My Notes on Ula&Smith (2009) Paper

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### Abstract

Consumers or firms contemplating purchasing a new product or adopting a new technology are often plagued by uncertainty: Will the benefits outweigh the costs? Should we buy now or wait and gather more information? In this paper, we study a dynamic programming model of this technology adoption problem. In each period, the consumer decides whether to adopt the technology, reject it, or wait and gather additional information by observing a signal about the technology's benefit. The technology's actual benefit may be constant or changing stochastically over time. The dynamic programming state variable is a probability distribution that describes the consumer's beliefs about the benefits of the technology. We allow general probability distributions on benefits and general signal processes and assume that the consumer updates her beliefs over time using Bayes' rule. We are interested in structural properties of this model. We show that improving the technology's benefit need not make the consumer better off and that first-order stochastic dominance improvements in the consumer's distribution on benefits need not increase the consumer's value function. Nevertheless, the model possesses a great deal of structure. For example, we obtain monotonic value functions and policies if we order distributions using likelihoodratio dominance rather than first-order stochastic dominance. We also examine convexity properties and provide many comparative statics results.

# Model Implementation ( - Part2.1 )

#### 2.1 Importing the required library

library(tidyverse)

### 2.2 Model parameters definition

 $\theta$ :

the Net Present Value of Stream of benefits

- If she decides to reject the technology, she receives nothing and no longer gathers information about the technology.
- If she decides to *adopt* the technology, she pays a fixed adoption cost K and receives a net expected benefit:

$$\int_{\theta}\theta\pi(\theta)d\theta-K$$

 $\theta$  has the density  $\pi$ .

• Now, if consumer choose to gather additional information, she pays c in that period and observe the signal x, drawn with likelihood function

$$L(x|\theta)$$

Having observed signal x, the consumer then updates her prior  $\pi$  using Baye's rule:

$$\prod(\theta; \pi, x) = \frac{L(x|\theta)\pi(\theta)}{f(x; \pi)}$$

where the  $f(x;\pi)$  is the predictive ditribution for signals x,

$$f(x;\pi) = \int_{\theta} L(x|\theta)\pi(\theta)d\theta$$

The consumer then continues into the next period, starting with a new prior distribution that is equal to her posterior distribution from this stage.

Because our dynamic programming state variable is distribution itself , we will frequently supress the domain of the distribution and write the posterior as

The consumer's optimal value function with k periods remaining

$$\begin{split} v_0^*(\pi) &= 0 \\ v_k^*(\pi) &= \max(0, \int_{\theta} \theta \pi(\theta) \theta - K, -c + \delta E[v_{k-1}^*(\Pi(\pi, x))]) \end{split}$$

where  $\delta$  (0< $\delta$ <1) is the discount factor and the expectation of the next period value function is taken over all possible random signals.

$$E[v_{k-1}^*(\Pi(\pi,\tilde{x})] = \int_x v_{k-1} * (\Pi(\pi,x)) f(x;\pi) dx$$

#### 2.3 Part 2.1-Example of

The model for  $\theta$  is defined as:

$$\theta = Ap^*$$

Consumer's uncertainty about p\* has a beta distribution

$$f(p^*) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} (p^*)^{\alpha-1} (1-p^*)^{\beta-1}$$

### Discussion on Beta Distribiution:

The pdf of the beta distribution, for 0 < p\*<1, and shape parameters  $\alpha, \beta>0$ , is a power function of the variable x and of its reflection (1-x) as follows:

$$\begin{split} f(x|\alpha,\beta) &= constant.x^{\alpha-1}(1-x)^{\beta-1} \\ &= \frac{x^{\alpha-1}(1-x)^{\beta-1}}{\int_0^1 u^{\alpha-1}(1-u)^{\beta-1}du} \\ &= \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)}x^{\alpha-1}(1-x)^{\beta-1} \\ &= \frac{1}{B(\alpha,\beta)}x^{\alpha-1}(1-x)^{\beta-1} \end{split}$$

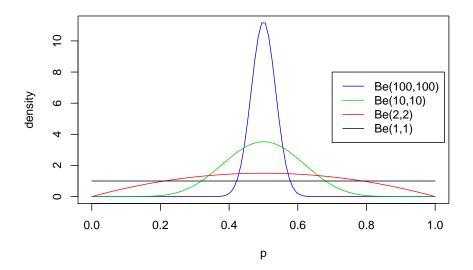
The beta function, B is a normalization constant to ensure that the total probability is 1.

#### 2.3.1 Example of Beta Distribution:

Example:

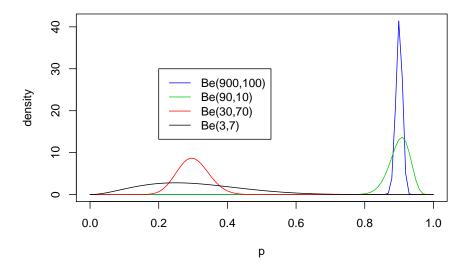
#### 2.3.1.1 equal alpha and beta:

```
p = seq(0,1,length.out = 100)
plot(p,dbeta(p,100,100),ylab = "density", type="l", col=4)
lines(p,dbeta(p,10,10), type = "l", col=3)
lines(p,dbeta(p,2,2), type = "l", col=2)
lines(p,dbeta(p,1,1), type = "l", col=1)
legend(0.7,8, c("Be(100,100)","Be(10,10)","Be(2,2)", "Be(1,1)"),col=c(4,3,2,1), lty = c(1,1,1,1)
```



### non-equal alpha and beta:

```
p = seq(0,1, length=100)
plot(p, dbeta(p, 900, 100), ylab="density", type ="l", col=4)
lines(p, dbeta(p, 90, 10), type ="l", col=3)
lines(p, dbeta(p, 30, 70), col=2)
lines(p, dbeta(p, 3, 7), col=1)
legend(0.2,30, c("Be(900,100)", "Be(90,10)", "Be(30,70)", "Be(3,7)"),lty=c(1,1,1,1),col=
```



Two points about the *Beta* distribution:

From these examples you should note the following:

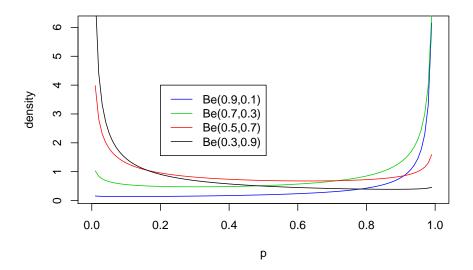
$$E[p] = \frac{\alpha}{\alpha + \beta}$$

- \* It turns out that the mean is exactly Thus the mean of the distribution is determined by the relative values of a and b.
  - The larger the values of a and b, the smaller the variance of the distribution about the mean.

$$Var[p] = E[(X - \mu)^2] = \frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)}$$

#### 2.3.1.2 alpha and beta <1

```
p = seq(0,1, length=100)
plot(p, dbeta(p, 0.9, 0.1), ylab="density", type ="l", col=4)
lines(p, dbeta(p, 0.7, 0.3), type ="l", col=3)
lines(p, dbeta(p, 0.5, 0.7), col=2)
lines(p, dbeta(p, 0.3, 0.9), col=1)
legend(0.2,4, c("Be(0.9,0.1)", "Be(0.7,0.3)", "Be(0.5,0.7)", "Be(0.3,0.9)"), col=c(4,3,2,1), lty=cell
```



#### 2.4 Beta Conjugate to Bernoulli Distibutaion

#### 2.4.1 Prior:

$$\begin{split} Beta(\alpha,\beta) &= \frac{\theta^{\alpha-1}(1-\theta)^{\beta-1}}{B(\alpha,\beta)} \\ &= const.\theta^{\alpha-1}(1-\theta)^{b-1} \end{split}$$

#### 2.4.2 Posterior Distribution:

$$P(\theta|data) = \frac{P(data|\theta)P(\theta)}{P(data)} \propto P(data|\theta)P(theta)$$

#### 2.5 Beta-Binomial Model

#### 2.5.1 Likelihood:

$$P(data|\theta) \propto \theta^z (1-\theta)^{n-z}$$
,

if we let z number of  $x_i$  with value 1, i.e, (seeing z success in n trial)

$$z = \sum_{i=1}^{n} X_i$$

#### 2.5.2 Derive Posterior

$$P(\theta|data) \propto (\theta^z(1-\theta)^{n-z})(\theta^{\alpha-1}(1-\theta)^{\beta-1})$$

$$\propto \theta^{z+\alpha-1} (1-\alpha)^{n-z+\beta-1}$$

Subsituing the following:

$$\alpha' = \alpha + z$$
$$\beta' = n + \beta - z$$

Now we have posterior:

$$\begin{split} P(\theta|data) &= \frac{\theta^{\alpha'-1}(1-\alpha)^{\beta'-1}}{B(\alpha',\beta')} \\ &= Beta(\alpha',\beta') \end{split}$$

#### 2.6 Example of the:

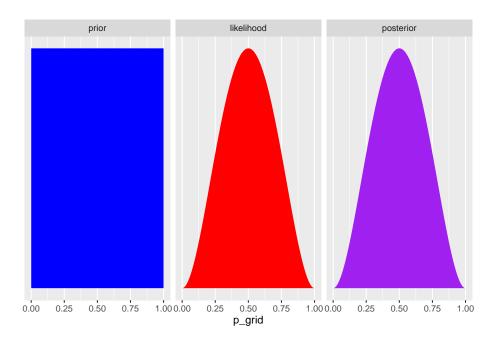
#### 2.6.1 Prior

Beta with  $\alpha=1$ ,  $\beta=1$ :

```
N <- 1000
alpha <- 1
beta <- 1
nsuccess <- 2
n_trials <- 4

data_1_1 <- tibble(p_grid = seq(from=0, to =1, length.out = N)) %>%
  mutate(prior=dbeta(p_grid,alpha,beta)) %>%
  mutate(likelihood = dbinom(nsuccess, size = n_trials, prob = p_grid)) %>%
  mutate(posterior = (likelihood * prior)) / sum(likelihood * prior))
```

```
data_1_1 %>%
  gather(key, value, -p_grid) %>%
  # this line allows us to dictate the order the panels will appear in
  mutate(key = factor(key, levels = c("prior", "likelihood", "posterior"))) %>%
  ggplot(aes(x = p_grid, ymin = 0, ymax = value, fill = key)) +
  geom_ribbon() +
  scale_fill_manual(values = c("blue", "red", "purple")) +
  scale_y_continuous(NULL, breaks = NULL) +
  theme(legend.position = "none") +
  facet_wrap(~key, scales = "free")
```



Analytically, the mean of the posterior should be:

$$Beta(\alpha + s, \beta + n - s)$$

with mean value of

$$E[X] = \frac{\alpha + s}{\alpha + \beta + n}$$

having  $\alpha$ =1,  $\beta$  = 1, s=2 and n=4 :

$$Beta(1+2,1+4-2)$$
 
$$Beta(3,3)$$

with mean value of

$$E[pgrid] = \frac{3}{3+6} = \frac{1}{3}$$

#### 2.7 Normal-Normal Conjugate:

$$\begin{split} P(\theta|Data) &= \frac{P(Data|\theta)P(\theta)}{P(Data)} \propto P(Data|\theta)P(\theta) \\ P(Data|\theta) &= \prod_{i=1}^{N} \frac{1}{\sqrt{2\pi\sigma_{x}^{2}}}exp(-\frac{(x_{i}-\theta)^{2}}{2\sigma_{x}^{2}}) \\ &\propto exp(-\frac{\sum_{i=1}^{N}(x_{i}-\theta)^{2}}{2\sigma_{x}^{2}}) \end{split}$$

Now posterior:

$$\begin{split} P(\theta|X) \propto exp(-\frac{\sum_{i=1}^{N}(x_i-\theta)^2}{2\sigma_x^2})exp(-\frac{(\theta-\theta_0)^2}{2\sigma_\theta^2})\\ = exp[- \end{split}$$

You can label chapter and section titles using {#label} after them, e.g., we can reference Chapter ??. If you do not manually label them, there will be automatic labels anyway, e.g., Chapter 4.

Figures and tables with captions will be placed in figure and table environments, respectively.

```
par(mar = c(4, 4, .1, .1))
plot(pressure, type = 'b', pch = 19)
```

Reference a figure by its code chunk label with the fig: prefix, e.g., see Figure 2.1. Similarly, you can reference tables generated from knitr::kable(), e.g., see Table 2.1.

```
knitr::kable(
  head(iris, 20), caption = 'Here is a nice table!',
  booktabs = TRUE
)
```

You can write citations, too. For example, we are using the **bookdown** package (Xie, 2020) in this sample book, which was built on top of R Markdown and **knitr** (?). change new line



Figure 2.1: Here is a nice figure!

Table 2.1: Here is a nice table!

Sepal.Length	Sepal.Width	Petal.Length	Petal.Width	Species
5.1	3.5	1.4	0.2	setosa
4.9	3.0	1.4	0.2	setosa
4.7	3.2	1.3	0.2	setosa
4.6	3.1	1.5	0.2	setosa
5.0	3.6	1.4	0.2	setosa
5.4	3.9	1.7	0.4	setosa
4.6	3.4	1.4	0.3	setosa
5.0	3.4	1.5	0.2	setosa
4.4	2.9	1.4	0.2	setosa
4.9	3.1	1.5	0.1	setosa
5.4	3.7	1.5	0.2	setosa
4.8	3.4	1.6	0.2	setosa
4.8	3.0	1.4	0.1	setosa
4.3	3.0	1.1	0.1	setosa
5.8	4.0	1.2	0.2	setosa
5.7	4.4	1.5	0.4	setosa
5.4	3.9	1.3	0.4	setosa
5.1	3.5	1.4	0.3	setosa
5.7	3.8	1.7	0.3	setosa
5.1	3.8	1.5	0.3	setosa

### Literature

3.1 To Do

# Methods

4.1 To DO

# **Applications**

5.1 To DO!

### Final Words

6.1 To Do!

# Bibliography

Xie, Y. (2020). bookdown: Authoring Books and Technical Documents with R Markdown. R package version 0.18.