



# Variance Reduction is an Antidote to Byzantine Workers: Better Rates, Weaker Assumptions and Communication Compression as a Cherry on the Top



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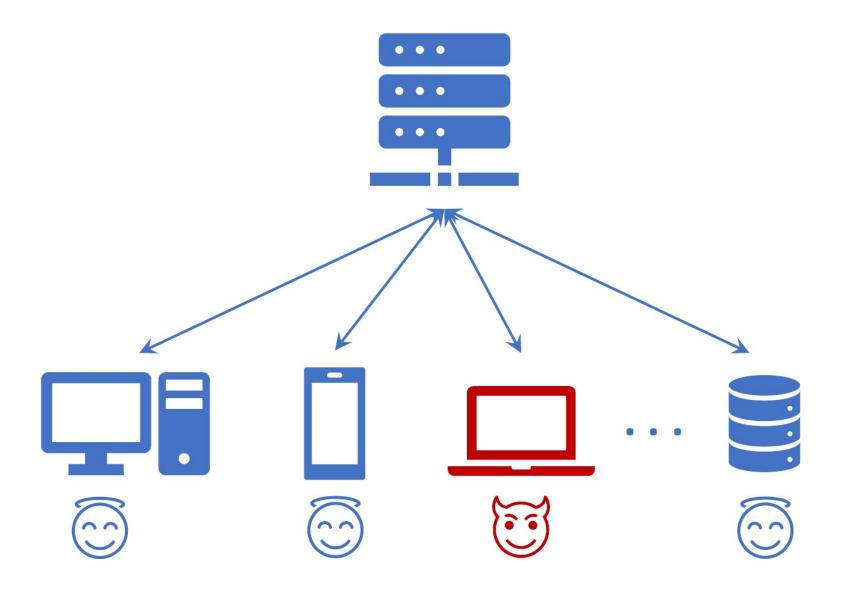
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# 1. Byzantine-Robust Optimization

#### Distributed optimization problem:

$$\min_{x \in \mathbb{R}^d} \left\{ f(x) = \frac{1}{\mathcal{G}} \sum_{i \in \mathcal{G}} f_i(x) \right\}, \quad f_i(x) = \frac{1}{m} \sum_{j=1}^m f_{i,j}(x) \quad \forall i \in \mathcal{G}$$

- $\bullet$   $\mathcal{G}$  is the set of good clients
- $\mathcal{B}$  is the set of  $Byzantine\ workers$  the workers that can arbitrarily deviate from the prescribed protocol (maliciously or not) and are assumed to be omniscient
- $\mathcal{G} \sqcup \mathcal{B} = [n]$  is the set of clients participating in training



#### Main difficulties in Byzantine-robust optimization:

- When functions are arbitrarily heterogeneous, the problem is impossible to solve
- Fraction of Byzantines  $\delta = B/n$  should be smaller than 1/2
- Standard approaches based on averaging are vulnerable
- Robust aggregation alone does not ensure robustness [1]

# 2. Robust Aggregation

#### Popular aggregation rules:

- $\operatorname{Krum}(x_1, \dots, x_n) := \operatorname{argmin}_{x_i \in \{x_1, \dots, x_n\}} \sum_{j \in S_i} ||x_j x_i||^2$  [7], where  $S_i \subseteq \{x_1, \dots, x_n\}$  are  $n - |\mathcal{B}| - 2$  closest vectors to  $x_i$
- Robust Fed. Averaging:  $RFA(x_1, ..., x_n) := \operatorname{argmin}_{x \in \mathbb{R}^d} \sum_{i=1}^n \|x x_i\|$
- Coordinate-wise Median:  $[CM(x_1, ..., x_n)]_t := \operatorname{argmin}_{u \in \mathbb{R}} \sum_{i=1}^n |u [x_i]_t|$ These defenses are vulnerable to Byzantine attacks [8,9] and do not satisfy the following definition.

#### Definition 1: $(\delta, c)$ -Robust Aggregator (modification of the definition from [1]

Assume that  $\{x_1, x_2, \ldots, x_n\}$  is such that there exists a subset  $\mathcal{G} \subseteq [n]$  of size  $|\mathcal{G}| = G \ge (1 - \delta)n$  for  $\delta < 0.5$  and there exists  $\sigma \geq 0$  such that  $\frac{1}{G(G-1)} \sum_{i,l \in \mathcal{G}} \mathbb{E}[\|x_i - x_l\|^2] \leq \sigma^2$  where the expectation is taken w.r.t. the randomness of  $\{x_i\}_{i\in\mathcal{G}}$ . We say that the quantity  $\hat{x}$  is  $(\delta, c)$ -Robust Aggregator  $((\delta, c)$ -RAgg) and write  $\hat{x} = \text{RAgg}(x_1, \dots, x_n)$  for some c > 0, if the following inequality holds:

$$\mathbb{E}\left[\|\widehat{x} - \overline{x}\|^2\right] \le c\delta\sigma^2,\tag{1}$$

where  $\overline{x} = \frac{1}{|\mathcal{G}|} \sum_{i \in \mathcal{G}} x_i$ . If additionally  $\widehat{x}$  is computed without the knowledge of  $\sigma^2$ , we say that  $\hat{x}$  is  $(\delta, c)$ -Agnostic Robust **Aggregator**  $((\delta, c)$ -ARAgg) and write  $\widehat{x} = ARAgg(x_1, \dots, x_n)$ .

One can robustify Krum, RFA, and CM using bucketing [1].

Algorithm Bucketing: Robust Aggregation using bucketing [1]

- 1: Input:  $\{x_1,\ldots,x_n\}$ ,  $s\in\mathbb{N}$  bucket size, Aggr aggregation
- 2: Sample random permutation  $\pi = (\pi(1), \dots, \pi(n))$  of [n]
- 3: Compute  $y_i=rac{1}{s}\sum_{k=s(i-1)+1}^{\min\{si,n\}}x_{\pi(k)}$  for  $i=1,\ldots,\lceil n/s 
  ceil$
- 4: **Return:**  $\widehat{x} = \operatorname{Aggr}(y_1, \dots, y_{\lceil n/s \rceil})$

#### 3. SGD and Variance Reduction

# **SGD**: $x^{k+1} = x^k - \gamma g^k$ , $g^k = \frac{1}{n} \sum_{i=1}^n \nabla f_{i,j_i^k}(x^k)$

- lacksquare Variances of the estimators  $\nabla f_{i,j^k}(x^k)$  do not go to zero
- X Byzantines can easily hide in the noise and create a large bias (even if the aggregation is robust)

**SAGA** [2]: 
$$x^{k+1} = x^k - \gamma g^k$$
,  $g^k = \frac{1}{n} \sum_{i=1}^n g_i^k$ ,

$$g_i^k = \nabla f_{j_i^k}(x^k) - \nabla f_{i,j_i^k}(w_{i,j_i^k}^k) + \frac{1}{m} \sum_{i=1}^m \nabla f_{i,i}(w_{i,j}^k)$$

- ✓ Variances of the estimators  $g_i^k$  go to zero
- $m{\times}$  Analysis relies on the unbiasedness:  $\mathbb{E}[g_i^k \mid x^k] = \nabla f_i(x^k)$

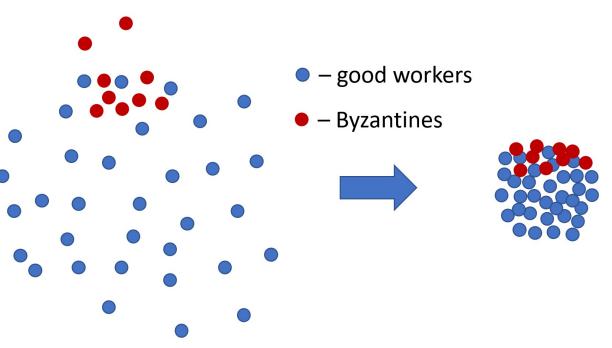
# SARAH/Geom-SARAH/PAGE [3,4,5]:

$$\overline{x^{k+1}} = x^k - \gamma g^k, \quad g^k = \frac{1}{n} \sum_{i=1}^n g_i^k,$$

$$g_i^k = \begin{cases} \nabla f_i(x^k), & \text{with prob. } p, \\ g_i^{k-1} + \nabla f_{i,j_i^k}(x^k) - \nabla f_{i,j_i^k}(x^{k-1}), & \text{with prob. } 1 - p \end{cases}$$

- $\checkmark$  Variances of the estimators  $g_i^k$  go to zero
- ✓ Analysis does not rely on the unbiasedness:  $\mathbb{E}[g_i^k \mid x^k] \neq \nabla f_i(x^k)$

How can variance reduction help? It leaves less space for Byzantines to hide in the noise.



#### Main Contributions

- ♦ New method: Byz-VR-MARINA. We make VR-MARINA (VR-method with compression) [6] applicable to Byzantinerobust learning using robust agnostic aggregation [1].
- ♦ New SOTA results under more general assumptions. Under quite general assumptions (no strong assumptions on the compression and second moment of the stochastic gradient; non-uniform sampling is supported), we prove new theoretical convergence results that are tight and outperform known ones when the target accuracy is small enough.

#### 4. Technical Preliminaries

#### Definition 2: Unbiased Compression

Stochastic mapping  $\mathcal{Q}: \mathbb{R}^d \to \mathbb{R}^d$  is called unbiased compressor/compression operator if there exists  $\omega \geq 0$  such that for any  $x \in \mathbb{R}^d$ 

$$\mathbb{E}\left[\mathcal{Q}(x)\right] = x, \quad \mathbb{E}\left[\|\mathcal{Q}(x) - x\|^2\right] \le \omega \|x\|^2. \tag{2}$$

#### Assumptions

- Smoothness and lower-boundedness:  $\forall x,y \in \mathbb{R}^d$  we have  $\|\nabla f(x) - \nabla f(y)\| \le L\|x - y\|$  and  $f_* = \inf_{x \in \mathbb{R}^d} f(x) > -\infty$
- $\zeta^2$ -heterogeneity:  $\frac{1}{G} \sum_{i \in \mathcal{G}} \|\nabla f_i(x) \nabla f(x)\|^2 \le \zeta^2 \quad \forall x \in \mathbb{R}^d$
- Global Hessian variance assumption:
- $\frac{1}{G} \sum_{i=1}^{n} \|\nabla f_i(x) \nabla f_i(y)\|^2 \|\nabla f(x) \nabla f(y)\|^2 \le L_{\pm}^2 \|x y\|^2$
- Local Hessian variance assumption:
- $\frac{1}{G} \sum_{i \in C} \mathbb{E} \|\widehat{\Delta}_i(x, y) \Delta_i(x, y)\|^2 \le \frac{\mathcal{L}_{\pm}^2}{b} \|x y\|^2$ , where  $\Delta_i(x, y) = 0$

 $\nabla f_i(x) - \nabla f_i(y)$  and  $\widehat{\Delta}_i(x,y)$  is an unbiased mini-batched estimator of  $\Delta_i(x,y)$  with batch size b

#### 5. New Method: Byz-VR-MARINA

### Algorithm Byz-VR-MARINA: Byzantine-tolerant VR-MARINA

- : **Input:** starting point  $x^0$ , stepsize  $\gamma$ , minibatch size b, probability  $p \in (0,1]$ , number of iterations K,  $(\delta,c)$ -ARAgg
- 2: **for**  $k = 0, 1, \dots, K 1$  **do**
- Get a sample from Bernoulli distribution with parameter p:  $c_k \sim$ Be(p). Broadcast  $g^k$ ,  $c_k$  to all workers
- 4: **for**  $i \in \mathcal{G}$  in parallel **do**
- 5:  $x^{k+1}=x^k-\gamma g^k$
- 6: Set  $g_i^{k+1}=egin{cases} \nabla f_i(x^{k+1}), & \text{if } c_k=1, \\ g^k+\mathcal{Q}\left(\widehat{\Delta}_i(x^{k+1},x^k)
  ight), & \text{otherwise,} \end{cases}$  where minibatched estimator  $\widehat{\Delta}_i(x^{k+1},x^k)$  of  $\nabla f_i(x^{k+1}) - \nabla f_i(x^k)$ ;  $\mathcal{Q}(\cdot)$  for  $i \in \mathcal{G}$  are computed independently
- end for
- $g^{k+1} = \mathtt{ARAgg}(g_1^{k+1}, \dots, g_n^{k+1})$
- 9: end for

# 6. Convergence in the Non-Convex Case

Let the introduced assumptions hold. Assume that 0 <  $\gamma \leq \frac{1}{L+\sqrt{A}}$ , where  $A = \frac{6(1-p)}{p} \left(\frac{4c\delta}{p} + \frac{1}{2G}\right) \left(\omega L^2 + \frac{(1+\omega)\mathcal{L}_{\pm}^2}{b}\right) + \frac{1}{2G}$  $\frac{6(1-p)}{p}\left(\frac{4c\delta(1+\omega)}{p}+\frac{\omega}{2G}\right)L_{\pm}^2$ . Then for all  $K\geq 0$  the point  $\widehat{x}^K$  chosen uniformly at random from the iterates  $x^0, x^1, \ldots, x^K$  produced by Byz-VR-MARINA satisfies

$$\mathbb{E}\left[\|\nabla f(\widehat{x}^K)\|^2\right] \le \frac{2\Phi_0}{\gamma(K+1)} + \frac{24c\delta\zeta^2}{p},\tag{3}$$

where  $\Phi_0 = f(x^0) - f_* + \frac{\gamma}{p} ||g^0 - \nabla f(x^0)||^2$  and  $\mathbb{E}[\cdot]$  denotes the full expectation.

• When  $\zeta = 0$  (homogeneous data) the method converges asymptotically to the exact solution with rate  $\mathcal{O}(1/K)$ 

# 7. Convergence in PŁ-case

#### Definition 3: Polyak-Łojasiewicz (PŁ) condition

Function f satisfies Polyak-Łojasiewicz (PŁ) condition with parameter  $\mu$  if for all  $x \in \mathbb{R}^d$  there exists  $x^* \in \operatorname{argmin}_{x \in \mathbb{R}^d} f(x)$  such

$$\|\nabla f(x)\|^2 \ge 2\mu \left( f(x) - f(x^*) \right). \tag{4}$$

#### Theorem 1

Let the introduced assumptions hold and function f satisfies  $\mu$ -PŁ condition. Assume that  $0 < \gamma \le \min\left\{\frac{1}{L+\sqrt{2A}}, \frac{p}{4u}\right\}$ , where A = 0 $\frac{6(1-p)}{p}\left(\frac{4c\delta}{p}+\frac{1}{2G}\right)\left(\omega L^2+\frac{(1+\omega)\mathcal{L}_{\pm}^2}{b}\right) + \frac{6(1-p)}{p}\left(\frac{4c\delta(1+\omega)}{p}+\frac{\omega}{2G}\right)L_{\pm}^2.$ Then for all  $K \geq 0$  the iterates produced by Byz-VR-MARINA satisfy

$$\mathbb{E}\left[f(x^{K}) - f(x^{*})\right] \leq (1 - \gamma \mu)^{K} \Phi_{0} + \frac{24c\delta\zeta^{2}}{\mu},$$
where  $\Phi_{0} = f(x^{0}) - f_{*} + \frac{2\gamma}{n} \|g^{0} - \nabla f(x^{0})\|^{2}.$ 
(5)

• When  $\zeta = 0$  (homogeneous data) the method converges linearly asymptotically to the exact solution

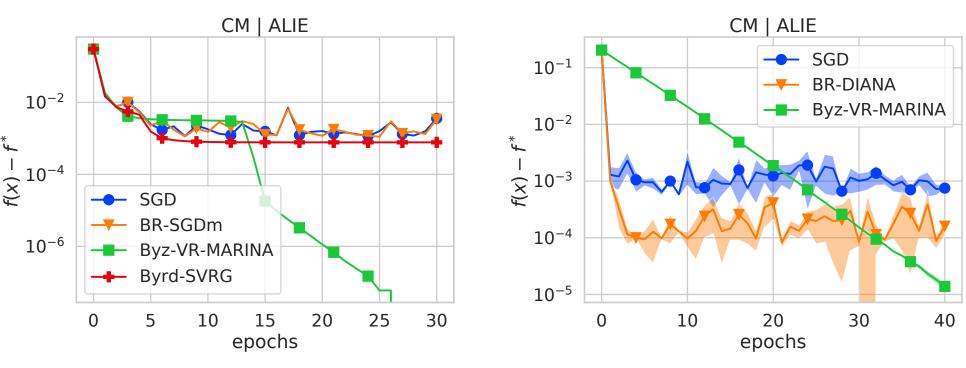
## 8. Comparison with Prior Work

|   | Setup                | Method                   | Complexity (NC)   | Complexity (PŁ)   |
|---|----------------------|--------------------------|---|---|
|   | Hom. data, no compr. | BR-SGDm [1]              | $\frac{1}{\varepsilon^2} + \frac{\sigma^2(c\delta + 1/n)}{b\varepsilon^4}$  | X   |
|   |                      | BR-MVR [1]               | $\frac{\frac{1}{\varepsilon^2} + \frac{\sigma\sqrt{c\delta + 1/n}}{\sqrt{b}\varepsilon^3}}{\frac{1}{\varepsilon^2} + \frac{n^2\delta\sigma^2}{Cb\varepsilon^2} + \frac{\sigma^2}{nb\varepsilon^4}}$ | X   |
|   |                      | BTARD-SGD [10]           | $\frac{1}{\varepsilon^2} + \frac{n^2 \delta \sigma^2}{Cb\varepsilon^2} + \frac{\sigma^2}{nb\varepsilon^4}$  | $\frac{1}{\mu} + \frac{\sigma^2}{nb\mu\varepsilon} + \frac{n^2\delta\sigma}{C\sqrt{b\mu\varepsilon}}$   |
|   |                      | $Byrd\text{-}SAGA\ [11]$ | X   | $rac{m^2}{b^2(1-2\delta)\mu^2}$  |
|   |                      | Byz-VR-MARINA            | $\frac{1+\sqrt{\frac{c\delta m^2}{b^3}+\frac{m}{b^2n}}}{\varepsilon^2}$   | $\frac{1+\sqrt{\frac{c\delta m^2}{b^3}+\frac{m}{b^2n}}}{\mu} + \frac{m}{b}$   |
|   | Het. data, no compr. | BR-SGDm [1]              | $\frac{1}{\varepsilon^2} + \frac{\sigma^2(c\delta + 1/n)}{b\varepsilon^4}$  | <b>X</b>  |
|   |                      | Byrd-SAGA $[11]$         | X   | $\frac{m^2}{b^2(1-2\delta)\mu^2}$   |
|   |                      | Byz-VR-MARINA            | $\frac{1+\sqrt{\frac{c\delta m^2}{b^2}(1+\frac{1}{b})+\frac{m}{b^2n}}}{\varepsilon^2}$  | $\frac{1+\sqrt{\frac{c\delta m^2}{b^2}(1+\frac{1}{b})+\frac{m}{b^2n}}}{\mu} + \frac{m}{b}$  |
|   | Het. data, compr.    | BR-CSGD [12]             | X   | $\frac{1}{\mu^2}$   |
|   |                      | BR-CSAGA [12]            | X   | $\frac{m^2}{b^2\mu^2(1-2\delta)^2}$   |
|   |                      | BROADCAST [12]           | X   | $\frac{\overline{b^2 \mu^2 (1{-}2\delta)^2}}{\frac{m^2 (1{+}\omega)^{3/2}}{b^2 \mu^2 (1{-}2\delta)}}$   |
|   |                      | Byz-VR-MARINA            | $\frac{1+\sqrt{c\delta(1+\omega)(1+\frac{1}{b})}}{\frac{p\varepsilon^2}{\sqrt{pn}\varepsilon^2}} + \frac{\sqrt{(1+\omega)(1+\frac{1}{b})}}{\sqrt{pn}\varepsilon^2}$                                 | $ \frac{1+\sqrt{c\delta(1+\omega)(1+\frac{1}{b})}}{\frac{p\mu}{\sqrt{pn}\mu}} + \frac{\frac{\sqrt{(1+\omega)(1+\frac{1}{b})}}{\sqrt{pn}\mu}}{+\frac{m}{b}+\omega} $ |
| L |                      | _                        | ,   | U   |

- Dependencies on numerical constants (and logarithms in PŁ setting), smoothness constants, and initial suboptimality are omitted
- $p = \min\{b/m, 1/(1+\omega)\}$  = probability of communication in Byz-VR-MARINA
- Analyses of BR-SGDm, BR-MVR, BTARD-SGD, BR-CSGD, BR-CSAGA rely on uniformly bounded variance assumption
- In the het. case, the methods converge only to the error  $\sim \zeta^2$
- The result for BROADCAST is derived for  $\omega \leq \frac{\mu^2(1-2\delta)^2}{56L^2(2-2\delta^2)}$

#### 9. Experiments

- We consider a logistic regression model with  $\ell_2$ -regularization and non-convex regularization  $\lambda \sum \frac{x_i^2}{1+x^2}$
- We have 4 good workers and 1 Byzantine worker
- A Little is enough (ALIE) attack [8] is considered: the Byzantine workers estimate the mean  $\mu_{\mathcal{G}}$  and standard deviation  $\sigma_{\mathcal{G}}$  of the good updates, and send  $\mu_{\mathcal{G}} - z\sigma_{\mathcal{G}}, z > 0$
- Byrd-SVRG a version of Byrd-SAGA with SVRG-estimator instead of SAGA-estimator
- BR-DIANA a version of BROADCAST with SGD-estimator instead of SAGA-estimator



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