

Branching Random Walks with Selection

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1 SPEED

Placeholder text.

In this essay we study Branching Random Walks (BRWs) with selection (also called N -branching random walks), which we can think of as a cloud of particles evolving on \mathbb{R} indexed by discrete time according to two mechanisms:

- 1 **branching** Each particle gives birth to its offspring around itself, according to some point process.
- 2 **selection** Out of all children of the current generation, the rightmost N are selected to form the next generation.

It will be convenient to think of BRWs and N -BRWs as stochastic processes taking values in the set \mathfrak{M} of counting measures μ on \mathbb{R} which put non-negative integer mass on every atom and further satisfy $\mu([x, \infty)) < \infty$ for all $x \in \mathbb{R}$. The latter condition is needed for the phrase ‘rightmost particles’ to be meaningful. We will write $\mathfrak{M}_N \subset \mathfrak{M}$ for measures which have total mass N and $\delta_{x_0} \in \mathfrak{M}_1$ for the unit mass at x_0 . The interpretation is that if μ is the value of the (N -)BRW at some time n , then there are exactly $\mu(\{x\})$ particles at position x at time n . There is a natural partial order on \mathfrak{M} : we say that $\mu \preceq \nu$ if $\mu([x, \infty)) \leq \nu([x, \infty))$ for all $x \in \mathbb{R}$. Naturally, for random elements \mathcal{L}, \mathcal{G} of \mathfrak{M} (such as BRWs) we say that $\mathcal{L} \preceq \mathcal{G}$ if there exists a coupling $(\mathcal{L}, \mathcal{G})$ such that $\mathcal{L} \preceq \mathcal{G}$ almost surely. For $\mathcal{L} \in \mathfrak{M}$ we also define $\{l \mid l \in \mathcal{L}\} := \{I \subset \mathbb{R} \mid \sum_{i \in I} \delta_i = \mathcal{L}\}$.

Using the notation introduced above, we can construct N -branching random walks in great generality. Suppose that \mathcal{L} is a random element of \mathfrak{M} and that $X := (X_n)_{n \geq 0}$ is an N -branching random walk evolving according to the law of \mathcal{L}_i . Then X is inductively constructed as follows: given $X_n \in \mathfrak{M}_N$ for $n \geq 0$, take N i.i.d. copies $(\mathcal{L}_i)_{i=1}^N$ of \mathcal{L} independently of X_n . Writing $X_k(1) \leq \dots \leq X_k(N)$ for the particles of X_k for all $k \in \mathbb{N}$, we let

$$\tilde{X}_{n+1} = \sum_{i=1}^N \sum_{l \in \mathcal{L}_i} \delta_{X_n(i)+l}, \quad (1.1)$$

and define X_{n+1} to be the rightmost N particles in \tilde{X}_{n+1} . This construction allows for a natural and important coupling between (N -)BRWs. This coupling was first described in [?], the way we present it here is more general and similar to [3] Lemma 4.1.

LEMMA 1.1 — *Let $1 \leq N_1 \leq N_2$ and $\mu_i \in \mathfrak{M}_{N_i}$ for $i = 1, 2$. Consider random elements $\mathcal{L}_i \in \mathfrak{M}_{N_i}$ for $i = 1, 2$. Then if $(X_n^{(i)})_{n \geq 0}$ is a(n) (N -)BRW which evolves according to the law of \mathcal{L}_i and starts from μ_i respectively, then there exists a coupling such that $X_n^{(1)} \preceq X_n^{(2)}$ almost surely for all $n \geq 0$.*

Sketch of proof. We construct the coupling inductively. The base case clearly holds. Given $X_n^{(1)} \preceq X_n^{(2)}$, independently take N_2 i.i.d. copies $\{(\mathcal{L}_i^{(1)}, \mathcal{L}_i^{(2)})\}_{i=1}^{N_2}$ of the coupling of \mathcal{L}_1 and \mathcal{L}_2 that witnesses the partial order. Using these, construct $\tilde{X}_{n+1}^{(1)}$ and $\tilde{X}_{n+1}^{(2)}$ as in (1.1). If the $X^{(i)}$ are regular BRWs just set $X_{n+1}^{(i)} = \tilde{X}_{n+1}^{(i)}$, if they are N -BRWs take the rightmost N -particles like before. Either way, we have $X_{n+1}^{(1)} \preceq X_{n+1}^{(2)}$ as desired. \square

1.1 BRANCHING RANDOM WALKS

Placeholder text.

In this section we describe the most important result in the study of BRWs and some of its consequences. At its heart is a probability change which enables us to say things about the BRW via its associated spine, a one-dimensional random walk. The idea was first introduced for BRWs by Lyons in [?]. In our exposition we follow [3] and [?, Section 4.7].

First let us introduce the notation that’s most commonly used in the literature of BRWs. For a BRW started from a single particle at zero, denote by \mathbb{T} the genealogical tree of the system and

write $(V(x))_{x \in \mathbb{T}}$ for the positions of the particles on the real line. Further let $|x|$ be the generation of x and write x_i for the ancestor of x in generation i so that $x_0 = \emptyset$ where \emptyset denotes the root of \mathcal{T} . Let (\mathbb{T}, V) be a BRW started from 0 and $\mathcal{L} \in \mathfrak{M}$ be the point process that governs its evolution. It is easy to see then that \mathbb{T} is a Galton-Watson tree with $\#\mathcal{L}$ as its reproduction law. For our results on N -BRWs we will assume

$$\mathbb{E}[\#\mathcal{L}] > 1 \quad \text{and} \quad \#\mathcal{L} \geq 1 \quad \text{almost surely,} \quad (1.2)$$

where we write $\#\mathcal{L}$ for the total mass of \mathcal{L} . The former ensures that \mathbb{T} is supercritical while the latter assumption is needed as we're ultimately interested in N -BRWs: if $\#\mathcal{L} = 0$ with nonzero probability then the N -BRW dies out almost surely. Consider now the logarithmic moment generating function

$$\psi(t) := \log \mathbb{E} \int_{\mathbb{R}} e^{tx} \mathcal{L}(dx) = \log \mathbb{E} \sum_{x \in \mathbb{T}: |x|=1} e^{tV(x)} \quad (1.3)$$

for $t \in \mathbb{R}$ where it is defined. Let $\zeta := \sup\{t > 0 \mid \psi(t) < \infty\}$. We will also assume that

$$0 < \zeta, \quad (1.4)$$

so that the BRW doesn't spread to the right too fast. Note that by standard results ψ is C^∞ on $(0, \zeta)$.

1.1.1 THE MANY-TO-ONE LEMMA AND AN AFFINE TRANSFORMATION

Suppose now that V satisfies

$$e^{\psi(1)} = \mathbb{E} \sum_{|x|=1} e^{V(x)} = 1. \quad (1.5)$$

Then we can define a random variable X by giving it's distribution function:

$$\mathbb{P}(X \leq x) = \mathbb{E} \sum_{|u|=1} \mathbb{1}_{\{V(u) \leq x\}} e^{V(u)}. \quad (1.6)$$

If $(S_n)_{n \geq 0}$ is a random walk with step distribution X started from some $S_0 = a \in \mathbb{R}$ then we have the following result:

LEMMA 1.2 (Many-to-one) — *Let V be as in (1.5) and take g measurable, $a \in \mathbb{R}$ and $n \geq 1$. If $V'(\cdot) = a + V(\cdot)$ and $S_0 = a$ almost surely, then provided the integrals exist*

$$\mathbb{E} \sum_{|x|=n} g(V'(x_1), \dots, V'(x_n)) = \mathbb{E} [e^{a-S_n} g(S_1, \dots, S_n)]. \quad (1.7)$$

Lemma 1.2 is a consequence of the Spinal Decomposition Theorem, possibly the most important tool in the study of BRWs. For more details see Chapter 4 of [?] and Section 4.7 in particular. $(S_n)_{n \geq 0}$ is sometimes called the random walk associated with the BRW (\mathbb{T}, V) . An application of the many-to-one formula with $a = 0$ and $g(x) = xe^{-x}$ yields $\mathbb{E}X = \mathbb{E} \sum_{|x|=1} V(x) \exp(V(x))$ while for the variance we get $\mathbb{V}(X) = \mathbb{E} \sum_{|x|=1} V(x)^2 \exp(V(x)) - (\mathbb{E}X)^2$ provided that the necessary integrability conditions hold.

One might ask themselves what use the Many-to-One lemma is if it relies on assumption (1.5). However, for any point process $\widehat{\mathcal{L}}$ which satisfies (1.4), there exists a deterministic, affine transformation of $\widehat{\mathcal{L}}$ under which it satisfies (1.5). We now restrict our attention to the class of point processes for which $\exists t^* \in (0, \zeta)$ such that

$$\psi(t^*) = t^* \psi'(t^*). \quad (1.8)$$

For a detailed discussion of when such t^* exists see the appendix of [?]. So suppose that $\widehat{\mathcal{L}}$ satisfies (1.4) and (1.8) and let $(\widehat{\mathbb{T}}, \widehat{V})$ be the corresponding BRW with logarithmic moment generating

function $\widehat{\psi}$. Let $\gamma_s : \mathcal{B}(\mathbb{R}) \rightarrow \mathcal{B}(\mathbb{R})$ to be the right shift operator on the space $\mathcal{B}(\mathbb{R})$ of Borel-measurable subsets of \mathbb{R} and define

$$\mathcal{L} := (t^* \widehat{\mathcal{L}}) \circ \gamma_{-\widehat{\psi}(t^*)}. \quad (1.9)$$

If (\mathbb{T}, V) is the BRW that corresponds to \mathcal{L} then it is easy to see that V and its logarithmic moment generator function ψ are given by

$$V(x) = t^* \widehat{V}(x) - |x| \widehat{\psi}(t^*) \quad \psi(t) = -t \widehat{\psi}(t^*) + t^* \widehat{\psi}'(tt^*). \quad (1.10)$$

for $x \in \mathbb{T}$. Notice that $\psi'(1) = 0$ so V satisfies (1.5). Recall the definition of the random variable X and the associated random walk $(S_n)_{n \geq 0}$. It follows from a simple calculation that then

$$\mathbb{E}X = \psi'(1) = 0 \quad \text{and} \quad \mathbb{V}(X) = \psi''(1) = (t^*)^2 \widehat{\psi}''(t^*). \quad (1.11)$$

1.1.2 KILLED BELOW A LINEAR BOUNDARY

Let $1 \leq m \leq \infty$ and call a sequence of vertices $(u_n)_{0 \leq n \leq m}$ a path if u_i is the parent of u_{i+1} for each $0 \leq i \leq m-1$. For $v \in \mathbb{R}$ we say that the vertex $u \in \mathbb{T}$ is (m, v) -good if there exists a path $(u_n)_{0 \leq n \leq m}$ with $u_0 := u$ such that $V(u_i) - V(u) \geq vi$ for all $i \in [m]$. This is essentially saying that there exists a path starting at u that stays to the right of the space-time line through $(u, V(u))$ with slope v , for at least m steps. Let (\mathbb{T}, V) be the transformed BRW governed by \mathcal{L} as described in the previous section, and let X be the associated centred random walk. Then we have the following two results:

THEOREM 1.3 ([2, Theorem 1.2]) — *Let $\rho(\infty, \epsilon)$ denote the probability that the root of (\mathbb{T}, V) is $(\infty, -\epsilon)$ -good. Then, as $\epsilon > 0$ goes to zero,*

$$\log \rho(\infty, \epsilon) \leq -\pi \sqrt{\frac{\mathbb{V}(X) + o(1)}{2\epsilon}}. \quad (1.12)$$

A similar result can be stated for the probability of observing a $(m, -\epsilon)$ -good root with m finite:

THEOREM 1.4 ([2, Consequence of proof of Theorem 1.2]) — *Let $\rho(m, \epsilon)$ denote the probability that the root of (\mathbb{T}, V) is $(m, -\epsilon)$ -good. For any $0 < \beta < \mathbb{V}(X)$, there exists $\theta > 0$ such that for all large m ,*

$$\log \rho(m, \epsilon) \leq -\pi \sqrt{\frac{\mathbb{V} - \beta}{2\epsilon}}, \quad \text{with } \epsilon := \theta m^{-2/3}.$$

1.2 EXPONENTIALLY DECAYING TAILS

1.2.1 CONSTRUCTION

The first variation of the N -branching random walk that we consider is very similar to the one studied in [1] by Bérard and Gouéré. However, we treat a slightly more general case where the number of offspring of each particle is random as opposed to being fixed at two. In this version of the N -branching random walk each particle dies and gives birth to a random number of offspring whose number is distributed like q . Given the position of the parent, say x , each child's position follows the law $p(\cdot - x)$ independently of the number and position of the other children.

Construction. Let $X = (X_n)_{n \geq 0} = (\sum_{i=1}^N \delta_{X_n(i)})_{n \geq 0}$ denote the \mathfrak{M}_N -valued discrete time Markov process defined by the branching-selection procedure detailed above. Note that we suppress the dependence on N in our notation for simplicity. We can construct X easily: Let $\mathcal{E}_N := (\epsilon_{l,i,j})_{l \geq 0, i \in [N], j \geq 1}$ and $\mathcal{M}_N := (\tau_{l,i})_{l \geq 0, i \in [N]}$ be i.i.d. collections of random variables distributed like p and q respectively, with the collections also independent from each other. Now, given the process up to time $n \geq 0$, we construct X_{n+1} as follows: define $Y_{n+1} := \sum_{i=1}^N \sum_{j=1}^{\tau_{n,i}} \delta_{X_n(i) + \epsilon_{n,i,j}}$ and take X_{n+1} to be given by the N rightmost particles of Y_{n+1} .

Let $\nu \in \mathfrak{M}$ be a random, finite counting measure with the same distribution as the offspring of a single particle at the origin in our branching-selection mechanism (the fact that $\nu \in \mathfrak{M}$ follows from Assumption ??). In other words, the number of atoms of ν has distribution q and each atom is placed independently at position drawn from p . Let us now define the logarithmic moment generation function of ν :

$$\psi(t) := \log \mathbb{E} \int_{\mathbb{R}} e^{tx} d\nu(x).$$

Note that in their analysis Bérard and Gouéré define a slightly different function $\Lambda(t) = \psi(t) - \log 2$, however the branching random walk literature usually uses our definition. We impose the following assumptions to gain access to the results of [2]:

Assumption 1. ψ is finite in some neighbourhood of 0.

Assumption 2. There exists $t^* > 0$ in the interior of the domain of ψ such that $t^* \psi'(t^*) = \psi(t^*)$.

Assumption ?? is in fact equivalent to the requirement that p have exponentially decaying tails, furthermore it implies that p has finite moments of all orders. The third assumption concerns the distribution q :

Assumption 3. q satisfies $q(0) = 0$ and $1 < \sum_{i=1}^{\infty} i^2 q(i) < \infty$.

The results that follow in this section are conditional upon Assumptions ??, ?? and ?? being satisfied. We now record a technical lemma that will help us later.

LEMMA 1.5 — *Let $\tau \in L^1$ be an \mathbb{N} -valued random variable and let $(\epsilon_n)_{n \geq 1}$ be an i.i.d. sequence of random variables with exponentially decaying tails, independent of τ . Then $M := \max_{1 \leq n \leq \tau} \epsilon_n$ has exponentially decaying tails.*

Proof. Let $C, \gamma, t_0 > 0$ be such that $\mathbb{P}(|\epsilon_1| \leq t) \geq 1 - Ce^{-\gamma t}$ for all $t > t_0$. Then for $t > t_0$ large enough, Bernoulli's inequality gives

$$\begin{aligned} \mathbb{P}(M > t) &\leq 1 - \mathbb{E}[\mathbb{P}(|\epsilon_1| \leq t)^\tau] \leq 1 - \mathbb{E}[(1 - Ce^{-\gamma t})^\tau] \\ &\leq 1 - \mathbb{E}[1 - Ce^{-\gamma t} \tau] = \underbrace{C \mathbb{E}[\tau]}_{< \infty} e^{-\gamma t}. \end{aligned}$$

Similarly, looking at the lower tail we get

$$\mathbb{P}(M < -t) \leq 1 - \mathbb{E}[\mathbb{P}(|\epsilon_1| \leq t)^\tau] \leq C \mathbb{E}[\tau] e^{-\gamma t}.$$

□

1.2.2 PROPERTIES OF THE MODEL

Denote by $\max X_n$ and $\min X_n$ the position of the right- and leftmost particle of X_n respectively. It is worth noting that $\min X_n$ and $\max X_n$ are integrable and hence finite by Assumptions ?? and ?? when started from any fixed $X_0 \in \mathfrak{M}_N$. Indeed, by independence we have

$$\mathbb{E}|\max X_n| \leq \mathbb{E} \left| \max X_0 + \sum_{l=0}^{n-1} \sum_{i=1}^N \sum_{j=1}^{\tau_{l,i}} \epsilon_{l,i,j} \right| \leq |\max X_0| + Nn \mathbb{E}[\tau_{0,1}] \mathbb{E}|\epsilon_{0,1,1}|. \quad (1.13)$$

Denote by $d(X_n) := \max X_n - \min X_n$ the diameter of X_n . We have the following result, analogous to Corollary 1 of [1]:

PROPOSITION 1.1 — *For any $N \geq 1$ and initial population $X_0 \in \mathfrak{M}_N$, we have*

$$\frac{d(X_n)}{n} \xrightarrow[n \rightarrow \infty]{a.s., L^1} 0.$$

Proof. Let $u \in \mathbb{N}_+$ and for $n \geq u$ consider the process X in the timeframe $\llbracket n - u, n \rrbracket$. Define $\mathcal{E} := \{\epsilon_{l,i,j} \mid l \in \llbracket n - u, n - 1 \rrbracket, i \in [N], j \in [\tau_{l,i}]\}$ and let $M := \max \mathcal{E}$, $m := \min \mathcal{E}$ noting that both have exponentially decaying tails by Lemma ???. Write $y := \max X_{n-u}$ for the rightmost particle's position at time $n - u$. Suppose that for each $k \in [u]$ we have $\min X_{n-u+k} < y + km$. As all steps during branching are $\geq m$, this implies in particular that the descendants of the particle 'y' survive all selection steps until time n . Therefore, on the event $A_u := \{\text{number of descendants of } y \text{ at time } n \text{ is } > N\}$ almost surely $\min X_{n-u+k} \geq y + k_0 m$ for some k_0 . By the definition of m this must also hold for all $k \in \llbracket k_0, u \rrbracket$, in particular for $k = u$. Noting that $\max X_n \leq y + uM$, it follows that

$$d(X_n) \mathbb{1}_{A_u} \leq u(M - m), \quad (1.14)$$

with probability one. A simple argument shows that $\mathbb{1}_{A_u} \rightarrow 1$ almost surely as $u \uparrow \infty$: take any path of length u started from 'y'. On A_u^c , along any such path the number of times that the corresponding particle has more than one child is less than N .

$$\mathbb{P}(A_u^c) \leq \sum_{k=0}^{N-1} \binom{u}{k} q(1)^{u-k} (1 - q(1))^k \leq Nu^{N-1} q(1)^{u-(N-1)} \rightarrow 0 \quad (1.15)$$

as $u \uparrow \infty$ since $q(1) < 1$. Fix $\epsilon > 0$ and take u large enough so that $\mathbb{P}(A_u^c) < \epsilon^2$. Consider the decomposition

$$\frac{d(X_n)}{n} = \frac{d(X_n)}{n} \mathbb{1}_{A_u} + \frac{d(X_n)}{n} \mathbb{1}_{A_u^c}. \quad (1.16)$$

Taking expectations and then taking n to infinity, the first term vanishes by (??). The second term is upper bounded by $(\mathbb{P}(A_u^c) \mathbb{E}[d(X_n)^2/n^2])^{1/2}$ using Hölder's inequality. A rough bound on $d(X_n)$ suffices now: at each branching step $l \geq 0$ take the maximum and the minimum of the $\sum_{j=1}^N \tau_{l,j}$ random walk steps. The diameter certainly grows by no more than the difference between these two at each step. By Lemma ??? this yields $\mathbb{E}[d(X_n)^2] = \mathcal{O}(n^2)$ which implies that the second term in ??? is $\mathcal{O}(\epsilon)$. Taking ϵ to zero concludes the proof of L^1 convergence. Almost sure convergence is a consequence of the proof of the next Proposition. \square

PROPOSITION 1.2 ([1, Proposition 2]) — *There exists $v_N = v_N(p) \in \mathbb{R}$ such that for any initial population $X_0 \in \mathfrak{M}_N$ the following holds almost surely and in L^1 :*

$$\lim_{n \rightarrow \infty} \frac{\min X_n}{n} = \lim_{n \rightarrow \infty} \frac{\max X_n}{n} = v_N. \quad (1.17)$$

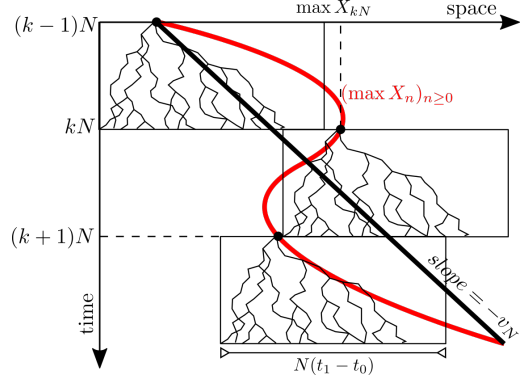
Proof. First we treat the case $X_0 = N\delta_0$. Recall the definition of \mathcal{E}_N and \mathcal{M}_N from the construction of X . For each $l \geq 0$ we define the process $(X_n^l)_{n \geq 0}$ by shifting the origin of time by l . More precisely, given the process up to time $n \geq 0$, define X_{n+1}^l to be given by the N rightmost particles of $\sum_{i=1}^N \sum_{j=1}^{\tau_{n+l,i}} \delta_{X_n^l(i) + \epsilon_{n+l,i,j}}$. It is clear that each $(X_n^l)_{n \geq 0}$ is distributed as the N -branching random walk with offspring law p . Start $(X_n^l)_{n \geq 0}$ from $N\delta_0$ for each $l \geq 0$ so that $(X_n^0)_{n \geq 0} = (X_n)_{n \geq 0}$ almost surely. From Lemma 1.1 it follows easily that

$$\max X_{n+m}^0 \leq \max X_n^0 + \max X_m^n \quad \forall n, m \geq 0. \quad (1.18)$$

For notational simplicity define $Y_{i,j} = \max X_{j-i}^i$ for $0 \leq i \leq j$. Then (??) reads $Y_{0,j} \leq Y_{0,i} + Y_{i,j}$ for all $0 \leq i \leq j$, which is familiar territory for Kingman's Subadditive Ergodic Theorem. We postpone showing that the conditions of the theorem hold to Lemma ???. Applying the theorem yields $\lim_{n \rightarrow \infty} n^{-1} \max X_n = \lim_{n \rightarrow \infty} \mathbb{E}[n^{-1} \max X_n] = \inf_n \mathbb{E}[n^{-1} \max X_n] = v_N \in \mathbb{R}$ where the first limit is almost sure. Noting that the process $(-X_n)_{n \geq 0}$ satisfies all the same assumptions as X , we can deduce from the identity $\min X_n = -\max(-X_n)$ that $\lim_{n \rightarrow \infty} n^{-1} \min X_n = \lim_{n \rightarrow \infty} \mathbb{E}[n^{-1} \min X_n] = \inf_n \mathbb{E}[n^{-1} \min X_n] = \tilde{v}_N \in \mathbb{R}$ exists too, where the first limit is almost sure. From the proof of Proposition ??? we immediately get $\tilde{v}_N = v_N$ by uniqueness of L^1 limits, which gives $\lim_{n \rightarrow \infty} n^{-1} d(X_n) = v_N - \tilde{v}_N = 0$ almost surely as claimed. The proof is complete in the case $X_0 = N\delta_0$. By translation invariance of the dynamics of the system the result also follows for initial conditions of the form $N\delta_{x_0}$ for any $x_0 \in \mathbb{R}$. Finally, for arbitrary $X_0 \in \mathfrak{M}_N$ note that the result is a consequence of Lemma 1.1 and a sandwiching argument between the initial configurations $N\delta_{\min X_0}$ and $N\delta_{\max X_0}$. \square

If we look at the previous proof, we see that the existence of v_N and \tilde{v}_N (the almost sure and L^1 limits of the left- and rightmost particles) when started from $X_0 = N\delta_0$ was shown without relying on Proposition ?? . We can in fact deduce Proposition ?? by an argument inspired by one of Prof. Berestycki's suggestions:

Alternative proof of Proposition ??. Let $Y = (Y_n)_{n \geq 0}$ be a branching random walk (without selection) with offspring distribution q and step distribution p . Start Y from δ_0 noting that initially there is only one particle. It is easy to see that the probability ρ_1 that the number of particles in Y_n has reached N by time N is strictly positive. Similar to the proof of Proposition ??, define $\mathcal{E} := \{\epsilon_j \mid 1 \leq i \leq \sum_{i=1}^{N^2} \tau_i, (\epsilon_j)_{j \geq 0} \stackrel{iid}{\sim} p \perp \tau_i \stackrel{iid}{\sim} q\}$ and $M := \min \mathcal{E}, m := \max \mathcal{E}$, where we should think of \mathcal{E} as the set of possible random walk steps that Y can take up to time N . By Lemma ?? we can choose $t_0 < t_1$ such that $\rho_2 := \mathbb{P}(t_0 \leq m \leq M \leq t_1) > 0$. We can now write



$$\begin{aligned} \mathbb{P}(\{Y_N \text{ has } N \text{ particles}\} \cap \{Nt_1 \geq \max Y_N \geq \min Y_N \geq Nt_0\}) &= \\ &= \mathbb{P}(Y_N \text{ has } N \text{ particles}) \mathbb{P}(Nt_1 \geq \max Y_N \geq \min Y_N \geq Nt_0 \mid Y_N \text{ has } N \text{ particles}) \\ &\geq \rho_1 \rho_2 > 0. \end{aligned} \quad (1.19)$$

Suppose that we couple X with $((Y_n^{(k)})_{0 \leq n \leq N})_{k \geq 0}$ which are independent copies of Y placed at the space-time points $(\max X_{kN}, kN)_{k \geq 0}$. By the second Borel-Cantelli lemma and (??) it follows that almost surely infinitely many of the $(Y_n^{(k)})_{0 \leq n \leq N}$ must have N particles by time N and have $Nt_0 \leq \min Y_N^{(k)} \leq \max Y_N^{(k)} \leq Nt_1$. This in turn implies that for infinitely many $k \geq 0$ the diameter $d(X_{kN})$ is less than $N(t_1 - t_0)$, which immediately yields $\tilde{v}_N = v_N$. \square

PROPOSITION 1.3 ([1, analogue of Proposition 3]) — *The sequence $(v_N)_{N \geq 1}$ is non-decreasing.*

Proof. This is again a consequence of Lemma 1.1. \square

Remark 1.1. From Proposition ?? we can deduce that v_N increases to a possibly infinite limit v_∞ as N goes to infinity. Assumption ?? implies that Λ is smooth on the interior of $\mathcal{D}(\Lambda)$ so that both quantities $v := \psi'(t^*)$ and $\chi := \frac{\pi^2}{2} t^* \psi''(t^*)$ are finite. In Section ?? we will see that v_∞ is in fact equal to v .

LEMMA 1.6 — *The random variables $Y_{i,j}$ as defined in the proof of Proposition ?? satisfy the hypothesis of Kingman's Subadditive Theorem.*

Proof. For each $k \geq 1$ the sequence $\{Y_{k,2k}, Y_{2k,3k}, \dots\} = \{\max X_k^k, \max X_k^{2k}, \dots\}$ is i.i.d. so stationary and ergodic. Clearly the distribution of $(Y_{i,i+k})_{k \geq 0} = (\max X_k^i)_{k \geq 0}$ is independent of i . $\mathbb{E}Y_{0,1}^+ = \mathbb{E}(\max X_1)^+ < \infty$ because $\max X_1 \in L^1$ by (??). Finally, $\mathbb{E}Y_{0,n} = \mathbb{E} \max X_n \geq n \mathbb{E} \min\{\epsilon_{0,i,j} \mid i \in [N], j \in [\tau_{0,i}]\}$ where the expectation is finite by Lemma ?? . \square

1.2.3 BRUNET-DERRIDA BEHAVIOUR

We are now ready to present and prove our main result in this section, the analogue of Bérard and Gouéré's Theorem 1:

THEOREM 1.7 — *As N goes to infinity,*

$$v_\infty - v_N = \frac{\chi}{(\log N)^2} + o((\log N)^{-2}).$$

First let us describe the coupling between the N -branching random walk and N independent branching random walks which allows us to relate Theorems 1.3 and 1.4 to the N -branching random walk. Let $(BRW_i)_{i \in [N]} = ((\mathcal{T}_i, \Phi_i))_{i \in [N]}$ be a set of N independent copies of the BRW with offspring distribution q and step distribution p . Define $\mathbb{T}_n := \bigsqcup_{i=1}^N \mathcal{T}_i(n)$ to be the disjoint union of vertices at depth n in the N BRWs, and fix an arbitrary (nonrandom) total order on \mathbb{T}_n for each n . We now inductively define a sequence $(G_n)_{n \geq 0}$ of random subsets of \mathbb{T}_n , each with exactly N elements. These random subsets will correspond to the particles alive in the coupled N -branching random walk at time n . Define $G_0 = \mathbb{T}_0$ and given G_n , define H_n to be the vertices in \mathbb{T}_{n+1} that descend from vertices in G_n . Finally, set G_{n+1} to be the set of N vertices in H_n with the greatest value, resolving ties via the fixed total order on \mathbb{T}_{n+1} . If we now define (with some abuse of notation) $\mathfrak{X}_n = \sum_{u, i: u \in G_n \cap \mathcal{T}_i} \delta_{\Phi_i(u)}$ then $(\mathfrak{X}_n)_{n \geq 0}$ has the same distribution as X started from $N\delta_0$. Going forward we will alternate between the notation of the two constructions of the N -branching random walk that we have given. Concretely, we will refer to \mathcal{T} , Φ , $\epsilon_{n,i,j}$ and $\tau_{n,i}$ without explicitly explaining the obvious relationships between these objects. Let us now record a technical lemma that will be used in the proof of the lower bound in Theorem ??.

LEMMA 1.8 ([4, Adapted by Bérard and Gouéré from Lemma 5.2]) — *Let $v_1 < v_2 \in \mathbb{R}$ and $1 \leq m \leq n \in \mathbb{N}$. Suppose $0 =: x_0, \dots, x_n$ is a sequence of real numbers such that $\max_{i \in \llbracket 0, n-1 \rrbracket} (x_{i+1} - x_i) \leq K$ for some $K > 0$, and define $I := \{i \in \llbracket 0, n-m \rrbracket \mid x_{i+j} - x_i \geq jv_1, \quad \forall j \in \llbracket 0, m \rrbracket\}$. If $x_n \geq v_2 n$, then $\#I \geq \frac{v_2 - v_1}{K - v_1} \frac{n}{m} - \frac{K}{K - v_1}$.*

Proof of lower bound in Theorem ??. As before, we first treat the case $X_0 = N\delta_0$. Our aim is to show $v_N := \lim_{n \rightarrow \infty} \mathbb{E}[n^{-1} \max X_n] \leq v_\infty - \chi/(\log N)^2 + o((\log N)^{-2})$. However, we shall show this with v_∞ replaced by v , which combined with the upper bound also proves that $v_\infty = v$. Recalling the proof of Proposition ??, we know by subadditivity that

$$v_N \leq \frac{\mathbb{E}[\max X_n]}{n} \quad \forall n \in \mathbb{N}. \quad (1.20)$$

In light of the desired correction term, we decompose (??) into

$$\mathbb{E}[\max X_n] = v - \frac{\chi}{(\log N)^2} + \mathbb{E}[\max X_n \mathbb{1}_{\{\max X_n \geq n(v - \chi(\log N)^{-2})\}}]. \quad (1.21)$$

However, this is not the form of the RHS that we will work with. Let $\gamma \in (0, 1)$ and $\epsilon = \epsilon(N)$ be such that approximately $\epsilon \sim \chi(\log N)^{-2}$ as $N \rightarrow \infty$. We will show that in

$$n^{-1} \mathbb{E}[\max X_n] = v - (1 - \gamma)\epsilon + n^{-1} \mathbb{E}[\max X_n \mathbb{1}_{\{\max X_n \geq n(v - (1 - \gamma)\epsilon)\}}] \quad (1.22)$$

the last term is $o((\log N)^{-2})$. This will yield the desired lower bound on $v - v_N$ if we take $\gamma \rightarrow 0$. We further decompose the problem: set $w = v - (1 - \gamma)\epsilon$ and $\delta > 0$. We have

$$n^{-1} \mathbb{E}[\max X_n \mathbb{1}_{\{\max X_n \geq nw\}}] \leq (v + \delta) \underbrace{\mathbb{P}(\max X_n \geq wn)}_{(I)} + \underbrace{\mathbb{E}[n^{-1} \max X_n \mathbb{1}_{\{(v + \delta)n \leq \max X_n\}}]}_{(II)}. \quad (1.23)$$

First we show that $(I) = o((\log N)^{-2})$, which requires careful scaling the variables involved. Set $n = \lceil N^\xi \rceil$ for some $0 < \xi < \gamma$ and $m = m(N) \leq n$ whose exact form we'll specify later. Let B be the number of vertices in $\bigsqcup_{i=0}^n G_i$ that are $(m, v - \epsilon)$ -good with respect to their corresponding BRWs. Let u_0, u_1, \dots, u_n be a path in \mathcal{T}_{i_0} for some $i_0 \in [N]$ such that $u_0 = \text{root}_{i_0}$ and $u_n \in G_n$ with $\Phi_{i_0}(u_n) = \max X_n$. In other words, let BRW_{i_0} be the random walk that the rightmost particle of the coupled N -branching random walk lives in at time n , and let u_0, \dots, u_n be the path connecting it to the root. Define the event $E := \{\max X_n \geq wn\}$, and apply Lemma ?? to the sequence of real numbers $(\Phi_{i_0}(u_i))_{i \in [n]}$. It follows that on E , for any $K > 0$, either one of the random walk steps along the path is $> K$ or $B \geq \frac{w - (v - \epsilon)}{K - (v - \epsilon)} \frac{n}{m} - \frac{K}{K - (v - \epsilon)}$. This yields

$$(I) = \mathbb{P}(E) \leq \mathbb{P}(M \geq K) + \mathbb{P}\left(B \geq \frac{w - (v - \epsilon)}{K - (v - \epsilon)} \frac{n}{m} - \frac{K}{K - (v - \epsilon)}\right), \quad (1.24)$$

where $M := \max\{\epsilon_{l,i,j} \mid l \in \llbracket 0, n-1 \rrbracket, i \in [N], j \in [\tau_{l,i}]\}$. Set $\beta \in (0, \chi)$ and let $\theta > 0$ be as in Theorem 1.4. Let $\lambda > 0$, and define

$$m := \left\lceil \theta^{3/2} \left(\frac{(1+\lambda) \log N}{(\chi-\beta)^{1/2}} \right)^3 \right\rceil, \quad (1.25)$$

and $\epsilon := \theta m^{-2/3}$. m is carefully chosen so that by Theorem 1.4,

$$\rho(m, \epsilon) \leq N^{-(1+\lambda)} \quad (1.26)$$

for all large N . Also let $K = \kappa \log N$ for $\kappa > 0$ and observe that $\frac{w-(v-\epsilon)}{K-(v-\epsilon)} \frac{n}{m} - \frac{K}{K-(v-\epsilon)} > 0$ for large enough N (independent of κ, γ). Thus (??) turns into

$$(I) = \mathbb{P}(E) \leq \underbrace{\mathbb{P}(M \geq K)}_{(a)} + \underbrace{\mathbb{P}(B \geq 1)}_{(b)}. \quad (1.27)$$

(a): $\mathbb{P}(M \geq K)$ is the probability of $S := \sum_{l=0}^{n-1} \sum_{i=1}^N \tau_{l,i} \in L^1$ i.i.d. variables with distribution p to be larger than K . Since p has exponentially decaying tails, we can take $C, \phi, t_0 > 0$ so that $p([t, \infty)) \leq C \exp(-\phi t)$ for all $t > t_0$. Then for $\kappa > t_0$ large enough for Bernoulli's inequality to apply, we have

$$\mathbb{P}(M \geq K) = 1 - \mathbb{E}[(1 - p([K, \infty)))^S] \leq 1 - \mathbb{E}[(1 - C \exp(-\phi K))^S] \quad (1.28)$$

$$= 1 - \mathbb{E}[(1 - C N^{-\phi \kappa})^S] \leq C N^{-\phi \kappa} \mathbb{E}[S] = \underbrace{C \mathbb{E}[\tau]}_{< \infty} N^{1+\xi-\phi \kappa}. \quad (1.29)$$

Thus, $\mathbb{P}(M \geq K) = o((\log N)^{-2})$ for large enough κ .

(b): Consider a vertex $u \in \mathcal{T}_{i_0}(d)$ for some $i_0 \in [N]$ at depth $d \in \llbracket 0, n \rrbracket$. The event $\{u \in G_d\}$ is measurable with respect to the sigma algebra generated by the random variables $\{\Phi_j(v) \mid j \in [N], \mathcal{T}_j \ni v \text{'s depth} \leq d\}$. On the other hand, the event $\{u \text{ is } (m, v_1)\text{-good}\}$ is determined by the variables $\{\Phi_{i_0}(v) - \Phi_{i_0}(u) \mid \mathcal{T}_{i_0} \ni v \text{'s depth} > d\}$, so that the two events are independent. We can write B as

$$B = \sum_{i=1}^N \sum_{u \in \mathcal{T}_i} \mathbb{1}_{\{u \text{ is } (n, v-\epsilon)\text{-good}\}} \mathbb{1}_{\{u \in G_d \text{ for some } d \in \llbracket 0, n \rrbracket\}}.$$

Taking expectations gives

$$\mathbb{E}[B] \leq N(n+1)\rho(m, \epsilon) = \mathcal{O}(N^{\xi-\lambda}) = o((\log N)^{-2}) \quad \text{as } N \text{ goes to infinity,} \quad (1.30)$$

where we used that G_n has N elements for all n . Applying Markov's inequality to B and combining with our estimate of (a) gives $(I) = o((\log N)^{-2})$ as desired.

We now turn to showing $(II) = o((\log N)^{-2})$. Consider the obvious inequality $\exp(t \max X_n) \leq \sum_{i \in [N], u \in \mathcal{T}_i(n)} \exp(t \Phi_i(u))$. If we set $\mathcal{G} := \sigma\{\tau_{l,i} \mid l \in \llbracket 0, n-1 \rrbracket, i \in [N]\}$, then we have

$$\mathbb{E}[\exp(t \max X_n)] \leq \mathbb{E} \sum_{i=1}^N \sum_{u \in \mathcal{T}_i(n)} \mathbb{E}[\exp(t \Phi_i(u)) \mid \mathcal{G}] = \mathbb{E} \sum_{i=1}^N \sum_{u \in \mathcal{T}_i(n)} \mathbb{E}_{\epsilon \sim p} [\exp(t \epsilon)]^n \quad (1.31)$$

$$= N \mathbb{E}_{\tau \sim q} [\tau]^n \mathbb{E}_{\epsilon \sim p} [\exp(t \epsilon)]^n, \quad (1.32)$$

where we used a telescoping sum along the path connecting u and the corresponding root for each vertex u in the sum. We can rewrite this in terms of $\psi(t)$:

$$N \mathbb{E}[\tau_{0,1}]^n \mathbb{E}_{\epsilon \sim p} [\exp(t \epsilon)]^n = N \mathbb{E}_{\epsilon \sim p \perp \tau \sim q} [\tau \exp(t \epsilon)]^n \quad (1.33)$$

$$= N \mathbb{E}_{\epsilon_j \sim p \perp \tau \sim q} \left[\sum_{j=1}^{\tau} \exp(t \epsilon_j) \right]^n = N \exp(n \psi(t)). \quad (1.34)$$

Recalling from Assumption ?? and Remark ?? that $\psi(t^*) = vt^*$, we obtain

$$\mathbb{E}[\exp(t^*(\max X_n - vn))] \leq N. \quad (1.35)$$

LEMMA 1.9 — Let $b > 0$. Then for all large enough a ,

$$x \mathbb{1}_{\{x \geq a\}} \leq \exp\left(b\left(x - \frac{a}{2}\right)\right), \quad \forall x \in \mathbb{R}. \quad (1.36)$$

Proof. Differentiate the map $f : x \mapsto \exp(b(x - a/2)) - x$ to find that for large enough a , f is increasing on $[a, \infty)$. Noting that $f(a) \geq 0$ for all large a concludes the proof. \square

Apply Lemma ?? with $X = \max X_n - vn$, $a = \delta n$, $b = t^*$ and take expectations to get

$$\mathbb{E}[(\max X_n - vn) \mathbb{1}_{\{\max X_n \geq (v+\delta)n\}}] \leq \mathbb{E}[\exp(t^*(X_n - vn - \delta n/2))],$$

which combined with (??) and a Chernoff bound gives

$$(II) = \mathbb{E}[\max X_n \mathbb{1}_{\{\max X_n \geq (v+\delta)n\}}] \leq N \exp(-t^* \delta n/2)(1 + |v|n) = o((\log N)^{-2}).$$

We have shown that for any choice of $\gamma \in (0, 1)$, $\beta \in (0, \chi)$ and $\lambda > \xi > 0$, for all N large enough

$$\mathbb{E}[\lceil N^\xi \rceil^{-1} \max X_{\lceil N^\xi \rceil}] \leq v - \frac{(1-\gamma)(\chi-\beta)}{(1+\lambda)^2(\log N)^2} + o((\log N)^{-2}). \quad (1.37)$$

Taking γ, β, λ and ξ to zero gives the desired result. \square

LEMMA 1.10 ([1, Lemma 3]) — Let $(M_n)_{n \geq 0}$ be a supercritical Galton-Watson process with square integrable offspring distribution started from $M_0 = 1$. Then there exist constants $r > 0$ and $\phi > 1$ such that $\mathbb{P}(M_n \geq \phi^n) \geq r$ for all $n \geq 0$.

Proof of upper bound in Theorem ??. Let $\tau \sim q$ and take $R < v$ such that $p([R, \infty)) > \mathbb{E}[\tau]^{-1}$. Define $M := (M_n)_{n \geq 0}$ to be a Galton-Watson process started from $M_0 = 1$ with offspring distribution \tilde{q} such that $\tilde{q}|\tau \sim \text{Binomial}(\tau, p([R, \infty)))$. Then M is supercritical and has square integrable offspring distribution. Hence by Lemma ?? there exist $r > 0$ and $\phi > 1$ such that $\mathbb{P}(M_n \geq \phi^n) \geq r$ for all $n \geq 0$.

Define $\lambda \in (0, 1)$ and let $\epsilon := \chi((1-\lambda)\log N)^{-2}$. Theorem 1.3 gives $\rho(\infty, \epsilon) = N^{\lambda-1+o(1)}$ as $N \rightarrow \infty$. Further define $s := \lceil \frac{\log N}{\log \phi} \rceil + 1$ and for $\eta \in (0, 1)$ define $m := \lceil \frac{(c-R)s}{\eta \epsilon} \rceil$ and finally set $n = s + m$. Consider a vertex u at depth m in a BRW $=(\mathcal{T}, \Phi)$ with offspring distribution q and step distribution p . The probability that there are at least ϕ^s distinct paths $u =: u_m, \dots, u_n$ with $\Phi(u_{i+1}) - \Phi(u_i) \geq R$ for all $i \in \llbracket m, n-1 \rrbracket$ is greater than r by Lemma ?? . Recall that the probability of the root being $(m, v - \epsilon)$ -good is $\rho(m, \epsilon)$. In light of the previous discussion, we see that the probability of observing a path $root = w_0, \dots, w_n$ in the BRW such that $\Phi(w_k) \geq k(v - \epsilon)$ for $k \in \llbracket 0, m \rrbracket$ and $\Phi(w_{k+1}) - \Phi(w_k) \geq R$ for $k \in \llbracket m, n-1 \rrbracket$ is at least $\rho(m, \epsilon)r$. By the choice of m and n , such a path must in fact be $(n, v - (1+\eta)\epsilon)$ -good. For $i \in [N]$ define A_j to be the event that BRW_i contains no more than ϕ^s distinct $(n, v - (1+\eta)\epsilon)$ -good paths starting at the root. By independence we get

$$\mathbb{P}\left(\bigcap_{i=1}^N A_i\right) \leq (1 - \rho(m, \epsilon))^N. \quad (1.38)$$

Denote $B := \{\min X_k < (v - (1+\eta)\epsilon) \text{ for all } k \in [n]\}$. On the event $B \cap [\cap_{i=1}^N A_i]^c$ one of the BRW_i s has $> \phi^s > N$ particles at time n that have stayed to the right of the space time line with slope $v - (1+\eta)\epsilon$ for all of $[n]$. By the definition of B this implies that there are $> N$ particles alive in the N -branching random walk which is a contradiction. Therefore we must have $B \subset \cap_{i=1}^N A_i$. Using the fact that $\rho(m, \epsilon) \leq \rho(\infty, \epsilon) = N^{\lambda-1+o(1)}$ and the inequality $1 - x \leq \exp(-x)$ for all $x \in \mathbb{R}$, we get

$$\mathbb{P}(B) \leq \mathbb{P}\left(\bigcap_{i=1}^N A_i\right) \leq \exp(-N^{\lambda+o(1)}) \quad (1.39)$$

LEMMA 1.11 ([1, Proposition 4]) — With the previous notations, for all N large enough,

$$v_N \geq (v - (1+\eta)\epsilon)(1 - n\mathbb{P}(B)) - n\mathbb{E}[|\Theta_n| \mathbb{1}_B], \quad (1.40)$$

where Θ_n is distributed as the minimum of $\sum_{l=0}^{n-1} \sum_{i=1}^N \tau_{l,i}$ i.i.d. random variables distributed like p , independent from $(\tau_{l,i})_{l \in \llbracket 0, n-1 \rrbracket, i \in [N]}$ which are i.i.d. with distribution q .

Proof. Identical to the one given in [1]. □

We immediately have $(v - (1 + \eta)\epsilon)n\mathbb{P}(B) = o((\log N)^{-2})$ by (??). To bound the other term, note that $\mathbb{E}[|\Theta_n|\mathbb{1}_B] \leq (\mathbb{E}[\Theta_n^2]\mathbb{P}(B))^{1/2}$ by Hölder's inequality □

REFERENCES

- [1] Jean Bérard and Jean-Baptiste Gouéré. Brunet-derrida behavior of branching-selection particle systems on the line. *Communications in Mathematical Physics*, 298(2):323–342, 2010.
- [2] Nina Gantert, Yueyun Hu, and Zhan Shi. Asymptotics for the survival probability in a supercritical branching random walk. *arXiv preprint arXiv:0811.0262*, 2008.
- [3] Bastien Mallein. N-branching random walk with α -stable spine. *Theory of Probability & Its Applications*, 62(2):295–318, 2018.
- [4] Robin Pemantle. Search cost for a nearly optimal path in a binary tree. *The Annals of Applied Probability*, pages 1273–1291, 2009.