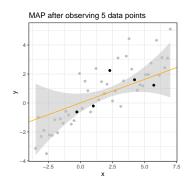
Introduction to Machine Learning

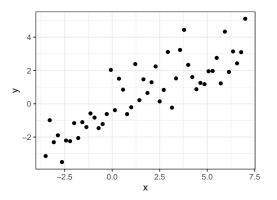
The Bayesian Linear Model



Learning goals

- Know the Bayesian linear model
- The Bayesian LM returns a (posterior) distribution instead of a point estimate
- Know how to derive the posterior distribution for a Bayesian LM

Let $\mathcal{D} = \{(\mathbf{x}^{(1)}, y^{(1)}), ..., (\mathbf{x}^{(n)}, y^{(n)})\}$ be a training set of i.i.d. observations from some unknown distribution.



Let $\mathbf{y} = (y^{(1)}, ..., y^{(n)})^{\top}$ and $\mathbf{X} \in \mathbb{R}^{n \times p}$ be the design matrix where the i-th row contains vector $\mathbf{x}^{(i)}$.

The linear regression model is defined as

$$y = f(\mathbf{x}) + \epsilon = \boldsymbol{\theta}^T \mathbf{x} + \epsilon$$

or on the data:

$$y^{(i)} = f(\mathbf{x}^{(i)}) + \epsilon^{(i)} = \boldsymbol{\theta}^T \mathbf{x}^{(i)} + \epsilon^{(i)}, \text{ for } i \in \{1, \dots, n\}$$

We now assume (from a Bayesian perspective) that also our parameter vector $\boldsymbol{\theta}$ is stochastic and follows a distribution. The observed values $y^{(i)}$ differ from the function values $f\left(\mathbf{x}^{(i)}\right)$ by some additive noise, which is assumed to be i.i.d. Gaussian

$$\epsilon^{(i)} \sim \mathcal{N}(0, \sigma^2)$$

and independent of \mathbf{x} and $\boldsymbol{\theta}$.

Let us assume we have **prior beliefs** about the parameter θ that are represented in a prior distribution $\theta \sim \mathcal{N}(\mathbf{0}, \tau^2 \mathbf{I}_p)$.

Whenever data points are observed, we update the parameters' prior distribution according to Bayes' rule

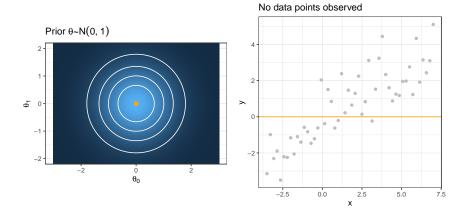
$$\underbrace{\rho(\boldsymbol{\theta}|\mathbf{X},\mathbf{y})}_{\text{posterior}} = \underbrace{\frac{\rho(\mathbf{y}|\mathbf{X},\boldsymbol{\theta})}{\rho(\mathbf{y}|\mathbf{X},\boldsymbol{\theta})} \underbrace{q(\boldsymbol{\theta})}_{\text{marginal}}^{\text{prior}}.$$

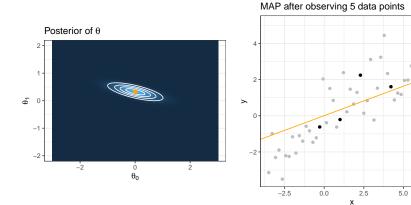
The posterior distribution of the parameter θ is again normal distributed (the Gaussian family is self-conjugate):

$$oldsymbol{ heta} \mid \mathbf{X}, \mathbf{y} \sim \mathcal{N}(\sigma^{-2} \mathbf{A}^{-1} \mathbf{X}^{ op} \mathbf{y}, \mathbf{A}^{-1})$$

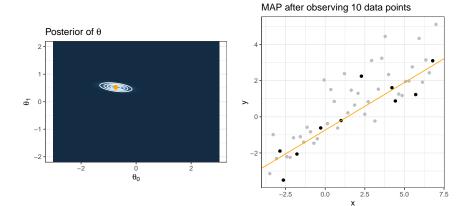
with
$$\mathbf{A} := \sigma^{-2} \mathbf{X}^{\top} \mathbf{X} + \frac{1}{\tau^2} \mathbf{I}_{p}$$
.

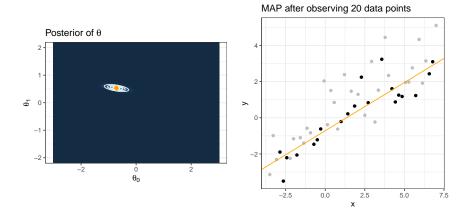
Note: If the posterior distribution $p(\theta \mid \mathbf{X}, \mathbf{y})$ are in the same probability distribution family as the prior $q(\theta)$ w.r.t. a specific likelihood function $p(\mathbf{y} \mid \mathbf{X}, \theta)$, they are called **conjugate distributions**. The prior is then called a **conjugate prior** for the likelihood. The Gaussian family is self-conjugate: Choosing a Gaussian prior for a Gaussian Likelihood ensures that the posterior is Gaussian.





7.5





Proof:

We want to show that

- for a Gaussian prior on $\theta \sim \mathcal{N}(\mathbf{0}, \tau^2 I_p)$
- for a Gaussian Likelihood $y \mid \mathbf{X}, \boldsymbol{\theta} \sim \mathcal{N}(\mathbf{X}^{\top} \boldsymbol{\theta}, \sigma^2 \mathbf{I}_n)$

the resulting posterior is Gaussian $\mathcal{N}(\sigma^{-2}\mathbf{A}^{-1}\mathbf{X}^{\top}\mathbf{y},\mathbf{A}^{-1})$ with $\mathbf{A}:=\sigma^{-2}\mathbf{X}^{\top}\mathbf{X}+\frac{1}{\tau^{2}}\mathbf{I}_{p}$. Plugging in Bayes' rule and multiplying out yields

$$\begin{split} \rho(\boldsymbol{\theta}|\mathbf{X},\mathbf{y}) & \propto & \rho(\mathbf{y}|\mathbf{X},\boldsymbol{\theta})q(\boldsymbol{\theta}) \propto \exp\left[-\frac{1}{2\sigma^2}(\mathbf{y}-\mathbf{X}\boldsymbol{\theta})^\top(\mathbf{y}-\mathbf{X}\boldsymbol{\theta}) - \frac{1}{2\tau^2}\boldsymbol{\theta}^\top\boldsymbol{\theta}\right] \\ & = & \exp\left[-\frac{1}{2}\left(\underbrace{\sigma^{-2}\mathbf{y}^\top\mathbf{y}}_{\text{doesn't depend on }\boldsymbol{\theta}} - 2\sigma^{-2}\mathbf{y}^\top\mathbf{X}\boldsymbol{\theta} + \sigma^{-2}\boldsymbol{\theta}^\top\mathbf{X}^\top\mathbf{X}\boldsymbol{\theta} + \tau^{-2}\boldsymbol{\theta}^\top\boldsymbol{\theta}\right)\right] \\ & \propto & \exp\left[-\frac{1}{2}\left(\sigma^{-2}\boldsymbol{\theta}^\top\mathbf{X}^\top\mathbf{X}\boldsymbol{\theta} + \tau^{-2}\boldsymbol{\theta}^\top\boldsymbol{\theta} - 2\sigma^{-2}\mathbf{y}^\top\mathbf{X}\boldsymbol{\theta}\right)\right] \\ & = & \exp\left[-\frac{1}{2}\boldsymbol{\theta}^\top\underbrace{\left(\sigma^{-2}\mathbf{X}^\top\mathbf{X} + \tau^{-2}\mathbf{I}_{\boldsymbol{\rho}}\right)}_{:=\mathbf{A}}\boldsymbol{\theta} + \sigma^{-2}\mathbf{y}^\top\mathbf{X}\boldsymbol{\theta}\right] \end{split}$$

This expression resembles a normal density - except for the term in red!

Note: We need not worry about the normalizing constant since its mere role is to convert probability functions to density functions with a total probability of one. We subtract a (not yet defined) constant c while compensating for this change by adding the respective terms ("adding 0"), emphasized in green:

$$\begin{split} \rho(\theta|\mathbf{X},\mathbf{y}) &\propto & \exp\left[-\frac{1}{2}(\theta-c)^{\top}\mathbf{A}(\theta-c) - c^{\top}\mathbf{A}\theta + \underbrace{\frac{1}{2}c^{\top}\mathbf{A}c}_{\text{doesn't depend on }\theta} + \sigma^{-2}\mathbf{y}^{\top}\mathbf{X}\theta\right] \\ &\propto & \exp\left[-\frac{1}{2}(\theta-c)^{\top}\mathbf{A}(\theta-c) - c^{\top}\mathbf{A}\theta + \sigma^{-2}\mathbf{y}^{\top}\mathbf{X}\theta\right] \end{split}$$

If we choose c such that $-c^{\top}\mathbf{A}\theta + \sigma^{-2}\mathbf{y}^{\top}\mathbf{X}\theta = 0$, the posterior is normal with mean c and covariance matrix \mathbf{A}^{-1} . Taking into account that \mathbf{A} is symmetric, this is if we choose

$$\sigma^{-2}\mathbf{y}^{\top}\mathbf{X} = c^{\top}\mathbf{A}$$

$$\Leftrightarrow \quad \sigma^{-2}\mathbf{y}^{\top}\mathbf{X}\mathbf{A}^{-1} = c^{\top}$$

$$\Leftrightarrow \quad c = \sigma^{-2}\mathbf{A}^{-1}\mathbf{X}^{\top}\mathbf{y}$$

as claimed.

Based on the posterior distribution

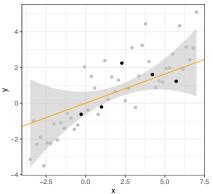
$$oldsymbol{ heta} \mid \mathbf{X}, \mathbf{y} \sim \mathcal{N}(\sigma^{-2} \mathbf{A}^{-1} \mathbf{X}^{ op} \mathbf{y}, \mathbf{A}^{-1})$$

we can derive the predictive distribution for a new observations \mathbf{x}_* . The predictive distribution for the Bayesian linear model, i.e. the distribution of $\boldsymbol{\theta}^{\top}\mathbf{x}_*$, is

$$\textit{y}_* \mid \mathbf{X}, \mathbf{y}, \mathbf{x}_* \sim \mathcal{N}(\sigma^{-2}\mathbf{y}^{\top}\mathbf{X}\mathbf{A}^{-1}\mathbf{x}_*, \mathbf{x}_*^{\top}\mathbf{A}^{-1}\mathbf{x}_*)$$

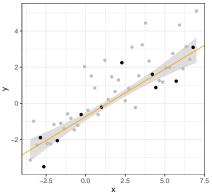
(applying the rules for linear transformations of Gaussians).





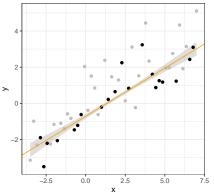
For every test input \mathbf{x}_* , we get a distribution over the prediction y_* . In particular, we get a posterior mean (orange) and a posterior variance (grey region equals +/- two times standard deviation).





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SUMMARY: THE BAYESIAN LINEAR MODEL

- By switching to a Bayesian perspective, we do not only have point estimates for the parameter θ , but whole **distributions**
- From the posterior distribution of θ , we can derive a predictive distribution for $y_* = \theta^\top \mathbf{x}_*$.
- ullet We can perform online updates: Whenever datapoints are observed, we can update the **posterior distribution** of eta

Next, we want to develop a theory for general shape functions, and not only for linear function.