

11.

5 def: $\pi(x)$, $Q(x, dy)$, $\alpha(x, y)$, $p(x, y)$, $P(x, A)$
 Reversibility: $p(x, y) \pi(x) = p(y, x) \pi(y)$
 Invariance: $\int_E P(x, A) \pi(x) dx = \pi(A)$

12.

- (y_i, x_i) , y_i continuous: $E[y_i], x_i$?
- Model
- Likelihood: $S^2, \hat{\beta}_{MLE} = (X^T X)^{-1} X^T y$, $(y - X\hat{\beta})^T (y - X\hat{\beta}) = \dots$
- σ^2 known: $(\tau X^T X + B_0^{-1})^{-1}$
- σ^2 unknown: $(X^T X + B_0^{-1})^{-1} = B_n$ — $\begin{cases} \beta | \sigma^2 \sim N_p(\hat{\beta}_{MLE}, \sigma^2 B_n) \\ \beta | \sigma^2 \sim N_p(\hat{\beta}_{MLE}, \sigma^2 B_n) \end{cases}$
- Jeffreys: $\sigma^2 (X^T X)^{-1}$

13.

- $P(m=j)$, M_j : likelihood, prior
- $\pi(\theta_j | y, M_j)$
- $P(m=j | y)$

14.

CPO_i = CONDITIONAL PREDICTIVE ORDINATE i $(x(y_i | y_{-i}))$
 LPML = LOG-PSEUDO-MARGINAL LIKELIHOOD

15.

- $z_i \Leftrightarrow y_i$: (y_i, x_i) : $y_i \in \{0, 1\}$
- $\beta | z, y \propto Z(z | \beta) \pi(\beta)$
- $[z | \beta, y] \propto Z(y | z) Z(z | \beta)$
- $\beta | - \sim N_p((X^T X + B_0^{-1})^{-1} (X^T \hat{\beta}_{MLE} + B_0^{-1} b_0), (X^T X + B_0^{-1})^{-1})$
- $\begin{cases} z_i | \beta, y_i = 1 & - \\ z_i | \beta, y_i = 0 & - \end{cases}$

16.

$L(\theta | D) = \prod_{i=1}^n [(f(y_i | \theta))^{d_i} (s_i(y_i | \theta))^{1-d_i}]$ (y_i, s_i) $y_i = \min(t_i, c_i)$, $d_i = \mathbb{1}(t_i \leq c_i)$
 Conditions: 1. $T_i \perp C_i$ 2. $C_i \perp \theta$ 3. all the subjects in the sample are \perp | parameter

17.

$$E[P(A)P(B)] = \frac{\alpha(A \cap B) - \alpha(A)\alpha(B)}{\alpha(A)+1}$$

$$Cov(P(A), P(B)) = \frac{\alpha_0(A \cap B) - \alpha_0(A)\alpha_0(B)}{\alpha+1}$$

$$E[P(A)] = \alpha_0(A)$$

$$Var(P(A)) = \frac{\alpha_0(A)\alpha_0(A^c)}{\alpha+1}$$

18.

$$Z(X_1, \dots, X_n) = Z(X_1)Z(X_2 | X_1) \dots \Rightarrow Z(X_1) = \alpha_0 \Rightarrow Z(X_2 | X_1) = \dots \Rightarrow Z(X_i | X_{i-1}, \dots, X_1) = \dots$$

$$\Rightarrow Z(X_1, \dots, X_n) = \alpha_0 \prod_{j=2}^n \left(\frac{\alpha}{\alpha+j-1} \alpha_0 + \frac{1}{\alpha+j-1} \sum_{i=1}^{j-1} d_{X_i} \right)$$

$$\Rightarrow Z(X_{n+1} | -) = \frac{\alpha}{\alpha+n} \alpha_0 + \frac{1}{\alpha+n} \sum_{j=1}^n d_{X_j}$$

19.

- Exchangeability
- Polya urn with z colors: $X_i \sim Be(\theta)$
- $IP(X_1, \dots, X_n)$
- De Finetti: $\lim \bar{X}_n \rightarrow Beta(w, b)$

20.

$$\left\{ \begin{array}{l} X_1, \dots, X_n | P \sim f(x)(w) = \int k(x, \theta) P(d\theta) \\ P \sim D_x \end{array} \right\} \parallel \left\{ \begin{array}{l} X_i | \theta_i \sim k(\cdot, \theta_i) \\ \theta_1, \dots, \theta_n | P \sim P \\ P \sim D_x \end{array} \right\} \quad \left\{ \begin{array}{l} X | \theta \perp P | \theta \end{array} \right.$$

21.

$$P|X \sim \int D_x + \sum_{j=1}^n \delta_{\theta_j} H(d\theta | X) : H(d\theta | X) \propto \left[\prod_{j=1}^n k(x_j, \theta_j) \right] \left[\alpha_0(d\theta) \prod_{j=1}^n \left[\alpha(d\theta) + \sum_{i=1}^{j-1} \delta_{\theta_i}(d\theta) \right] \right]$$

22./23.

$$p_n^* = \arg \min_{\hat{p}_n} E[L(p, \hat{p}_n) | X_1, \dots, X_n]$$

Binder: $L(p, \hat{p}_n) = \sum_{i,j} [C_1 \mathbb{1}(s_i = \hat{s}_i) \mathbb{1}(s_i \neq \hat{s}_j) + C_2 \mathbb{1}(s_i \neq \hat{s}_j) \mathbb{1}(\hat{s}_i = \hat{s}_j)]$

$$M_{ij} = IP(\hat{s}_i = \hat{s}_j | X_1, \dots, X_n), \quad j^* = \arg \min_j \sum_i | \mathbb{1}(\hat{s}_i = \hat{s}_j) - M_{ij} |$$

$$3. \quad X_1, \dots, X_n | \theta \sim P_\theta \Rightarrow P(\theta \in B | X_n = x) = \frac{\int_B f(x|\theta) \pi(d\theta)}{\int_{\Theta} f(x|\theta) \pi(d\theta)}$$

$$\theta \sim \pi(\theta)$$

proof. • I $\delta(A \times B) = \int_B P_\theta(A) \pi(d\theta) = \dots = \int_A \int_B f(x|\theta) \pi(d\theta) \lambda^{(n)}(dx)$

• $\mu_n(A) = \delta(A \times \Theta) = \dots = \int_A \mu_n(x) \lambda^{(n)}(dx)$

• II $\delta(\cdot \times B) \ll \mu_n(\cdot) \Rightarrow RN: \delta(A \times B) = \int_A \pi(x, B) \mu_n(dx) = \dots = \int_A \int_{\Theta} \pi(x, B) f(x|\theta) \pi(d\theta) \lambda^{(n)}(dx)$

• $\delta(A \times B) = P(X \in A, \theta \in B) = \int \dots \Rightarrow \pi(x, B) = \pi(\theta \in B | X_n = x)$

• I = II

$$5. \quad P(|\text{err}(\tau)| > c) = P\left(\frac{|\text{err}(\tau)|}{\sqrt{\sigma^2/T}} > \frac{c}{\sqrt{\sigma^2/T}}\right) = 2\left(1 - \Phi\left(\frac{c}{\sqrt{\sigma^2/T}}\right)\right) \approx 2\left(1 - \Phi\left(\frac{c}{\sqrt{\sigma^2 \alpha/T}}\right)\right)$$
