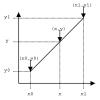
Interpolation and data fitting

Commonly, typically in experiments, we know $f(x_i)$ at various x_i , not necessarily equally spaced. Additionally, we may not quite know the analytical expression of f(x) that describes the data and allow to calculate its value at arbitrary $x \neq x_i$.

If the desired x is/are between largest and smallest x_i 's, the problem is of *interpolation*. If outside, then it is *extrapolation*.

Extrapolation is terribly risky endeavor unless backed up by solid theoretical idea(s). Here we study only *interpolation*.

Interpolating (x, y) between (x_0, y_0) and (x_1, y_1) with a straight line,



$$\frac{y-y_0}{x-x_0} = \frac{y_1-y_0}{x_1-x_0} \implies y = y_0 + \frac{x-x_0}{x_1-x_0} (y_1-y_0)$$



Generalize this idea by considering a polynomial of degree N to approximate a function from a set of N+1 points

$$P_N(x_i) \equiv f(x_i) = y_i$$
 where $i = 0, 1, ..., N$

The general form of $P_N(x)$ is

$$P_{N}(x) = a_{0} + a_{1}(x - x_{0}) + a_{2}(x - x_{0})(x - x_{1}) + \dots + a_{N}(x - x_{0})(x - x_{2}) \dots (x - x_{N-1})$$

$$P_{0}(x_{0}) = a_{0} = y_{0} = f(x_{0})$$

$$P_{1}(x_{1}) = a_{0} + a_{1}(x_{1} - x_{0}) = y_{1} = f(x_{1})$$

$$P_{2}(x_{2}) = a_{0} + a_{1}(x_{2} - x_{0}) + a_{2}(x_{2} - x_{0})(x_{2} - x_{1}) = y_{2} = f(x_{2}) \text{ etc.}$$

For linear interpolation N = 1,

$$a_0 = y_0, \ a_1 = \frac{y_1 - y_0}{x_1 - x_0} \Rightarrow P_1(x) = y(x) = y_0 + \frac{x - x_0}{x_1 - x_0} (y_1 - y_0)$$

A particularly useful form, although looks complicated, is

$$P_1(x) = y_0 - \frac{x - x_0}{x_1 - x_0} y_0 + \frac{x - x_0}{x_1 - x_0} y_1 = \frac{x - x_1}{x_0 - x_1} y_0 + \frac{x - x_0}{x_1 - x_0} y_1$$

This is most suitable for iterative processes.



Generalized interpolation formula by Lagrange

$$P_N(x) = \sum_{i=0}^{N} \prod_{k \neq i} \frac{x - x_k}{x_i - x_k} y_i$$

As an example, show that (DIY)

$$P_2(x) = \frac{(x-x_1)(x-x_2)}{(x_0-x_1)(x_0-x_2)} y_0 + \frac{(x-x_0)(x-x_2)}{(x_1-x_0)(x_1-x_2)} y_1 + \frac{(x-x_0)(x-x_1)}{(x_2-x_0)(x_2-x_1)} y_2$$

Ex 1. Consider the following data table and show the f(x = 4) = 20

X	2	3	5	8	12
f(x)	10	15	25	40	60

Ex 2. To understand what is going on, take a smaller data set

X	0	10	20	30
f(x)	-250	0	50	-100

Each term in Lagrange formula i.e. coefficient of yi

$$i = 1 \frac{(x-10)(x-20)(x-30)}{0-10)(0-20)(0-30)} = -\frac{x^3}{6000} + \frac{x^2}{100} - \frac{11x}{60} + 1$$

$$i = 2 \frac{(x-0)(x-20)(x-30)}{(10-0)(10-20)(10-30)} = \frac{x^3}{2000} - \frac{x^2}{40} + \frac{3x}{10}$$

$$i = 3 \frac{(x-0)(x-10)(x-30)}{20-0)(20-10)(20-30)} = -\frac{x^3}{2000} + \frac{x^2}{50} - \frac{3x}{20}$$

$$i = 4 \frac{(x-0)(x-10)(x-20)}{(30-0)(30-10)(30-20)} = \frac{x^3}{6000} - \frac{x^2}{200} + \frac{x}{30}$$

Multiply each line with the corresponding $f(x_i)$ and add together the terms of like power to obtain

$$y(x) \equiv P_4(x) = -x^2 + 35x - 250$$

Lagrange's formula gives a simple quadratic description of the data.

If required, we can calculate, say, f(x = 15) either using the quadratic formula or directly from Lagrange's formula.

DIY: Find both way f(x = 1969) from the data

X	1951	1961	1971
f(x)	2.8	3.2	4.5



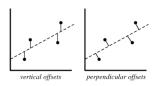
Least square fitting

Set of data points generated (x_i, y_i) are expected to be described by a **known** function f(x) but with undetermined coefficients.

Examples – stress-strain data points within elastic limit, voltage-current data points. Task is to determine modulus of elasticity and Resistance of the material under consideration.

$$V = \mathbf{R}i$$
 and stress = $\mathbf{k} \times \text{strain}$

Least square fitting is a way to do this and for that we need sum of squares of the offset (residual) of the data points from the curve f(x). Offsets can be either vertical or perpendicular



Vertical offsets allow uncertainties of data points along x, y-axes to be independent of each other.



Why squares and not absolute values?

Squares permit offsets to be treated as continuous differentiable quantities.

Absolute values result in discontinuous derivatives.

Fit N data points (x_i, y_i, σ_i) to a model $f(x_i; a_1, a_2, \ldots, a_M)$ with M parameters (M < N). Least square suggests to minimize squares of weighted offsets with respect to the parameters a_k 's

$$\chi^2 = \sum_{i=1}^N \left(\frac{y_i - f(x_i; a_1, a_2, \dots, a_M)}{\sigma_i} \right)^2 \rightarrow \frac{\partial \chi^2}{\partial a_j} = 0$$

Linear regression: model the data points with a straight line

$$f(x) = a_1 + a_2 x \implies \chi^2 = \sum_{i=1}^{N} \left(\frac{y_i - a_1 - a_2 x_i}{\sigma_i} \right)^2$$

Minimizing χ^2 w.r.t. a_1, a_2 yields

$$\frac{\partial \chi^2}{\partial a_1} = -2 \sum_{i=1}^N \frac{y_i - a_1 - a_2 x_i}{\sigma_i^2} = 0$$

$$\frac{\partial \chi^2}{\partial a_2} = -2 \sum_{i=1}^N \frac{x_i (y_i - a_1 - a_2 x_i)}{\sigma_i^2} = 0$$

Introducing a shorthand symbol $\sum_{i=1}^{N} \equiv \sum_{i}$, we ended up with

$$\sum_{i} \frac{y_{i}}{\sigma_{i}^{2}} = a_{1} \sum_{i} \frac{1}{\sigma_{i}^{2}} + a_{2} \sum_{i} \frac{x_{i}}{\sigma_{i}^{2}} \qquad \mathcal{S}_{y} = a_{1} \mathcal{S} + a_{2} \mathcal{S}_{x}$$

$$\equiv$$

$$\sum_{i} \frac{x_{i} y_{i}}{\sigma_{i}^{2}} = a_{1} \sum_{i} \frac{x_{i}}{\sigma_{i}^{2}} + a_{2} \sum_{i} \frac{x_{i}^{2}}{\sigma_{i}^{2}} \qquad \mathcal{S}_{xy} = a_{1} \mathcal{S}_{x} + a_{2} \mathcal{S}_{xx}$$

The solution of these equations are absolutely straight forward,

$$\textbf{a}_1 = \frac{\mathcal{S}_{xx}\mathcal{S}_y - \mathcal{S}_x\mathcal{S}_{xy}}{\Delta}, \ \ \textbf{a}_2 = \frac{\mathcal{S}_{xy}\mathcal{S} - \mathcal{S}_x\mathcal{S}_y}{\Delta} \text{ where } \Delta = \mathcal{S}\mathcal{S}_{xx} - \mathcal{S}_x^2$$

Estimating the errors σ_{a_i} on the parameters a_i determined above

$$\sigma_{a_{i}}^{2} = \sum_{i} \sigma_{i}^{2} \left(\frac{\partial a_{i}}{\partial y_{i}}\right)^{2}$$

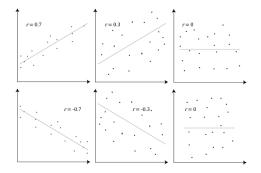
$$\Rightarrow \frac{\partial a_{1}}{\partial y_{i}} = \frac{S_{xx} - S_{x} x_{i}}{\Delta \sigma_{i}^{2}}, \quad \frac{\partial a_{2}}{\partial y_{i}} = \frac{S x_{i} - S_{x}}{\Delta \sigma_{i}^{2}}$$

$$\Rightarrow \sigma_{a_{1}}^{2} = \frac{S_{xx}}{\Delta} \quad \text{and} \quad \sigma_{a_{2}}^{2} = \frac{S}{\Delta}$$

To this end, in order to get some idea about how *good* is the **linear** fit we define an estimate called *Pearson's correlation coefficient* r

$$r^2 = \frac{\mathcal{S}_{xy}}{\mathcal{S}_{xx}\mathcal{S}_{yy}}$$
 where $0 \le r^2 \le 1$

As $r^2 \rightarrow 1$ fit gets better.



NB: But is the straight line model itself good to fit the data? To address this requires ideas of *goodness of fit*, *confidence level*, testing some *null hypothesis* against χ^2 probability distribution and so on.

The straight line model is generic enough for use in a few other models which can be reduced to straight line form usually by taking logarithms,

exponential: $f(x) = a e^{bx} \rightarrow \log f(x) = \log a + bx$

logarithm: $f(x) = a + b \log x$

power law: $f(x) = ax^b \rightarrow \log f(x) = \log a + b \log x$

Least square polynomial fitting

Polynomial model $f(x) = \sum_{k=0}^{n} a_k x^k$ can also be subjected to linear fitting. For simplified expressions, we put all the errors $\sigma_i^2 = 1$ without any loss of generality. In such case, $\chi^2 \to R^2$

$$R^{2} = \sum_{i=1}^{n} \left[y_{i} - \left(a_{0} + a_{1}x_{i} + \dots + a_{k}x_{i}^{k} \right) \right]^{2}$$

$$\frac{\partial R^{2}}{\partial a_{0}} = -2 \sum_{i=1}^{n} \left[y_{i} - \left(a_{0} + a_{1}x_{i} + \dots + a_{k}x_{i}^{k} \right) \right] = 0$$

$$\frac{\partial R^{2}}{\partial a_{1}} = -2 \sum_{i=1}^{n} \left[y_{i} - \left(a_{0} + a_{1}x_{i} + \dots + a_{k}x_{i}^{k} \right) \right] x_{i} = 0$$

$$\dots$$

$$\frac{\partial R^{2}}{\partial a_{k}} = -2 \sum_{i=1}^{n} \left[y_{i} - \left(a_{0} + a_{1}x_{i} + \dots + a_{k}x_{i}^{k} \right) \right] x_{i}^{k} = 0$$

Set of equations for ai's are,

$$a_{0}n + a_{1} \sum_{i} x_{i} + \dots + a_{k} \sum_{i} x_{i}^{k} = \sum_{i} y_{i}$$

$$a_{0} \sum_{i} x_{i} + a_{1} \sum_{i} x_{i}^{2} + \dots + a_{k} \sum_{i} x_{i}^{k+1} = \sum_{i} x_{i} y_{i}$$

$$\dots$$

$$a_{0} \sum_{i} x_{i}^{k} + a_{1} \sum_{i} x_{i}^{k+1} + \dots + a_{k} \sum_{i} x_{i}^{2k} = \sum_{i} x_{i}^{k} y_{i}$$

which when written in matrix form becomes a problem of matrix inversion, and you are free to use your favorite inverter.

$$\begin{pmatrix} n & \sum x_i & \cdots & \sum x_i^k \\ \sum x_i & \sum x_i^2 & \cdots & \sum x_i^{k+1} \\ \vdots & \vdots & \ddots & \vdots \\ \sum x_i^k & \sum x_i^{k+1} & \cdots & \sum x_i^{2k} \end{pmatrix} \begin{pmatrix} a_0 \\ a_1 \\ \vdots \\ a_k \end{pmatrix} = \begin{pmatrix} \sum y_i \\ \sum x_i y_i \\ \vdots \\ \sum x_i^k y_i \end{pmatrix}$$

$$\Rightarrow \mathbf{X} \cdot \mathbf{a} = \mathbf{Y} \rightarrow \mathbf{a} = \mathbf{X}^{-1} \mathbf{Y}$$

DIY fitting

- 1. A time versus angular velocity data set on deceleration of a rotating disc is in the file fit1.dat. Fit it with the following functions (i) $\omega(t) = \omega_0 + \omega_c t$ and (ii) $\omega(t) = \omega_0 \, e^{-\omega_c t}$. Determine ω_0 , ω_c and the quality of fit for both the functions by *Pearson's r*.
- 2. Distance (r) versus height (h) of the trajectory of a test missile is given in the datafile fit2.dat. Try quadratic fit $h = a_0 + a_1 r + a_2 r^2$ and determine the highest point reached by the missile.