

Linear model for regression

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In this term, we introduce classical linear regression model.

- ① Least square
- ② Subset Selection
 - ① Best subset selection
 - ② Forward subset selection
 - ③ Backward subset selection
- ③ Shrinkage method
 - ① Lasso regression
 - ② Ridge regression
 - ③ Least angle regression
- ④ Derived input directions
 - ① Principal component regression
 - ② Partial least square

Linear regression models and least square

Input vector: $X^T = (x_1, \dots, x_p)$.

The linear regression model has the form:

$$f(X) = \beta_0 + \sum_{j=1}^p x_j \beta_j$$

For a set of training data: $(x_1, y_1), \dots, (x_N, y_N)$ to estimate β ,
where $x_i = (x_{i1}, \dots, x_{ip})^T$

$$RSS(\beta) = \sum_{i=1}^N (y_i - f(x_i))^2 \quad (1)$$

$$= \sum_{i=1}^N (y_i - \beta_0 - \sum_{j=1}^p x_{ij} \beta_j)^2 \quad (2)$$

In vector notation, we have

$$\hat{\beta} = (X^T X)^{-1} X^T y$$

$$\hat{y} = X\hat{\beta} = X(X^T X)^{-1} X^T y$$

Gauss Markov Theorem: the least squares estimates of the parameters β have the smallest variance among all linear unbiased estimates.

Benifits: Better prediction accuracy and better interpretation.

- 1 With subset selection we retain only a subset of the variables, and eliminate the rest from the model.
- 2 Least squares regression is used to estimate the coefficients of the inputs that are retained.

Subset selection

Algorithm 6.1 *Best subset selection*

1. Let \mathcal{M}_0 denote the *null model*, which contains no predictors. This model simply predicts the sample mean for each observation.
 2. For $k = 1, 2, \dots, p$:
 - (a) Fit all $\binom{p}{k}$ models that contain exactly k predictors.
 - (b) Pick the best among these $\binom{p}{k}$ models, and call it \mathcal{M}_k . Here *best* is defined as having the smallest RSS, or equivalently largest R^2 .
 3. Select a single best model from among $\mathcal{M}_0, \dots, \mathcal{M}_p$ using cross-validated prediction error, C_p (AIC), BIC, or adjusted R^2 .
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图 1: Best subset selection

- ① simple and conceptually appealing
- ② computational limitations

Subset selection

Algorithm 6.2 *Forward stepwise selection*

1. Let \mathcal{M}_0 denote the *null* model, which contains no predictors.
 2. For $k = 0, \dots, p - 1$:
 - (a) Consider all $p - k$ models that augment the predictors in \mathcal{M}_k with one additional predictor.
 - (b) Choose the *best* among these $p - k$ models, and call it \mathcal{M}_{k+1} . Here *best* is defined as having smallest RSS or highest R^2 .
 3. Select a single best model from among $\mathcal{M}_0, \dots, \mathcal{M}_p$ using cross-validated prediction error, C_p (AIC), BIC, or adjusted R^2 .
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2: forward stepwise subset selection

Algorithm 6.3 *Backward stepwise selection*

1. Let \mathcal{M}_p denote the *full* model, which contains all p predictors.
 2. For $k = p, p - 1, \dots, 1$:
 - (a) Consider all k models that contain all but one of the predictors in \mathcal{M}_k , for a total of $k - 1$ predictors.
 - (b) Choose the *best* among these k models, and call it \mathcal{M}_{k-1} . Here *best* is defined as having smallest RSS or highest R^2 .
 3. Select a single best model from among $\mathcal{M}_0, \dots, \mathcal{M}_p$ using cross-validated prediction error, C_p (AIC), BIC, or adjusted R^2 .
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3: backward stepwise subset selection

Ridge regression

Ridge regression:

$$\hat{\beta}^{\text{ridge}} = \underset{\beta}{\operatorname{argmin}} \left\{ \frac{1}{2} \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^p x_{ij} \beta_j \right)^2 + \lambda \sum_{j=1}^p \beta_j^2 \right\}.$$

Equivalent version:

$$\hat{\beta}^{\text{ridge}} = \underset{\beta}{\operatorname{argmin}} \left\{ \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^p x_{ij} \beta_j \right)^2 \right\},$$

subject to $\sum_{j=1}^p \beta_j^2 \leq t.$

Solution: $\hat{\beta}^{\text{ridge}} = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y}$

Lasso regression

Lasso regression:

$$\hat{\beta}^{\text{lasso}} = \underset{\beta}{\operatorname{argmin}} \left\{ \frac{1}{2} \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^p x_{ij} \beta_j \right)^2 + \lambda \sum_{j=1}^p |\beta_j| \right\}$$

Equivalent version:

$$\hat{\beta}^{\text{lasso}} = \underset{\beta}{\operatorname{argmin}} \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^p x_{ij} \beta_j \right)^2$$

subject to $\sum_{j=1}^p |\beta_j| \leq t.$

Lasso vs Ridge

| Estimator | Formula |
|-------------------------|---|
| Best subset (size M) | $\hat{\beta}_j \cdot I(\hat{\beta}_j \geq \hat{\beta}_{(M)})$ |
| Ridge | $\hat{\beta}_j / (1 + \lambda)$ |
| Lasso | $\text{sign}(\hat{\beta}_j)(\hat{\beta}_j - \lambda)_+$ |

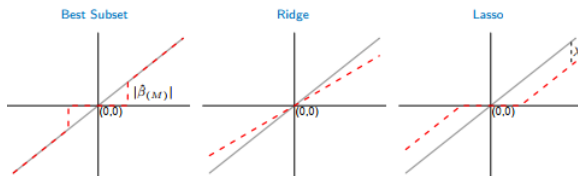


图 4: Enter Caption

Least angle regression

A kind of “democratic” version of forward stepwise regression.

Algorithm 3.2 *Least Angle Regression.*

1. Standardize the predictors to have mean zero and unit norm. Start with the residual $\mathbf{r} = \mathbf{y} - \bar{\mathbf{y}}$, $\beta_1, \beta_2, \dots, \beta_p = 0$.
 2. Find the predictor \mathbf{x}_j most correlated with \mathbf{r} .
 3. Move β_j from 0 towards its least-squares coefficient $\langle \mathbf{x}_j, \mathbf{r} \rangle$, until some other competitor \mathbf{x}_k has as much correlation with the current residual as does \mathbf{x}_j .
 4. Move β_j and β_k in the direction defined by their joint least squares coefficient of the current residual on $(\mathbf{x}_j, \mathbf{x}_k)$, until some other competitor \mathbf{x}_l has as much correlation with the current residual.
 5. Continue in this way until all p predictors have been entered. After $\min(N - 1, p)$ steps, we arrive at the full least-squares solution.
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图 5: Least angle regression

Lasso and Least angle regression

Algorithm 3.2a *Least Angle Regression: Lasso Modification.*

- 4a. If a non-zero coefficient hits zero, drop its variable from the active set of variables and recompute the current joint least squares direction.
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6: Lasso modification

Least angle regression:

$$x_j^T (y - X\beta) = r \cdot s_j$$

Lasso regression:

$$R(\beta) = \frac{1}{2} \|y - X\beta\|_2^2 + \lambda \|\beta\|_1$$

$$x_j^T (y - X\beta) = \lambda \cdot \text{sign}(\beta_j)$$

Principal component regression

Input matrix $X \in R^{N \times p}$

$$X = UDV^T$$

$$X^T X = VD^2 V^T$$

The first principal component has the property $z_1 = Xv_1$

$$\text{Var}(z_1) = \text{Var}(Xv_1) = \frac{d_1^2}{N}$$

$z_1 = Xv_1 = u_1 d_1$, u_1 is normalized first principal component.

Partial least square

Unlike Principal component regression, PLS also consider input's relationship with output.

Consider m predictors X_1, \dots, X_m , p response Y_1, \dots, Y_p

The first principal component T_1, U_1 is linear combination of $X = (X_1, \dots, X_m), Y = (Y_1, \dots, Y_p)$.

$$T_1 = w_{11}X_1 + \dots, w_{1m}X_m = w_1'X$$

$$U_1 = v_{11}Y_1 + \dots + v_{1p}Y_p = v_1'Y$$

Partial least square

The score of T_1, U_1 denote as t_1, u_1 , where

$$t_1 = X_0 w_1 = \begin{bmatrix} x_{11} & x_{12} & \cdots & x_{1m} \\ x_{21} & x_{22} & \cdots & x_{2m} \\ \vdots & \vdots & \ddots & \vdots \\ x_{n1} & x_{n2} & \cdots & x_{nm} \end{bmatrix} \begin{bmatrix} w_{11} \\ w_{12} \\ \vdots \\ w_{1m} \end{bmatrix} = \begin{bmatrix} t_{11} \\ t_{21} \\ \vdots \\ t_{n1} \end{bmatrix}$$
$$u_1 = Y_0 v_1 = \begin{bmatrix} y_{11} & y_{12} & \cdots & y_{1p} \\ y_{21} & y_{22} & \cdots & y_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ y_{n1} & y_{n2} & \cdots & y_{np} \end{bmatrix} \begin{bmatrix} v_{11} \\ v_{12} \\ \vdots \\ v_{1p} \end{bmatrix} = \begin{bmatrix} u_{11} \\ u_{21} \\ \vdots \\ u_{n1} \end{bmatrix}$$

Partial least square

$$\max \langle t_1, u_1 \rangle = \langle X_0 w_1, Y_0 v_1 \rangle = w_1^T X_0^T Y_0 v_1$$

$$\text{subject to } \|w_1\|^2 = \|v_1\|^2 = 1$$

By lagrange:

$$L = w_1^T X_0^T Y_0 v_1 - \frac{\lambda}{2}(\|w_1\|^2 - 1) - \frac{\theta}{2}(\|v_1\|^2 - 1)$$

$$\begin{cases} \frac{\partial L}{\partial w_1} = X_0^T Y_0 v_1 - \lambda w_1 = 0 \\ \frac{\partial L}{\partial v_1} = Y_0^T X_0 w_1 - \theta v_1 = 0 \end{cases} \Rightarrow \begin{cases} Y_0^T X_0 X_0^T Y_0 v_1 = \lambda^2 v_1 \\ X_0^T Y_0 Y_0^T X_0 w_1 = \lambda^2 w_1 \end{cases}$$

Partial least square

Then construct regression function of $X_1, \dots, X_m, Y_1, \dots, Y_p$ to T_1 ,

$$\begin{cases} X_0 = t_1 \alpha'_1 + E_1 \\ Y_0 = t_1 \beta'_1 + F_1 \end{cases}$$

where E_1, F_1 are residue matrix of size

$$n \times m, n \times p. \alpha'_1 = (\alpha_{11}, \dots, \alpha_{1m}), \beta'_1 = (\beta_{11}, \dots, \beta_{1p}).$$

By least square,
$$\begin{cases} \alpha_1 = X_0^T t_1 / \|t_1\|^2 \\ \beta_1 = Y_0^T t_1 / \|t_1\|^2 \end{cases}$$

Repeat r times, we get

$$X_0 = t_1 \alpha'_1 + \dots + t_r \alpha'_r + E_r,$$

$$Y_0 = t_1 \beta'_1 + \dots + t_r \beta'_r + F_r.$$

$$\begin{aligned} & \max_{\alpha} \text{Corr}^2(y, X\alpha) \text{Var}(X\alpha) \\ & \text{subject to } \|\alpha\| = 1, \alpha^T S \hat{\varphi}_{\ell} = 0, \ell = 1, \dots, m-1. \end{aligned}$$