

## Econometrics: Methods and Applications



### Peer-graded Assignment: Test Exercise 6

Goals and skills being used:

- Experience the process of practical application of time series analysis.
- Get hands-on experience with the analysis of time series.
- Give correct interpretation of outcomes of the analysis.

### Questions

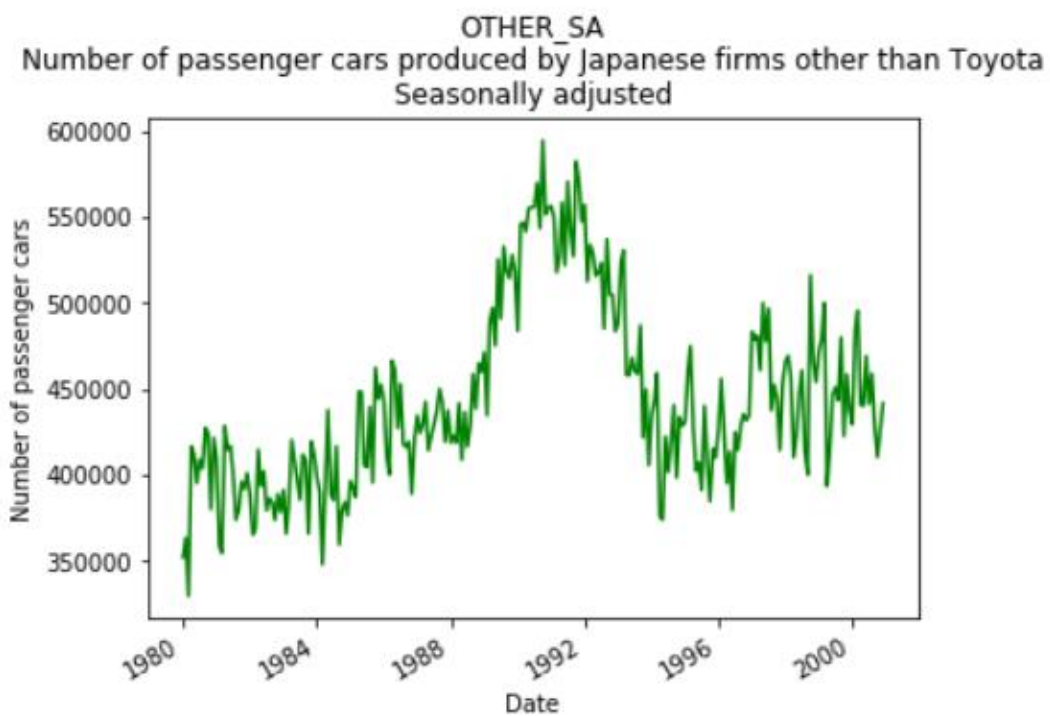
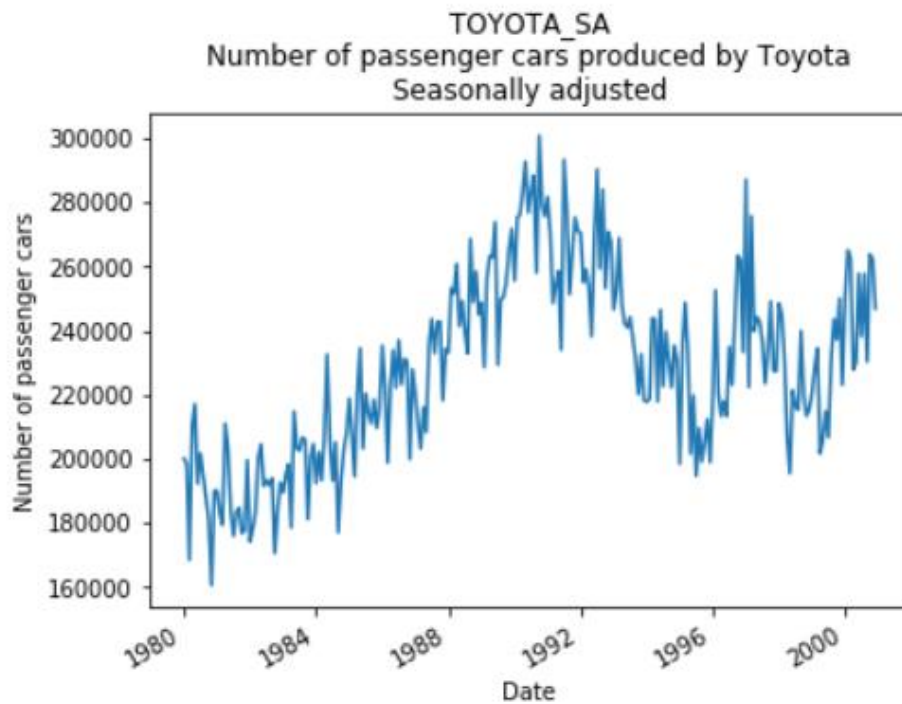
This test exercise uses data that are available in the data file TestExer6. The question of interest is to model monthly production of Toyota passenger cars and to investigate whether the monthly production of all other brands of Japanese passenger cars has predictive power for the production of Toyota. Monthly production data are available from January 1980 until December 2000. The data for January 1980 until December 1999 are used for specification and estimation of models, and the data for 2000 are left out for forecast evaluation purposes.

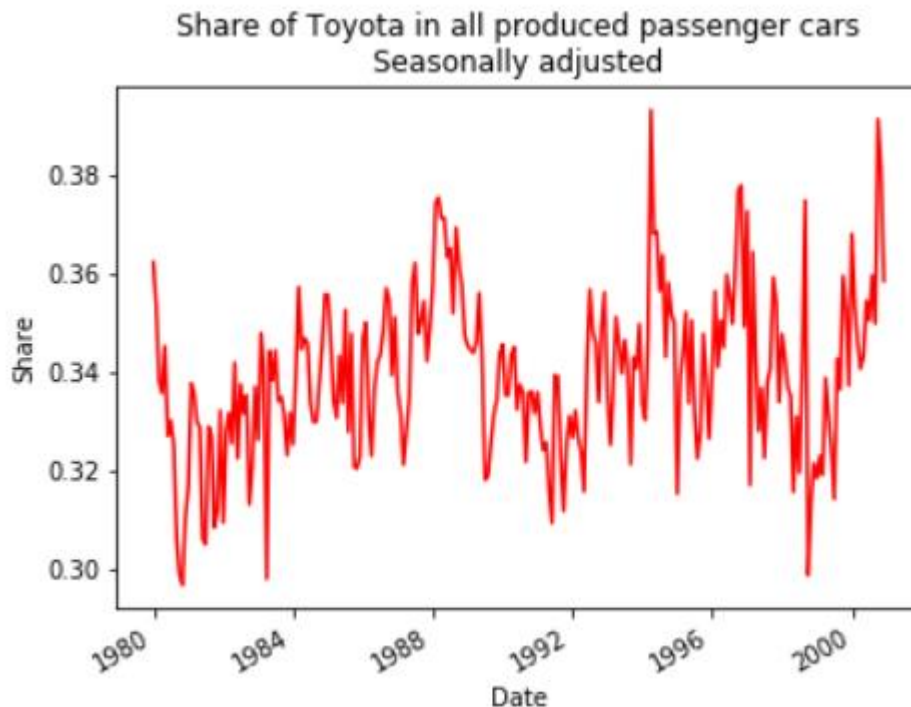
In answering the questions below, you should use the seasonally adjusted production data denoted by '**toyota-sa**' and '**other-sa**'. We will denote these variables by  $y = \text{toyota-sa}$  and  $x = \text{other-sa}$ .

- Make time series plots of the variables  $y_t$  and  $x_t$ , and also of the share of Toyota in all produced passenger cars, that is  $y_t / (y_t + x_t)$ . What conclusions do you draw from these plots?

series plots of the variables  $y_t$  and  $x_t$  ,

$y = \text{TOYOTA\_SA}$       and  
 $x = \text{OTHER\_SA}$





### Answer (a)

The share of passage cars produced by Toyota in all produced passenger cars is relatively stable. It fluctuates between 30% and 38%.

- (b) (i) Perform the Augmented Dickey-Fuller (ADF) test for  $y_t$ . In the ADF test equation, include a constant ( $\alpha$ ) and three lags of  $\Delta y_t$ , as well as the variable of interest,  $y_{t-1}$ . Report the coefficient of  $y_{t-1}$  and its standard error and t-value and draw your conclusion.

### Calculation (b)i

Augmented Dickey-Fuller (ADF) test for  $y_t$

$$\text{ADF: } \Delta y_t = \alpha + \beta t + \rho y_{t-1} + \sum_{j=1}^3 \gamma_j \Delta y_{t-j} + \varepsilon_t$$

A time series is said to be “stationary” if it has no trend, exhibits constant variance over time, and has a constant autocorrelation structure over time.

One way to test whether a time series is stationary is to perform an augmented Dickey-Fuller test, which uses the following null and alternative hypotheses:

$H_0$ : The time series is non-stationary. In other words, it has some time-dependent structure and does not have constant variance over time.

$H_A$ : The time series is stationary.

If the p-value from the test is less than some significance level (e.g.,  $\alpha = .05$ ), then we can reject the null hypothesis and conclude that the time series is stationary.

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Results: Ordinary least squares
=====
Model:                OLS                Adj. R-squared:    0.293
Dependent Variable:    delta_y            AIC:              5221.9067
Date:                 2023-01-04 15:50    BIC:              5239.2258
No. Observations:     236                Log-Likelihood:   -2606.0
Df Model:              4                 F-statistic:      25.34
Df Residuals:          231                Prob (F-statistic): 2.05e-17
R-squared:             0.305              Scale:           2.3331e+08
=====
              Coef.      Std.Err.    t      P>|t|    [0.025    0.975]
-----
const        19281.8945  8430.4101   2.2872  0.0231  2671.5701 35892.2190
y_lag1       -0.0832     0.0368  -2.2623  0.0246  -0.1557   -0.0107
delta_y_lag1  -0.5630     0.0699  -8.0565  0.0000  -0.7007   -0.4253
delta_y_lag2  -0.3243     0.0745  -4.3546  0.0000  -0.4710   -0.1776
delta_y_lag3  -0.0639     0.0650  -0.9835  0.3264  -0.1920    0.0641
=====
Omnibus:            0.924          Durbin-Watson:      2.013
Prob(Omnibus):      0.630          Jarque-Bera (JB):   0.736
Skew:               0.132          Prob(JB):           0.692
Kurtosis:           3.072          Condition No.:      1948470
=====
* The condition number is large (2e+06). This might indicate
strong multicollinearity or other numerical problems.

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coefficient of  $y_{t-1} = -0.0832$       std error: 0.0368  
    t-value: -2.2623  
    p-value: 0.0246

### Answer (b)i

We reject  $H_0$  and conclude that the time series  $y_t$  is stationary.

(ii) Perform a similar ADF test for  $x_t$ .

### Calculation (b)ii

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Results: Ordinary least squares
=====
Model:                OLS                Adj. R-squared:      0.262
Dependent Variable:    y                AIC:                5469.0478
Date:                 2023-01-04 18:10    BIC:                5486.3670
No. Observations:      236              Log-Likelihood:      -2729.5
Df Model:              4                F-statistic:         21.86
Df Residuals:          231              Prob (F-statistic):  2.58e-15
R-squared:             0.275             Scale:              6.6487e+08
=====
              Coef.      Std.Err.    t      P>|t|      [0.025    0.975]
-----
xlag1         -0.0696      0.0331   -2.1057  0.0363    -0.1348    -0.0045
delta_x_lag1  -0.5112      0.0675   -7.5701  0.0000    -0.6443    -0.3781
delta_y_lag2  -0.3614      0.0703   -5.1387  0.0000    -0.5000    -0.2228
delta_y_lag3  -0.1030      0.0645   -1.5961  0.1118    -0.2301     0.0241
const       31540.3617 14808.7693  2.1298  0.0342  2362.8412 60717.8822
=====
Omnibus:          2.655             Durbin-Watson:       2.050
Prob(Omnibus):    0.265             Jarque-Bera (JB):    2.495
Skew:             -0.125            Prob(JB):            0.287
Kurtosis:         3.438             Condition No.:       3959566
=====
* The condition number is large (4e+06). This might indicate
strong multicollinearity or other numerical problems.

```

coefficient of  $x_{t-1} = -0.0696$       std error: 0.0331  
    t-value: -2.1057  
    p-value: 0.0363

### Answer (b)ii

The p-value of  $x_{t-1}$  is lower than the threshold value of 0.05. We reject  $H_0$  and conclude that the time series  $x_t$  is stationary.

- (c) Perform the two-step Engle-Granger test for cointegration of the time series  $y_t$  and  $x_t$ . In step 1, regress  $y_t$  on a constant and  $x_t$ . In step 2, perform a regression of the residuals  $e_t$  in the model  $\Delta e_t = \alpha + \rho e_{t-1} + \beta_1 \Delta e_{t-1} + \beta_2 \Delta e_{t-2} + \beta_3 \Delta e_{t-3} + \omega_t$ . What is your conclusion?

Step 2: ADF:  $\Delta e_t = \alpha + \rho e_{t-1} + \beta_1 \Delta e_{t-1} + \beta_2 \Delta e_{t-2} + \beta_3 \Delta e_{t-3} + \omega_t$

Model:	OLS	Adj. R-squared:	0.252			
Dependent Variable:	y	AIC:	5146.6454			
Date:	2023-01-04 19:05	BIC:	5163.9646			
No. Observations:	236	Log-Likelihood:	-2568.3			
Df Model:	4	F-statistic:	20.75			
Df Residuals:	231	Prob (F-statistic):	1.26e-14			
R-squared:	0.264	Scale:	1.6961e+08			
-----						
	Coef.	Std.Err.	t	P> t	[0.025	0.975]
-----						
e_lag1	-0.2930	0.0680	-4.3057	0.0000	-0.4270	-0.1589
delta_e_lag1	-0.2858	0.0785	-3.6396	0.0003	-0.4406	-0.1311
delta_e_lag2	-0.1416	0.0754	-1.8794	0.0614	-0.2901	0.0068
delta_e_lag3	-0.0960	0.0657	-1.4607	0.1454	-0.2254	0.0335
const	24.9917	847.8255	0.0295	0.9765	-1645.4676	1695.4510
-----						
Omnibus:	12.771		Durbin-Watson:		2.010	
Prob(Omnibus):	0.002		Jarque-Bera (JB):		25.834	
Skew:	-0.214		Prob(JB):		0.000	
Kurtosis:	4.563		Condition No.:		19350	
=====						
* The condition number is large (2e+04). This might indicate strong multicollinearity or other numerical problems.						

- if t-value of  $e_{t-1} < -3.8 \rightarrow$  cointegrated
- if t-value of  $e_{t-1} > -3.8 \rightarrow$  not cointegrated

5. Januar 2023

- (d) Construct the first twelve sample autocorrelations and sample partial autocorrelations of  $\Delta y_t$  and use the outcomes to motivate an AR(12) model for  $\Delta y_t$ . Check that only the lagged terms at lags 1 to 5, 10, and 12 are significant, and estimate the following model:
- $$\Delta y_t = \alpha + \sum_{j=1}^5 \beta_j \Delta y_{t-j} + \beta_6 \Delta y_{t-10} + \beta_7 \Delta y_{t-12} + \varepsilon_t$$
- (Recall that the estimation sample is Jan 1980 - Dec 1999).

### Calculation (d)

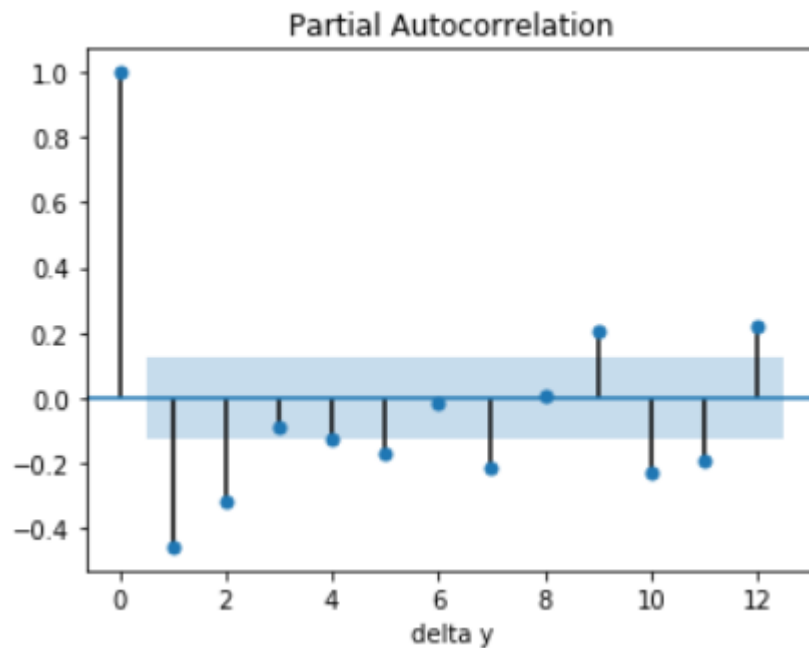
$n = 240$                       estimation sample is Jan 1980 - Dec 1999  
(monthly data, 20 years)

Significance of the first order sample autocorrelation coefficient. One way to select the orders of an AR model is to test whether the (partial) correlations differ significantly from zero. The null hypothesis of no autocorrelation ( $\rho_1 = 0$ ) can be tested against the alternative  $\rho_1 \neq 0$ . At (approximate) 5 per cent level, the null hypothesis is rejected if  $|r_1| > \frac{2}{\sqrt{n}}$ . In this case the first order autocorrelation is significant. (Textbook p.564)

$\frac{2}{\sqrt{n}} = 0.1291$                       Significance Threshold estimation sample

These correlations (denoted by  $r_k$ ) can be estimated from the sample by the sample partial autocorrelation function (SPACF). To select the orders, we can plot the correlations  $r_k$  against the time lag  $k$ . (Textbook p.548)

The plot of  $r_k$  is called the correlogram.



Lag	$ r_k $	$ r_k  > \frac{2}{\sqrt{n}}$ Significance Threshold
1	0.45808902	> 0.1291
2	0.31796111	> 0.1291
3	0.08675872	
4	0.12326431	
5	0.17378258	> 0.1291
6	0.01803195	
7	0.21442809	> 0.1291
8	0.00786458	
9	0.20881609	> 0.1291
10	0.23168022	> 0.1291
11	0.19678128	> 0.1291
12	0.22153186	> 0.1291

### Answer (d)

Only the lagged terms at lags 1, 2, 5, 7 and 9 to 12 are significant.



- (e) Extend the model of part (d) by adding the Error Correction (EC) term  $(y_t - 0.45x_t)$ , that is, estimate the ECM 
$$\Delta y_t = \alpha + \gamma(y_{t-1} - 0.45x_{t-1}) + \sum_{j=1}^5 \beta_j \Delta y_{t-j} + \beta_6 \Delta y_{t-10} + \beta_7 \Delta y_{t-12} + \varepsilon_t$$
 (estimation sample is Jan 1980- Dec 1999). Check that the EC term is significant at the 5% level, but not at the 1% level.

### Calculation (e)

#### Error correction model

The ADL model can be rewritten in terms of changes of the variables—that is, in terms of the first differences

$$\Delta y_t = y_t - y_{t-1} \text{ and } \Delta x_t = x_t - x_{t-1}$$

We consider this reformulation first for the ADL(1,1) model—that is, the model (7.32) with  $p = 1$ ,  $r = 1$ , and  $q = 0$ . By subtracting  $y_{t-1}$  from both sides of the equation (7.32), we can write this model as

$$\Delta y_t = \beta_0 \Delta x_t - (1 - \phi)(y_{t-1} - \lambda x_{t-1} - \delta) + \varepsilon_t, \quad (7.33)$$

with  $\delta = \alpha / (1 - \phi)$  and where  $\lambda = (\beta_0 + \beta_1) / (1 - \phi)$  is the long-run multiplier. (Textbook p.639)

## Error correction model

```

Results: Ordinary least squares
=====
Model:                OLS                Adj. R-squared:    0.436
Dependent Variable:   delta_y            AIC:              4977.7723
Date:                2023-01-05 12:04    BIC:              5008.5968
No. Observations:    227                Log-Likelihood:   -2479.9
Df Model:             8                  F-statistic:      22.87
Df Residuals:         218                Prob (F-statistic): 3.18e-25
R-squared:            0.456              Scale:           1.8797e+08
=====
              Coef.    Std.Err.    t    P>|t|    [0.025    0.975]
-----
const        4728.0072  2133.7034   2.2159  0.0277  522.6791 8933.3353
delta_lag1   -0.5223    0.0706  -7.4009  0.0000  -0.6614 -0.3832
delta_lag2   -0.1866    0.0833  -2.2403  0.0261  -0.3508 -0.0224
delta_lag3   -0.1581    0.0809  -1.9552  0.0518  -0.3175  0.0013
delta_lag4   -0.1847    0.0743  -2.4860  0.0137  -0.3311 -0.0383
delta_lag5   -0.1331    0.0611  -2.1785  0.0304  -0.2535 -0.0127
delta_lag10  -0.2737    0.0520  -5.2649  0.0000  -0.3762 -0.1712
delta_lag12   0.2516    0.0542   4.6402  0.0000   0.1448  0.3585
EC_term      -0.1503    0.0696  -2.1599  0.0319  -0.2875 -0.0132
=====
Omnibus:            0.012                Durbin-Watson:      2.043
Prob(Omnibus):      0.994                Jarque-Bera (JB):   0.095
Skew:               -0.004                Prob(JB):           0.954
Kurtosis:           2.900                Condition No.:      76171
=====
* The condition number is large (8e+04). This might indicate
strong multicollinearity or other numerical problems.

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coefficient EC\_term = -0.1503

std error: 0.0696

t-value: -2.1599

p-value: **0.0319**

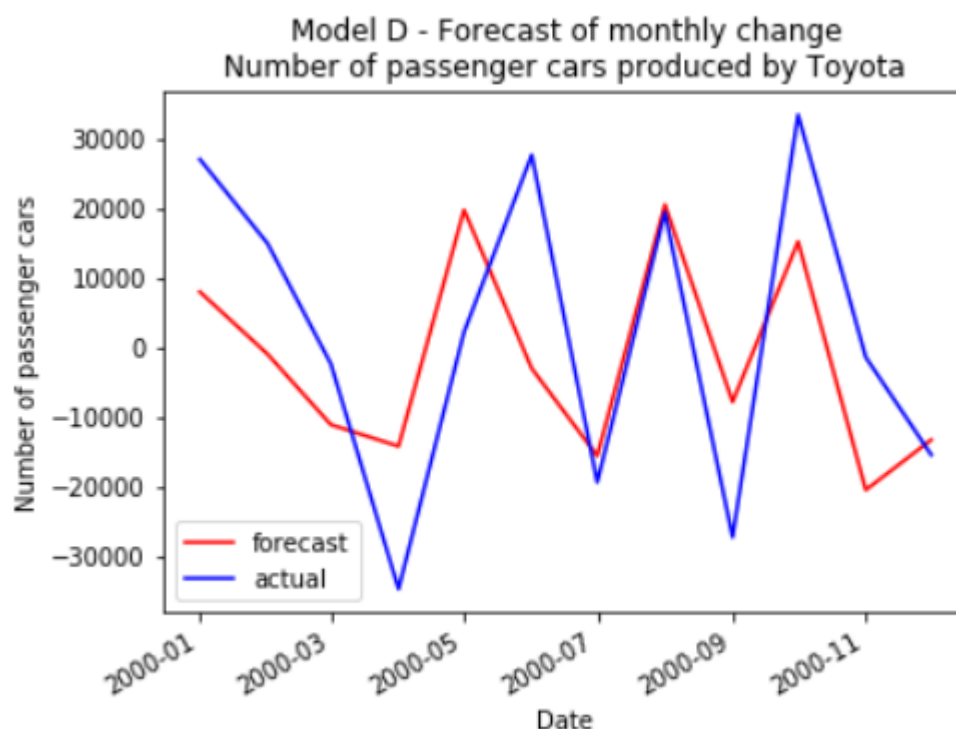
### Answer (e)

The EC term is significant at the 5% level, but not at the 1% level. The p-value of **0.0319** is below 0.05 (95%-Confidence Level), but above 0.01 (99%-Confidence Level).

- (f) Use the models of parts (d) and (e) to make two series of 12 forecasts of monthly changes in production of Toyota passenger cars in 2000. At each month, you should use the data that are then available, for example, to forecast production for September 2000 you can use the data up to and including August 2000. However, do not re-estimate the model and use the coefficients as obtained in parts (d) and (e). For each of the two forecast series, compute the values of the root mean squared error (RMSE) and of the mean absolute error (MAE). Finally, give your interpretation of the outcomes.

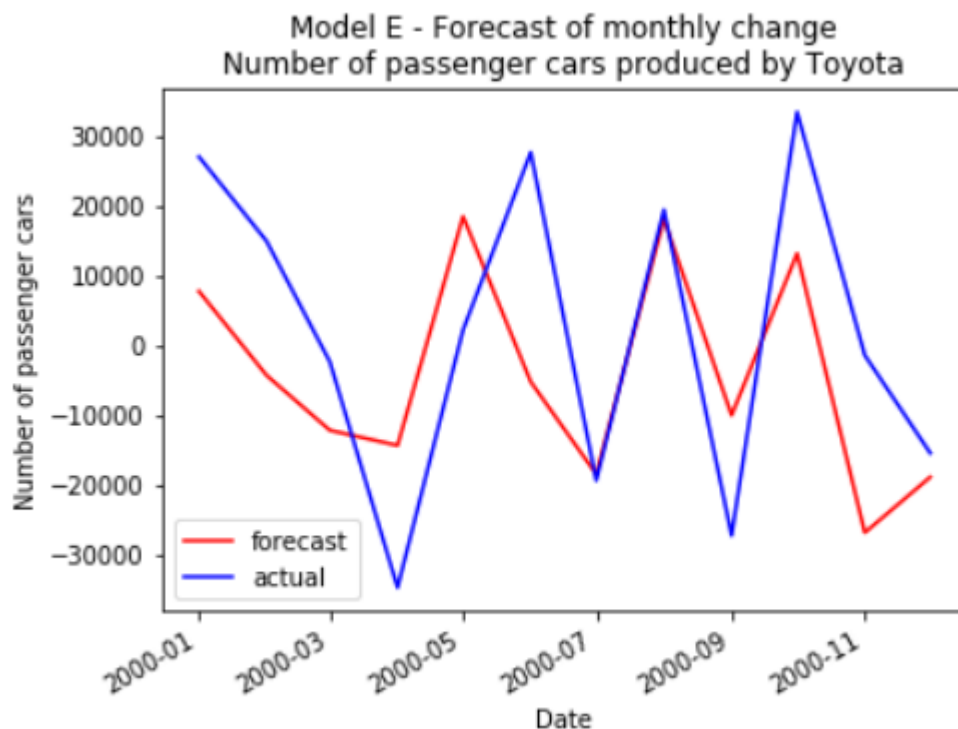
### Calculation (f)

Part (d) restricted AR(12) model for  $\Delta y_t$  with lag 1,2,3,4,5,10,12



RMSE = 16991.804      root mean squared error  
MAE = 176438.711      mean absolute error

Part (e) Error correction model for  $\Delta y_t$  with lag 1,2,3,4,5,10,12



RMSE = 18204.759  
MAE = 18668.0201

root mean squared error  
mean absolute error

### Answer (f)

The restricted AR(12) model (Model D) produces forecasts with a better fit than the Error correction model (Model E).