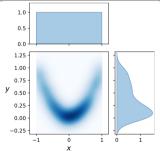
Standard Random Variables

Notation: P(X) / P(X), E[X] / E(X), assume *independent* and identical distribution (iid). Python: np.random

| | PMF Prob. Mass Function Valid i. $\forall x_i$. $P\{X = x_i\} \ge 0$ ii. $\sum_{i=0}^{\infty} P\{X = x_i\} = 1$ (density sum to 1) | CDF Cumulative Distribution $F_X(x) = P\{X \le \lfloor x \rfloor\}, \ 1 - P\{X > x\}, \\ P\{X = K\} = P\{X \le k\} - P\{X \le k - 1\}$ | $E[X] = \sum_{i=0}^{\infty} x_i P\{X = x_i\}$ | $\begin{aligned} & \textit{Var}[X] = \\ & E[X^2] - E[X]^2 \\ & \textit{LOTUS}, E[g(X)] = \\ & \sum g(x) \textit{P}\{X = k\} \end{aligned}$ |
|--|--|---|--|---|
| Bernoulli trial $X \sim Bern(p)$ | $P\{X\} = p, \ P\{\overline{X}\} = q$ | q = (1 - p) | p | pq |
| Binomial with replace $X \sim Bin(n,p)$ #successes in n Bern (p) trials $X \sim Bin(1,p)$, $(0-1)$ distribution if $n=1$ | $P\{X = k\} = \binom{n}{k} p^k q^{n-k}$ $P\{X = k\} = p^k q^{1-k}$ | Normal Approximation Poisson $n \to \infty$, $p \to 0$, $\lambda = np$ is moderate | пр | npq |
| Geometric / Negative Binomial $X \sim Geom(p), \ X \sim NegBin(r, p)$ in $n \ Bern(p)$ trials until $1^{st}/r$ successes | $P\{X = k\} = q^{k}p, \qquad k = \#failures$ $P\{X = k\} = \binom{n-1}{r-1}q^{k-1}p^{r}, \ k = \#trials$ | $1 - q^{k+1}, x > 0$ $[Exp(\lambda)]$ Approximation | $\frac{q}{p}$, $\frac{1 \cdot r}{p}$ | $\frac{(1-p)\cdot r}{p^2}$ |
| Poisson $X \sim Pois(\lambda), \ \lambda = np > 0$ memoryless #events in a fixed interval of time t | $P\{X = k\} = \frac{e^{-\lambda} \lambda^k}{k!}$ $Pois(\lambda t) \ given \ Exp(\lambda) \ as \ waiting \ time \ interval$ | by def | λ | λ |
| [Negative] HyperGeometric no replace $X \sim NHGemo(w, b, n)$, total $N = w + b$ #successes in n draws / until n failures | $P\{X = k\} = \frac{\binom{w}{k}\binom{b}{n-k}}{\binom{w+b}{n}}$ | / | $np = n\frac{w}{N}$ $n\frac{w}{b+1}$ | $\left rac{N-n}{N-1}npq ight $ |
| Joint Prob | $P_{ij} = P\left\{X = x_i, Y = y_j\right\}$ | $F(x,y) = \sum_{0}^{\lfloor x_i \rfloor} \sum_{0}^{\lfloor y_j \rfloor} P_{ij}$ = $P\{X \le x_i, Y \le y_j\}$ | $\sum_{i=0}^{\infty} \sum_{j=0}^{\infty} P_{ij} = 1$ | |
| Marginal Prob marginalize over another variable | $P\{X = x_i\} = \sum_{y} P\{X = x_i, Y = y_j\} $ = $\sum_{j=1}^{\infty} P_{ij}$ | $F_X(x) = F(x, \infty)$ = $\sum_{x_i \le x} \sum_{j=0}^{\infty} P_{ij}$ | $\forall i, j. \ P_{ij} \ge 0$ | |

| | | Υ | | | | Marginals for X |
|------------------------|------|------|------|------|------|---------------------------------|
| | | 1 | 2 | 3 | | $g(x) = \sum f(x, y)$ |
| | 1 | 0.32 | 0.03 | 0.01 | 0.36 | $\frac{1}{y}$ |
| Х | 2 | 0.06 | 0.24 | 0.02 | 0.32 | |
| | 3 | 0.02 | 0.03 | 0.27 | 0.32 | ΣΣ (() . |
| Marginals for | | 0.40 | 0.30 | 0.30 | 1 4 | $\sum_{X} \sum_{Y} f(x, y) = 1$ |
| $h(y) = \sum_{x} f(x)$ | x, y | | | | | • |



the joint density function

| Joint Prob | $f(x,y) = \frac{\partial}{\partial x \partial y} F(x,y)$ | $F(x,y) = \int_{-\infty}^{x} \int_{-\infty}^{y} f(u, v) du dv$ $= \iint_{B} f(x,y) dx dy$ | $\frac{\text{Valid}}{F(-\infty,\infty)} = 1$ | |
|---|--|--|---|---|
| Marginal Prob | $f_X(x) = \int_{y=-\infty}^{\infty} f(x,y) dy$ | $F_X(x) = F(x, \infty), y \to \infty$ = $\int_{-\infty}^x f_X(x) dx$ | $\forall x, y. \ f(x, y) \ge 0$ | |
| $\frac{Continuous}{X \in R}$ Distribution | PDF Prob Density Function Valid $i \cdot \forall x. \ f_x(x) \ge 0$ $ii. \int_{-\infty}^{\infty} f_X(x) dx = 1$ (density sum to 1) | CDF Cumulative Distribution $F_X(x) = \int_{-\infty}^x f_X(t) dt \text{, complement, LoTP}$ $f_X(x) = \frac{d}{dx} F_X(x)$ | $E[X] = \int_{-\infty}^{\infty} x \cdot f(x) dx$ | $Var[X]$ $LOTUS, E[g(X)] = \int_{-\infty}^{\infty} g(x)f(x)dx$ |
| Uniform $X \sim U(a,b) = e^{-\lambda Exp(\lambda)}$ a completely random point in $[low, high]$ | $f_X(x) = \begin{cases} \frac{1}{b-a}, & x \in [a, b] \\ 0, & otherwise \end{cases}$ | $F_X(x) = \begin{cases} 0, & x < a \\ \frac{x-a}{b-a}, & x \in [a, b] \\ 1, & x > b \end{cases}$ | $\frac{a+b}{2}$ | $\frac{(b-a)^2}{12}$ |
| Exponential $X \sim Exp(\lambda)$, $rate \ \lambda = \frac{1}{\theta} > 0$ memoryless waiting time between 2 successive events | $f_X(x) = \begin{cases} \lambda e^{-\lambda x}, & x > 0 \\ 0, & otherwise \end{cases}$ | $F_X(x) = \begin{cases} 1 - e^{-\lambda x}, & x > 0 \\ 0, & otherwise \end{cases}$ | $\theta = \frac{1}{\lambda}$ integrate by part | $\theta^2 = \frac{1}{\lambda^2}$ / tabular |
| Pareto $X \sim Pareto(shape = \alpha) = e^{Exp(\lambda)}$ cascade events, wealth, $x_m = 1$. <i>Loglinearity</i> | $f_X(x) = \begin{cases} \alpha x_m^{\alpha} x^{-(\alpha+1)}, & x \ge x_m \\ 0, & x < x_m \end{cases}$ | $F_X(x) = \begin{cases} 1 - \left(\frac{x_m}{x}\right)^{\alpha}, & x \ge x_m \\ 0, & x < x_m \end{cases}$ | $\begin{cases} \frac{\alpha x_m}{\alpha - 1}, & \alpha > 1 \\ \infty, & \alpha \le 1 \end{cases}$ | $\begin{cases} valid, \ \alpha > 2 \\ \infty, \ \alpha \le 2 \end{cases}$ |
| Normal / Gaussian $X \sim N(0,1)$ Standard $X \sim N(\mu, \sigma^2)$ | $f_X(x) = ce^{-\frac{x^2}{2}}, x \in R, c = \frac{1}{\sqrt{2\pi}\sigma}$ $f_X(x) = ce^{-\frac{(x-\mu)^2}{2\sigma^2}} = ce^{-\frac{(\frac{x-\mu}{\sigma})^2}{2}}$ | $\Phi_X(x) = c \int_{-\infty}^x e^{-\frac{x^2}{2}} dx$ | μ | σ^2 |
| $normal(loc = \mu, scale = \sigma)$ | $f_X(x) = ce^{-\frac{1}{2\sigma^2}} = ce^{-\frac{1}{2}}$ | $\Phi_{\frac{X-\mu}{\sigma}}(z)$ CLT | | |
| Beta $X \sim Beta(\alpha, \beta) = \frac{\Gamma(\alpha, 1)}{\Gamma(\alpha, 1) + \Gamma(\beta + 1)}$ As prior for Bayesian | $f_X(x) = \begin{cases} \frac{1}{B} x^{\alpha - 1} (1 - x)^{\beta - 1}, & x \in [0, 1] \\ 0, & otherwise \end{cases}$ | $B = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)} = \frac{(\alpha-1)!(\beta-1)!}{(\alpha+\beta-1)!}$ | $\frac{\alpha}{\alpha+\beta}$ | $\frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$ |
| Gamma $X \sim \Gamma(k, 1)$, shape $k > 0$ $X \sim \Gamma(k, \lambda)$, scale= $1/\lambda > 0$ $Y X \sim N(\mu, \frac{1}{\lambda X})$ | $f_{X}(x) = \frac{1}{\Gamma(k)} x^{k-1} e^{-x} , x > 0$ $f_{X}(x) = \frac{1}{\Gamma(k)} \frac{(\lambda x)^{k}}{x} e^{-\lambda x} , x > 0$ | $\Gamma(k) = (k-1)! , k \in \mathbb{N}$ $\Gamma(k) = \int_0^\infty e^{-x} x^{k-1} dx, k \in \mathbb{R}^+$ $\Gamma(1,\lambda) = Exp(\lambda), -\frac{1}{\lambda} \log(U)$ | <u>k</u> λ | $\frac{k}{\lambda^2}$ |