factor(Compositional$Sampletype, labels = c(0,1))
202
203
  summary(glm(Sample_dummy~A + B + C + D, family = binomial))
204
   summary(glm(Sample\_dummy^A + C + D + E, family = binomial))
205
  summary(glm(Sample_dummy~A + B + D + E, family = binomial))
206
   summary(glm(Sample_dummy~A + B + C + E, family = binomial))
208
  compositional_model <- glm(Sample_dummy~A+ B + C + D, family =</pre>
209
   summary(compositional_model_A)
210
211
  compositional_model_E <- glm(Sample_dummy~ B+ C+D+E, family = binomial)</pre>
212
   summary(compositional_model_E)
213
  #Predicted values
  Pred sample <- compositional model A$fit > 0.5
216
217
```

```
#Confusion matrix
218
   table (Pred_sample, Sampletype)
219
220
   ##Cross-Validation
221
   test_num <- as.integer(nrow(Compositional)/3)</pre>
222
   index <- sample(nrow(Compositional), test_num)</pre>
223
   Sample_dummy_cross_val <- factor(Sampletype[-index],labels = c(0,1))
224
   compos cross val model <- qlm("factor(Sampletype, labels = c(0,1)) ~
225
    \rightarrow A+B+C+D",
                                     data = Compositional[-index,], family =
226
                                      → binomial)
   Pred_sample_cross_val <- predict(compos_cross_val_model,</pre>
       Compositional[index,c("A","B","C","D","E")],
                                         type = "response")>0.5
228
   table (Pred_sample_cross_val, Sampletype[index])
229
230
231
   ##Cross-Validation multiple loops
232
   accuracy_score_list <- c()</pre>
233
234
   for (i in seq(1,10)) {
235
      test_num <- as.integer(nrow(Compositional)/3)</pre>
236
      index <- sample(nrow(Compositional), test_num)</pre>
237
      Sample\_dummy\_cross\_val <- factor(Sampletype[-index],labels = c(0,1))
238
      compos_cross_val_model <- glm("factor(Sampletype, labels = c(0,1)) ~</pre>
239
       \rightarrow A+B+C+D",
                                        data = Compositional[-index,], family =
240
                                         Pred_sample_cross_val <- predict(compos_cross_val_model,</pre>
241
          Compositional[index,c("A", "B", "C", "D", "E")],
                                           type = "response")>0.5
242
     correct_obs <-
243
       sum(diag(table(Pred_sample_cross_val,Sampletype[index])))
     accuracy_score_list <-
244
          append(accuracy_score_list, (correct_obs/test_num))
245
246
    }
247
248
   summary(1-accuracy_score_list)
249
250
   1-c(1,2,3)
251
   predict (compos cross val model,
252
        Compositional[index,c("A", "B", "C", "D", "E")],
            type = "response")
253
```

254 detach(Compositional)