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# *Redes de Computadores*

## **Delay Models in Computer Networks**

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- 
- » *What are the common multiplexing strategies?*
  - » *What is a Poisson process?*
  - » *What is the Little theorem?*
  - » *What is a queue?*
  - » *What is the meaning of service time  $1/\mu$  in a queue of packets?*
  - » *What is the meaning of traffic intensity  $\rho$  in a queue model?*
  - » *What is the probability of a M/M/1 queue being in a given state  $n$  ?*
  - » *What is the mean number of clients in a M/M/1 queue? What is the mean waiting time in a M/M/1 queue? What is the relationship between  $N$  and  $\rho$  in a M/M/1 queue?*
  - » *What are the differences between M/M/1 and M/G/1 queues? How to estimate mean number of packets and mean delay in a M/G/1 queue?*
  - » *How to model a network of transmission lines? How to calculate the mean number of packets and mean delay in this case?*
  - » *What is a Jackson Network? Why is it important?*

# *Multiplexing Traffic on a Link*

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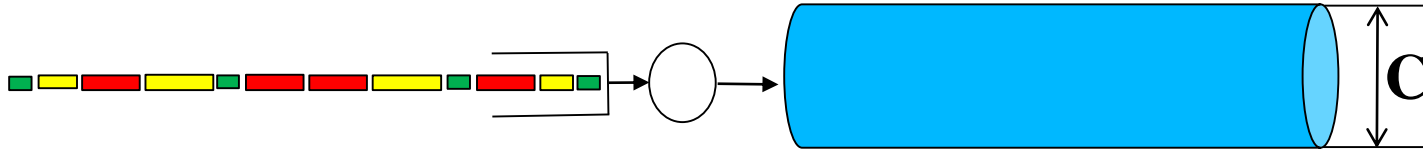


- ♦ Communication link
  - » Bit pipe with a given capacity  $C$  (bit/s)
  - » Link capacity  $\rightarrow$  rate at which bits are transmitted to the link
  - » Link may transport multiplexed traffic streams
  
- ♦ Multiplexing strategies
  - » Statistical Multiplexing
  - » Frequency Division Multiplexing
  - » Time Division Multiplexing
  
- ♦ Multiplexing strategy affects traffic delay

# *Statistical Multiplexing*

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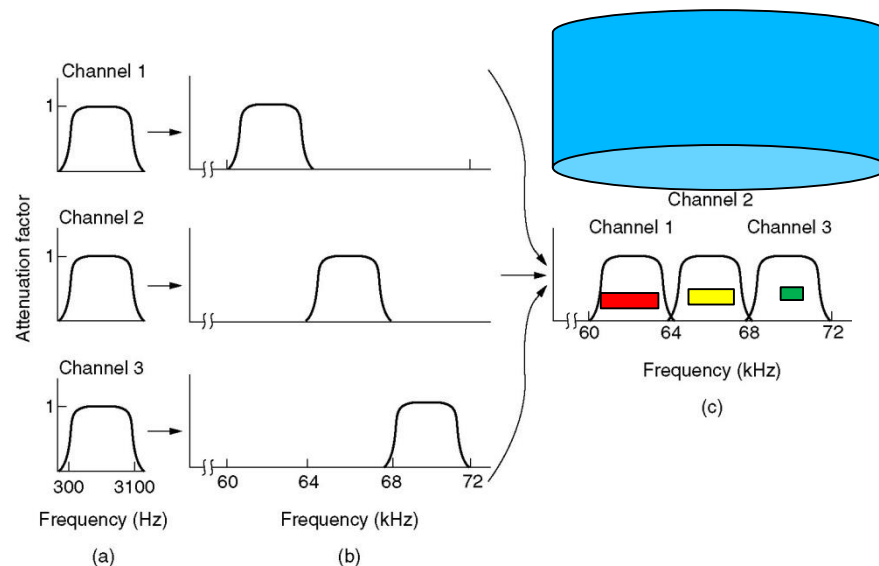
- ♦ Packets of all traffic streams merged in a single queue
- ♦ Packets transmitted on a first-come first-served basis
- ♦ Time required to transmit a packet of length  $L \rightarrow T_{\text{frame}} = L/C$



# *FDM – Frequency Division Multiplexing*

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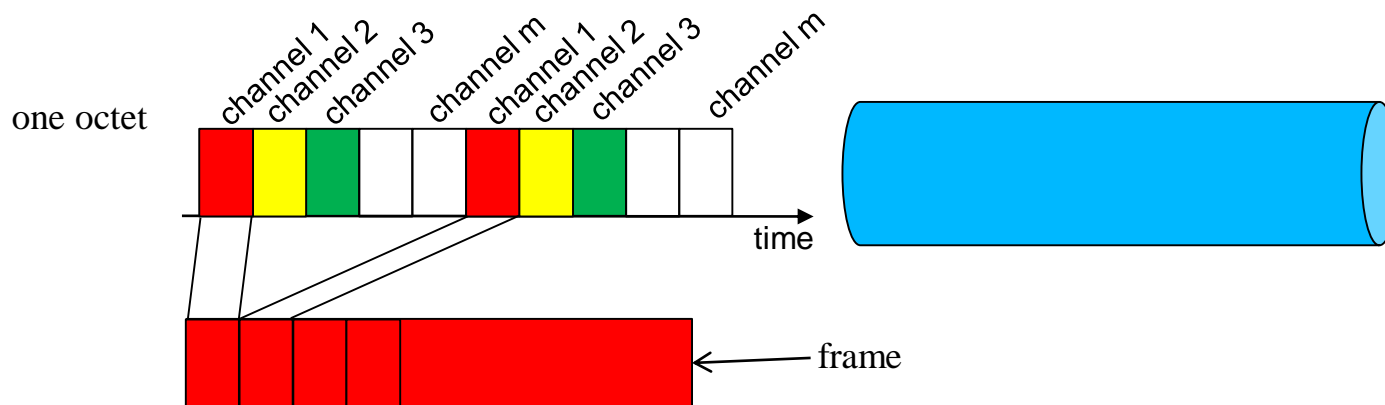
- ♦ Link capacity  $C$  subdivided into  $m$  portions
- ♦ Channel bandwidth  $W$  subdivided into  $m$  channels of  $W/m$  Hz
- ♦ Capacity of each channel  $\rightarrow C/m$
- ♦ Time required to transmit a packet of length  $L \rightarrow T_{\text{frame}} = Lm/C$



# *TDM – Time Division Multiplexing*

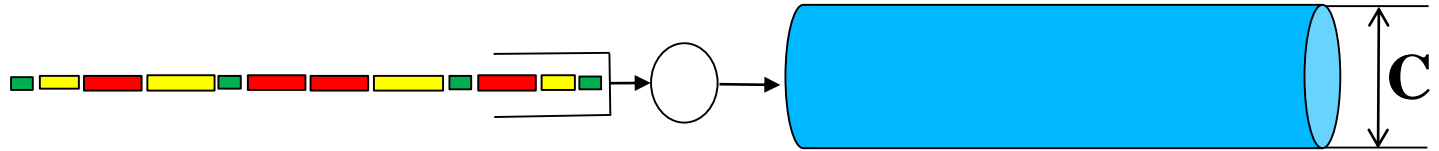
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- ♦ Time axis divided into  $m$  slots of fixed length  
(usually one octet long)
- ♦ Communication  $\rightarrow$   $m$  channels with capacity  $C/m$
- ♦ Time required to transmit a packet of length  $L \rightarrow T_{\text{frame}} = Lm/C$



# Delay on Computer Networks

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## ♦ Delay

- » Important performance parameter in computer networks
- » Characterized using queue models

## ♦ Queue model

- » Customers arrive at random times to obtain service
- » Customer → packet to be transmitted through a link
- » Serve a packet = transmit a packet
- » Service time → **packet transmission time** =  $T_{\text{pac}(\text{frame})} = L/C$

## ♦ Queue models enable the quantification of

- » Average number of customers/packets in the network
- » Average delay per packet → waiting plus service times

# Computer Networks Modeled as Queue Networks

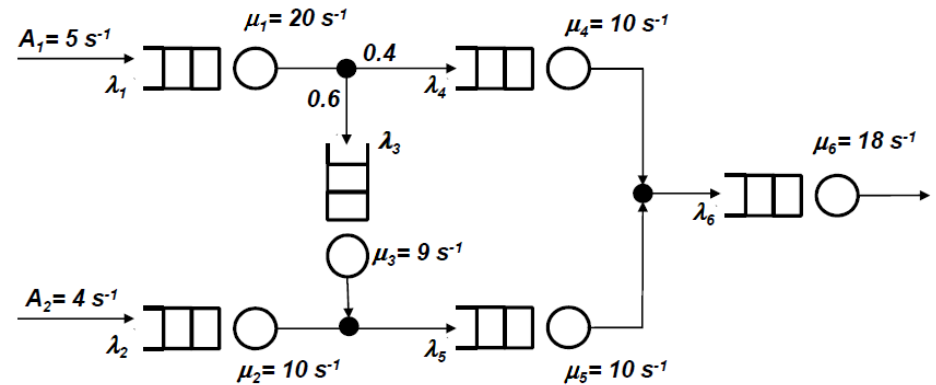
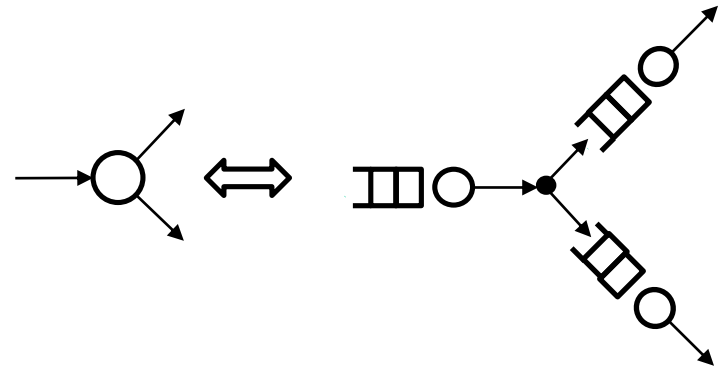
Mobile network

Global ISP

Home network

Regional ISP

Institutional network





# Poisson Distribution and Poisson Process

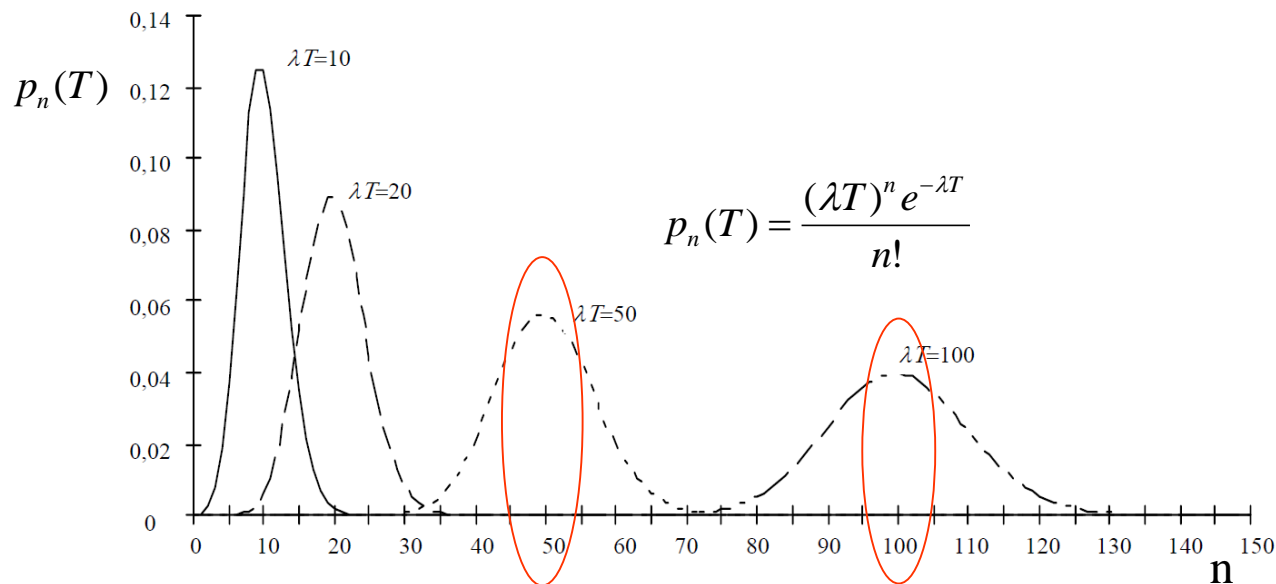
- ◆ **Poisson distribution** with parameter  $m$

$$P[N = n] = p_n = \frac{m^n e^{-m}}{n!}, \quad n = 0, 1, \dots \quad E[N] = \text{Var}[N] = m$$

- ◆ **Poisson process**

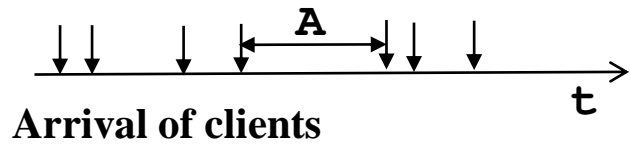
»  $\lambda T = m$  , (e.g.  $\lambda \rightarrow$  arrivals/s )

»  $P[\text{n arrivals in interval } T] = p_n(T) = p_n = \frac{(\lambda T)^n e^{-\lambda T}}{n!} \quad E[N] = \text{Var}[N] = \lambda T$



# Inter-Arrival Interval $A$ – Statistical Characterization

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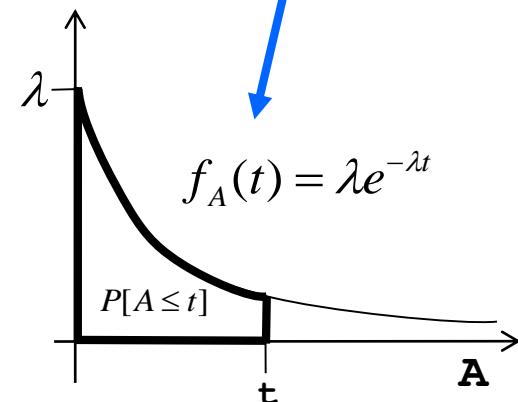
$A$  – time interval between the  
arrival of consecutive clients

$$F_A(t) = P[A \leq t] = 1 - P[A > t] = 1 - p_0(t) = 1 - e^{-\lambda t}$$

$$f_A(t) = pdf = \frac{\partial F_A(t)}{\partial t} = \lambda e^{-\lambda t} \quad \leftarrow \text{Exponential distribution}$$

$$E[A] = 1/\lambda$$

$$Var[A] = 1/\lambda^2$$



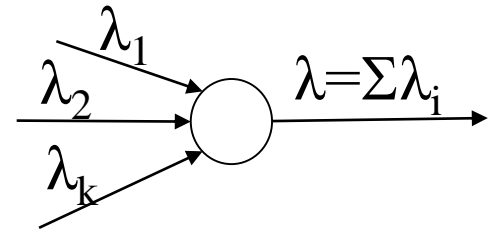
- 
- ♦ What is the difference between deterministic and Poisson arrivals?

# Markov Process - Properties

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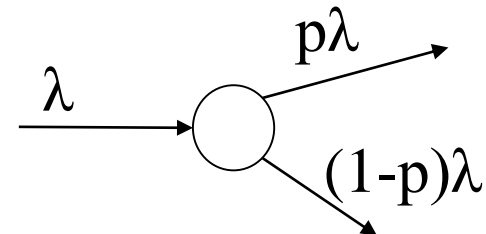
## ♦ Merging Property

- »  $A_1, A_2, \dots, A_k$  are independent Poisson Processes with rates  $\lambda_1, \lambda_2, \dots, \lambda_k$
- »  $A = \sum A_i$  still is a Poisson process, with rate  $\lambda = \sum \lambda_i$



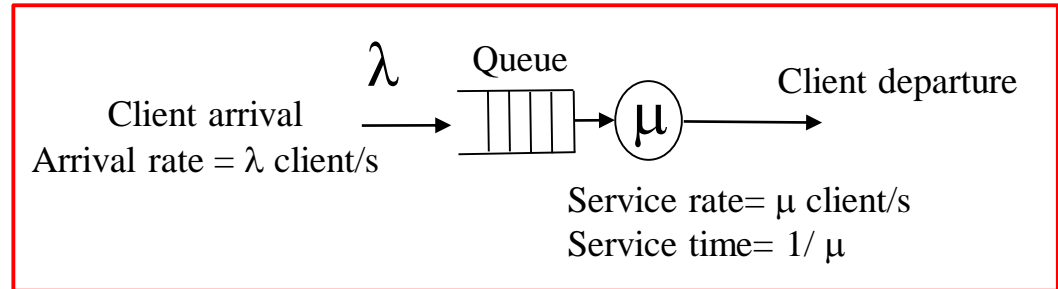
## ♦ Splitting property

- » Packets arrive to a router according to a Poisson Process  $(A, \lambda)$
- » They are routed randomly to two output lines with probabilities **p** and **1-p**
- » Packets leaving the router still are Poisson Processes, characterized by  $(A, p\lambda)$  and  $(A, (1-p)\lambda)$



# Queue Model

- ♦ Queue – model used for
  - » Customers waiting in line
  - » Packets in a network
- ♦ Used to determine
  - » Average number of clients in the system  $\rightarrow N$
  - » Average delay experienced by a client  $\rightarrow T$
- ♦ Queue characterized in terms of
  - »  $\lambda$  - arrival rate of client (average number of clients per time unit)
  - »  $\mu$  - service rate (average number of clients the server processes per time unit)
  - »  $\rho = \lambda / \mu$  – traffic intensity (occupation of the server)
- ♦ Kendall notation  $\rightarrow \mathbf{A/S/s/K}$ 
  - » A – arrival statistical process
  - » S – service statistical process
  - » s – number of servers
  - » K – capacity of the system in buffers



# *Little's Theorem*

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## ◆ $N = \lambda T$

- »  $N$  - average number of clients in a system
- »  $T$  - average amount of time a client spends in the system
- »  $\lambda$  - arrival rate of clients to the system

## ◆ $T = T_w + T_s$

- »  $T_w$  - time a client waits in the queue for being served
- »  $T_s$  - service time

## ◆ $N = N_w + N_s$

- »  $N_w$  - number of clients waiting in the queue for being served
- »  $N_s$  - number of clients being served

## ◆ $N_w = \lambda T_w$

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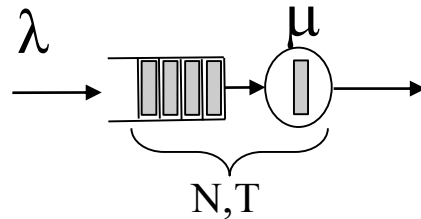
$$N_w = \lambda T_w \rightarrow T_w = N_w / \lambda$$

- ♦ The (mean) time a client has to wait before being served ( $T_w$ ) depends on the number of clients waiting ( $N_w$ ) and on the arrival rate of clients ( $\lambda$ )
- ♦ No dependence on the service rate?!
- ♦ Can you explain it?

# *Little's Theorem*

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- ♦ Can be applied to a single Queue



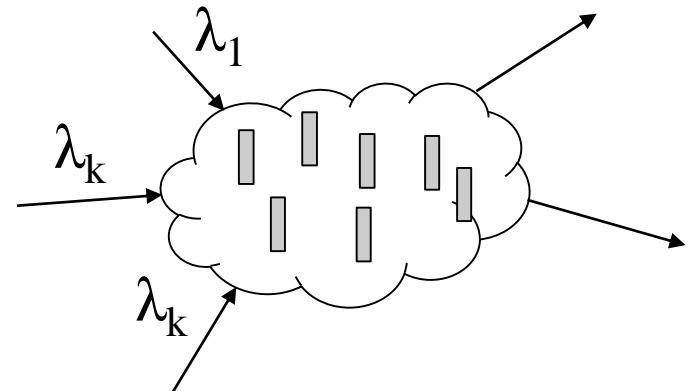
- ♦ Can be applied to a complex system

- » For each stream  $i \rightarrow N_i = \lambda_i T_i$

- » For the system:

$$\lambda = \sum \lambda_i \quad N = \sum N_i$$

$$T = (\sum N_i) / (\sum \lambda_i) \rightarrow T = N / \lambda$$



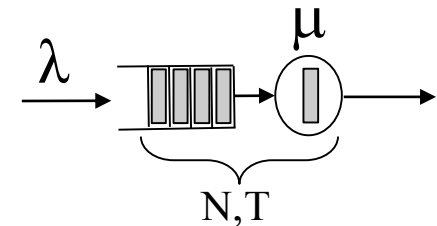


# M/M/1 Queue

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## ♦ M/M/1

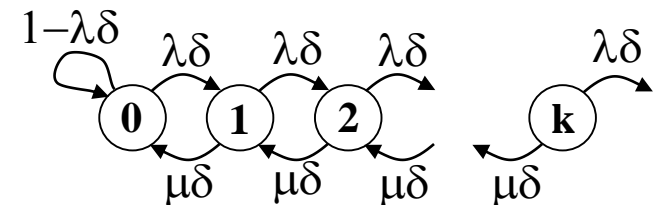
- » Poisson arrival, exponential service time



## ♦ Modeled by a Markov Chain

- » State  $\mathbf{k}$  -  $k$  clients in the queue
- »  $p(i, j)$  – probability of transition from state  $i$  to state  $j$
- » When  $\delta \rightarrow 0$

$$\begin{aligned} p(i, i+1) &= \lambda\delta & p(i, i-1) &= \mu\delta \\ p(i, i) &= 1 - \lambda\delta - \mu\delta & p(0, 0) &= 1 - \lambda\delta \\ p(i, j) &= 0 \text{ for other values } i, j \end{aligned}$$



## » Birth-death chain

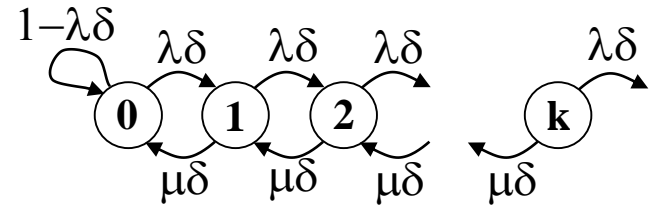
- Transitions between adjacent states
- $\lambda\delta$  and  $\mu\delta$  become flow rates between states

$$\begin{aligned} p(i, i+1) &= p_1(\delta) = (\lambda\delta)e^{-\lambda\delta} \approx \lambda\delta \\ p(0, 0) &= p_0(\delta) = e^{-\lambda\delta} \approx 1 - \lambda\delta \end{aligned}$$

# M/M/1 Queue – Equilibrium Analysis

- ♦  $P(j)$  – probability of the Markov chain be in state  $j$
- ♦ Markov Chain - global balance equations

$$P(j) \sum_{i \neq j}^{\infty} p(j, i) = \sum_{i \neq j}^{\infty} P(i) p(i, j)$$



- ♦ In the case of M/M/1

$$P(0)\lambda\delta = P(1)\mu\delta \Rightarrow P(1) = \rho P(0)$$

$$P(2) = \rho P(1) = \rho^2 P(0)$$

$$P(n) = \rho^n P(0)$$

$$\sum_{i=0}^{\infty} P(i) = 1$$

$$\sum_{i=0}^{\infty} \rho^i P(0) = \frac{P(0)}{1 - \rho} = 1$$

$$P(0) = 1 - \rho$$

$$P(n) = \rho^n (1 - \rho)$$

# *M/M/1 Queue*

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- ♦ Average Queue size N

$$N = \sum_{n=0}^{\infty} nP(n) = \sum_{n=0}^{\infty} n\rho^n(1-\rho) = \frac{\rho}{1-\rho} \quad N = \sum_{n=0}^{\infty} nP(n) = \frac{\rho}{1-\rho} = \frac{\lambda/\mu}{1-\lambda/\mu} = \frac{\lambda}{\mu-\lambda}$$

- ♦ Average amount of time the client spends in the system, T

» Little's formula,  $T=N/\lambda \quad \rightarrow \quad T = \frac{1}{\mu-\lambda}$

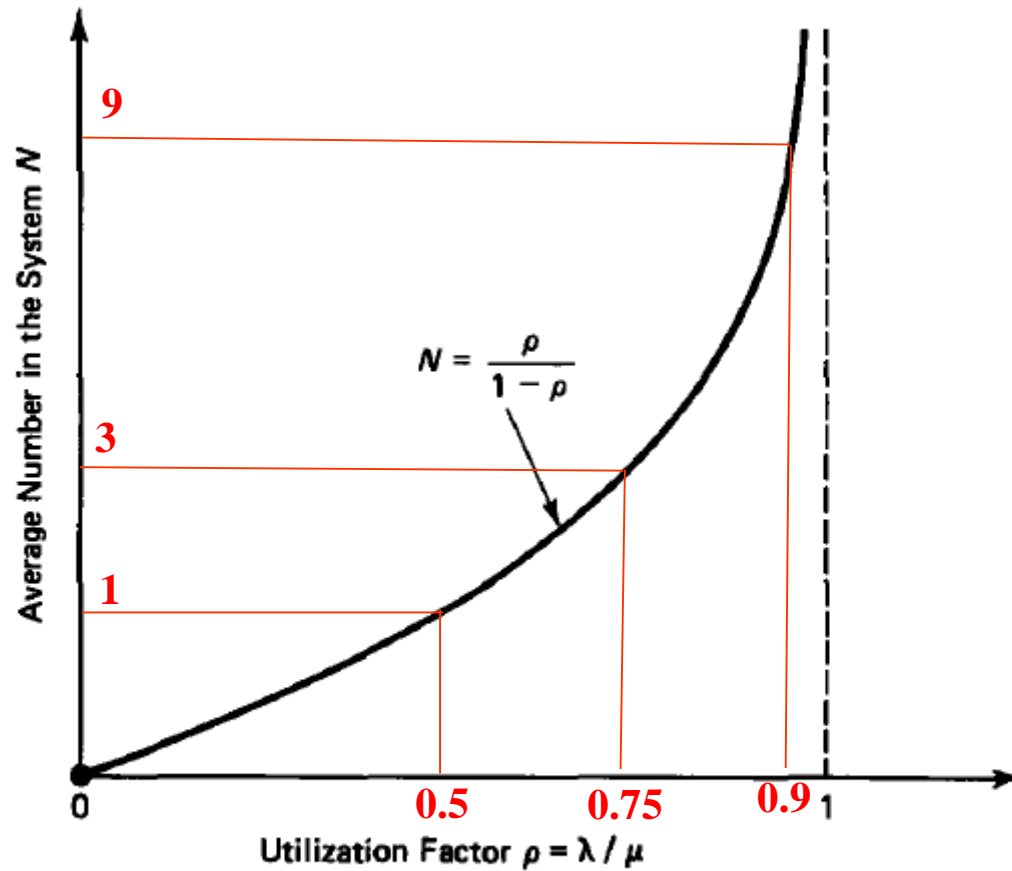
- ♦ Average waiting time  $T_w \rightarrow T_w = T - T_s = \frac{1}{\mu-\lambda} - \frac{1}{\mu} = \frac{\rho}{\mu(1-\rho)}$

- ♦ Average number of clients waiting in the queue,  $N_w$

$$N_w = T_w \lambda = \frac{\lambda}{\mu-\lambda} - \frac{\lambda}{\mu} = N - \rho$$

## *M/M/1 Queue – $N=f(\rho)$*

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**Figure 3.6** The average number in the system versus the utilization factor in the *M/M/1* system. As  $\rho \rightarrow 1$ ,  $N \rightarrow \infty$ .

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♦ M/M/1:  $\rho=0.9 \rightarrow N=9$

♦ Why have clients to wait if the server is busy only 90% of his time?

♦ What would happen for D/D/1,  $\rho=0.9$ ?

# *Packet Length, Service Time, Speed*

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- » 100 packet/s are required to be transmitted through a link
- » Packets arrive according to a Poisson process
- » Packet lengths are exponentially distributed  $\rightarrow E[L]=10^4$  bit/packet
- » Link has capacity  $C=10$  Mbit/s

## ◆ Then

- » Arrival rate:  $\lambda=100$  packet/s
- » Service rate:  $\mu=C/E[L]=10^7/10^4=10^3$  packet/s
- »  $\rho=\lambda/\mu=0.1$ ,  $N=\rho/(1-\rho)=1/9$ ,  $T=N/\lambda=1/900$  s

## ◆ Assume now: $\lambda'=10\lambda$ and $C'=10C \rightarrow \mu'=10C/E[L]=10\mu$

- » Then  $\rho'=\rho$  and  $N'=N$  but  $T'=N'/\lambda'=T/10$

The speed of the system increases!

## *M/M/1/B Queue*

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- ♦ M/M/1 queue has limited capacity (B buffers)
  - » Packets can be lost
  - » Probability of packet being lost =  $P(B)$  → Queue is full
- ♦ Analysis similar to M/M/1

$$\sum_{i=0}^B P(i) = 1 \qquad P(n) = \rho^n P(0)$$

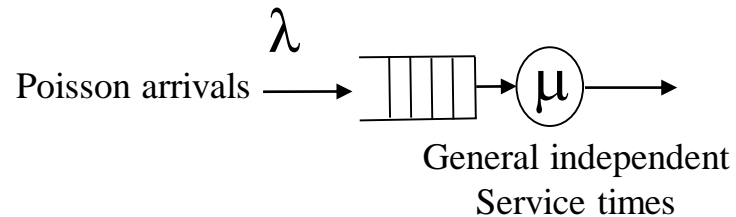
$$P(0) = \frac{1 - \rho}{1 - \rho^{B+1}} \qquad P(B) = \frac{(1 - \rho)\rho^B}{1 - \rho^{B+1}}$$

- ♦ Particular cases

$$\rho = 1, \quad P(B) = \frac{1}{B+1} \qquad \rho \gg 1, \quad P(B) \approx \frac{\rho - 1}{\rho} = \frac{\lambda - \mu}{\lambda}$$

# *M/G/1 Queue*

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- ♦ Poisson arrivals at rate  $\lambda$
- ♦ Service time has arbitrary distribution with given  $E[X]$  and  $E[X^2]$ 
  - » Service times Independent and Identically Distributed (IID)
  - » Independent of arrival times
  - »  $E[\text{service time}] = E[X] = 1/\mu$
  - » Single Server queue



# *M/G/1 Queue – Pollaczek-Khinchin (P-K) Formula*

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$$T_w = \frac{\lambda E[X^2]}{2(1-\rho)}$$

- ♦ where  $\rho = \lambda/\mu = \lambda E[X]$  = line utilization
- ♦ From Little's Theorem
  - »  $N_w = \lambda T_w$
  - »  $T = T_w + E[X] = T_w + 1/\mu$
  - »  **$N = \lambda T = \lambda(T_w + 1/\mu) = N_w + \rho$**

# *M/G/1 Queue – Proof of (P-K) Formula*

$$T_w = \frac{\lambda E[X^2]}{2(1-\rho)}$$

## ◆ Let

- $T_w(i)$  - waiting time in queue of  $i^{\text{th}}$  arrival
- $R(i)$  – residual service time seen by the  $i^{\text{th}}$  arrival
- $N_w(i)$  – number of clients found in queue by the  $i^{\text{th}}$  arrival
- $X(i)$  – service time of the  $i^{\text{th}}$  arrival

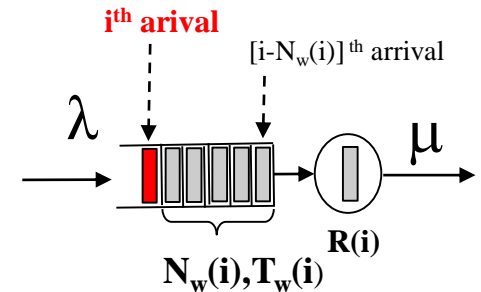
$$T_w(i) = \sum_{j=i-N_w(i)}^{i-1} X(j) + R(i)$$

$$E[T_w(i)] = T_w = E[N_w(i)] \times E[X(i)] + E[R(i)] = \frac{N_w}{\mu} + E[R(i)]$$

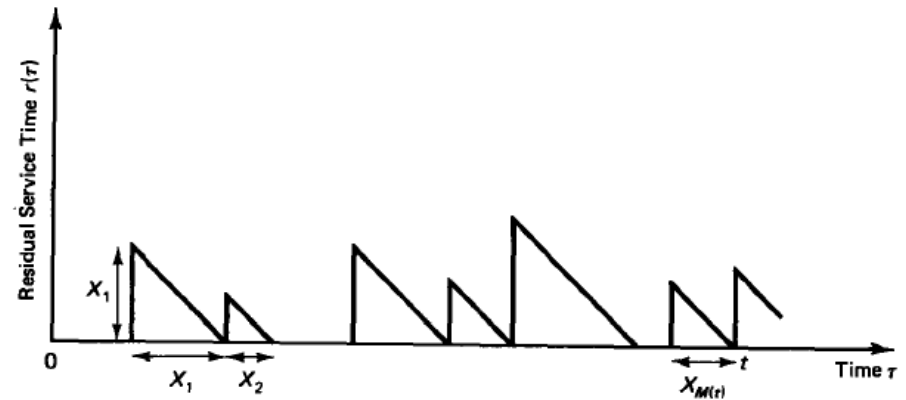
» Using Little's formula

$$T_w = \frac{\lambda T_w}{\mu} + E[R(i)]$$

$$T_w = \frac{E[R(i)]}{1-\rho}$$



# *M/G/1 Queue – Proof of (P-K) Formula*



**Figure 3.10** Derivation of the mean residual service time. During period  $[0, t]$ , the time average of the residual service time  $r(\tau)$  is

$M(t)$  – number of clients served by time  $t$

$$E[R(i)] = R_t = \frac{1}{t} \int_0^t r(\tau) d\tau = \frac{1}{t} \sum_{i=1}^{M(t)} \frac{X_i^2}{2} = \frac{M(t)}{2t} \sum_{i=1}^{M(t)} \frac{X_i^2}{M(t)}$$

$$t \rightarrow \infty, \quad \frac{M(t)}{t} = \lambda = \text{arrival rate} = \text{departure rate}$$

$$E[R(i)] = \frac{\lambda}{2} \sum_{i=1}^{M(t)} \frac{X_i^2}{M(t)} = \frac{\lambda}{2} \times E[X^2]$$

$$T_w = \frac{E[R(i)]}{1 - \rho}$$

$$T_w = \frac{\lambda E[X^2]}{2(1 - \rho)}$$

## *M/G/1 Examples*

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- ♦ Case M/M/1

- »  $E[X] = 1/\mu$  ;  $E[X^2] = 2/\mu^2$

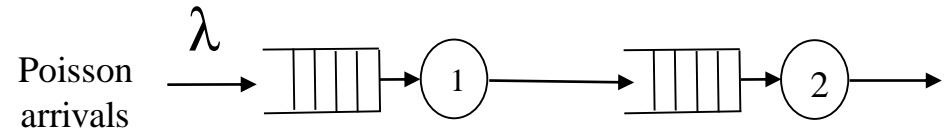
$$T_w = \frac{\lambda}{\mu^2(1-\rho)} = \frac{\rho}{\mu(1-\rho)}$$

- ♦ Case M/D/1

- » Deterministic, constant service time  $1/\mu$

- »  $E[X] = 1/\mu$  ;  $E[X^2] = 1/\mu^2$

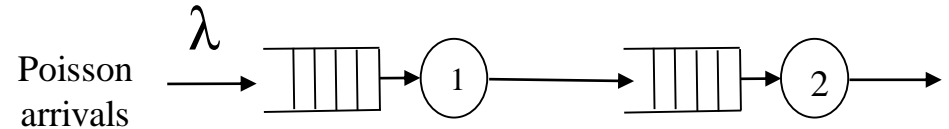
$$T_w = \frac{\lambda}{2\mu^2(1-\rho)} = \frac{\rho}{2\mu(1-\rho)}$$



- ♦ Assume Queue 1 is M/D/1.
- ♦ Can the arrival of packets to Queue 2 be described as a Poisson process?

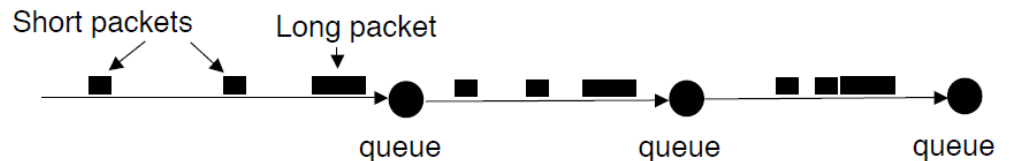
# Networks of Transmission Lines - Problems

## ♦ Case 1



- » Arrival to  $Q_1 \rightarrow$  Poisson,  $\lambda$
- » Assume constant packet length  $\rightarrow Q_1 = M/D/1$
- » Arrival to  $Q_2$  is not Poisson;  $\lambda_2 < \mu_2 \rightarrow 1/\lambda_2 > 1/\mu_2$   
 $\rightarrow$  no waiting at  $Q_2$

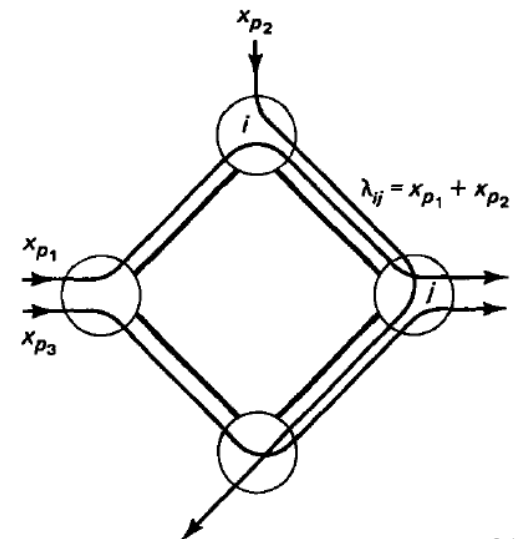
## ♦ Case 2



- »  $Q_1 = M/M/1$
- » arrival to  $Q_2$  strongly related to packet length
- » long packets require long service at each node
- » shorter packets will catch up long packets  $\rightarrow$  interarrival times change  
 $\rightarrow Q_2$  cannot be modeled as  $M/M/1$

# Kleinrock Independence Approximation

- ◆ Merging several packet streams on a transmission line  
restores independence of interarrival times and packet lengths
- ◆ M/M/1 can be used to model each communication link
- ◆ Approximation good for
  - » systems involving Poisson stream arrivals at the entry points
  - » packet lengths nearly exponentially distributed
  - » densely connected networks
  - » Moderate to heavy traffic loads



# Kleinrock Independence Approximation

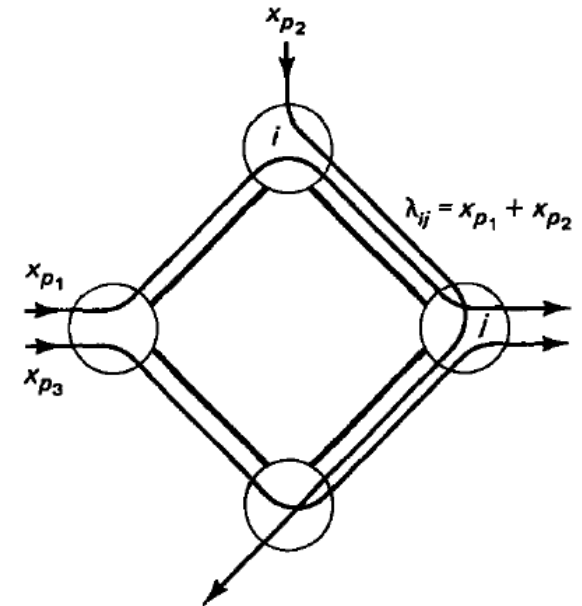
- ♦ Let
  - »  $x_p$  = arrival rate of packets along path  $p$
  - »  $\lambda_{ij}$  = arrival rate of packets to link  $(i,j)$
  - »  $\mu_{ij}$  = service rate on link  $(i,j)$
- ♦ Link queues  $\rightarrow$  independent M/M/1 queues

$$\lambda_{ij} = \sum_{\text{all } p \text{ traversing link}(i,j)} x_p \quad \rho_{ij} = \frac{\lambda_{ij}}{\mu_{ij}} \quad N_{ij} = \frac{\rho_{ij}}{1 - \rho_{ij}}$$

- ♦ And
  - »  $N$  = Average number of packets in network
  - »  $T$  – Average packet delay in network

$$N = \sum_{i,j} N_{ij} \quad \lambda = \sum_{\text{all path } p} x_p = \text{total external arrival rate}$$

$$T = \frac{N}{\lambda}$$

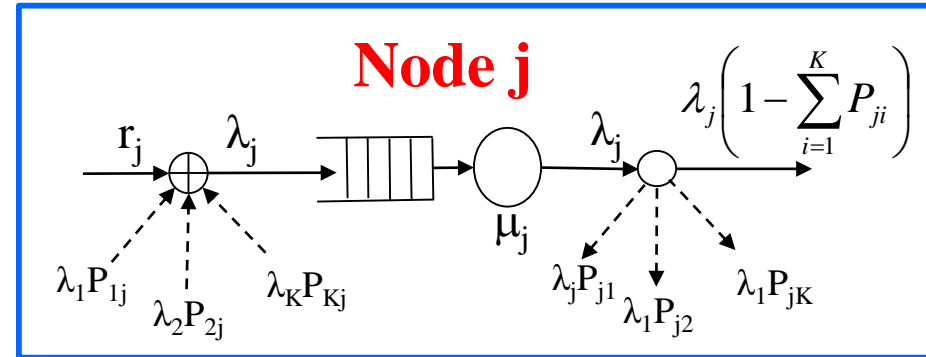




# Jackson Networks

- ◆ Arrival rate at node  $j$

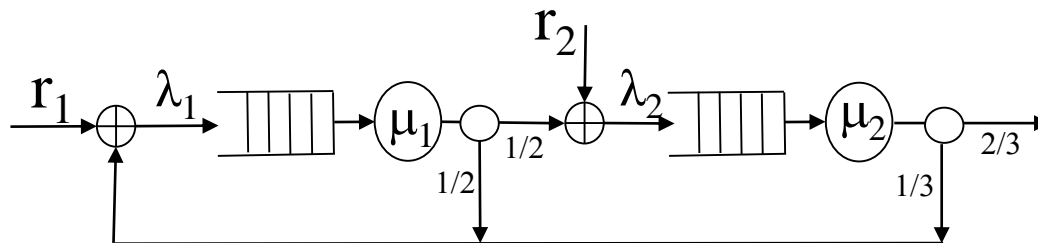
$$\lambda_j = r_j + \sum_{i=1}^K \lambda_i P_{ij} \quad , j = 1, 2, \dots, K$$



- ◆ Independent routing of packets

- » When a packet leaves node  $i$  it comes to node  $j$  with probability  $P_{ij}$
- » Packets can loop inside network
- » Packet leaves the system at node  $j$  with probability

$$P = 1 - \sum_{i=1}^K P_{ji}$$



# *Jackson Networks*

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- ♦ Let the state of the system be defined by  $\vec{n} = (n_1, n_2, \dots, n_K)$   
 $n_j$  – number of clients in  $Q_j$
- ♦ Jackson's theorem:  $P(\vec{n}) = \prod_{j=1}^K P_j(n_j) = \prod_{j=1}^K \rho_j^{n_j} (1 - \rho_j)$ , where  $\rho_j = \frac{\lambda_j}{\mu_j}$ 
  - » State of  $Q_j$  ( $n_j$ ) is independent  $\left( \prod_{j=1}^K \right)$  of state of other queues
  - » Similar to independent M/M/1 queues!
  - » Similar to Kleinrock's independence
- ♦ Again, by Little's theorem

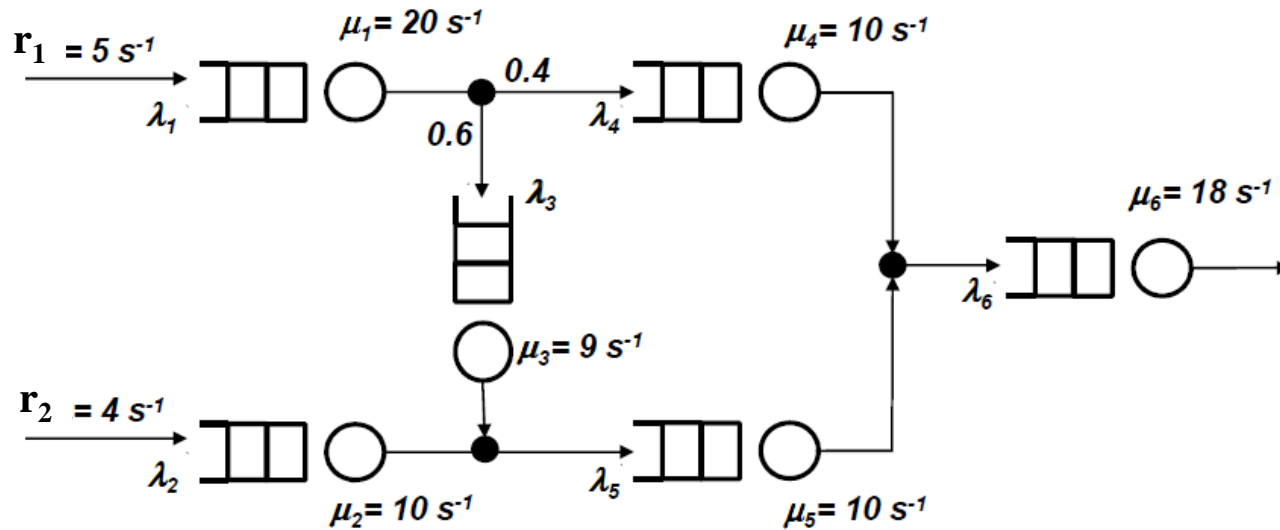
$$N_j = \frac{\rho_j}{1 - \rho_j} \quad N = \sum_{j=1}^K N_j \quad \lambda = \sum_{j=1}^K r_j \quad T = \frac{N}{\lambda}$$

---

Networks of transmission lines / Jackson networks

Give examples of networks that could be described by these models

# Jackson Network - Example



$$\lambda = \sum_{i=1}^6 r_i = 9 \text{ s}^{-1}$$

$$N = \sum_{i=1}^6 N_i = 5.08$$

$$T = \frac{N}{\lambda} = \frac{5.08}{9} = 0.56 \text{ s}$$

Queue $i$	$\mathbf{r}_i$ ( $\text{s}^{-1}$ )	$\lambda_i$ ( $\text{s}^{-1}$ )	$\mu_i$ ( $\text{s}^{-1}$ )	$\rho_i = \lambda_i/\mu_i$	$\mathbf{N}_i = \rho_i/(1 - \rho_i)$
1	5	5	20	0.25	0.33
2	4	4	10	0.40	0.67
3	-	3	9	0.33	0.50
4	-	2	10	0.20	0.25
5	-	7	10	0.70	2.33
6	-	9	18	0.50	1

# *Homework*

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1. Review slides
2. Read *Bertsekas&Gallager*
  - » Sections 3.1, 3.2, 3.3, 3.5, 3.6, 3.8
3. Answer questions at moodle