

Model Summary

State model

$$z_{i,t} \sim \text{binom}(\phi_{i,t}, z_{i,t-1})$$
$$\text{logit}(\phi_{i,t}) = \mu_\phi + \beta_\phi * \text{TIME}_t + \epsilon_{\phi t}$$

where z is the true state of the chicks (i.e., 0,1,2) and TIME is really just relative chick age. The is assuming a single year and a single site. We are estimating two general parameters here (μ_ϕ and β_ϕ) as well as an ϵ_ϕ for each time step. So there is replication at the nest level here (not time step level).

Observation model

$$y_{i,t} \sim \text{binom}(p_{i,t} * w_{i,t}, z_{i,t})$$
$$\text{logit}(p_{i,t}) = \mu_p + \beta_p * \text{TIME}_t + \epsilon_{pi}$$

where y is the observed state (i.e., 0,1,2,) and w is a binary for day/night (actual day/night fed into the model). That way each time step can be modeled but the w index will ensure that the p isn't impacted because it was simply night and there's no way to see the penguins (Chris came up with that). We need to model each time step because survival is from one hour to the next. We are estimating two general parameters here (μ_p and β_p) as well as an ϵ_p for each time nest So there is replication at the time step level here (not nest level).

Priors

$$\mu_p = \text{logit}(m_p)$$
$$\mu_\phi = \text{logit}(m_\phi)$$
$$m_p \sim \text{beta}(1.5, 1.5)$$
$$m_\phi \sim \text{beta}(1.5, 1.5)$$
$$\beta_\phi \sim N(0, 100)T(0, 1)$$
$$\beta_p \sim N(0, 10)T(0, 1)$$
$$\epsilon_{pi} \sim N(0, \tau_p)T(-10, 10)$$
$$\epsilon_{\phi i} \sim N(0, \tau_\phi)T(-10, 10)$$
$$\tau_p = \frac{1}{\sigma_p^2}$$
$$\tau_\phi = \frac{1}{\sigma_\phi^2}$$
$$\sigma_p \sim \text{unif}(0, 10)$$
$$\sigma_\phi \sim \text{unif}(0, 10)$$

Next steps

When adding covariates I think it should look something like below. I think the β_1 should stay in there, because it's forcing the model to account for chick age when measuring survival due to those other factors:

$$\begin{aligned} z_{i,t} &\sim \text{binom}(\phi_{i,t}, z_{i,t-1}) \\ \text{logit}(\phi_{i,t}) &= \mu_\phi + \beta_{1\phi} * \text{TIME}_t + \beta_{2\phi} * \text{SIC}_t + \beta_{3\phi} * \text{KRILL}_t + \epsilon_{\phi t} \\ y_{i,t} &\sim \text{binom}(p_{i,t} * w_{i,j}, z_{i,t}) \\ \text{logit}(p_{i,t}) &= \mu_p + \beta_p * \text{TIME}_t + \epsilon_{pi} \end{aligned}$$

When modeling a single year but multiple sites, all nests could be modeled together, with the μ_ϕ , $\beta_{1\phi}$, $\beta_{2\phi}$, $\beta_{3\phi}$, and $\epsilon_{\phi t}$ parameters modeled with random site effects. I don't think the parameters associated with p should get a random effect, because detection is already being modeled with nest identity as a random effect - no need to add on a site level as well.

I suppose it's a little more complicated than this, since we don't have covariates at the hourly time scale (except maybe temperature from the cameras). The effect of covariates on survival would have to be modeled at the day/week level (I guess hierarchically where each day or week is a 'group'?). Or would it be better just to replicate the KRILL value for every hour within a week, for instance?

When modeling multiple years and multiple sites, I guess the μ_ϕ , $\beta_{1\phi}$, $\beta_{2\phi}$, $\beta_{3\phi}$, and $\epsilon_{\phi t}$ parameters should be modeled with random effect for site AND a random effect for year?