Introduction

Euklidian Norm: $\|x\|_2 = \sqrt{\sum_{i=1}^{n} x_i^2} = \sqrt{x^T x}$

$\|x\|_2^2 = x^T \cdot x$

Weighting Eukl. Norm: $\|x\|_Q^2 = x^T Q \cdot x$

Frobenius Norm: $\|x\|_F^2 = trace(AA^T) = \sum_{i=1}^{n} \sum_{j=1}^{m} A_{ij} A_{ij}$

Jacobian: $\nabla f(x) = \frac{\partial f}{\partial x}(x)$ in $\mathbb{R}^{n \times m}$ Hessian: $\nabla^2 f(x)$

Error in variables: $\hat{R}_{EV}(N) = \frac{\frac{1}{N}\sum_{k=1}^{N} u(k)}{\frac{1}{N}\sum_{k=1}^{N} i(k)}$

Simple Approach: $\hat{R}_{SA}(N) = \frac{1}{N} \cdot \sum_{k=1}^{N} \frac{u(k)}{i(k)}$

Least Squares: $\hat{R}_{LS}(N) = \underset{R \in \mathbb{R}}{arg\,min} \sum_{k=1}^{N} (R \cdot i(k) - u(k))^2$

$$= \frac{\frac{1}{N}\sum_{k=1}^{N} u(k) \cdot i(k)}{\frac{1}{N}\sum_{k=1}^{N} i(k)^2}$$

Matrix derivates: $\frac{d(c^T x)}{dx} = c \qquad \frac{d(x^T A x)}{dx} = (A^T + A)x$

Linear and non-linear models:
- linear if parameters linear i.e. $(\theta_1 x^2 + \theta_2 x + \theta_3)$
- nonliniar if i.e $(\sin(\theta_1)x + \theta_2)$ or derivatives in other orders than 1

Table of Derivatives:

$\mathbf{f(x)}$	$\mathbf{f'(x)}$
$g(x) \cdot h(x)$	$g'(x) \cdot h(x) + g(x) \cdot h'(x)$
$g(h(x))$	$g'(h(x)) \cdot h'(x)$
$\sin(x)$	$-\cos(x)$
$\cos(x)$	$\sin(x)$
$\tan(x) = \frac{\sin(x)}{\cos(x)}$	$\frac{1}{\cos^2(x)} = \sec^2(x)$
e^{kx}	$\frac{1}{k} e^{kx}$
$\ln(x)$	$\frac{1}{x}$
$\log_a x$	$\frac{1}{\ln a}(x \ln x - x)$

Random Variables and Probability

Dependent Probability: $P(A \vee B) = P(A) + P(B)$

Independent Prob.: $P(A, B) = P(A \wedge B) = P(A) \cdot P(B)$

Conditional Prob.: $P(A|B) = \frac{P(A|B)}{P(B)} = \frac{P(B|A) \cdot P(A)}{P(B)}$

$P(X \in [a, b]) = \int_a^b p_X(x)dx \qquad p(x|y) = \frac{p(x, y)}{p(y)}$

Mean/Expectation value: $\mathbb{E}\{\mu_X\} := \mu_X = \int_{-\infty}^{\infty} x \cdot p_X(x)dx$

$\mathbb{E}\{a + bX\} := a + b\mathbb{E}\{X\}$

Variance: $\sigma_X^2 := \mathbb{E}\{(X - \mu_X)^2\} = \mathbb{E}\{X^2\} - \mu_X^2$

Standard deviation: $\sigma_X = \sqrt{\sigma_X^2}$

Distributions

Uniform distribution: $P_y(x) = \begin{cases} \frac{1}{b-a} & if \quad x \in [a, b] \\ 0 & else \end{cases}$

Mean: $\mu_X = \int_{-\infty}^{\infty} x\, p_X(x)dx = \int_a^b \frac{1}{b-a} \cdot xdx = \frac{a+b}{2} =: \mu_X$

Normal distribution: $X \sim \mathcal{N}(\mu, \sigma^2) \quad \hat{\theta}_{LS} \sim \mathcal{N}(\theta_0, \Sigma_{\hat{\theta}})$

Gauss.png

$p(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \cdot exp(-\frac{(x-\mu)^2}{2\sigma^2})$

Multidimensional Normal Distribution:

$$p(x) = \frac{1}{\sqrt{(2\pi)^n \cdot det(\Sigma)}} \cdot exp(-\frac{1}{2} \cdot (x - \mu)^T \cdot \Sigma^{-1} \cdot (x - \mu))$$

Weibull distribution: $F(x) = 1 - e^{-(\lambda \cdot x)^k}$

Laplace distribuation: $f(x|\mu, b) = \frac{1}{2b} \cdot exp\left(-\frac{|x-\mu|}{b}\right)$

Useful statistic definitions

Covariance and Correlaton: $\sigma(X, Y) := \mathbb{E}(X - \mu_X)(Y - \mu_Y)$

$$= \int_{-\infty}^{\infty}\int_{-\infty}^{\infty} (X - \mu_X)(y - \mu_Y) \cdot p_{X,Y}(x, y)\, dx\, dy$$

Covariance Matrix: $\Sigma_x = cov(X) = \mathbb{E}\{XX^T\} - \mu_x \mu_x^T$ is PSD

$\Sigma = \begin{bmatrix} \sigma_x^2 & \sigma_{yx} \\ \sigma_{xy} & \sigma_y^2 \end{bmatrix} \qquad \sigma_{xy} = \sigma yx = \rho_{xy} \cdot \sigma_x \cdot \sigma_y$ where ρ is correlation

Multidimensional Random Variables:

$\mathbb{E}f(X) = \int_{\mathbb{R}^n} f(x) p_X(x) d^n x$

$cov(X) = \mathbb{E}\{(X - \mu_X)(X - \mu_X)^T\}$

$cov(X) = \mathbb{E}\{XX^T\} - \mu_X \mu_X^T$

$cov(Y) = \Sigma_y = A\Sigma_x A^T \quad for \quad y = A \cdot x$

$\mathbb{E}\{AX\} = A \cdot \mathbb{E}\{X\}$

Rules for variance:

$var(aX) = a^2 \cdot var(X)$

$var(X + Y) = var(X) + var(Y) + 2 \cdot cov(X, Y)$

Verschiebesatz: $var(X) = \mathbb{E}((X - \mathbb{E}(X))^2) = \mathbb{E}(X^2) - (\mathbb{E}(X))^2$

Bayes Theorem:
$P(A|B) = \frac{P(A,B)}{P(B)} = \frac{P(B|A)P(A)}{P(B)}$

Correlation:
uncorrelated if $\rho(X, Y) = 0$, $\rho(X, Y) := \frac{cov(X,Y)}{\sigma_x \sigma_y}$

Statistical estimators:

Biased- and Unbiasedness An estimator $\hat{\theta}_N$ is called unbiased iff $\mathbb{E}\{\hat{\theta}_N(y_N)\} = \theta_0$, where θ_0 is the true value of a parameter. Otherwise, is called biased.

Asymptotic Unbiasedness An estimator $\hat{\theta}_N$ is called asymptotically unbiased iff $\lim_{n \to \infty} \mathbb{E}\{\hat{\theta}_N(y_N)\} = \theta_0$

Consistency An estimator $\hat{\theta}_N(y_N)$ is called consistent if, for any $\epsilon > 0$, the probability $P(\hat{\theta}_N(y_N) \in [\theta_0 - \epsilon, \theta_0 + \epsilon])$ tends to one as $N \to \infty$.

Unconstrainded Optimization

Theorem 1: (First Order Necessary Conditions)
If $x^* \in D$ is local minimizer of $f : D \to \mathbb{R}$ and $f \in C^1$ then $\nabla f(x^*) = 0$ Definition (Stationary Point) A point $\bar{x}$ with $\nabla f(\bar{x}) = 0$ is called a stationary point of f.

Theorem 2: (Second Order Necessary Conditions)
If $x^* \in D$ is local minimizer of $f : D \to R$ and $f \in C^2$ then $\nabla^2 f(x^*) \succeq 0$

Theorem 3: (Second Order Sufficient Conditions and Stability under Perturbations)
Assume that $f : D \to R$ is C^2. If $x^* \in D$ is a stationary point and $\nabla^2 f(x^*) \succ 0$ then x^* is a strict local minimizer of f. In addition, this minimizer is locally unique and is stable against small perturbations of f, i.e. there exists a constant C such that for sufficiently small $p \in \mathbb{R}^n$ holds

$$\|x^* - \underset{x}{arg\,min}(f(x) + p^T x)\| \leq C\|p\|$$

Linear Least Squares Estimation

Preliminaries: i.i.d. and Gaussian noise

Overall Model: $y(k) = \phi(k)^T \theta + \epsilon(k)$

LS cost function as sum: $\sum_{k=1}^{N} (y(k) - \phi(k)^T \theta)^2$

LS cost function: $f(\theta) = \|y_N - \Phi_N \theta\|_2^2$

Unique minimizers: $\hat{\theta}_{LS} = \underset{\theta \in \mathbb{R}}{arg\,min} f(\theta) \theta^* = \underbrace{(\Phi^T \Phi)^{-1} \Phi^T}_{\Phi^+} y$

Pseudo Inverse: $\Phi^+ = (\Phi^T \Phi)^{-1} \Phi^T$

Weighted Least Squares (unitless)

For i.i.d noise: Unweight Least Squares is optimal: $W = I$

$$f_{WLS}(\theta) = \sum_{k=1}^{N} \frac{(y(k) - \phi(k)^T \theta)^2}{\sigma_\epsilon^2(k)} = \|y_N - \Phi_N \theta\|_W^2$$

$$= (Y_N - \Phi \cdot \theta)^T \cdot W \cdot (Y_N - \Phi \cdot \theta)$$

Solution for WLS:

$\hat{\theta}_{WLS} = \tilde{\Phi}^+ \tilde{y} \qquad$ mit $\tilde{\Phi} = W^{\frac{1}{2}} \Phi$ und $\tilde{y} = W^{\frac{1}{2}} y$

$$= \underset{\theta \in \mathbb{R}}{arg\,min} f_{WLS}(\theta) = (\Phi^T W \Phi)^{-1} \Phi^T W y$$

Ill-Posed Least Squares

Singular Value Decomposition: $A = USV^T \in \mathbb{R}^{mxn}$
with $U \in \mathbb{R}^{mxm}$, $V \in \mathbb{R}^{nxn}$ and $S \in \mathbb{R}^{mxn}$ where S is a Matrix with non-negative elements $(\sigma_1, \ldots, \sigma_r, 0, \ldots, 0)$ on the diagonal and 0 everywhere else.

Moore Penrose Pseudi Inverse:

$\Phi^+ = VS^+U^T = V(S^T S + \alpha I)^{-1} S^T U^T$

Φ^+ therefore selects $\theta^* \in S^*$ with minimal norm.

Regularization for Least Squares:

$\lim_{a \to 0} (\Phi^T \Phi + \alpha I)^{-1} \Phi^T = \Phi^+ \quad$ with Φ^+ $MPPI$

$\theta^* = (\Phi^T \Phi + \alpha \mathbb{I})^{-1} \Phi^T y$

Statistical Analysis of WLS

Expectation of Least Squares Estimator:

$$E\{\hat{\theta}_{WLS}\} = E\{(\Phi_N^T W \Phi_N)^{-1} \Phi_N^T W y_N\} = \theta_0$$

Covariance of the least squares estimator:

$$cov(\hat{\theta}_{WLS}) = (\Phi_N^T W \Phi_N)^{-1} = (\Phi_N^T \Sigma_{\in N}^{-1} \Phi_N)^{-1}$$

$$cov(\hat{\theta}_{WLS}) \succeq (\Phi_N^T W \Phi_N)^{-1}$$

Example LLS

Example of the Linear Least Square Estimator for: $N = 2$

$\varepsilon(1) \sim \mathcal{N}(0|\sigma_1^2)$ $\qquad$ $\varepsilon(2) \sim \mathcal{N}(0|\sigma_2^2)$

$$N = 2; \quad \Sigma_{\varepsilon_N} = \begin{bmatrix} \sigma_1^2 & 0 \\ 0 & \sigma_2^2 \end{bmatrix} \qquad W^{OPT} = \Sigma_{\varepsilon_N}^{-1} = \begin{bmatrix} \frac{1}{\sigma_1^2} & 0 \\ 0 & \frac{1}{\sigma_2^2} \end{bmatrix}$$

$$cov(\hat{\theta}_{WLS}) = (Y_N - \Phi_N \theta)^T \cdot W \cdot (Y_N - \Phi_N \theta)$$

$$= \sum_{k=1}^{2} (y(k) - \phi(k)^T \theta) \cdot \frac{1}{\sigma_k^2} \cdot (y(k) - \phi(k)^T \theta)$$

Measuring the goodness of Fit using: R^2 $\quad (0 \leq R^2 \leq 1)$

$$R^2 = 1 - \frac{\|y_N - \Phi_N \hat{\theta}\|_2^2}{\|y_N\|_2^2} = 1 - \frac{\|\epsilon_N\|_2^2}{\|y_N\|_2^2}$$

$$= \frac{\|y_N\|_2^2 - \|\epsilon_N\|_2^2}{\|y_N\|_2^2} = \frac{\|\hat{y}_N\|_2^2}{\|y_N\|_2^2}$$

Residual: $\epsilon_N \uparrow \rightarrow \quad R^2 \rightarrow 0 \ (\Rightarrow bad)$

Estimating the Covariance with the Single Experiment:

$$\hat{\sigma}_\varepsilon^2 := \frac{1}{N-d} \sum_{k=1}^{N} (y(k) - \phi(k)^T \hat{\theta}_{LS})^2 = \frac{\|y_N - \phi_N \hat{\theta}_{LS}\|_2^2}{N-d}$$

$$\hat{\Sigma}_{\hat{\theta}} := \hat{\sigma}_\varepsilon^2 (\phi_N^T \phi_N)^{-1} = \sigma_\epsilon^2 (\Phi_N^+ \Phi_N^{+T}) = \frac{\|y_N - \phi_N \hat{\theta}_{LS}\|_2^2}{N-d} \cdot (\phi_N^T \phi_N)^{-1}$$

Bayesian Estimation and the Maximum a Posteriori Estimate

Assumptions:

- Measurement: $\quad y_N \in \mathbb{R}^N \quad$ has i.i.d. noise
- Linear Model: $\quad M(\theta) = \phi_N \cdot \theta \quad$ and $\theta \in \mathbb{R}$

$$p(\theta|y_N) = \frac{p(y_N, \theta)}{p(y_N)} = \frac{p(y_N|\theta) \cdot p(\theta)}{p(y_N)}$$

$$\hat{\theta}_{MAP} = \underset{\theta \in \mathbb{R}}{\arg\min} \{-\log(p(y_N|\theta)) - \log(p(\theta))\}$$

MAP Example: Regularised Least Squares

$\theta = \bar{\theta} \pm \sigma_\theta \quad$ with $\quad \bar{\theta} = \theta_{\text{a-priori}}$

$$\hat{\theta}_{MAP} = \underset{\theta \in \mathbb{R}}{\arg\min} \frac{1}{2} \cdot \frac{1}{\sigma_{\epsilon^2}} \cdot \|y_N - \Phi_N \cdot \theta\|_2^2 + \frac{1}{2} \cdot \frac{1}{\sigma_\theta^2} \cdot (\theta - \bar{\theta})^2$$

Maximum Likelihood Estimation

L_2 **Estimation: Maximum Likelihood Estimation (ML):**

- Measurement Errors assumed to be Normally distributed
- Model described by a non-linear function M(θ)
- Every unbiased estimator needs to satisfy the Cramer-Rau inequality, which gives a lower bound on the covariance matrix

Model: $y = M(\theta) + \epsilon$

$$P(y|\theta) = C \prod_{i=1}^{N} exp\left(\frac{-(y_i - M_i(\theta))^2}{2 \cdot \sigma_i^2}\right) \qquad C = \prod_{i=1}^{N} \frac{1}{\sqrt{2 \cdot \pi \sigma_i^2}}$$

Positive log-Likelihood: Logarithm makes a product to a sum!

$$\log p(y|\theta) = \log(C) + \sum_{i=1}^{N} - \frac{(y_i - M_i(\theta))^2}{2 \cdot \sigma_i^2}$$

Negative log-Likelihood:

$$\hat{\theta}_{ML} = \underset{\theta \in \mathbb{R}^d}{arg\,max}\ p(y|\theta) = \underset{\theta \in \mathbb{R}^d}{\arg\min} \sum_{i=1}^{N} \frac{(y_i - M_i(\theta))^2}{2\sigma_i^2}$$

$$= \underset{\theta}{\arg\min} \frac{1}{2} \sum_{i=1}^{N} \left(\frac{y_i - M_i(\theta)}{\sigma_i}\right)^2$$

$$= \underset{\theta \in \mathbb{R}^d}{\arg\min} \frac{1}{2} \|S^{-1}(y - M(\theta))\|_2^2 \qquad \text{mit: } S = \begin{bmatrix} \sigma_1 & & \\ & \ddots & \\ & & \sigma_N \end{bmatrix}$$

L_1 **Estimation:**

- Measurement Errors assumed to be Laplace distributed and more robust against outliers.

$$\underset{\theta}{min} \|y - M(\theta)\|_1 = \underset{\theta}{min} \sum_{i=1}^{N} |y_i - M_i(\theta)|$$

$$\Rightarrow \text{ median of } \{Y_1, \cdots, Y_N\}$$

Recursive Linear Least Squares

$$\theta_{ML}(N) = \underset{\theta \in \mathbb{R}}{\arg\min} \frac{1}{2} \|y_N - \Phi_N \cdot \theta\|_2^2 \qquad \text{(forgetting factor: } \alpha)$$

$$\hat{\theta}_{ML}(N+1) = \underset{\theta \in \mathbb{R}^d}{\arg\min} \left(\alpha \cdot \frac{1}{2} \cdot \|\theta - \hat{\theta}_{ML}(N)\|Q_N^2 + \frac{1}{2} \cdot \|y(N+1) - \varphi(N+1)^T \cdot \theta\|_2^2\right)$$

Q_0 given, and $\hat{\theta}_{ML}(0)$ given

$$Q_{N+1} = \alpha \cdot Q_N + \varphi(N+1) \cdot \varphi(N+1)^T$$

$$\hat{\theta}_{ML}(N+1) = \hat{\theta}_{ML}(N) + Q_{N+1}^{-1} \cdot \varphi(N+1) \cdot [y(N+1) - \varphi(N+1)^T \cdot \hat{\theta}_{ML}(N)]$$

Cramer-Rao-Inequality (Fisher information Matrix M)

$$\Sigma_{\hat{\theta}} \succeq M^{-1} = (\Phi_N^T \cdot \Sigma^{-1} \cdot \Phi)^{-1} \quad M = \int_{y_n} \nabla_\theta^2 L(\theta_0, y_n) \cdot p(y_n|\theta_0) dy_n$$

Assumptions:

- Minimising a Linear Model
- Gaussian Noise: $X \sim \mathcal{N}(0, \Sigma)$)

$$L(\theta, y_N) = -\log(p(y_N|\theta))$$

$$= \frac{1}{2} \cdot (\Phi_N \cdot \theta - y_N)^T \cdot \Sigma^{-1} \cdot (\Phi_N \cdot \theta - y_N)$$

$$M = \mathbb{E}\{\nabla_\theta^2 L(\theta, y_N)\} = \nabla_\theta^2 L(\theta, y_N) = \Phi_N^T \cdot \Sigma^{-1} \cdot \Phi_N$$

$\Rightarrow W = \Sigma^{-1}$ is the optimal weighting Matrix for WLS.

Continuous Time Systems

Ordinary Differential Equations (ODE):

$$\dot{x} = f(x(t), u(t), \epsilon(t), p)$$

Differential Algebraic Equations(DAE):

$$\dot{x} = f(x(t), u(t), \epsilon(t), p)$$

$$0 = g(x, z).$$

LTI Sytem (ODE):

$$\dot{x} = Ax + Bu \quad y = Cx + Du$$

$$G(s) = C(sI - A)^{-1} B + D$$

Numerical Integration Methods

Euler Integration Step

$$\tilde{x}(t; x_{init}, u_{const}) = x_{init} + tf(x_{init}, u_{const}), \quad t \in [0, \Delta t]$$

$$\tilde{x}_{j+1} = \tilde{x}_j + hf(\tilde{x}_j, u_{const}), \quad j = 0, ..., M-1$$

- Approximation becomes better by decreasing the step size h.
- Concistency Error: h^2
- Total Number of steps: $\Delta t / h$
- Error in the final step of order $h \Delta t$
- Linear in step size $\rightarrow$ order one
- Taking more steps is more accurate but needs more computional time

Runge-Kutta Method of Order Four

$$k_1 = f(\tilde{x}_j, u_{const})$$

$$k_2 = f(\tilde{x}_j, \frac{h}{2} k_1, u_{const})$$

$$k_3 = f(\tilde{x}_j, \frac{h}{2} k_2, u_{const})$$

$$k_4 = f(\tilde{x}_j, hk_3, u_{const})$$

$$\tilde{x}_{j+1} = \tilde{x}_j + \frac{h}{6}(k_1 + 2k_2 + 2k_3 + k_4)$$

One Step of RK4 is thus as expensive as four steps of euler
accurency of final approximation is of order $h^4 \Delta$ t
$\rightarrow$ rk4 needs fewer functions to obtain the same accuracy level as euler

Discrete Time Systems

Det. Model as State Space	Stoch. Model as State Space
Det. Model as Input-Output	Stoch. Model as Input-Output

State Space Model

$x_{k+1} = f_k(x_k, u_k), k = 0, 1, \ldots, N-1$ with input vector u_k and state vector x_k

Input-Output Model

$y(k) = h(u(k), \ldots, u(k-n), y(k-1), \ldots, y(k-n))$

LTI system as State-Space Model:

$$x_{k+1} = Ax_k + Bu_k, \quad k = 0, 1, ..., N-1.$$

LTI system as Input-Output Model:

$$G(s) = \frac{b_0 + b_1 s + \ldots + b_n s^n}{a_0 + a_1 s + \ldots + a_{n-1} s^{n-1} + s^n} \quad | \cdot s = z^{-1}$$

$$G(z) = \frac{b_0 + b_1 z^{-1} + \ldots + b_n z^{-n}}{a_0 + a_1 z^{-1} + \ldots + a_n z^{-n}}$$

$$= \frac{b_0 z^n + b_n z^{n-1} + \ldots + b_n}{a_0 z^n + a_1 z^{n-1} + \ldots + a_n} \quad \Rightarrow \text{Also called "polynomial model".}$$

Deterministic Model

The output of the system can be obtained with absolute certainty. The Output y or the state x, depend on the known inputs $u(1), \dots, u(N)$, the previous Outputs $y(1), \dots, y(N)$ or state $x(n-1)$ and initial conditions.

State Space Model

$$x(t+1) = f(x(k), u(k))$$
$$y(k) = g(x(k), u(k))$$

Initial conditions: $x(1) = x_{init}$

Input-Output Model

$$y(k) = h(u(k), \dots, u(k-n), y(k-1), \dots, y(k-n))$$

Initial conditions: $y(1) = y_1, \dots, y(n) = y_n \quad u(1) = u_1, \dots, u(n) = u_n$

Finite Impulse Response (FIR):

$$y(k) = b_0 u(k) + \dots + b_{n_b} u(k-n_b)$$

$$G(z) = b_0 + b_1 z^{-1} + \dots + b_{n_b} z^{-n_b} \quad | \cdot \frac{z^{n_b}}{z^{n_b}}$$
$$= \frac{b_0 z^{n_b} + b_1 z^{n_b - 1} + \dots + b_{n_b}}{z^{n_b}}$$

Auto Regressive Models with Exogenous Inputs (ARX):

$$a_0 y(k) + \dots + a_{n_a} y(k-n_a) = b_0 u(k) + \dots + b_{n_b} u(k-n_b)$$

$$G(z) = \frac{b_0 z^n + b_1 z^{n-1} + \dots + b_n}{a_0 z^n + a_1 z^{n-1} + \dots + a_n}$$

The next output depends on the previous output. Also called **IIR** (infinite impulse response)

Stochastic Model

Real systems are far from deterministic.

- there is stochastic noise $\epsilon(k)$
- there are constant and unknown parameters p
- measured outputs depend $y(k)$ depend in both, $\epsilon(k)$ and p

Assumptions: noise is **i.i.d** and enters the model like a normal input, but as a random variable

State Space Model

$$x(t+1) = f(x(k), u(k), \epsilon(k))$$
$$y(k) = g(x(k), u(k), \epsilon(k))$$

Input-Output Model

$$y(k) = h(u(k), \dots, u(k-n), y(k-1), \dots, y(k-n), \epsilon(k), \dots, \epsilon(k-n))$$
$$\text{for} \quad k = n+1, n+2, \dots$$

Measurement Noise (Output Error Model)

$$y(k) = M(k; U, x_{init}, p) + \epsilon(k)$$

Stochastic Disturbance (Equation Errors)

$$y(k) = h(p, u(k), \dots, u(k-n), y(k-1), \dots, y(k-n)) + \epsilon(k)$$
$$for \quad k = n+1, n+2, \dots$$

Linear In the Parameters models (LIP):

$$y(k) = \sum_{i=1}^{d} \theta_i \phi_i(u(k) \dots, y(k-1), \dots) + \epsilon(k)$$
$$y(k) = \varphi(k)^T \theta + \epsilon(k) \quad \text{where } \varphi = (\phi_1(\cdot), \dots, \phi_d(\cdot))$$

LIP-LTI Models with Equation Errors (ARX)

combining best of two worlds (LTI and LIP)

$$a_0 y(k) + \dots + a_{n_a} y(k-n_a) = b_0 u(k) + \dots + b_{n_b} u(k-n_b) + \epsilon(k)$$

Auto-regressive moving average with eXogeneous input (ARMAX):

$$a_0 y(k) + \dots + a_{n_a} y(k-n_a)$$
$$=$$
$$b_0 u(k) + \dots + b_{n_b} u(k-n_b) + \epsilon(k) + c_1 \epsilon(k-1) + \dots + c_{n_x} \epsilon(k-n_c)$$

Auto-regressive moving average without inputs (ARMA):

$$a_0 y(k) + \dots + a_{n_a} y(k-n_a)$$
$$=$$
$$\epsilon(k) + c_1 \epsilon(k-1) + \dots + c_{n_x} \epsilon(k-n_c)$$

Where c_i represent the noise coefficient Have to use non-linear leas squares with the unknown noise terms$\epsilon(k-i)$

Difference Deterministic and Stochastic Models

$$-stochasticnoise\epsilon(k)$$
$$-unknownbutconstantparameterp$$
$$-measuredoutputy(k)dependonboth, \epsilon(k)andp$$

Example for State Space Model

$\ddot{a} = m \cdot \dot{a} + g \cdot a + c \cdot u$

$y = \dot{a}$

$$x = \begin{bmatrix} a \\ \dot{a} \end{bmatrix} \quad \dot{x} = \begin{bmatrix} \dot{a} \\ \ddot{a} \end{bmatrix} \quad \dot{x} = Ax + Bu \quad y = Cx + Du$$

$$A = \begin{bmatrix} 0 & 1 \\ g & m \end{bmatrix} \quad B = \begin{bmatrix} 0 \\ c \end{bmatrix} \quad C = \begin{bmatrix} 0 & 1 \end{bmatrix} \quad D = \begin{bmatrix} 0 \end{bmatrix}$$

check it

Pure Output Error (OE) Minimization

Assume: i.i.d. gaussian noise only affecting output using non-linear least squares

$$\theta_{ML} = \min_{\theta} \sum_{k=1}^{N} (y(k) - M(k; U, x_{init} p))^2$$

Output Error Minimization for FIR Models: lead to convex problems, therefore global minimum can be found

$$y(k) = (u(k), u(k-1), \dots, u(k-n_{n_b})) \cdot \theta + \varepsilon(k)$$
$$= \min_{\theta} \sum_{k=n_b+1}^{N} (y(k) - \underbrace{(u(k), u(k-1), \dots, u(k-n_{n_b}))}_{\text{Deterministic part is also } M(k;U,x_{init},p)} \cdot \theta)^2$$

They often need a very high dimension n_b to obtain a reasonable fit. As a consequence ARX models are usually used instead.

Equation Error Minimization: Assume: i.i.d. $\epsilon(k)$ noise enters the input-output equation as additive disturbance

$$y(k) = h(p, u(k), \dots, u(k-n), y(k-1), \dots, y(k-n)) + \epsilon(k)$$
$$\text{for} \quad k = n+1, n+2$$

if the i.i.d noise is gaussian, a maximum likelihood formulation to estimate the unknown parameter vector $\theta = p$ is given:

$$\theta_{ML} = \min_{\theta} \sum_{k=n+1}^{N} (y(k) - h(p, u(k), \dots, y(k-1), \dots)))^2$$

u and k are known input and output measurements, and the algorithm minimises the so called **equation errors** or **prediction errors**.
This problem is also known as **Prediction error minimisation(PEM)**
Such a problem is convex if p enters linearly in f, i.e. if the model is **linear-in-the-parameters (LIP)**

PEM of LIP Models

$$y(k) = \varphi(k)^T \theta + \epsilon(k)$$
$$\text{where } \varphi = (\phi_1(\cdot), \dots, \phi_d(\cdot))^T \text{are the regressor variables}$$

considering this last expression, the prediction error minimisation(PEM) problem can be formulated as:

$$\min_{\theta} \underbrace{\sum_{k=1}^{N} (y(k) - \varphi(k)^{\mathrm{T}} \theta)^2}_{= \| y_N - \Phi_N \theta \|_2^2}$$

Which can be solved using LLS $\theta^* = \Phi_N^+ y_N$

Special Case: PEM of LIP-LTI Models with Equation Errors(ARX) General ARX model equation

$$a_0 y(k) + \dots + a_{n_a} y(k-n_a) = b_0 u(k) + \dots + b_{n_b} u(k-n_b) + \epsilon(k)$$

In order to have a determined estimation problem, a_0 has to be fixed, otherwise the number of optimal solutions would be infinitive. Therefore we sually fix $a_0 = 1$ and use $\theta = (a_1, \dots, a_{n_a}, b_0, \dots, b_{n_b})^{\mathrm{T}}$ as the parameter estimation vector. The regressor vector is given by:

$$\varphi = (-y(k-1), \dots, -y(k-n_a), u(k), \dots, u(k-n_b))^{\mathrm{T}}$$

leading to the optimal solution provided by LLS:

$$y(k) = \varphi(k)^{\mathrm{T}} \theta + \epsilon(k)$$

Pure Output Error (OE) Minimization

Models with Input and Output Errors:

$$y(k) = M(k; U + \varepsilon_N^u, x_{init}, p) + \epsilon^y(k)$$

model.pdf

Assume: i.i.d. gaussian noise on both input and output with variance σ_u^2 for the input and σ_y^2 for the output

$$\underset{\theta}{arg\,min} \sum_{k-1}^{N} \frac{1}{\sigma_y^2}(y(k) - M(k; U + \epsilon_N^u, x_{init}, p))^2 + \frac{1}{\sigma_u^2}(\epsilon_u(k))^2$$

$$\underset{\theta}{arg\,min} \sum_{k-1}^{N} \frac{1}{\sigma_y^2}(y(k) - M(k; \tilde{U}, x_{init}, p))^2 + \frac{1}{\sigma_u^2}(u(k) - \tilde{u}(k))^2$$

Fourier Transformation

FT:

$$F\{F\}(\omega) = \int_{-\infty}^{\infty} f(t)e^{-j\omega t} dt$$

iFT:

$$f(t) = F^{-1}\{F\}(t) = \frac{1}{2\pi}\int_{-\infty}^{\infty} F(\omega)e^{-j\omega t} d\omega$$

DFT:

$$U(m) := \sum_{k=0}^{N-1} u(t)e^{-j\frac{2\pi mk}{N}}$$

iDFT:

$$u(n) := \sum_{k=0}^{N-1} U(k)e^{j\frac{2\pi kn}{N}}$$

Useful frequency things

$\omega = 2\pi f = \frac{2\pi}{T}$ $\quad f_s > 2f_{max}$ $\quad T = N\Delta t = \frac{N}{f_s}$

Aliasing and Leakage Errors

Aliasing Error: Due to sampling of continous signal to discrete signal. Avoid with Nyquist Theoreme:

$$f_{Nyquist} = \frac{1}{2\Delta t}[\text{Hz}] \quad or \quad \omega_{Nyquist} = \frac{2\pi}{2\Delta t}[\text{rad/s}]$$

Leackage Error: Due to windowing.

$$\omega_{base} := \frac{2\pi}{N \cdot \Delta t} = \frac{2\pi}{T} \rightarrow \omega = m\frac{2\pi}{N \cdot \Delta t}$$

Crest Factor = Scheitelfaktor

$$\text{Crest Factor} = \frac{u_{max}}{u_{rms}} \quad \text{with} : u_{rms} := \sqrt{\frac{1}{T}\int_0^T u(t)^2\, dt}$$

$$\text{and} \quad u_{max} := \underset{t\in[0,T]}{max} |u(t)|$$

Optimising Multisine for optimal crest factor

Frequency: Choose frequencies in logarithmic manner as multiples of the base frequency. $\omega_{k+1}/\omega_k \approx 1.05$

Phase: To prevent high peaks (Crest Factor) in the Signal, the phases of the different frequencies are modulated accordingly. (Positive interference)

Multisine Identification Implementation procedure

Window Length: Integer multiple of sampling time: $T = N \cdot \delta t$

Harmonics of base frequency: Are contained in multisine $\omega_{base} = \frac{2\pi}{T}$

Highest contained Frequency: Is **half** of Nyquist frequency: $\omega_{Nyquist} = \frac{2\pi}{4\Delta T}$

Experiment and Analysis: (Step 2): Insert Multisine periodically. Drop first Periods (till transients died out). Record M Periods, each with N samples, of input and output data. Average all the M periods and make the DFT (or vice versa). Finally build transfer function: $\hat{G}_{j\omega_k} = \frac{\hat{Y}(k(p))}{\hat{U}(k(p))}$

Nonparametric and Frequency Domain Identification Models

Impulse response and transfer function:

$$y(t) = \int_0^{\infty} g(\tau)u(t-\tau)\,\delta t$$

$$Y(s) = G(s) \cdot U(s)$$

$$G(s) = \int_0^{\infty} \epsilon^{-st} g(t)\, dt$$

Bode diagram from frequency sweeps:

$$u(t) = A \cdot sin(\omega \cdot t), \quad y(t) = \|G(j \cdot \omega)\| A \cdot sin(\omega \cdot t + \alpha)$$

Bode Diagramm

Magnitude = Amplitude $|G(j\omega)|$

Phase $arg\, G(j\omega)$

Hier sollten wir glaub noch bisschen was rein machen.

Recursive Least Squares

New Inverse Covariance: $Q_K = Q_{k-1} + \phi_K \phi_K^T$

Kalman Filter

Valid for Discrete and Linear!

If recursive least squares: $x_{k+1} = A_k \cdot x_k$

$x_{k+1} = A_k \cdot x_k + \omega_k$ and $y_k = C_k \cdot x_k + v_k$

Steps of Kalman Filter

1 Prediction

$$\hat{x}_{[k|k-1]} = A_{k-1} \cdot \hat{x}_{[k-1|k-1]}$$

$$P_{[k|k-1]} = A_{k-1} \cdot P_{[k-1|k-1]} \cdot A_{k-1}^T \cdot W_{k-1}$$

If RLS, without: W_{k-1}

2 Innovation update

$$P_{[k|k]} = (P_{[k|k-1]}^{-1} + C_k^T \cdot V^{-1} \cdot C_k)^{-1}$$

$$\hat{x}_{[k|k]} = \hat{x}_{[k|k-1]} + P_{[k|k]} \cdot C_k^T \cdot V^{-1} \cdot (y_k - C_k \cdot \hat{x}_{[k|k-1]})$$