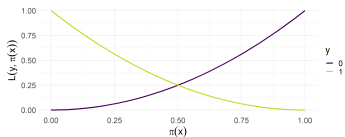# Introduction to Machine Learning

Brier Score

Learning goals

- Know the Brier score
- Derive the risk minimizer
- Derive the optimal constant model
- Understand the connection between Brier score and Gini splitting

BRIER SCORE

The binary Brier score is defined on probabilities $\pi(\mathbf{x}) \in [0, 1]$ and 0-1-encoded labels $y \in \{0, 1\}$ and measures their squared distance (*L2* loss on probabilities).

$$L\left(y, \pi(\mathbf{x})\right) = \left(\pi(\mathbf{x}) - y\right)^2$$
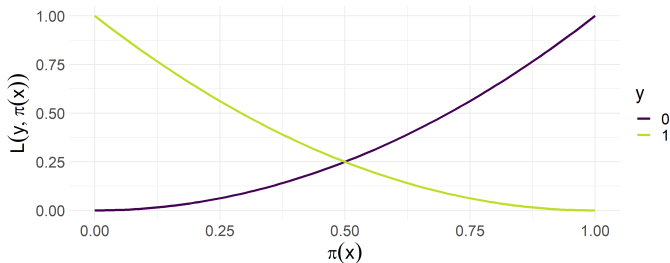
BRIER SCORE: RISK MINIMIZER

The risk minimizer for the (binary) Brier score is

$$\pi^*(\mathbf{x}) = \eta(\mathbf{x}) = \mathbb{P}(y \mid \mathbf{x} = \mathbf{x}),$$

which means that the Brier score will reach its minimum if the prediction equals the "true" probability of the outcome.

The risk minimizer for the multiclass Brier score is

$$\pi^*(\mathbf{x}) = \mathbb{P}(y = k \mid \mathbf{x} = \mathbf{x}).$$

Proof: We only show the proof for the binary case. We need to minimize

$$\mathbb{E}_x \left[L(1, \pi(\mathbf{x})) \cdot \eta(\mathbf{x}) + L(0, \pi(\mathbf{x})) \cdot (1 - \eta(\mathbf{x})) \right],$$

BRIER SCORE: RISK MINIMIZER

which we do point-wise for every **x**. We plug in the Brier score

$$\arg\min_c L(1, c)\eta(\mathbf{x}) + L(0, c)(1 - \eta(\mathbf{x}))$$
$$= \arg\min_c \ (c - 1)^2 \eta(\mathbf{x}) + c^2(1 - \eta(\mathbf{x}))$$
$$= \arg\min_c \ (c - \eta(\mathbf{x}))^2.$$

The expression is minimal if $c = \eta(\mathbf{x}) = \mathbb{P}(y = 1 \mid \mathbf{x} = \mathbf{x})$.

BRIER SCORE: OPTIMAL CONSTANT MODEL

The optimal constant probability model $\pi(\mathbf{x}) = \theta$ w.r.t. the Brier score for labels from $\mathcal{Y} = \{0, 1\}$ is:

$$
\begin{aligned}
\min_{\theta} \mathcal{R}_{\mathsf{emp}}(\theta) &= \min_{\theta} \sum_{i=1}^{n} \left(y^{(i)} - \theta \right)^2 \\
\Leftrightarrow \frac{\partial \mathcal{R}_{\mathsf{emp}}(\theta)}{\partial \theta} &= -2 \cdot \sum_{i=1}^{n} (y^{(i)} - \theta) = 0 \\
\hat{\theta} &= \frac{1}{n} \sum_{i=1}^{n} y^{(i)}.
\end{aligned}
$$

This is the fraction of class-1 observations in the observed data.
(This also directly follows from our *L2* proof for regression).

Similarly, for the multiclass brier score the optimal constant is
$\hat{\theta}_k = \frac{1}{n} \sum\limits_{i=1}^{n} [y = k]$.

BRIER SCORE MINIMIZATION = GINI SPLITTING

When fitting a tree we minimize the risk within each node $\mathcal{N}$ by risk minimization and predict the optimal constant. Another approach that is common in literature is to minimize the average node impurity $\text{Imp}(\mathcal{N})$.

Claim: Gini splitting $\text{Imp}(\mathcal{N}) = \sum_{k=1}^{g} \pi_k^{(\mathcal{N})} \left(1 - \pi_k^{(\mathcal{N})} \right)$ is equivalent to the Brier score minimization.

Note that $\pi_k^{(\mathcal{N})} := \frac{1}{n_{\mathcal{N}}} \sum_{(\mathbf{x},y) \in \mathcal{N}} [y = k]$

Proof: We show that the risk related to a subset of observations $\mathcal{N} \subseteq \mathcal{D}$ fulfills

$$\mathcal{R}(\mathcal{N}) = n_{\mathcal{N}} \text{Imp}(\mathcal{N}),$$

where Imp is the Gini impurity and $\mathcal{R}(\mathcal{N})$ is calculated w.r.t. the (multiclass) Brier score

$$L(y, \pi(\mathbf{x})) = \sum_{k=1}^{g} \left([y = k] - \pi_k(\mathbf{x}) \right)^2 .$$

BRIER SCORE MINIMIZATION = GINI SPLITTING

$$
\mathcal{R}(\mathcal{N}) = \sum_{(\mathbf{x},y)\in\mathcal{N}}\sum_{k=1}^{g}([y=k]-\pi_k(\mathbf{x}))^2 = \sum_{k=1}^{g}\sum_{(\mathbf{x},y)\in\mathcal{N}}\left([y=k]-\frac{n_{\mathcal{N},k}}{n_{\mathcal{N}}}\right)^2,
$$

by plugging in the optimal constant prediction w.r.t. the Brier score ($n_{\mathcal{N},k}$ is defined as the number of class k observations in node $\mathcal{N}$):

$$
\hat{\pi}_k(\mathbf{x}) = \pi_k^{(\mathcal{N})} = \frac{1}{n_{\mathcal{N}}}\sum_{(\mathbf{x},y)\in\mathcal{N}}[y=k] = \frac{n_{\mathcal{N},k}}{n_{\mathcal{N}}}.
$$

We split the inner sum and further simplify the expression

$$
= \sum_{k=1}^{g}\left(\sum_{(\mathbf{x},y)\in\mathcal{N}:\ y=k}\left(1-\frac{n_{\mathcal{N},k}}{n_{\mathcal{N}}}\right)^2 + \sum_{(\mathbf{x},y)\in\mathcal{N}:\ y\neq k}\left(0-\frac{n_{\mathcal{N},k}}{n_{\mathcal{N}}}\right)^2\right)
$$

$$
= \sum_{k=1}^{g} n_{\mathcal{N},k}\left(1-\frac{n_{\mathcal{N},k}}{n_{\mathcal{N}}}\right)^2 + (n_{\mathcal{N}}-n_{\mathcal{N},k})\left(\frac{n_{\mathcal{N},k}}{n_{\mathcal{N}}}\right)^2,
$$

since for $n_{\mathcal{N},k}$ observations the condition $y=k$ is met, and for the remaining $(n_{\mathcal{N}}-n_{\mathcal{N},k})$ observations it is not.

BRIER SCORE MINIMIZATION = GINI SPLITTING

We further simplify the expression to

$$
\begin{aligned}
\mathcal{R}(\mathcal{N}) &= \sum_{k=1}^{g} n_{\mathcal{N},k} \left(\frac{n_{\mathcal{N}} - n_{\mathcal{N},k}}{n_{\mathcal{N}}} \right)^2 + (n_{\mathcal{N}} - n_{\mathcal{N},k}) \left(\frac{n_{\mathcal{N},k}}{n_{\mathcal{N}}} \right)^2 \\
&= \sum_{k=1}^{g} \frac{n_{\mathcal{N},k}}{n_{\mathcal{N}}} \frac{n_{\mathcal{N}} - n_{\mathcal{N},k}}{n_{\mathcal{N}}} \left(n_{\mathcal{N}} - n_{\mathcal{N},k} + n_{\mathcal{N},k} \right) \\
&= n_{\mathcal{N}} \sum_{k=1}^{g} \pi_k^{(\mathcal{N})} \cdot \left(1 - \pi_k^{(\mathcal{N})} \right) = n_{\mathcal{N}} \mathsf{Imp}(\mathcal{N}).
\end{aligned}
$$