732A96 Advanced Machine Learning Graphical Models and Hidden Markov Models

Jose M. Peña IDA, Linköping University, Sweden

Lecture 3: Parameter Learning

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Literature

Main sources

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R resources

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Parameter Learning for BNs: Maximum Likelihood

• Given a sample $d_{1:N}$, the log likelihood function is

$$\begin{split} \log p(d_{1:N}|\theta_G,G) &= \log [\prod_{l} p(d_l|\theta_G,G)] = \log [\prod_{l} \prod_{i} p(d_l[X_i]|d_l[pa_G(X_i)],\theta_G)] \\ &= \log [\prod_{l} \prod_{i} \theta_{X_i = d_l[X_i]|pa_G(X_i) = d_l[pa_G(X_i)]}] = \log [\prod_{i} \prod_{j} \prod_{k} \theta_{X_i = k|pa_G(X_i) = j}^{N_{ijk}}] \\ &= \sum_{i} \sum_{j} \sum_{k} N_{ijk} \log \theta_{X_i = k|pa_G(X_i) = j} \end{split}$$

• To maximize the log likelihood function subject to the constraint $\sum_k \theta_{X_i=k|pa_G(X_i)=j} = 1$ for all i and j, we maximize

$$\sum_{i} \sum_{j} \sum_{k} \textit{N}_{ijk} \log \theta_{\textit{X}_i = k \mid \textit{pa}_{\textit{G}}(\textit{X}_i) = j} + \sum_{i} \sum_{j} \lambda_{ij} \bigl(\sum_{k} \theta_{\textit{X}_i = k \mid \textit{pa}_{\textit{G}}(\textit{X}_i) = j} - 1\bigr)$$

where λ_{ii} are called Lagrange multipliers.¹

• Setting to zero the derivative with respect to $\theta_{X_i=k|pa_G(X_i)=j}$ gives

$$\theta_{X_i=k|pa_G(X_i)=j} = -N_{ijk}/\lambda_{ij}$$

• Replacing this into the constraint gives $\lambda_{ij} = -N_{ij}$ and, thus, $\theta_{X:=k|nac(X_i)=i}^{ML} = N_{ijk}/N_{ij}$.

¹Any stationary point of the Lagrangian function is a stationary point of the original function subject to the constraints. Moreover, the log likelihood function is concave.

Parameter Learning for BNs: Maximum A Posteriori

- Alternatively, we can choose the parameter values θ_G with maximum posterior probability

$$p(\theta_G|d_{1:N},G) = p(d_{1:N}|\theta_G,G)p(\theta_G|G)/p(d_{1:N}|G) \propto p(d_{1:N}|\theta_G,G)p(\theta_G|G)$$

where $p(d_{1:N}|\theta_G, G)$ is the likelihood function, $p(\theta_G|G)$ is a prior probability distribution, and $p(d_{1:N}|G)$ is a normalization constant.

• Assuming that $p(\theta_G|G) = \prod_i \prod_j p(\theta_{X_i|pa_G(X_i)=j}|G)$ and $p(\theta_{X_i|pa_G(X_i)=j}|G) \sim \textit{Dirichlet}(\alpha_{ij1}, \dots, \alpha_{ijk_i})$, we have that

$$p(\theta_{X_i|pa_G(X_i)=j}|G) \propto \prod_k \theta_{X_i=k|pa_G(X_i)=j}^{\alpha_{ijk}-1}$$

and thus

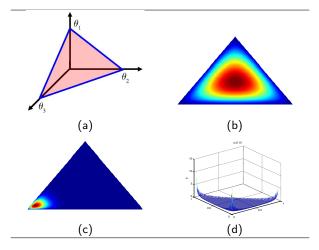
$$p(d_{1:N}|\theta_G,G)p(\theta_G|G)$$

$$\propto \prod_{i} \prod_{j} \prod_{k} \theta_{X_i=k|pa_G(X_i)=j}^{N_{ijk}} \prod_{i} \prod_{j} \prod_{k} \theta_{X_i=k|pa_G(X_i)=j}^{\alpha_{ijk}-1} = \prod_{i} \prod_{j} \prod_{k} \theta_{X_i=k|pa_G(X_i)=j}^{N_{ijk}+\alpha_{ijk}-1}$$

The posterior probability distribution is maximized when

$$\theta_{X_i=k|pa_G(X_i)=j}^{MAP} = (N_{ijk} + \alpha_{ijk} - 1)/(N_{ij} + \alpha_{ij} - k_i)$$

Parameter Learning for BNs: Maximum A Posteriori



(a) The Dirichlet distribution over a 3-valued random variable is defined over the simplex represented by the triangular surface. Points in this surface satisfy $0 \le \theta_i \le 1$ and $\sum_i \theta_i = 1$. (b) Dirichlet(2,2,2). (c) Dirichlet(20,2,2). (d) Dirichlet(0.1,0.1,0.1). Source: Murphy (2012).

- Let $d_{1:N}$ be an **incomplete sample**, i.e. $d_I[X_i] = ?$ for some i and I. Let $o_{1:N}$ denote the observed part of $d_{1:N}$, and $u_{1:N}$ the unobserved part.
- ▶ The log likelihood function over $o_{1:N}$ is

$$\log p(o_{1:N}|\theta_{G}, G) = \log \prod_{l} \sum_{u_{l}} p(o_{l}, u_{l}|\theta_{G}, G) = \sum_{l} \log \sum_{u_{l}} p(o_{l}, u_{l}|\theta_{G}, G)$$

• To maximize it subject to the constraint $\sum_k \theta_{X_i=k|pa_G(X_i)=j}=1$ for all i and j, we maximize

$$\sum_{l} \log \sum_{u_l} p(o_l, u_l | \theta_G, G) + \sum_{i} \sum_{j} \lambda_{ij} \left(\sum_{k} \theta_{X_i = k | pa_G(X_i) = j} - 1 \right)$$

• Its derivative with respect to $\theta_{X_i=k|pa_G(X_i)=j}$ is

$$\begin{split} & \sum_{l} \frac{\sum_{u_{l}:c_{l}[X_{i}]=k,c_{l}[pa_{G}(X_{i})]=j} \prod_{i'} \theta_{X_{i'}=c_{l}[X_{i'}]|pa_{G}(X_{i'})=c_{l}[pa_{G}(X_{i'})]}{\theta_{X_{i}=k|pa_{G}(X_{i})=j} \sum_{u_{l}} p(o_{l},u_{l}|\theta_{G},G)} + \lambda_{ij} \\ & = \sum_{l} \sum_{u_{l}:c_{l}[X_{i}]=k,c_{l}[pa_{G}(X_{i})]=j} \frac{p(u_{l}|o_{l},\theta_{G},G)}{\theta_{X_{i}=k|pa_{G}(X_{i})=j}} + \lambda_{ij} = M_{ijk}/\theta_{X_{i}=k|pa_{G}(X_{i})=j} + \lambda_{ij} \end{split}$$

where $c_l = \{o_l, u_l\}$ and $M_{ijk} = \sum_l \sum_{u_l: c_l[x_i] = k, c_l[pa_G(x_i)] = i} p(u_l|o_l, \theta_G, G)$.

Setting the derivative to zero gives

$$\theta_{X_i=k|pa_G(X_i)=j}=-M_{ijk}/\lambda_{ij}$$

• Replacing this into the constraint gives $\lambda_{ij} = -M_{ij}$ and, thus, $\theta_{X_i=k|pa_G(X_i)=j}^{ML} = M_{ijk}/M_{ij}$. No closed form solution but ...

```
EM algorithm

Set \theta_G to some initial values
Repeat until \theta_G does not change

Compute p(U_l|o_l,\theta_G,G) for all l /* E step */

Compute M_{ijk}

Set \theta_G = M_{ijk}/M_{ij} /* M step */
```

▶ Note that computing $p(U_I|o_I, \theta_G, G)$ requires inference.

- As shown before, maximizing the log likelihood function over O is inefficient as no closed form solution exists.
- Moreover, it is ineffective due to multimodality, i.e. each completion of the data defines a unimodal function but their sum may be multimodal.
- Consider instead maximizing its expectation

$$\begin{split} & \mathrm{E}_{U_{1:N}} \big[\log p(o_{1:N}, U_{1:N} | \theta_G, G) \big] = \sum_{l} \sum_{u_l} p(u_l | o_l, \theta_G, G) \log p(o_l, u_l | \theta_G, G) \\ & = \sum_{l} \sum_{u_l} p(u_l | o_l, \theta_G, G) \sum_{i} \log \theta_{X_i = c_l[X_i] | pa_G(X_i) = c_l[pa_G(X_i)]} \end{split}$$

where $c_l = \{o_l, u_l\}$. Then

$$\mathbb{E}_{U_{1:N}}[\log p(o_{1:N}, U_{1:N} | \theta_G, G)] = \sum_i \sum_j \sum_k M_{ijk} \log \theta_{X_i = k | pa_G(X_i) = j}$$

where $M_{ijk} = \sum_{l} \sum_{u_{l}:c_{l}[X_{i}]=k,c_{l}[pa_{G}(X_{i})]=j} p(u_{l}|o_{l},\theta_{G},G)$.

▶ Then, $\theta_{X_i=k|pa_G(X_i)=j}^{ML} = M_{ijk}/M_{ij}$. No closed form solution but it suggests the EM algorithm too.

Another way to motivate the EM algorithm is as follows:

$$\log p(o_{1:N}|\theta_G,G) = L(q,\theta_G) + KL(q||p)$$

where

- $L(q, \theta_G) = \sum_{u_1 \cdot N} q(u_{1:N}) \log \left[p(o_{1:N}, u_{1:N} | \theta_G, G) / q(u_{1:N}) \right]$
- $\mathsf{KL}(q||p) = -\sum_{u_{1:N}}^{n} q(u_{1:N}) \log \left[p(u_{1:N}|o_{1:N}, \theta_G, G) / q(u_{1:N}) \right]$
- $q(U_{1:N})$ is a probability distribution.
- ▶ To see it, note that $L(q, \theta_G)$
 - $= \sum_{u_{1:N}} q(u_{1:N}) [\log p(u_{1:N}|o_{1:N},\theta_G,G) + \log p(o_{1:N}|\theta_G,G)] \sum_{u_{1:N}} q(u_{1:N}) \log q(u_{1:N})$
 - $= \log p(o_{1:N}|\theta_G, G) + \sum_{u_{1:N}} q(u_{1:N}) \log p(u_{1:N}|o_{1:N}, \theta_G, G) \sum_{u_{1:N}} q(u_{1:N}) \log q(u_{1:N})$

$$= \log p(o_{1:N}|\theta_G,G) - KL(q||p)$$

- Note that $KL(q||p) \ge 0$ and, thus, $\log p(o_{1:N}|\theta_G, G) \ge L(q, \theta_G)$ for any $q(U_{1:N})$.
- ► E step: Maximize the lower bound $L(q, \theta_G)$ by setting $q(U_{1:N}) = p(U_{1:N}|o_{1:N}, \theta_G, G)$, since then KL(q||p) = 0.
- Note that now $L(q, \theta_G) = \mathbb{E}_{U_{1:N}}[\log p(o_{1:N}, U_{1:N}|\theta_G, G)] + constant.$
- M step: Maximize the lower bound $L(q, \theta_G)$ with respect to θ_G .
- The last step may introduce a non-zero KL(q||p), resulting in an iterative process: The EM algorithm.

Parameter Learning for MNs: Iterative Proportional Fitting Procedure

• Given a complete sample $d_{1:N}$, the log likelihood function is

$$p(d_{1:N}|\theta_G, G) = \sum_{K \in CI(G)} \sum_k N_k \log \varphi(k) - N \log Z$$

where N_K is the number of instances in $d_{1:N}$ where K takes value k. Then

$$p(d_{1:N}|\theta_G, G)/N = \sum_{K \in CI(G)} \sum_k p_e(k) \log \varphi(k) - \log Z$$

where $p_e(X)$ is the empirical probability distribution obtained from $d_{1:N}$.

▶ Let $Q \in CI(G)$. The derivative with respect to $\varphi(q)$ is

$$\frac{\partial p(d_{1:N}|\theta_G,G)/N}{\partial \varphi(q)} = \frac{p_e(q)}{\varphi(q)} - \frac{1}{Z} \frac{\partial Z}{\partial \varphi(q)}$$

▶ Let $Y = X \setminus Q$. Then

$$\frac{\partial Z}{\partial \varphi(q)} = \sum_{y} \prod_{K \in Cl(G) \setminus Q} \varphi(k, \overline{k}) = \frac{Z}{\varphi(q)} \sum_{y} \prod_{K \in Cl(G) \setminus Q} \varphi(k, \overline{k}) \frac{\varphi(q)}{Z} = \frac{Z}{\varphi(q)} p(q | \theta_G, G)$$

where \overline{k} denotes the elements of q corresponding to the elements of $K \cap Q$.

Putting together the results above, we have that

$$\frac{\partial p(d_{1:N}|\theta_G,G)/N}{\partial \varphi(q)} = \frac{p_e(q)}{\varphi(q)} - \frac{p(q|\theta_G,G)}{\varphi(q)}$$

Parameter Learning for MNs: Iterative Proportional Fitting Procedure

Setting the derivative to zero gives ²

$$\varphi^{ML}(k) = \varphi(k)p_e(k)/p(k|\theta_G,G)$$

for all $K \in Cl(G)$. No closed form solution but ...

IPFP

Set p(X) to some initial probability distribution Compute $\phi(K)$ for all $K \in Cs(G)$ as shown before Set $\psi(K) = \exp \phi(K)$ for all $K \in Cs(G)$ Compute $\varphi(K)$ for all $K \in Cl(G)$ as shown before Repeat until convergence Set $\varphi(K) = \varphi(K)p_e(K)/p(K|\theta_G,G)$ for all $K \in Cl(G)$

- Iterative coordinate ascend method.
- Note that computing $p(K|\theta_G, G)$ in the last line requires inference. Moreover, the multiplication and division are elementwise.
- Initializing the factors φ(K) from a probability distribution instead of directly ensures that Z = 1 and, thus, it does not need to be computed, which is NP-hard.

²The log likelihood function is concave.

Parameter Learning for MNs: Iterative Proportional Fitting Procedure

- ▶ We now show that the IPFP preserves Z across iterations.
- Let the factor updated in the current iteration have the superscript t+1, whereas the rest of the factors have the superscript t. Let $Q \in Cl(G)$ and $Y = X \setminus Q$. Then

$$\rho^{t+1}(Q|\theta_{G},G) = \sum_{w} \rho^{t+1}(w,Q|\theta_{G},G) = \sum_{y} \varphi^{t+1}(Q) \frac{1}{Z^{t+1}} \prod_{K \in Cl(G) \setminus Q} \varphi^{t}(k) \\
= \sum_{y} \varphi^{t}(Q) \frac{p_{e}(Q)}{p^{t}(Q|\theta_{G},G)} \frac{1}{Z^{t+1}} \prod_{K \in Cl(G) \setminus Q} \varphi^{t}(k) = \frac{p_{e}(Q)}{p^{t}(Q|\theta_{G},G)} \frac{Z^{t}}{Z^{t+1}} \sum_{w} p^{t}(w,Q|\theta_{G},G) \\
= \frac{p_{e}(Q)}{p^{t}(Q|\theta_{G},G)} \frac{Z^{t}}{Z^{t+1}} p^{t}(Q|\theta_{G},G) = p_{e}(Q) \frac{Z^{t}}{Z^{t+1}}$$

Then

$$1 = \sum_{q} p^{t+1}(q|\theta_G, G) = \frac{Z^t}{Z^{t+1}} \sum_{q} p_e(q) = \frac{Z^t}{Z^{t+1}}$$

 Exercise. Sketch how to perform parameter learning for MNs from an incomplete sample.