



Exercises on Estimation of reliability parameters from experimental data

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α Percentile values of the $\chi^2(f)$ distribution

α	0.905	0.925	0.950	0.975	0.990	0.995	0.999
f							
1	0.0001	0.0001	0.0001	0.0001	0.0001	0.0001	0.0001
2	0.0100	0.0501	0.1013	0.2107	0.3387	0.4753	0.7173
3	0.0781	0.2148	0.3527	0.5844	0.8788	1.2128	1.9238
4	0.2048	0.4844	0.7143	1.0645	1.4868	1.9238	2.7768
5	0.4844	0.8538	1.1483	1.6023	2.0783	2.4759	3.3527
6	0.6757	1.2367	1.5708	2.2027	2.7928	3.1526	4.1017
7	0.8788	1.6759	2.0345	2.6758	3.3074	3.6781	4.6026
8	1.1344	2.1790	2.5748	3.1526	3.7722	4.1512	5.0088
9	1.3844	2.7003	3.1147	3.6965	4.3028	4.6821	5.4878
10	1.6354	3.2462	3.6965	4.2929	4.8783	5.2093	5.9894
11	1.8864	3.8162	4.3028	4.9563	5.4922	5.8233	6.5126
12	2.1374	4.4097	4.9348	5.6413	6.1311	6.4621	7.0574
13	2.3884	5.0267	5.5914	6.3523	6.7941	7.1251	7.6238
14	2.6394	5.6672	6.2723	7.0893	7.4801	7.8111	8.2112
15	2.8904	6.3312	6.9772	7.8517	8.1881	8.5171	8.8196
16	3.1414	7.0187	7.7062	8.6393	8.9181	9.2421	9.4486
17	3.3924	7.7297	8.4592	9.4523	9.6701	9.9841	10.0986
18	3.6434	8.4642	9.2362	10.2903	10.4421	10.7501	10.7696
19	3.8944	9.2222	10.0372	11.1523	11.2341	11.5421	11.4516
20	4.1454	10.0037	10.8612	12.0383	12.0461	12.3641	12.1536
21	4.3964	10.8087	11.7072	12.9483	12.8781	13.2061	12.8756
22	4.6474	11.6372	12.5752	13.8823	13.7301	14.0681	13.6176
23	4.8984	12.4892	13.4652	14.8393	14.6021	14.9401	14.3796
24	5.1494	13.3647	14.3772	15.8193	15.4941	15.8221	15.1516
25	5.4004	14.2627	15.3012	16.8213	16.4061	16.7241	15.9336
26	5.6514	15.1832	16.2462	17.8453	17.3381	17.6461	16.7356
27	5.9024	16.1262	17.2122	18.8913	18.2901	18.5881	17.5576
28	6.1534	17.0917	18.1992	19.9593	19.2621	19.5501	18.3996
29	6.4044	18.0797	19.2072	21.0493	20.2541	20.5321	19.2616
30	6.6554	19.0892	20.2362	22.1613	21.2661	21.5341	20.1436
35	7.4664	21.5597	23.3612	25.1853	24.2901	24.5581	23.0456
40	8.2674	24.4332	26.7512	29.1913	28.2961	28.5641	26.8776
45	9.0684	27.7067	30.4412	33.1873	32.2821	32.5501	30.7096
50	9.8694	31.5262	34.4812	37.1833	36.2481	36.5161	34.5416

Exercise 1

A feedwater pump of an energy production plant is characterized by a constant failure rate. In order to assess the performance of the pump, a right censored test of the first type is carried out on 10 identical pumps. The duration of the test is $t_0 = 500$ hours for each pump. Table 1 reports the observed failure times.

Table 1. Results of the reliability tests on the 10 pumps.

Pump 1	Pump 2	Pump 3	Pump 4	Pump 5	Pump 6	Pump 7	Pump 8	Pump 9	Pump 10
205	99	No failure during the test time	91	458	No failure during the test time	No failure during the test time	No failure during the test time	43	35

You are required to:

- 1) estimate the pump failure rate using the method of maximum likelihood;
- 2) what is the 95% two-sided confidence interval of the pump mean time to failure.

Exercise 3

The number of defective rivets, D , on an airplane wing can be assumed to have a Poisson distribution with parameter λ , i.e.,

$$P(D=d) = \frac{\lambda^d e^{-\lambda}}{d!}, \quad d = 0, 1, 2, \dots$$

A random sample of n wings is observed and (d_1, d_2, \dots, d_n) defective rivets are found.

1. What is $\hat{\lambda}_{ML}$, the maximum likelihood estimator of λ ?
2. Is this estimator unbiased?
3. Find the method-of-moments estimator of λ .

$T = TTT$ (Total Time on Test)
 $r =$ number of failures

	I, fixed t_0	II, fixed r
one-sided (lower)	$\hat{\theta}_1 = \frac{2T}{\chi^2_{\alpha}(2r+2)}$ Weil degrees of freedom percentile	$\hat{\theta}_1 = \frac{2T}{\chi^2_{\alpha}(2r)}$
two-sided (lower and upper)	$(\hat{\theta}_1, \hat{\theta}_2) = \left(\frac{2T}{\chi^2_{1-\alpha/2}(2r+2)}, \frac{2T}{\chi^2_{\alpha/2}(2r+2)} \right)$ $P(\hat{\theta}_1 < MITF < \hat{\theta}_2) = \alpha$	$(\hat{\theta}_1, \hat{\theta}_2) = \left(\frac{2T}{\chi^2_{1-\alpha/2}(2r)}, \frac{2T}{\chi^2_{\alpha/2}(2r)} \right)$

Exercise 2

The failure time of a new type of industrial filter is an exponential random variable having an unknown value λ . A group of twenty filters are being monitored and, at present, their failure times are (in weeks):

1.2, 1.8*, 2.2, 4.1, 5.6, 8.4, 11.8*, 13.4*, 16.2, 21.7, 29*, 41, 42*, 42.4*, 49.3, 60.5, 61*, 94, 98, 99.2,

where an * next to the data means that the filter is still working, whereas an unstarred data point means that the filter failed at that time.

1. What is the maximum likelihood estimate $\hat{\lambda}_{ML}$ of λ .
2. What is the 90% two-sided confidence interval of λ .

Exercise 4

Let p be the probability of failure on demand of a new type of relief valve used in energy production plants. Considering past experience on similar relief valves, an expert suggests that p can have only three values: $p_1 = 5 \cdot 10^{-4}$, $p_2 = 1 \cdot 10^{-3}$, $p_3 = 5 \cdot 10^{-3}$. Furthermore, the expert has observed the operation of similar valves in energy production plants for a long period of time and he proposes to use the following prior distribution for p :

$$\begin{aligned} P(p=p_1) &= 0.2 \\ P(p=p_2) &= 0.6 \\ P(p=p_3) &= 0.2 \end{aligned}$$

The new type of relief valve is then used for 1 year and 2 failures to start over 500 demands are observed. You are required to:

- a) update the probability distribution of p ;
- b) Compute the probability that the new type of valve will have 0 failures out of 3 demands

Exercise 5

The time to failure T (years) of a certain item is an exponential random variable with probability density function:

$$p(t) = \lambda e^{-\lambda t}, t > 0$$

From prior experience we are led to believe that λ is a value of an exponential random variable Λ with probability density function:

$$\pi'(\lambda) = 2e^{-2\lambda}, \lambda > 0$$

1. Estimate the item reliability at a time $t = 1$ year.
2. if we have a sample of 3 item failure times: $(t_1, t_2, t_3) = (1; 2.8; 2.2)$, find the posterior distribution of Λ , the new estimation of the item reliability at time $t = 1$ year and the 95% upper confidence limit of λ (numerical solution of the integral is not required).
3. What is the maximum likelihood estimation of λ and its 95% upper confidence limit using Frequentist statistics?

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Exercise 6

Two independent and identical cables feed an important node in a power distribution network. Assume that the failure time of each cable is an exponential random variable with the (same) failure rate λ . Assuming that the following failure times have been data collected during a right censored test of the first type performed on ten cables with test duration equal to 15 years:

9.1 7.7 5.5 11.7 10.6 7.4 9.1 10.7 8.9 12.5 [in years]

You are required to:

- Q1a) estimate the parameter λ using the method of maximum likelihood;
- Q1b) estimate the 95% two-sided confidence interval of the cable failure rate;
- Q1c) repeat Q1a) following a Bayesian approach. Assume that according to an expert, the prior distribution of λ is:

$$P'(\lambda) = 0.1e^{-0.1\lambda}$$

- Q1d) Consider the following maintenance strategy: each time a cable fails, a maintenance intervention (repair) is immediately performed. During maintenance, the power distribution is interrupted. The cost of the repair is 10000 €, whereas the cost of the power distribution interruption is 30000 €. You are required to estimate the average overall cost of the cables maintenance strategy in one year for the grid owner.

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EXPONENTIAL
- GAMMA

$f(\cdot)$, $E[\cdot]$, $\text{Var}(\cdot)$

#1 (#9)

$$T \sim E(\lambda) : f_T(t) = \lambda e^{-\lambda}, F_T(t) = 1 - e^{-\lambda}, R_T(t) = e^{-\lambda}$$

$$1. L(\lambda) = \prod_{i=1}^6 f(t_i) \prod_{j=1}^4 R(t_j) = (\lambda^6 e^{-(\sum_{i=1}^6 t_i)\lambda}) (e^{-\sum_{j=1}^4 t_j \lambda}) = \lambda^6 e^{-\lambda(\sum_{i=1}^6 t_i + 4t_0)}$$

$$\ell(\lambda) = \log(L(\lambda)) = 6 \log(\lambda) - \lambda \pi \pi$$

$$\frac{\partial \ell(\lambda)}{\partial \lambda} = \frac{6}{\lambda} - \pi \pi = 0 \Rightarrow \lambda = \frac{6}{\pi \pi} = \frac{6}{2931} = 2.047 \cdot 10^{-3}$$

$$2. (\lambda_1^{-1}, \lambda_2^{-1}) = \left[\frac{2\pi}{\chi^2_{1+\alpha}(2r+2)}, \frac{2\pi}{\chi^2_{1-\alpha}(2r)} \right]$$

$$T = \pi \pi = 2931$$

$$\alpha = 0.95$$

$$r = \# \text{ failed} = 6$$

$$\chi^2_{0.975}(14) = 26.1$$

$$\chi^2_{0.025}(12) = 4.40$$

$$\Rightarrow [224.598, 1332.273]$$

$$\leq \text{MTTF} \leq$$

$$\Rightarrow \lambda^{-1} : 7.506 \cdot 10^{-4} \leq \lambda \leq 4.4524 \cdot 10^{-3}$$

#2

$$n=20, r=n-8=12$$

$$\lambda = \frac{r}{\pi \pi} = \frac{12}{702.8} = 0.01707$$

$$\alpha = 0.90$$

$$\chi^2_{0.95}(26) = 38.9$$

$$\chi^2_{0.05}(24) = 13.8$$

$$\Rightarrow \left[\frac{13.8}{2 \cdot 702.8}, \frac{38.9}{2 \cdot 702.8} \right] = [0.00981787, 0.027675]$$

#3

$$L(\lambda) = \prod_{i=1}^n \left(\frac{\lambda^{d_i} e^{-\lambda}}{d_i!} \right) = \frac{\lambda^{\sum d_i} e^{-n\lambda}}{\prod_{i=1}^n (d_i!)}$$

$$1. \log L(\lambda) \propto (\sum d_i) \log(\lambda) - n\lambda \rightarrow \frac{\partial}{\partial \lambda} : \frac{\sum d_i}{\lambda} - n = 0 \Rightarrow \lambda_{MUE} = \frac{\sum d_i}{n}$$

$$2. E[\lambda_{MUE}] = \frac{1}{n} \sum E[d_i] = \frac{1}{n} n\lambda = \lambda \quad \checkmark$$

$$3. E[D] = \frac{\sum d_i}{n} = \text{first moment} = \lambda_{MOM}$$

#4

$$P(p_i | \text{data}) = \frac{P(\text{data} | p_i) P(p_i)}{P(\text{data})}$$

$$P(X \sim \text{Bin}(n=500, p=p_i) = 2)$$

a)

	p	P(p)	P(data p)	P(data p) P(p)	P(p')
p1	5 · 10 ⁻⁴	0.2	0.0243	0.0048	0.0471
p2	10 ⁻³	0.6	0.0758	0.0454	0.4468
p3	5 · 10 ⁻³	0.2	0.2570	0.0514	0.5060

$$b) P(0 \text{ failures out of 3 demands}) =$$

$$= 0.0471 \underbrace{P(X \sim \text{Bi}(3, p_1) = 0)}_{(1-p_1)^3} + 0.4468 \cdot \underbrace{P(X \sim \text{Bi}(3, p_2) = 0)}_{(1-p_2)^3} + 0.5060 \underbrace{P(X \sim \text{Bi}(3, p_3) = 0)}_{(1-p_3)^3}$$

$$= 0.98088$$

#5

$$1. R(1) = P(T > 1) = \int_0^{\infty} P(T > 1 | \lambda) \pi(\lambda) d\lambda = \int_0^{\infty} e^{-\lambda} 2e^{-2\lambda} d\lambda = \frac{2}{3}$$

2. posterior \propto likelihood \cdot prior

$$\propto (\lambda^3 e^{-\lambda \sum t_i}) (2e^{-2\lambda}) = \lambda^3 e^{-8\lambda}$$

$$\text{posterior} = \text{Gamma}(\lambda=8, \alpha=4) \quad : \quad \pi(\lambda | \text{data}) = \frac{8^4 \lambda^{4-1} e^{-8\lambda}}{\Gamma(4)} = \frac{8(8\lambda)^3 e^{-8\lambda}}{\Gamma(4)}$$

$$R(1) = P(T > 1) = \int_0^{\infty} P(T > 1 | \lambda') \pi(\lambda') d\lambda' = \int_0^{\infty} (e^{-\lambda'}) \left(\frac{8^4}{6} \lambda'^3 e^{-8\lambda'} \right) d\lambda' = 0.6243$$

Bayesian statistic upper bound:

$$\int_0^{\lambda_{0.95}} \pi(\lambda') d\lambda' = 0.95 \Rightarrow \lambda_{0.95} = 0.9676$$

$$3. \lambda_{MUE} = \frac{n}{\sum T_i} = \frac{3}{6} = \frac{1}{2}$$

$$95\% \text{ for } \sum T_i : \frac{2\tau}{\chi_{0.95}^2(6)} = \frac{12}{12.6} \Rightarrow \lambda_{0.95} = \frac{12.6}{12} = 1.05$$

#6 (TDE)

$$a) \text{ exponential} \rightarrow \lambda_{MUE} = \frac{r}{\sum T_i} = \frac{10}{93.2} = 0.107296137$$

$$b) \alpha = 0.95 \rightarrow \lambda \in \left[\frac{\chi_{1-\alpha}^2(2r)}{2 \cdot \sum T_i}, \frac{\chi_{1+\alpha}^2(2r+2)}{2 \cdot \sum T_i} \right] = [0.059012876, 0.183476395]$$

numbers are possibly wrong

$$c) \pi(\lambda) = 0.1 e^{-0.1\lambda}$$

posterior \propto likelihood \cdot prior

$$\propto (\lambda^{10} e^{-\lambda \cdot 11}) (0.1 e^{-0.1\lambda})$$

$$\propto 0.1 \lambda^{10} e^{-\lambda \cdot 93.3} \rightarrow \text{Gamma}(\lambda=93.3, \alpha=11)$$

$$\pi(\lambda | \text{data}) = \frac{(93.3)^{11} \lambda^{10} e^{-93.3\lambda}}{\Gamma(11)}$$

$$d) \text{ For a cable } \lambda_{MUE} = 0.107296137 \quad (\sim P(\lambda))$$

Since the system has two cables:

$$\lambda_{\text{system}} = 2 \lambda_{MUE} = 0.214592275 \quad (\sim P(\lambda))$$

If we have k failures the cost will be: $k(10'000 + 30'000)$,
since k obeys a Poisson the expectation is

$$\lambda = \lambda_{\text{system}}$$

So the overall cost in one year is:

$$\lambda_{\text{system}} (10'000 + 30'000) \cdot 1 = 8583.690987$$