

Model description

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A series y_t is decomposed into:

$$y_t = \mu_t + \gamma_t \quad (1)$$

where μ_t gives a trend and γ_t a seasonal component.

Probabilities for states $S_t = 0$ and $S_t = 1$ are governed by

$$S_t^* = \beta_0 + \beta_1 S_{t-1} + \eta_t \quad \eta_t \sim i.i.d.N(0, 1). \quad (2)$$

The Probit model discriminates between regimes whether S_t^* is non-negative. This gives the identities:

$$\begin{aligned} S_t = 0 & \quad \forall \quad \eta_t < -\beta_0 - \beta_1 \\ S_t = 1 & \quad \forall \quad \eta_t \geq -\beta_0 - \beta_1. \end{aligned} \quad (3)$$

The components themselves in 1 are assumed to satisfy:

$$\mu_t = \mu_{t-1} + \nu_{S_t} + \xi_t \quad \xi_t \sim i.i.d.N(0, \sigma_{\xi, S_t}^2) \quad (4)$$

with

$$\nu_{S_t} = \nu_0 + S_t \nu_1, \quad (5)$$

$$\sigma_{\xi, S_t}^2 = \sigma_{\xi, 0}^2 (1 - S_t) + \sigma_{\xi, 1}^2 S_t. \quad (6)$$

γ_t nets out seven-day (-period) seasonality:

$$\gamma_t = \sum_{j=1}^6 \gamma_{t-j} + \omega_t \quad \omega_t \sim i.i.d.N(0, \sigma_{\omega_t}^2). \quad (7)$$

The stochastic component ω_t allows the seasonal coefficients to vary and assumes measurement errors. By their nature, ξ_t and ω_s are uncorrelated for all t and s .

Endogeneity between the trend and regime processes arises from the joint density function of all innovations in the system:

$$\begin{bmatrix} \xi_{t,0} \\ \xi_{t,1} \\ \omega_t \\ \eta_t \end{bmatrix} \sim i.i.d.N(0, \Sigma), \quad \Sigma = \begin{bmatrix} \sigma_{\xi,0}^2 & 0 & 0 & \rho_0 \\ 0 & \sigma_{\xi,1}^2 & 0 & \rho_1 \\ 0 & 0 & \sigma_\omega^2 & 0 \\ \rho_0 & \rho_1 & 0 & 1 \end{bmatrix}. \quad (8)$$

Note, that the covariance ρ_{S_t} between innovations to the regime process and trend innovations is itself regime dependent. This is a function of the regime induced heteroskedasticity in 6. If a certain regime innovation triggers a switch from $S_{t-1} = 0$ to $S_t = 1$, $\sigma_{\xi,0}$ follows suit and changes to $\sigma_{\xi,1}$ in t . Since η_t has a constant variance by definition, ρ_0 is forced to change in proportion to $\sigma_{\xi,1}$ and becomes ρ_1 . However, the correlation between ξ_{t,S_t} and η_t stays constant and is thus uniquely defined as

$$\varrho = \frac{\rho_0}{\sigma_{\xi,0}} = \frac{\rho_1}{\sigma_{\xi,1}}, \quad \sigma_{\xi,0}, \sigma_{\xi,1} \neq 0. \quad (9)$$

I originally employed the approach to model log COVID-19 infections for my Bachelor thesis. Similar models were used in Chang, Choi, and Park 2017; Kim, Piger, and Startz 2008; Sinclair 2009; Luginbuhl and Vos 2003 and many more.

In state space form the measurement equation is defined as:

$$\underset{1 \times 1}{y_t} = \underset{1 \times 8}{Z} \underset{8 \times 1}{\alpha_t} \quad (10)$$

with the transition equation:

$$\underset{8 \times 1}{\alpha_{t+1}} = \begin{bmatrix} \mu_{t+1} \\ \nu_0 \\ \gamma_{t+1} \\ \vdots \\ \gamma_{t+1-5} \end{bmatrix} = \underset{8 \times 8}{T} \underset{8 \times 1}{\alpha_t} + \underset{8 \times 1}{\lambda_{S_t}} + \underset{8 \times 3}{R} \underset{3 \times 1}{u_{t+1}} \quad u_t \sim N(0, \underset{3 \times 3}{Q_{S_t}}) \quad (11)$$

where

$$Z = [1 \quad 0 \quad 1 \quad 0 \quad \dots \quad 0],$$

$$T = \text{diag} \left(\begin{bmatrix} 1 & 1 \\ 0 & 1 \end{bmatrix}, \begin{bmatrix} -1 & -1 & \dots & -1 & -1 \\ 1 & 0 & \dots & 0 & 0 \\ 0 & 1 & \dots & 0 & 0 \\ \vdots & \vdots & \ddots & \vdots & \vdots \\ 0 & 0 & \dots & 1 & 0 \end{bmatrix} \right),$$

$$\lambda_{S_t} = \begin{bmatrix} S_t \nu_1 \\ 0 \\ 0 \\ \vdots \\ 0 \end{bmatrix}, \quad R = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & 1 \\ 0 & 0 & 0 \\ \vdots & \vdots & \vdots \\ 0 & 0 & 0 \end{bmatrix}, \quad Q_{S_t} = \begin{bmatrix} \sigma_{\xi, S_t}^2 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & \sigma_\omega^2 \end{bmatrix}.$$

References

- Chang, Yoosoon, Yongok Choi, and Joon Y Park (2017). “A new approach to model regime switching”. In: *Journal of Econometrics* 196.1, pp. 127–143.
- Kim, Chang-Jin, Jeremy Piger, and Richard Startz (2008). “Estimation of Markov regime-switching regression models with endogenous switching”. In: *Journal of Econometrics* 143.2, pp. 263–273.
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