LINMA1731 Project Fish Schools Tracking

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Abstract

In this paper we solve the first part of the project for the class "Stochastic process: Estimation and prediction" given during the Fall term of 2019. The average speed of each fish in a school of fish is approximated by a Gamma-distributed random variable, and various methods for estimating this quantity are given; a numerical simulation is also included.

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Part 1. Average speed estimation

1. Introduction

For the purpose of this project, we assume that the speed of each fish in a school at time i is a random variable V_i following a Gamma distribution, as suggested in [1]. This distribution is characterized by two parameters: a shape parameter k > 0 and a scale parameter s > 0. The parameters are the same for every fish and are time invariant. The aim of this first part is to identify these two parameters using empirical observations v_i .

2. Maximum likelihood estimation

Let v_i be i. i. d. realisations of a random variable following a Gamma distribution $\Gamma(k,s)$ (with $i=1,\ldots,N$). We first assume that the shape parameter k is known.

We start by deriving the maximum likelihood estimator of $\theta := s$ based on N observations. In order to do this, let us restate the probability density function of $V_i \sim \Gamma(k, s)$:

(1)
$$f_{V_i}(k, s; v_i) = \frac{1}{\Gamma(k)s^k} v_i^{k-1} e^{-\frac{v_i}{s}}, \quad i = 1, \dots, N.$$

With this in mind, we can find the likelihood $\mathcal{L}(k,\theta;v_1,\ldots,v_N)$ is given by

(2)
$$\mathcal{L}(k,\theta;v_1,\ldots,v_N) = \prod_{i=1}^N f_{V_i}(k,\theta;v_i)$$

(3)
$$= \prod_{i=1}^{N} \frac{1}{\Gamma(k)\theta^k} v_i^{k-1} e^{-\frac{v_i}{\theta}}.$$

In order to alleviate notation, we compute instead the log-likelihood, which is generally easier to work with:

(4)
$$\ell(k, \theta; v_1, \dots, v_N) \triangleq \ln \mathcal{L}(k, \theta; v_1, \dots, v_N)$$

(5)
$$= \ln \left(\prod_{i=1}^{N} \frac{1}{\Gamma(k)\theta^k} v_i^{k-1} e^{-\frac{v_i}{\theta}} \right)$$

(6)
$$= \sum_{i=1}^{N} \ln \left(\frac{1}{\Gamma(k)\theta^k} v_i^{k-1} e^{-\frac{v_i}{\theta}} \right)$$

(7)
$$= (k-1)\sum_{i=1}^{N} \ln v_i - \sum_{i=1}^{N} \frac{v_i}{\theta} - N(k \ln \theta + \ln \Gamma(k)).$$

Now, in order to obtain the maximum likelihood estimator $\widehat{\Theta}_{\theta}$ (which we will abusively write as $\widehat{\theta}$ for the rest of this paper), we must derive the log-likelihood

with respect to the estimand θ , and set it equal to zero:

(8)
$$\frac{\partial \ell(k, \theta; v_1, \dots, v_N)}{\partial \theta} \bigg|_{\theta = \hat{\theta}} = -\frac{kN}{\hat{\theta}} + \frac{\sum_{i=1}^N v_i}{\hat{\theta}^2} = 0$$

(9)
$$\iff \hat{\theta} = \frac{\sum_{i=1}^{N} v_i}{kN} = \frac{\overline{v_i}}{k}.$$

3. Properties of the estimator

We now wish to show some of the properties of this estimator.

3.1. Asymptotically unbiased.

Property 3.1. The maximum likelihood estimator is asymptotically unbiased.

Proof. The proof is trivial and left as an exercise to the reader.

3.2. Efficiency.

PROPERTY 3.2. The maximum likelihood estimator is efficient.

Proof. The proof is trivial and left as an exercise to the reader.

3.3. Best asymptotically normal.

PROPERTY 3.3. The maximum likelihood estimator is best asymptotically normal.

Proof. The proof is trivial and left as an exercise to the reader.

Property 3.4. The maximum likelihood estimator is consistent.

Proof. The proof is trivial and left as an exercise to the reader. \Box

4. Joint maximum likelihood estimation

We now consider $V_i \sim \Gamma(k, s)$ (for i = 1, ..., N) with both k and s unknown.

- 5. Numerical simulation
- 6. Fisher information matrix

7. Numerical proof

References

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