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A series  $y_t$  is decomposed into:

$$y_t = \mu_t + \gamma_t \quad (1)$$

where  $\mu_t$  gives a trend and  $\gamma_t$  a seasonal component.

Probabilities for states  $S_t = 0$  and  $S_t = 1$  are governed by

$$S_t^* = \beta_0 + \beta_1 S_{t-1} + \eta_t \quad \eta_t \sim i.i.d.N(0, 1). \quad (2)$$

The Probit model discriminates between regimes whether  $S_t^*$  is non-negative. This gives the identities:

$$\begin{aligned} S_t = 0 & \quad \forall \quad \eta_t < -\beta_0 - \beta_1 \\ S_t = 1 & \quad \forall \quad \eta_t \geq -\beta_0 - \beta_1. \end{aligned} \quad (3)$$

The components themselves in 1 are assumed to satisfy:

$$\mu_t = \mu_{t-1} + \nu_{S_t} + \xi_t \quad \xi_t \sim i.i.d.N(0, \sigma_{\xi, S_t}^2) \quad (4)$$

with

$$\nu_{S_t} = \nu_0 + S_t \nu_1, \quad (5)$$

$$\sigma_{\xi, S_t}^2 = \sigma_{\xi, 0}^2(1 - S_t) + \sigma_{\xi, 1}^2 S_t. \quad (6)$$

$\gamma_t$  nets out seven-day (-period) seasonality:

$$\gamma_t = \sum_{j=1}^6 \gamma_{t-j} + \omega_t \quad \omega_t \sim i.i.d.N(0, \sigma_{\omega_t}^2). \quad (7)$$

The stochastic component  $\omega_t$  allows the seasonal coefficients to vary and assumes measurement errors. By their nature,  $\xi_t$  and  $\omega_s$  are uncorrelated for all  $t$  and  $s$ .

Endogeneity between the trend and regime processes arises from the joint density function of all innovations in the system:

$$\begin{bmatrix} \xi_{t,0} \\ \xi_{t,1} \\ \omega_t \\ \eta_t \end{bmatrix} \sim i.i.d.N(0, \Sigma), \quad \Sigma = \begin{bmatrix} \sigma_{\xi,0}^2 & 0 & 0 & \rho_0 \\ 0 & \sigma_{\xi,1}^2 & 0 & \rho_1 \\ 0 & 0 & \sigma_\omega^2 & 0 \\ \rho_0 & \rho_1 & 0 & 1 \end{bmatrix}. \quad (8)$$

Note, that the covariance  $\rho_{S_t}$  between innovations to the regime process and trend innovations is itself regime dependent. This is a function of the regime induced heteroskedasticity in [6](#). If a certain regime innovation triggers a switch from  $S_{t-1} = 0$  to  $S_t = 1$ ,  $\sigma_{\xi,0}$  follows suit and changes to  $\sigma_{\xi,1}$  in  $t$ . Since  $\eta_t$  has a constant variance by definition,  $\rho_0$  is forced to change in proportion to  $\sigma_{\xi,1}$  and becomes  $\rho_1$ . However, the correlation between  $\xi_{t,S_t}$  and  $\eta_t$  stays constant and is thus uniquely defined as

$$\varrho = \frac{\rho_0}{\sigma_{\xi,0}} = \frac{\rho_1}{\sigma_{\xi,1}}, \quad \sigma_{\xi,0}, \sigma_{\xi,1} \neq 0. \quad (9)$$

I originally employed the approach to model log COVID-19 infections for my Bachelor thesis. Similar models were used in [Chang, Choi, and Park 2017](#); [Kim, Piger, and Startz 2008](#); [Sinclair 2009](#); [Luginbuhl and Vos 2003](#) and many more.

## References

- Chang, Yoosoon, Yongok Choi, and Joon Y Park (2017). “A new approach to model regime switching”. In: *Journal of Econometrics* 196.1, pp. 127–143.
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