

# A Flexible Approach for Predictive Biomarker Discovery

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## Background

- Predictive biomarkers are treatment effect modifiers.
- In high dimensions, these biomarkers are discovered using interpretable conditional average treatment effect estimators, like the modified covariates procedures of Tian et al. (2014).
- These methods make simplifying assumptions about the datagenerating process, resulting in a lack of Type-I error rate control.
- High false discovery rates lead to wasted resources, negatively affecting patient outcomes.

## Variable Importance Parameter

Consider n identically and independently distributed (i.i.d) full-data random vectors  $X = (W, A, Y^{(0)}, Y^{(1)}) \sim P_X$ .

- W: A p-length random vector of centered pretreatment biomarkers with nonzero variance.
- A: A random binary indicator of treatment assignment.
- $Y^{(0)}, Y^{(1)}$ : Continuous potential outcomes under assignment to the control and treatment allocations, respectively.

Our causal variable importance parameter is  $\Psi^F(P_X) = (\Psi_1^F(P_X), \dots, \Psi_p^F(P_X))$ , where

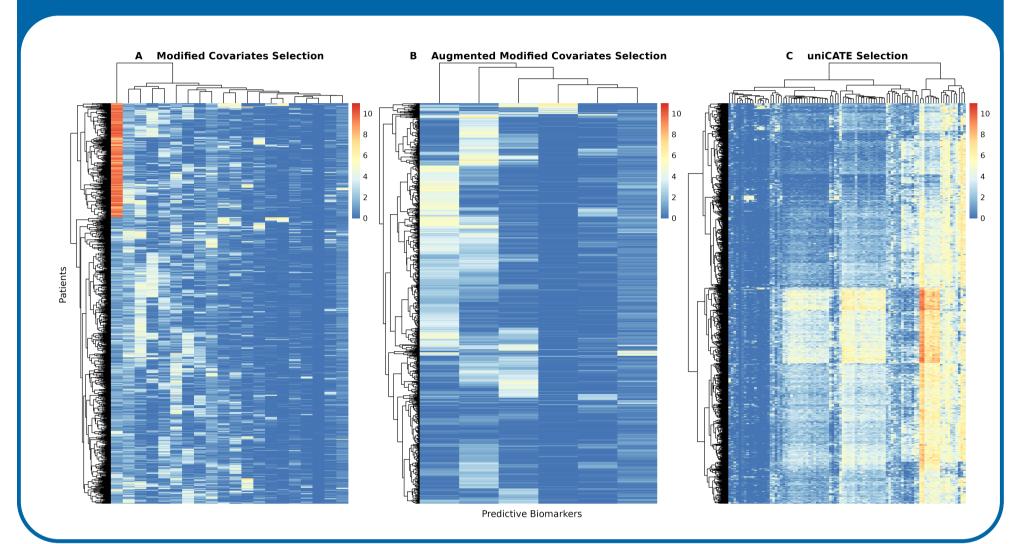
$$\Psi_j^F(P_X) \equiv \frac{\mathbb{E}_{P_X} \left[ \left( Y^{(1)} - Y^{(0)} \right) W_j \right]}{\mathbb{E}_{P_X} \left[ W_j^2 \right]}.$$

Given access instead to n i.i.d. censored random observations O = (W, A, Y) where  $Y = AY^{(1)} + (1 - A)Y^{(0)}$ ,  $\Psi^F(P_X)$  is identified under the assumptions of unmeasured confounding and positivity by  $\Psi(P_0) = (\Psi_1(P_0), \dots, \Psi_p(P_0))$ . Here,

$$\Psi_{j}(P_{0}) \equiv \frac{\mathbb{E}_{P_{0}} \left[ \left( \bar{Q}_{0}[A=1,W] - \bar{Q}_{0}[A=0,W] \right) W_{j} \right]}{\mathbb{E}_{P_{0}} \left[ W_{j}^{2} \right]},$$

where  $\bar{Q}_0[A, W] = \mathbb{E}_{P_0}[Y|A, W]$ .

## Application to IMmotion 150 and 151



## Inference

Let  $g_0(W) = \mathbb{P}[A = 1|W]$ , and let  $\hat{g}$  and  $\hat{Q}$  be estimators of  $g_0$  and  $\bar{Q}_0$ , respectively. Define the Augmented Inverse Probability Weighted outcome difference as

$$\begin{split} T(O) \equiv \left(\frac{I(A=1)}{\hat{g}(W)} - \frac{I(A=0)}{1-\hat{g}(W)}\right) \left(Y - \hat{\bar{Q}}(A,W)\right) \\ + \hat{\bar{Q}}(1,W) - \hat{\bar{Q}}(0,W). \end{split}$$

We derive from the efficient influence function of  $\Psi_j(P_0)$ ,  $D_j(O)$ , the double-robust one-step estimator

$$\hat{\Psi}_j(P_n) \equiv \frac{\sum_{i=1}^n T(O_i) W_{ij}}{\sum_{i=1}^n W_{ij}^2},$$

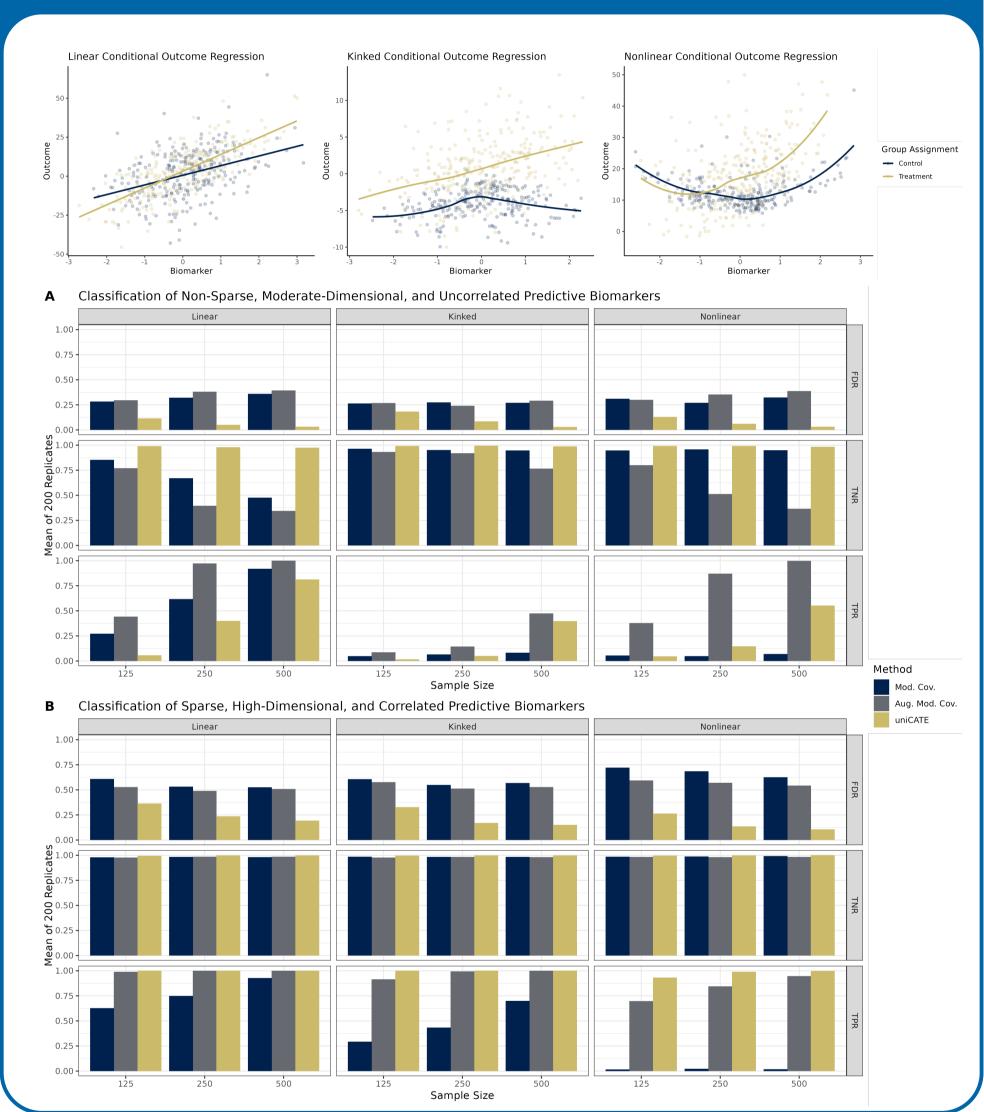
where  $\sum_{i} W_{ij} = 0$  for all j and  $P_n$  is the empirical distribution. If  $\hat{g}$  and  $\hat{\bar{Q}}$  are trained using sample splitting techniques, and we assume that  $\|\hat{g} - g_0\|_2 \|\hat{\bar{Q}} - \bar{Q}_0\|_2 = o_p(n^{-1/2})$ , then

$$\sqrt{n}\left(\hat{\Psi}_j(P_n) - \Psi_j(P_0)\right) \stackrel{D}{\to} N\left(0, \mathbb{V}_{P_0}\left[D_j(O)\right]\right).$$

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#### Randomized Control Trial Simulations



#### Conclusion

Predictive biomarker discovery benefits from formal statistical inference procedures that control false discovery rates. This estimator is implemented in the uniCATE R package available at github.com/insightsengineering/uniCATE.

#### Reference

Tian et al. (2014) A Simple Method for Estimating Interactions Between a Treatment and a Large Number of Covariates, Journal of the American Statistical Association, 109:508, 1517-1532, DOI: 10.1080/01621459.2014.951443