

Branching Random Walks with Selection

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1 EFFECT OF SELECTION ON ANCESTRY: AN EXACTLY SOLUBLE CASE AND ITS PHENOMENOLOGICAL GENERALIZATION

Placeholder text.

2 THE SPACE D

This section is based on Sections 3,11 and 12 of [?]. Consider $D(I, \mathbb{R})$ where I is a closed interval. Let Λ be the space of continuous bijections from $[0, 1]$ to itself which are zero at zero. The J_1 topology is defined by the metric

$$d_{J_1}(f, g) = \inf_{\lambda \in \Lambda} \{\|f \circ \lambda\|_\infty \vee \|\lambda - \mathbb{1}\|_\infty\} \quad (2.1)$$

The M_1 topology is defined by the same metric but on the completed graph. These definitions extend to $I = [0, \infty)$ by saying that $x_n \in D$ converges to $x \in D$ if convergence happens in $D([0, t])$ for all continuity points t of x . The topology this defines is metrisable (see page 83 of [?]). The topologies they induce are called strong and sometimes written SJ_1 and SM_1 . We talk about weak topologies when we consider $D(I, \mathbb{R}^k)$ with $k > 1$ and they are equal to the product topology, but they are not important to our discussion.

The $(S)J_2$ and $(S)M_2$ topologies are the ones induced by applying the Hausdorff metric to the functions graphs and completed graphs respectively.

3 BACKGROUND

THEOREM 3.1 ([?, Theorem 7.4.1.]) — *Let $\{X_{m,n} \mid 0 \leq m < n\}$ be a family of random variables satisfying*

- (i) $X_{l,n} \leq X_{l,m} + X_{m,n}$ for all $l < m < n$
- (ii) The distribution of $\{X_{m+k,n+k} \mid 0 \leq m < n\}$ does not depend on $k \in \mathbb{N}$
- (iii) $\mathbb{E}[X_{0,1}^+] < \infty$ and there exists $\gamma > -\infty$ such that $\mathbb{E}[X_{0,n}] > \gamma n$ for all $n \in \mathbb{N}$.

Then

4 EXACTLY SOLUBLE

The model we consider is constructed as follows: The individual at position $x_0 \in \mathbb{R}$ has offspring according to a poisson process with intensity measure that has density $x \mapsto \exp(-(x - x_0))$ with respect to the Lebesgue measure. Out of the offspring of all individuals, the N right-most are selected to form the population of the next generation. By additivity of Poisson processes, the location of the offspring is described by a Poisson process with density $\Psi(x) \sum_{i=1}^N \exp(-(x - x_i))$ where x_i denotes the position of the i 'th individual in the current generation.

Let us find the joint density of the $N + 1$ rightmost particles. Let x_1, \dots, x_{N+1} denote them, with $X_{N+1} < X_i$ for $i \in [N]$ but with no particular order on X_1, \dots, X_N . Then weby the mean value theorem for integrals we have:

$$\mathbb{P} \left(\bigcap_{i=1}^{N+1} \{X_i \in [x_i, x_i + \Delta_i]\} \right) = \frac{1}{N!} \exp \left(- \int_{x_{N+1}}^{\infty} \Psi(u) du \right) \prod_{i=1}^{N+1} \{\Delta_i \Psi(x_i)\} (1 + o(1)).$$

Thus, dividing by $\Delta_1 \times \dots \times \Delta_{N+1}$ and taking the Δ_i go to zero, we get the density

$$f_{X_1, \dots, X_{N+1}}(x_1, \dots, x_{N+1}) = \frac{1}{N!} \exp \left(- \int_{x_{N+1}}^{\infty} \Psi(u) du \right) \prod_{i=1}^{N+1} \Psi(x_i).$$

Now we can marginalise to obtain the density of X_{N+1} :

$$f_{X_{N+1}}(x) = \frac{1}{N!} \Psi(x) \left(\int_x^{\infty} \Psi(u) du \right)^N \exp \left(- \int_x^{\infty} \Psi(u) du \right).$$

From here it is easy to see that conditional on X_{N+1} , the X_i are independent with density

$$f(x) = \frac{\Psi(x)}{\int_{X_{N+1}}^{\infty} \Psi(u) du} \mathbb{1}_{\{X_{N+1} < x\}}.$$

5 SPEED

Placeholder text.

In generality, the models that we will discuss in this essay can be described as evolving according to two mechanisms: ...

5.1 EXPONENTIALLY DECAYING TAILS

5.1.1 CONSTRUCTION

The first variation of the N -branching random walk that we consider is the one studied in [1] by Bérard and Gouéré. Suppose that at timestep $n \geq 0$ there is a particle at $x \in \mathbb{R}$. During the branching step the particle dies giving birth to two children, whose positions independently (from each other and the past) follow a distribution with cumulative distribution function $p(\cdot - x)$. Out of all $2N$ children, the population at time $n + 1$ is then formed by the N rightmost particles.

Construction. Let $X = (X_n)_{n \geq 0} = (\sum_{i=1}^N \delta_{X_n(i)})_{n \geq 0}$ denote the \mathcal{C}_N -valued discrete time Markov process defined by the branching-selection procedure detailed above. Note that we suppress the dependence on N in our notation for simplicity. We can construct X easily: Let $\mathcal{E}_N := (\epsilon_{l,i,j})_{l \geq 0, i \in \llbracket 1, N \rrbracket, j=1,2}$ be an i.i.d. collection of random variables distributed like p . For $n \geq 0$ define $Y_{n+1} := \sum_{i=1}^N \sum_{j=1,2} \delta_{X_n(i) + \epsilon_{n,i,j}}$ and take X_{n+1} to be the counting measure supported on the rightmost N atoms of Y_{n+1} . This construction gives rise to an important monotonicity property that we record in the following Lemma:

LEMMA 5.1 ([1, Corollary 2]) — *For any $1 \leq N_1 \leq N_2$ and $\mu_i \in \mathcal{C}_{N_i}$ with $i = 1, 2$ such that $\mu_1 \preceq \mu_2$, there exists a coupling $(X_n^{(1)}, X_n^{(2)})_{n \geq 0}$ between two versions of the branching-selection particle system started from μ_1 and μ_2 respectively satisfying $X_n^{(1)} \preceq X_n^{(2)}$ almost surely for all $n \geq 0$.*

Proof. The proof is a straightforward consequence of our construction. The idea is to take an i.i.d. family $\mathcal{E}_{N_2} = (\epsilon_{l,i,j})_{l \geq 0, i \in \llbracket 1, N_2 \rrbracket, j=1,2}$ of random variables with law given by p and to use it to construct both $(X^{(1)})_n$ and $(X_n^{(2)})_{n \geq 0}$. \square

5.1.2 PROPERTIES OF THE MODEL

Let us now define the logarithmic moment generation function of p :

$$\Lambda(t) := \log \int_{\mathbb{R}} \exp(tx) dp(x).$$

In their analysis, Bérard and Gouéré impose some conditions on the domain $\mathcal{D}(\Lambda) := \{t \mid \Lambda(t) < \infty\}$ of Λ in order for the results of [2] to be applicable.

Assumption 1. Λ is finite in some neighbourhood of 0.

Assumption 2. There exists $t^* > 0$ in the interior of $\mathcal{D}(\Lambda)$ such that $t^* \Lambda'(t^*) - \Lambda(t^*) = \log 2$.

Assumption 1 is in fact equivalent to the requirement that p have exponentially decaying tails, furthermore it implies that p has finite moments of all orders. The results that follow in this section are conditional upon Assumptions 1 and 2 being satisfied.

Denote by $\max X_n$ and $\min X_n$ the right- and leftmost atom of X_n respectively. It is worth noting that $\min X_n$ and $\max X_n$ are integrable and hence finite by Assumption 1 when started from any fixed $X_0 \in \mathcal{C}_N$. Denote by $d(X_n) := \max X_n - \min X_n$ the diameter of X_n .

PROPOSITION 5.1 ([1, Corollary 1]) — *For any $N \geq 1$ and initial population $X_0 \in \mathcal{C}_N$, we have*

$$\frac{d(X_n)}{n} \xrightarrow[n \rightarrow \infty]{a.s., L^1} 0.$$

PROPOSITION 5.2 ([1, Proposition 2]) — *There exists $v_N = v_N(p) \in \mathbb{R}$ such that for any initial population $X_0 \in \mathcal{C}_N$ the following holds almost surely and in L^1 :*

$$\lim_{n \rightarrow \infty} \frac{\min X_n}{n} = \lim_{n \rightarrow \infty} \frac{\max X_n}{n} = v_N. \quad (5.1)$$

Proof. Recall the definition of \mathcal{E}_N from the construction of X . For each $l \geq 0$ we define the process $(X_n^l)_{n \geq 0}$ by shifting the origin of time by l . More precisely, given the process up to time $n \geq 0$, define X_{n+1}^l to be given by the N rightmost atoms of $\sum_{i=1}^N \sum_{j=1,2} \delta_{X_n^l(i) + \epsilon_{n+l,i,j}}$. It is clear that each $(X_n^l)_{n \geq 0}$ is distributed as the N -branching random walk with offspring law p . Suppose that for each $l \geq 0$ we start $(X_n^l)_{n \geq 0}$ from $N\delta_0$ and notice that $(X_n^0)_{n \geq 0} = (X_n)_{n \geq 0}$ almost surely, provided that $X_0 = N\delta_0$ also. From Lemma 5.1 it follows easily that

$$\max X_{n+m}^0 \leq \max X_n^0 + \max X_m^n \quad \forall n, m \geq 0. \quad (5.2)$$

Provided that the conditions of the theorem hold, applying Kingman's Subadditive Ergodic theorem yields $\lim_{n \rightarrow \infty} n^{-1} \max X_n = \lim_{n \rightarrow \infty} \mathbb{E}[n^{-1} \max X_n] = \inf_n \mathbb{E}[n^{-1} \max X_n] = v_N \in \mathbb{R}$ where the first limit is almost sure and also in L^1 . Using the fact that p has exponentially decaying tails (Assumption 1) and the independence of \mathcal{E}_N , the conditions are easily verified.

From Proposition 5.1 we immediately get $\lim_{n \rightarrow \infty} n^{-1} \min X_n = v_N$, so the proof is complete in the case $X_0 = N\delta_0$. By translation invariance of the dynamics of the system the result also follows for initial conditions of the form $N\delta_{x_0}$ for any $x_0 \in \mathbb{R}$. Finally, for arbitrary $X_0 \in \mathcal{C}_N$ note that the result is a consequence of Lemma 5.1 and a sandwiching argument between the initial configurations $N\delta_{\min X_0}$ and $N\delta_{\max X_0}$. \square

PROPOSITION 5.3 ([1, Proposition 3]) — *The sequence $(v_N)_{N \geq 1}$ is non-decreasing.*

Proof. This is again a consequence of Lemma 5.1. \square

Remark 5.1. From Proposition 5.3 we can deduce that v_N increases to a possibly infinite limit v_∞ as N goes to infinity. Assumption 1 implies that Λ is smooth on the interior of $\mathcal{D}(\Lambda)$ so that both quantities $v := \Lambda'(t^*)$ and $\chi := \frac{\pi^2}{2} t^* \Lambda''(t^*)$ are finite. In Section 5.1.4 we will see that v_∞ is in fact equal to v .

5.1.3 KILLED BRANCHING RANDOM WALKS

Following the notation used in [1], we formally define a Branching Random Walk (BRW) to be a pair (\mathcal{T}, Φ) , where \mathcal{T} is a rooted binary tree and Φ is a random map assigning a random variable $\Phi(u)$ to each vertex $u \in \mathcal{T}$. Φ must be such that $\Phi(\text{root}) = 0$ and $\{\Phi(v) - \Phi(u) \mid u \text{ is the parent of } v\}$ is i.i.d. with common distribution p . We call $\Phi(u)$ the value of the BRW at vertex u and write $\mathcal{T}(n)$ for the set of vertices in \mathcal{T} at distance n from the root. We say a sequence of vertices u_1, u_2, \dots is a path if u_{i+1} is the parent of u_i for each $i \geq 1$.

Suppose that we have a BRW (\mathcal{T}, Φ) and take $v \in \mathbb{R}$ and $m \geq 1$. We say that vertex u is (m, v) -good if there exists a path $u = u_0, u_1, \dots, u_m$ such that $\Phi(u_i) - \Phi(u) \geq vi$ for all $i \in \llbracket 0, m \rrbracket$. This is

essentially saying that there exists a path started from u that stays to the right of the space-time line through $(u, \Phi(u))$ with slope v , for at least m steps. The definition of an (∞, v) -good vertex is analogous. We now state two results from [2] that we will need to prove Theorem 5.4. Recall the definitions of v and χ from Remark ??.

THEOREM 5.2 ([2, Theorem 1.2]) — *Let $\rho(\infty, \epsilon)$ denote the probability that the root of the BRW with offspring distribution p is $(\infty, v - \epsilon)$ -good. Then, as $\epsilon > 0$ goes to zero,*

$$\rho(\infty, \epsilon) \leq \exp \left(- \left(\frac{\chi + o(1)}{\epsilon} \right)^{1/2} \right). \quad (5.3)$$

A similar result can be stated for the probability of observing a $(m, v - \epsilon)$ -good root with m finite:

THEOREM 5.3 ([1, Theorem 3]) — *Let $\rho(m, \epsilon)$ denote the probability that the root of the BRW with offspring distribution p is $(m, v - \epsilon)$ -good. For any $0 < \beta < \chi$, there exists $\theta > 0$ such that for all large m ,*

$$\rho(m, \epsilon) \leq \exp \left(- \left(\frac{\chi - \beta}{\epsilon} \right)^{1/2} \right), \quad \text{with } \epsilon := \theta m^{-2/3}. \quad (5.4)$$

5.1.4 BRUNET-DERRIDA BEHAVIOUR

We are now ready to present the main result of Bérard and Gouéré on N -branching random walks with exponentially decaying tails:

THEOREM 5.4 — *As N goes to infinity,*

$$v_\infty - v_N = \frac{\chi}{(\log N)^2} + o((\log N)^{-2}). \quad (5.5)$$

Let us describe the coupling between the N -branching random walk and N independent branching random walks, which will allow us to relate Theorems 5.2 and 5.4 to the N -branching random walk. Let $(\text{BRW}_i)_{i \in [1, N]} = ((\mathcal{T}_i, \Phi_i))_{i \in [1, N]}$ be a set of N independent copies of the BRW with offspring distribution p . Define $\mathbb{T}_n := \bigsqcup_{i=1}^N \mathcal{T}_i(n)$ to be the disjoint union of vertices at depth n in the N BRWs, and fix an arbitrary (nonrandom) total order on \mathbb{T}_n for each n . We now inductively define a sequence $(G_n)_{n \geq 0}$ of random subsets of \mathbb{T}_n , each with exactly N elements. These random subsets will correspond to the particles alive in the coupled N -branching random walk at time n . Define $G_0 = \mathbb{T}_0$ and given G_n , define H_n to be the vertices in \mathbb{T}_{n+1} that descend from vertices in G_n . Finally, set G_{n+1} to be the set of N vertices in H_n with the greatest value, resolving ties via the fixed total order on \mathbb{T}_{n+1} . If we now define (with some abuse of notation) $\mathfrak{X}_n = \sum_{u, i: u \in G_n \cap \mathcal{T}_i} \delta_{\Phi_i(u)}$ then $(\mathfrak{X}_n)_{n \geq 0}$ has the same distribution as X started from $N\delta_0$. Let us record a technical lemma that will be used in the proof of the lower bound in Theorem 5.4.

LEMMA 5.5 ([3, Adapted by Bérard and Gouéré from Lemma 5.2]) — *Let $v_1 < v_2 \in \mathbb{R}$ and $1 \leq m \leq n \in \mathbb{N}$. Suppose $0 =: x_0, \dots, x_n$ is a sequence of real numbers such that $\max_{i \in [0, n-1]} (x_{i+1} - x_i) \leq K$ for some $K > 0$, and define $I := \{i \in [0, n-m] \mid x_{i+j} - x_i \geq jv_1, \quad \forall j \in [0, m]\}$. If $x_n \geq v_2 n$, then $|I| \geq \frac{v_2 - v_1}{K - v_1} \frac{m}{n} - \frac{K}{K - v_1}$.*

Proof of lower bound in Theorem 5.4. As before, set $X_0 = N\delta_0$. Our aim is to show $v_N := \lim_{n \rightarrow \infty} \mathbb{E}[n^{-1} \max X_n] \leq v_\infty - \chi/(\log N)^2 + o((\log N)^{-2})$. However, we shall show this with v_∞ replaced by v , which combined with the upper bound also proves that $v_\infty = v$. Let $\beta \in (0, \chi)$ and let $\theta > 0$ be as in Theorem 5.4. Let $\lambda > 0$, and define

$$m := \left\lceil \theta^{3/2} \left(\frac{(1 + \lambda) \log N}{(\chi - \beta)^{1/2}} \right)^3 \right\rceil, \quad (5.6)$$

and $\epsilon := \theta m^{-2/3}$. The scale of ϵ and m is carefully chosen so that by Theorem 5.4,

$$\rho(m, \epsilon) \leq N^{-(1+\lambda)} \quad \text{for all large } N. \quad (5.7)$$

Take $\gamma \in (0, 1)$ and define $v_1 = v - \epsilon$ and $v_2 = v - (1 - \gamma)\epsilon$ noting that $v_1 < v_2 < v$. Finally, let $n = \lceil N^\xi \rceil$ for some $0 < \xi < \lambda$ and consider the following inequality with $\delta > 0$:

$$\begin{aligned} \mathbb{E} [n^{-1} \max X_n] &= \mathbb{E} [n^{-1} \max X_n [\mathbb{1}_{\{\max X_n < nv_2\}} + \mathbb{1}_{\{nv_2 \leq \max X_n < n(v+\delta)n\}} + \mathbb{1}_{\{(v+\delta)n \leq \max X_n\}}]] \\ &\leq v_2 + (v + \delta) \underbrace{\mathbb{P}(\max X_n \leq v_2 n)}_{(I)} + \underbrace{\mathbb{E} [n^{-1} \max X_n \mathbb{1}_{\{(v+\delta)n \leq \max X_n\}}]}_{(II)}. \end{aligned} \quad (5.8)$$

The strategy for the proof is to show that both (I) and (II) are $o((\log N)^{-2})$. The result then follows, as $v_2 = v - (1 - \gamma)(\chi - \beta)(1 + \lambda)^{-2}(\log N)^{-2}$ where γ, β, λ can be taken arbitrarily small.

Let B_n be the number of vertices in $\sqcup_{i=1}^n G_i$ that are (m, v_1) -good with respect to their respective BRWs. Define $K = \kappa \log(2Nn)$ for some $\kappa > 0$ and notice that the quantity $\frac{v_2 - v_1}{K - v_1} \frac{m}{n} - \frac{K}{K - v_1} = \Theta(N^\xi (\log N)^{-4})$ so that for large enough N it is positive. Let u_0, u_1, \dots, u_n be a path in \mathcal{T}_{i_0} for some $i_0 \in \llbracket 1, N \rrbracket$ such that $u_0 = \text{root}_{i_0}$ and $u_n \in G_n$ with $\Phi_{i_0}(u_n) = \max X_n$. In other words, let u_0, \dots, u_n be the path from the root to the rightmost particle at time n of the coupled N -branching random walk. On the event $E := \{\max X_n \geq v_2 n\}$, we apply Lemma 5.5 to the sequence of real numbers $(\Phi_{i_0}(u_i))_{i \in \llbracket 1, n \rrbracket}$ to see that either there is an (m, v_1) -good vertex among the u_i or one of the random walk steps along the path is $\geq K$. These events are respectively included in the events that $B_n \geq 1$ and that $M := \max\{\epsilon_{l,i,j} \mid l \in \llbracket 0, n-1 \rrbracket, i \in \llbracket 1, N \rrbracket, j = 1, 2\} \geq K$. We can use this to bound the probability of E :

$$\mathbb{P}(E) \leq \mathbb{P}(M \geq K) + \mathbb{P}(B_n \geq 1). \quad (5.9)$$

Consider a vertex $u \in \mathcal{T}_{i_0}(d)$ for some $i_0 \in \llbracket 1, N \rrbracket$ at depth $d \in \llbracket 0, n \rrbracket$. The event $\{u \in G_d\}$ is measurable with respect to the sigma algebra generated by the random variables $\{\Phi_j(v) \mid j \in \llbracket 1, N \rrbracket, \mathcal{T}_j \ni v \text{'s depth} \leq d\}$. On the other hand, the event $\{u \text{ is } (m, v_1)\text{-good}\}$ is determined by the variables $\{\Phi_{i_0}(v) - \Phi_{i_0}(u) \mid \mathcal{T}_{i_0} \ni v \text{'s depth} > d\}$, so that the two events are independent. We can write B_n as

$$B_n = \sum_{i \in \llbracket 1, N \rrbracket, u \in \mathcal{T}_i} \mathbb{1}_{\{u \text{ is } (m, v_1)\text{-good}\}} \mathbb{1}_{\{u \in G_d \text{ for some } d \in \llbracket 0, n \rrbracket\}}.$$

Taking expectations gives

$$\mathbb{E}[B_n] \leq N(n+1)\rho(m, \epsilon) = \mathcal{O}(N^{\xi-\lambda}) \quad \text{as } N \text{ goes to infinity,}$$

where we used that G_n has N elements for all n . Recall that the distribution p has exponentially decaying tails, so that there exist $C, \gamma > 0$ such that $\mathbb{P}_{X \sim p}(X > t) \leq C \exp(-\gamma t)$ for all large t . This gives $\mathbb{P}(M \geq K) \leq 1 - (1 - \exp(-\gamma \kappa \log(2Nn)))^N = \mathcal{O}(N(2Nn)^{-\gamma \kappa}) = \mathcal{O}(N^{1-(\xi+1)\gamma \kappa})$. Thus we can take κ large enough so that we get

$$\mathbb{P}(E) \leq \mathcal{O}(N^{\xi-\lambda}), \quad \text{which proves (I) is } o((\log N)^{-2}). \quad (5.10)$$

To show that (II) is $o((\log N)^{-2})$ first consider the obvious inequality $\exp(t \max X_n) \leq \sum_{i \in \llbracket 1, N \rrbracket, u \in \mathcal{T}_i(n)} \exp(t \Phi_i(u))$. Taking expectations gives $\mathbb{E}[\exp(t \max X_n)] \leq N 2^n \exp(n \Lambda(t))$, where we used a telescoping sum along the path connecting the root and u and the fact that $\#\mathcal{T}_i(n) = 2^n$ for each i . Recalling from Assumption 2 and Remark 5.1 that $\Lambda(t^*) = vt^* - \log 2$, we obtain

$$\mathbb{E}[\exp(t^* \max X_n)] \leq N \exp(vnt^*). \quad (5.11)$$

LEMMA 5.6 — Let $b > 0$. Then for all large enough a ,

$$x \mathbb{1}_{\{x \geq a\}} \leq \exp\left(b\left(x - \frac{a}{2}\right)\right), \quad \forall x \in \mathbb{R}. \quad (5.12)$$

Proof. Differentiate the map $f : x \mapsto \exp(b(x - a/2)) - x$ to find that for large enough a , f is increasing on $[a, \infty)$. Noting that $f(a) \geq 0$ for all large a concludes the proof. \square

Apply Lemma 5.6 with $X = \max X_n - vn$, $a = \delta n$, $b = t^*$ and take expectations to get

$$\mathbb{E}[(\max X_n - vn) \mathbb{1}_{\{\max X_n \geq (v+\delta)n\}}] \leq \mathbb{E}[\exp(t^*(X_n - vn - \delta n/2))],$$

which combined with 5.11 and a Chernoff bound gives

$$(II) = \mathbb{E}[\max X_n \mathbb{1}_{\{\max X_n \geq (v+\delta)n\}}] \leq N \exp(-\delta n/2)(1 + |v|n) = o((\log N)^{-2}).$$

We have shown that for any choice of $\gamma \in (0, 1)$, $\beta \in (0, \chi)$ and $\lambda > \xi > 0$, for all N large enough

$$\mathbb{E}[\lceil N^\xi \rceil^{-1} \max X_{\lceil N^\xi \rceil}] \leq v - (1 - \gamma) \frac{\chi - \beta}{(1 + \lambda)^2 (\log N)^2} + o((\log N)^{-2}). \quad (5.13)$$

Recall from the proof of Proposition 5.2 that $v_N = \inf_n n^{-1} \mathbb{E}[\max X_n]$, so the left hand side in 5.13 can be replaced by v_N . Taking γ, β, λ and ξ to zero gives the desired result. \square

5.2 GENERALISED BERNOULLI

Placeholder text.

Define the *censored* Galton Watson process $(X_n)_{n \geq 0}$ with offspring distribution \mathcal{X} to be the discrete time \mathbb{N} valued stochastic process satisfying

$$(i) \quad X_0 = N$$

$$(ii) \quad X_{n+1} = \min \left\{ N, \sum_{i=1}^{X_n} \mathcal{X}_{i,n+1} \right\} \text{ for all } n \geq 0$$

where $(\mathcal{X}_{i,n})_{i \geq 0, n \geq 0}$ is an i.i.d. collection of random variables with distribution \mathcal{X} . Complete at some later point, since this is really easy material.

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