732A96 Advanced Machine Learning Graphical Models and Hidden Markov Models

Jose M. Peña IDA, Linköping University, Sweden

Lecture 3: Parameter Learning

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Literature

- Main source
 - Bishop, C. M. Pattern Recognition and Machine Learning. Springer, 2006.
 Chapters 8 and 9.
- Additional source
 - Koski, T. J. T. and Noble, J. M. A Review of Bayesian Networks and Structure Learning. *Mathematica Applicanda* 40, 51-103, 2012.

Parameter Learning for BNs: Maximum Likelihood

• Given a sample $d_{1:N}$, the log likelihood function is

$$\begin{split} \log p(d_{1:N}|\theta_G,G) &= \log [\prod_{l} p(d_l|\theta_G,G)] = \log [\prod_{l} \prod_{i} p(d_l[X_i]|d_l[pa_G(X_i)],\theta_G)] \\ &= \log [\prod_{l} \prod_{i} \theta_{X_i = d_l[X_i]|pa_G(X_i) = d_l[pa_G(X_i)]}] = \log [\prod_{i} \prod_{j} \prod_{k} \theta_{X_i = k|pa_G(X_i) = j}^{N_{ijk}}] \\ &= \sum_{i} \sum_{j} \sum_{k} N_{ijk} \log \theta_{X_i = k|pa_G(X_i) = j} \end{split}$$

• To maximize the log likelihood function subject to the constraint $\sum_k \theta_{X_i=k|pa_G(X_i)=j} = 1$ for all i and j, we maximize

$$\sum_{i} \sum_{j} \sum_{k} \textit{N}_{ijk} \log \theta_{\textit{X}_i = k \mid \textit{pa}_{\textit{G}}(\textit{X}_i) = j} + \sum_{i} \sum_{j} \lambda_{ij} (\sum_{k} \theta_{\textit{X}_i = k \mid \textit{pa}_{\textit{G}}(\textit{X}_i) = j} - 1)$$

where λ_{ij} are called Lagrange multipliers.¹

• Setting to zero the derivative with respect to $\theta_{X_i=k|pa_G(X_i)=j}$ gives

$$\theta_{X_i=k|pa_G(X_i)=j} = -N_{ijk}/\lambda_{ij}$$

• Replacing this into the constraint gives $\lambda_{ij} = -N_{ij}$ and, thus, $\theta_{X:=k|nac(X_i)=i}^{ML} = N_{ijk}/N_{ij}$.

¹Any stationary point of the Lagrangian function is a stationary point of the original function subject to the constraints. Moreover, the log likelihood function is concave.

- Let d_{1:N} be an incomplete sample, i.e. d_i[X_i] =? for some i and I. Let o_{1:N} denote the observed part of d_{1:N}, and u_{1:N} the unobserved part.
- ▶ The log likelihood function over $o_{1:N}$ is

$$\log p(o_{1:N}|\theta_{G}, G) = \log \prod_{l} \sum_{u_{l}} p(o_{l}, u_{l}|\theta_{G}, G) = \sum_{l} \log \sum_{u_{l}} p(o_{l}, u_{l}|\theta_{G}, G)$$

• To maximize it subject to the constraint $\sum_k \theta_{X_i=k|pa_G(X_i)=j}=1$ for all i and j, we maximize

$$\sum_{l} \log \sum_{u_l} p(o_l, u_l | \theta_G, G) + \sum_{i} \sum_{j} \lambda_{ij} \left(\sum_{k} \theta_{X_i = k | pa_G(X_i) = j} - 1 \right)$$

• Its derivative with respect to $\theta_{X_i=k|pa_G(X_i)=j}$ is

$$\sum_{l} \frac{\sum_{u_{l}:c_{l}[X_{i}]=k,c_{l}[pa_{G}(X_{i})]=j} \prod_{i'} \theta_{X_{i'}=c_{l}[X_{i'}]|pa_{G}(X_{i'})=c_{l}[pa_{G}(X_{i'})]}}{\theta_{X_{i}=k|pa_{G}(X_{i})=j} \sum_{u_{l}} p(o_{l},u_{l}|\theta_{G},G)} + \lambda_{ij}}$$

$$=\sum_{l}\sum_{u_{l}:c_{l}[X_{i}]=k,c_{l}[pa_{G}(X_{i})]=j}\frac{p(u_{l}|o_{l},\theta_{G},G)}{\theta_{X_{i}=k|pa_{G}(X_{i})=j}}+\lambda_{ij}=M_{ijk}/\theta_{X_{i}=k|pa_{G}(X_{i})=j}+\lambda_{ij}$$

where $c_l = \{o_l, u_l\}$ and $M_{ijk} = \sum_l \sum_{u_l: c_l [X_i] = k, c_l [pa_G(X_i)] = j} p(u_l | o_l, \theta_G, G)$.

Setting the derivative to zero gives

$$\theta_{X_i=k|pa_G(X_i)=j}=-M_{ijk}/\lambda_{ij}$$

• Replacing this into the constraint gives $\lambda_{ij} = -M_{ij}$ and, thus, $\theta_{X_i=k|pa_G(X_i)=j}^{ML} = M_{ijk}/M_{ij}$. No closed form solution but ...

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EM algorithm

Set \theta_G to some initial values
Repeat until \theta_G does not change
Compute p(U_l|o_l,\theta_G,G) for all l /* E step */
Compute M_{ijk}
Set \theta_G = M_{ijk}/M_{ij} /* M step */
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▶ Note that computing $p(U_I|o_I, \theta_G, G)$ requires inference.

- As shown before, maximizing the log likelihood function over O is inefficient as no closed form solution exists.
- Moreover, it is ineffective due to multimodality, i.e. each completion of the data defines a unimodal function but their sum may be multimodal.
- Consider instead maximizing its expectation

$$\begin{split} & \mathrm{E}_{U_{1:N}} \big[\log p(o_{1:N}, U_{1:N} | \theta_G, G) \big] = \sum_{l} \sum_{u_l} p(u_l | o_l, \theta_G, G) \log p(o_l, u_l | \theta_G, G) \\ & = \sum_{l} \sum_{u_l} p(u_l | o_l, \theta_G, G) \sum_{i} \log \theta_{X_i = c_l[X_i] | pa_G(X_i) = c_l[pa_G(X_i)]} \end{split}$$

where $c_l = \{o_l, u_l\}$. Then

$$\mathbb{E}_{U_{1:N}}[\log p(o_{1:N}, U_{1:N} | \theta_G, G)] = \sum_i \sum_j \sum_k M_{ijk} \log \theta_{X_i = k | pa_G(X_i) = j}$$

where $M_{ijk} = \sum_{l} \sum_{u_l:c_l[X_i]=k,c_l[pa_G(X_i)]=j} p(u_l|o_l,\theta_G,G)$.

▶ Then, $\theta_{X_i=k|pa_G(X_i)=j}^{ML} = M_{ijk}/M_{ij}$. No closed form solution but it suggests the EM algorithm too.

Another way to motivate the EM algorithm is as follows:

$$\log p(o_{1:N}|\theta_G,G) = L(q,\theta_G) + KL(q||p)$$

where

- $L(q, \theta_G) = \sum_{u_1 : N} q(u_{1:N}) \log \left[p(o_{1:N}, u_{1:N} | \theta_G, G) / q(u_{1:N}) \right]$
- $KL(q||p) = -\sum_{u_{1:N}}^{n} q(u_{1:N}) \log \left[p(u_{1:N}|o_{1:N}, \theta_G, G) / q(u_{1:N}) \right]$
- $q(U_{1:N})$ is a probability distribution.
- ▶ To see it, note that $L(q, \theta_G)$
 - $= \sum_{u_{1:N}} q(u_{1:N}) [\log p(u_{1:N}|o_{1:N},\theta_G,G) + \log p(o_{1:N}|\theta_G,G)] \sum_{u_{1:N}} q(u_{1:N}) \log q(u_{1:N})$
 - $= \log p(o_{1:N}|\theta_G, G) + \sum_{u_{1:N}} q(u_{1:N}) \log p(u_{1:N}|o_{1:N}, \theta_G, G) \sum_{u_{1:N}} q(u_{1:N}) \log q(u_{1:N})$

$$= \log p(o_{1:N}|\theta_G,G) - KL(q||p)$$

- Note that $KL(q||p) \ge 0$ and, thus, $\log p(o_{1:N}|\theta_G,G) \ge L(q,\theta_G)$ for any $q(U_{1:N})$.
- ► E step: Maximize the lower bound $L(q, \theta_G)$ by setting $q(U_{1:N}) = p(U_{1:N}|o_{1:N}, \theta_G, G)$, since then KL(q||p) = 0.
- Note that now $L(q, \theta_G) = \mathbb{E}_{U_{1:N}}[\log p(o_{1:N}, U_{1:N}|\theta_G, G)] + constant.$
- M step: Maximize the lower bound $L(q, \theta_G)$ with respect to θ_G .
- The last step may introduce a non-zero KL(q||p), resulting in an iterative process: The EM algorithm.

Parameter Learning for MNs: Iterative Proportional Fitting Procedure

• Given a complete sample $d_{1:N}$, the log likelihood function is

$$p(d_{1:N}|\theta_G, G) = \sum_{K \in CI(G)} \sum_k N_k \log \varphi(k) - N \log Z$$

where N_K is the number of instances in $d_{1:N}$ where K takes value k. Then

$$p(d_{1:N}|\theta_G, G)/N = \sum_{K \in CI(G)} \sum_k p_e(k) \log \varphi(k) - \log Z$$

where $p_e(X)$ is the empirical probability distribution obtained from $d_{1:N}$.

▶ Let $Q \in Cl(G)$. The derivative with respect to $\varphi(q)$ is

$$\frac{\partial p(d_{1:N}|\theta_G,G)/N}{\partial \varphi(q)} = \frac{p_e(q)}{\varphi(q)} - \frac{1}{Z} \frac{\partial Z}{\partial \varphi(q)}$$

▶ Let $Y = X \setminus Q$. Then

$$\frac{\partial Z}{\partial \varphi(q)} = \sum_{y} \prod_{K \in Cl(G) \setminus Q} \varphi(k, \overline{k}) = \frac{Z}{\varphi(q)} \sum_{y} \prod_{K \in Cl(G) \setminus Q} \varphi(k, \overline{k}) \frac{\varphi(q)}{Z} = \frac{Z}{\varphi(q)} p(q | \theta_G, G)$$

where \overline{k} denotes the elements of q corresponding to the elements of $K \cap Q$.

Putting together the results above, we have that

$$\frac{\partial p(d_{1:N}|\theta_G,G)/N}{\partial \varphi(q)} = \frac{p_e(q)}{\varphi(q)} - \frac{p(q|\theta_G,G)}{\varphi(q)}$$

Parameter Learning for MNs: Iterative Proportional Fitting Procedure

Setting the derivative to zero gives ²

$$\varphi^{ML}(k) = \varphi(k)p_e(k)/p(k|\theta_G,G)$$

for all $K \in Cl(G)$. No closed form solution but ...

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IPFP

Initialize \varphi(K) for all K \in CI(G)

Repeat until convergence

Set \varphi(K) = \varphi(K)p_e(K)/p(K|\theta_G,G) for all K \in CI(G)
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- Iterative coordinate ascend method.
- Note that computing $p(K|\theta_G, G)$ in the last line requires inference. Moreover, the multiplication and division are elementwise.
- Note also that Z needs to be computed in each iteration, which is computationally hard. This can be avoided by a careful initialization.

²The log likelihood function is concave.