

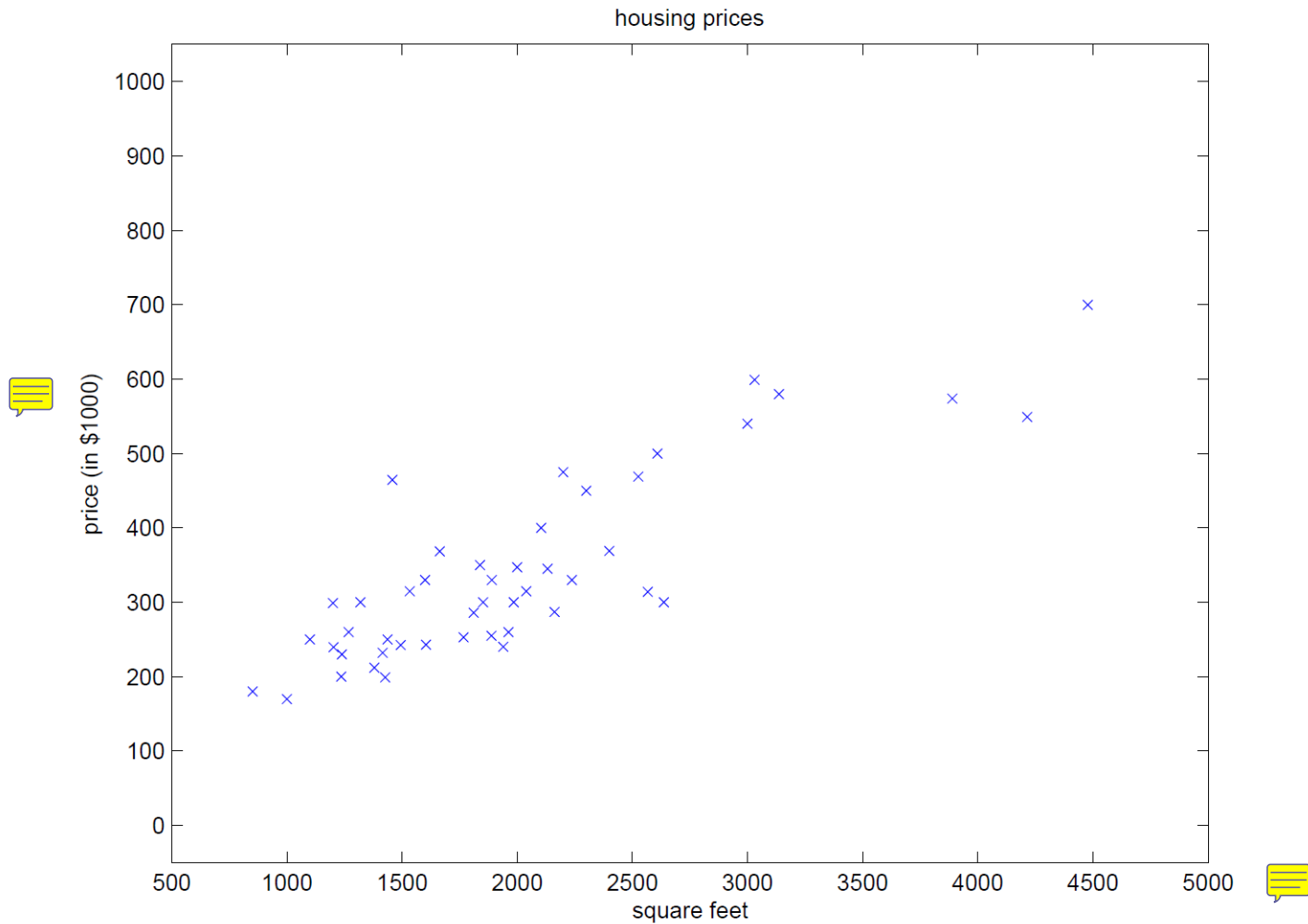
LEC003 Demand Forecasting

VG441 SS2020

Cong Shi
Industrial & Operations Engineering
University of Michigan

Linear Regression

- “Best fitting line”



Linear Regression

- (Linear) Hypothesis Function:

$$h_{\theta}(x) = \theta_0 x_0 + \theta_1 x_1 + \cdots + \theta_n x_n$$

- Minimizing the least-squares cost:

$$J(\theta_{0\dots n}) = \frac{1}{2m} \sum_{i=1}^m \left(h_{\theta}(x^{(i)}) - y^{(i)} \right)^2$$

Matrix Derivation

- Turn everything into matrix notation

$$h_{\theta}(x) = \theta^T x \quad \theta = \begin{pmatrix} \theta_0 \\ \theta_1 \\ \dots \\ \theta_n \end{pmatrix} \in \mathbb{R}^{n+1} \quad x = \begin{pmatrix} x_0 \triangleq 1 \\ x_1 \\ \dots \\ x_n \end{pmatrix} \in \mathbb{R}^{n+1}$$

- Design matrix

$$X = \begin{bmatrix} x^{(1)} \\ x^{(2)} \\ \vdots \\ x^{(m-1)} \\ x^{(m)} \end{bmatrix} = \begin{bmatrix} 1 & x_1^{(1)} & \dots & \dots & \dots & x_n^{(1)} \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ 1 & x_1^{(m)} & \dots & \dots & \dots & x_n^{(m)} \end{bmatrix} \quad \text{}$$

Normal Equation

- Cost function

$$J(\theta) = \frac{1}{2m} (X\theta - y)^T (X\theta - y)$$

- Throwing out the constant

$$J(\theta) = \theta^T X^T X \theta - 2(X\theta)^T y + y^T y$$

- The “famous” normal equation

$$\frac{\partial J}{\partial \theta} = 2X^T X \theta - 2X^T y = 0$$

$$X^T X \theta = X^T y$$

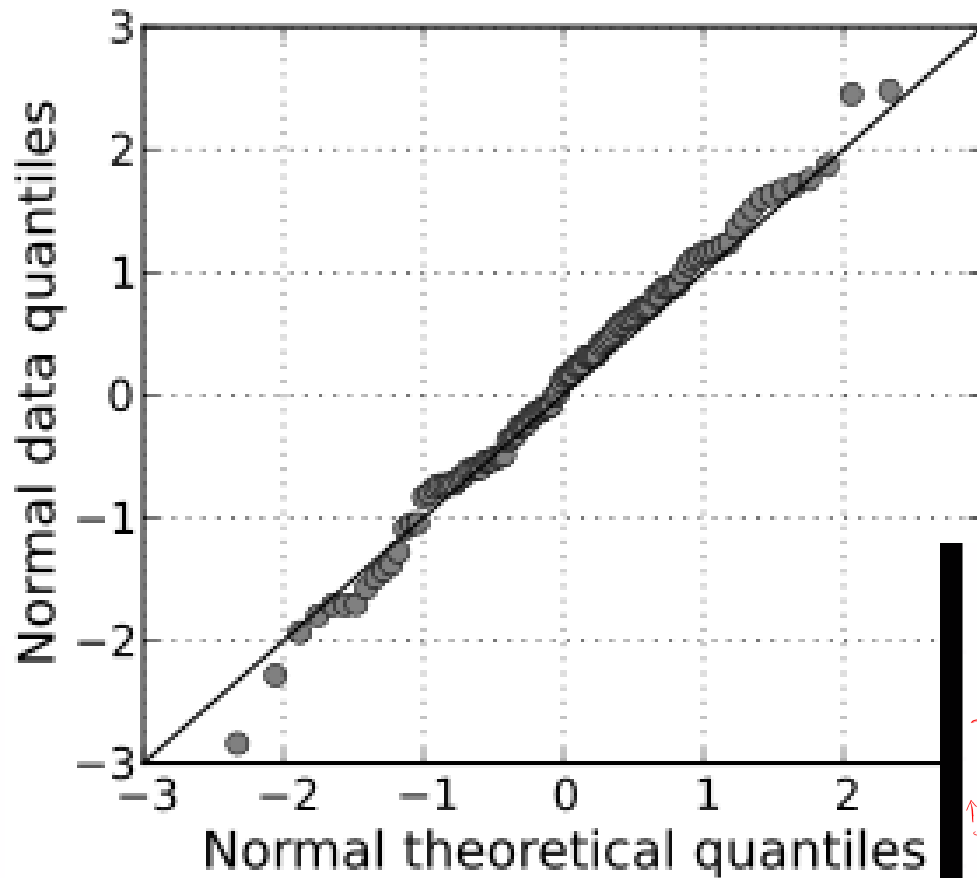
$$\theta = (X^T X)^{-1} X^T y$$

$$\left\{ \begin{array}{l} \frac{\partial x^T B}{\partial x} = B \\ \frac{\partial x^T b}{\partial x} = b \\ \frac{\partial x^T x}{\partial x} = 2x \end{array} \right. \quad \left\{ \begin{array}{l} \frac{\partial x^T B x}{\partial x} = 2Bx \end{array} \right. \quad \begin{array}{l} \text{Symmetric} \\ B = B^T \end{array}$$

Residuals

- The residual should be normally distributed

$$y^{(i)} = \theta^T x^{(i)} + \epsilon^{(i)}$$



WTO: R^2 not sig.

$$F = \frac{[SS(\text{mean}) - SS(fit)] / (\# \text{ param} - 1)}{SS(fit) / (n - \# \text{ param})}$$

variance that
→ can be explained
by features.

Variance not explained by features.

F large \rightarrow small p-value \rightarrow ~~more likely~~
 \rightarrow more likely to reject H_0

lin. reg. Two most important questions

11) R^2 r-squared $\in [0, 1]$

(2) Is R^2 significant? $SS(\text{fit}) = a_1^2 + a_2^2 + a_3^2 + a_4^2$

$R^2: y$
 $y = (\theta_0 + \theta_1 x)$
 $\text{Var}(f) = \frac{SS(fit)}{n}$
 \uparrow knowledge in SCM
 x explains R^2 amount of variation time spent on VG44
 $\text{var}(\text{mean}) = \frac{SS(\text{mean})}{n}$ (without considering x)
 $R^2 = \frac{\text{var}(\text{mean}) - \text{var}(f)}{\text{var}(\text{mean})}$

Probabilistic Interpretation

- Deriving the log-likelihood function

$$\epsilon^{(i)} = y^{(i)} - \theta^T x^{(i)} \sim \mathcal{N}(0, \sigma^2)$$

$$p(\epsilon^{(i)}) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(\epsilon^{(i)})^2}{2\sigma^2}\right)$$

$$p(y^{(i)} | x^{(i)}; \theta) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(y^{(i)} - \theta^T x^{(i)})^2}{2\sigma^2}\right)$$



Probabilistic Interpretation

- Deriving the log-likelihood function

$$L(\theta) = L(\theta; X, \vec{y}) = p(\vec{y}|X; \theta)$$

$$\begin{aligned} L(\theta) &= \prod_{i=1}^m p\left(y^{(i)}|x^{(i)}; \theta\right) \\ &= \prod_{i=1}^m \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{\left(y^{(i)} - \theta^T x^{(i)}\right)^2}{2\sigma^2}\right) \end{aligned}$$

Probabilistic Interpretation

Out[12]: OLS Regression Results

Dep. Variable:	HOUSING PRICE	R-squared:	0.959			
Model:	OLS	Adj. R-squared:	0.958			
Method:	Least Squares	F-statistic:	891.1			
Date:	Tue, 12 May 2020	Prob (F-statistic):	0.00			
Time:	23:16:28	Log-Likelihood:	-1523.8			
No. Observations:	506	AIC:	3074			
Df Residuals:	493	BIC:	3129			
Df Model:	13					
Covariance Type:	nonrobust					
	coef	std err	t	P> t	[0.025	0.975]
CRIM	-0.0916	0.034	-2.675	0.008	-0.159	-0.024
ZN	0.0487	0.014	3.379	0.001	0.020	0.077
INDUS	-0.0038	0.064	-0.059	0.953	-0.130	0.123
CHAS	2.8564	0.904	3.160	0.002	1.080	4.633
NOX	-2.8808	3.359	-0.858	0.392	-9.481	3.720
RM	5.9252	0.309	19.168	0.000	5.318	6.533
AGE	-0.0072	0.014	-0.523	0.601	-0.034	0.020

- Maximum Likelihood Estimator (M

$$\ell(\theta) = \log L(\theta)$$

$$= \log \prod_{i=1}^m \frac{1}{\sqrt{2\pi}\sigma} \exp \left(-\frac{(y^{(i)} - \theta^T x^{(i)})^2}{2\sigma^2} \right)$$

$$= \sum_{i=1}^m \log \frac{1}{\sqrt{2\pi}\sigma} \exp \left(-\frac{(y^{(i)} - \theta^T x^{(i)})^2}{2\sigma^2} \right)$$

$$= m \log \frac{1}{\sqrt{2\pi}\sigma} - \frac{1}{\sigma^2} \cdot \frac{1}{2} \sum_{i=1}^m (y^{(i)} - \theta^T x^{(i)})^2$$

96%

test

$H_0: \theta_1 = 0$ (Crim rate does not matter)

$H_A: \theta_1 \neq 0$

Prob is small

→ reject H_0


Python Time!

- `from sklearn import linear_model`
- `from statsmodels.api import OLS`



- AR(p) 

$$\underline{X_t} = \mu + \sum_{i=1}^p \underline{\varphi_i} \underline{X_{t-i}} + \varepsilon_t$$

- MA(q) 

$$X_t = \mu + \underline{\varepsilon_t} + \sum_{i=1}^q \theta_i \varepsilon_{t-i}$$


- ARMA(p,q) 

$$X_t = \mu + \varepsilon_t + \sum_{i=1}^p \varphi_i X_{t-i} + \sum_{i=1}^q \theta_i \varepsilon_{t-i}$$

ARIMA(p,d,q)



- d is the degree of difference, e.g., d = 1


$$Y_t = X_t - X_{t-1}$$

$$Y_t = \mu + \varepsilon_t + \sum_{i=1}^p \varphi_i Y_{t-i} + \sum_{i=1}^q \theta_i \varepsilon_{t-i}$$

- How about d = 2?



```
df = pd.read_csv('PowerUsage.csv', names=['value'], header=0)
df.head()

Out[2]:
  value
0    88
1    84
2    85
3    85
4    84

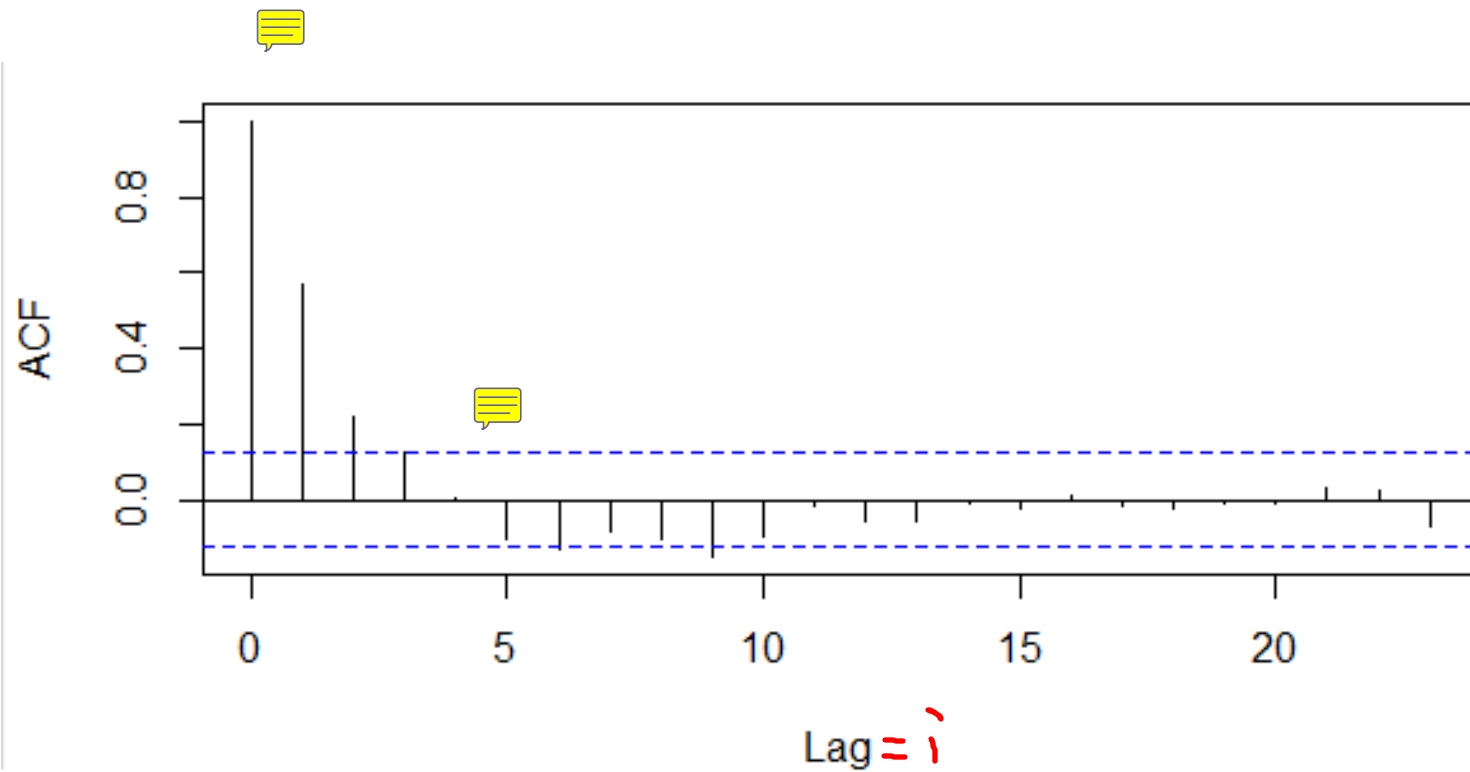
In [3]: from statsmodels.tsa.stattools import adfuller
        from numpy import log
        result = adfuller(df.value.dropna())
        print('ADF Statistic: %f' % result[0])
        print('p-value: %f' % result[1])

ADF Statistic: -2.464240
p-value: 0.124419

In [ ]: # Original Series
fig, axes = plt.subplots(3, 2, sharex=True)
axes[0, 0].plot(df.value); axes[0, 0].set_title('Original Series')
print('ADF Statistic: %f' % result[0])
```

hypo test for stationarity
Augmented Dickey-Fuller test.
H₀: t.s. not stationary
H₁: t.s. stationary
Large
fail to reject H₀

ACF/PACF

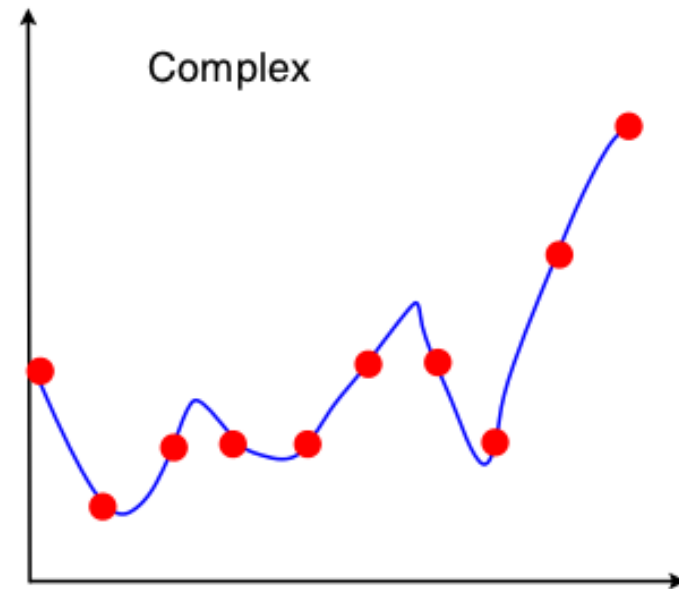
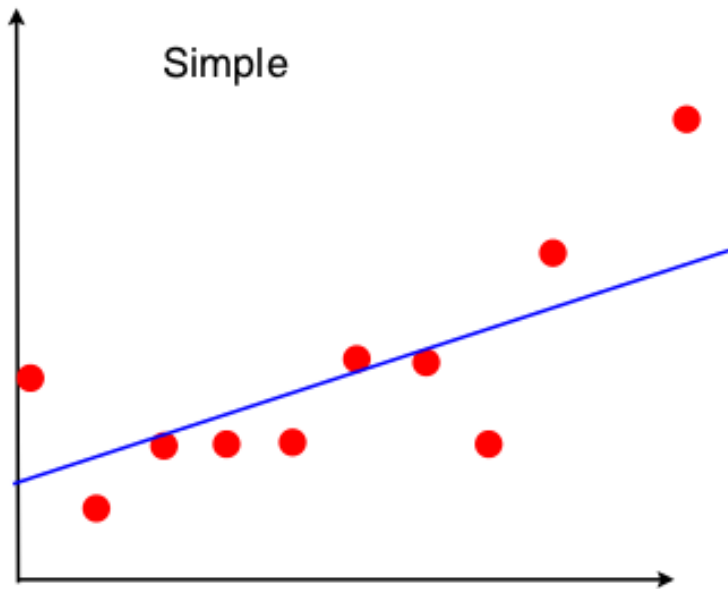


Python Time!

- `statsmodels.tsa.arima_model`










Bias-Variance Tradeoff



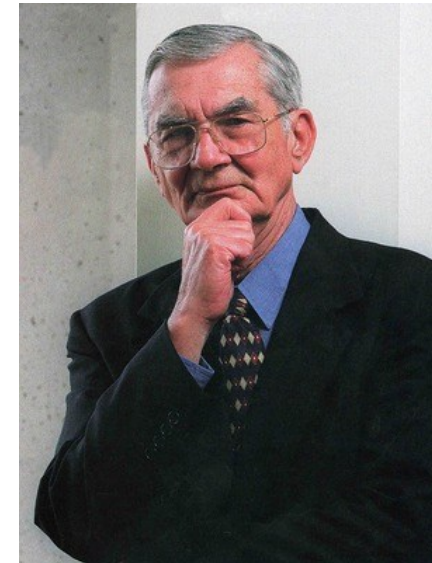
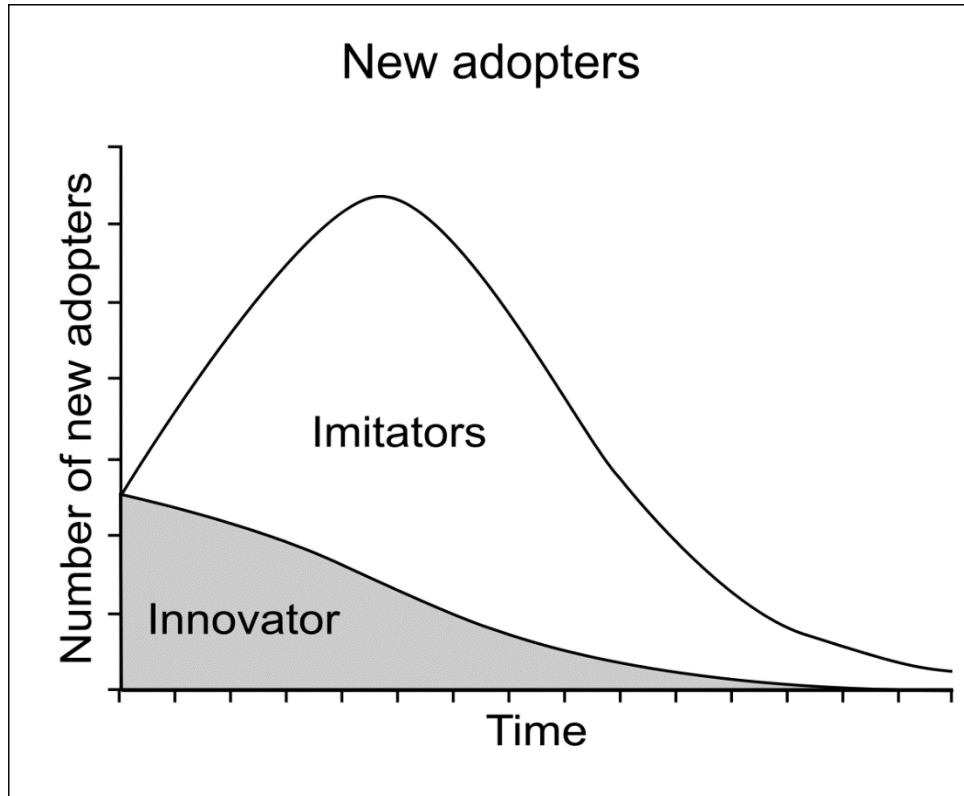
Kaggle Competition

All Competitions

Active Completed InClass		All Categories ▾ Default Sort ▾	
	Jigsaw Multilingual Toxic Comment Classification Use TPUs to identify toxicity comments across multiple languages Featured • a month to go • Code Competition • 862 Teams		\$50,000
	M5 Forecasting - Accuracy Estimate the unit sales of Walmart retail goods Featured • 2 months to go • 3589 Teams		\$50,000
	M5 Forecasting - Uncertainty Estimate the uncertainty distribution of Walmart unit sales. Featured • 2 months to go • 389 Teams		\$50,000
	University of Liverpool - Ion Switching Identify the number of channels open at each time point Research • 16 days to go • 2333 Teams		\$25,000
	TReNDS Neuroimaging Multiscanner normative age and assessments prediction with brain function, structure, and connectivity Research • 2 months to go • 275 Teams		\$25,000
	ALASKA2 Image Steganalysis Detect secret data hidden within digital images Research • 2 months to go • 237 Teams		\$25,000
	Prostate cANcer graDe Assessment (PANDA) Challenge Prostate cancer diagnosis using the Gleason grading system Featured • 2 months to go • Code Competition • 309 Teams		\$25,000

Bass Diffusion Model

- How about projecting sales of new products?



Frank Bass

Bass Diffusion Model

Cumulative purchase probability of a random customer $F(t)$

Purchase probability at time t $f(t) = F'(t)$

The rate of purchase at time t (given no purchase so far)

$$\frac{f(t)}{1 - F(t)} = p + qF(t)$$

Coefficient of innovation

Coefficient of imitation

Bass Solution

$$\frac{dF/dt}{1-F} = p + qF$$

$$\frac{dF}{dt} = p + (q-p)F - qF^2$$

$$\int \frac{1}{p + (q-p)F - qF^2} dF = \int dt$$

$$\begin{aligned} \frac{1}{(p+qF)(1-F)} &= \frac{A}{p+qF} + \frac{B}{1-F} \\ &= \frac{A - AF + pB + qFB}{(p+qF)(1-F)} \\ &= \frac{A + pB + F(qB - A)}{(p+qF)(1-F)} \end{aligned}$$

$$A = q/(p+q)$$

$$B = 1/(p+q)$$

Bass Solution

$$\int \frac{1}{(p + qF)(1 - F)} dF = \int dt$$

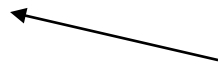
$$\int \left(\frac{A}{p + qF} + \frac{B}{1 - F} \right) dF = t + c_1$$

$$\int \left(\frac{q/(p + q)}{p + qF} + \frac{1/(p + q)}{1 - F} \right) dF = t + c_1$$

$$\frac{1}{p + q} \ln(p + qF) - \frac{1}{p + q} \ln(1 - F) = t + c_1$$

$$\frac{\ln(p + qF) - \ln(1 - F)}{p + q} = t + c_1$$

Boundary Condition



$$t = 0 \Rightarrow F(0) = 0$$

$$t = 0 \Rightarrow c_1 = \frac{\ln p}{p + q}$$

$$F(t) = \frac{p(e^{(p+q)t} - 1)}{pe^{(p+q)t} + q}$$

Bass Solution



Calibration

- Sales in any period are $s(t) = mf(t)$
- Cumulative sales up to time t are $S(t) = mF(t)$

$$\frac{s(t)/m}{1 - S(t)/m} = p + qS(t)/m$$

$$s(t) = [p + qS(t)/m][m - S(t)]$$

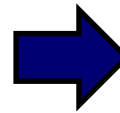
$$s(t) = \beta_0 + \beta_1 S(t) + \beta_2 S(t)^2 \quad (BASS)$$

$$\beta_0 = pm$$

$$\beta_1 = q - p$$

$$\beta_2 = -q/m$$

Conduct a linear regression!



$$m = \frac{-\beta_1 \pm \sqrt{\beta_1^2 - 4\beta_0\beta_2}}{2\beta_1}$$

$$p = \frac{\beta_0}{m}; \quad q = -m\beta_2$$

Python Time!

- `from sklearn import linear_model`
- `from statsmodels.api import OLS`



Discrete Choice Model

- Multinomial Logit Model

$$U_{ni} = V_{ni} + \epsilon_{ni}$$

Gumbel

$$f(x) = e^{-x} e^{-e^{-x}}$$

$$F(x) = e^{-e^{-x}}$$



$$\begin{aligned} P_{ni} &= \mathbb{P}(U_{ni} > U_{nj} \quad \forall j \neq i) \\ &= \mathbb{P}(V_{ni} + \epsilon_{ni} > V_{nj} + \epsilon_{nj} \quad \forall j \neq i) \end{aligned}$$

Discrete Choice Model

- Multinomial Logit Model

Gumbel $f(x) = e^{-x} e^{-e^{-x}}$
 $F(x) = e^{-e^{-x}}$

$$\begin{aligned} P_{ni} &= \int (P_{ni} | \epsilon_{ni}) f(\epsilon_{ni}) d\epsilon_{ni} \\ &= \int (P_{ni} | \epsilon_{ni}) e^{-\epsilon_{ni}} e^{-e^{-\epsilon_{ni}}} d\epsilon_{ni} \\ &= \int \left(\prod_{j \neq i} e^{-e^{-(\epsilon_{ni} + V_{ni} - V_{nj})}} \right) e^{-\epsilon_{ni}} e^{-e^{-\epsilon_{ni}}} d\epsilon_{ni} \end{aligned}$$



$$P_{ni} = \frac{e^{V_{ni}}}{\sum_j e^{V_{nj}}}$$