

Bayes 决策与概率密度估计习题

姓名：甘云冲

学号：2101213081

1.

(1)

$$P(X|\text{Positive}) = \frac{P(\text{Positive}|X)P(X)}{P(\text{Positive}|X)P(X) + P(\text{Positive}|\neg X)P(\neg X)} = \frac{100\% * 0.5\%}{100\% * 0.5\% + 5\% * 99.5\%} = 9.1\%$$

该人患有疾病X的概率为9.1%。

(2)

$$\begin{aligned} P(X|\text{Positive, Positive}) &= \frac{P(\text{Positive, Positive}|X)P(X)}{P(\text{Positive, Positive}|X)P(X) + P(\text{Positive, Positive}|\neg X)P(\neg X)} \\ &= \frac{P(\text{Positive}|X)^2 P(X)}{P(\text{Positive}|X)^2 P(X) + P(\text{Positive}|\neg X)^2 P(\neg X)} \\ &= \frac{(100\%)^2 * 0.5\%}{(100\%)^2 * 0.5\% + (5\%)^2 * 99.5\%} \\ &= 66.8\% \end{aligned}$$

该人患有疾病X的概率为66.8%。

$$66.8\% * 10 > (1 - 66.8\%) * 1$$

医生应当认为患者有病。

2.

(1)

$$\begin{aligned}
P(\omega_1|x) &= P(\omega_2|x) \\
\Leftrightarrow \frac{P(x|\omega_1)P(\omega_1)}{P(x|\omega_1)P(\omega_1) + P(x|\omega_2)P(\omega_2)} &= \frac{P(x|\omega_2)P(\omega_2)}{P(x|\omega_1)P(\omega_1) + P(x|\omega_2)P(\omega_2)} \\
\Leftrightarrow \frac{P(x|\omega_1)}{P(x|\omega_1) + P(x|\omega_2)} &= \frac{P(x|\omega_2)}{P(x|\omega_1) + P(x|\omega_2)} \\
\Leftrightarrow P(x|\omega_1) &= P(x|\omega_2)
\end{aligned}$$

令 $x = (a_1 + a_2)/2$, 可得到:

$$P(x|\omega_1) = \frac{1}{\pi b} \frac{1}{1 + (\frac{a_2 - a_1}{2b})^2} = P(x|\omega_2)$$

得证结论。

(2)

$$\begin{aligned}
\lim_{x \rightarrow +\infty} P(\omega_1|x) &= \lim_{x \rightarrow +\infty} \frac{P(x|\omega_1)P(\omega_1)}{P(x|\omega_1)P(\omega_1) + P(x|\omega_2)P(\omega_2)} \\
&= \lim_{x \rightarrow +\infty} \frac{(1 + (\frac{x-a_2}{b})^2)P(\omega_1)}{(1 + (\frac{x-a_2}{b})^2)P(\omega_1) + (1 + (\frac{x-a_1}{b})^2)P(\omega_2)} \\
&= \frac{P(\omega_1)}{P(\omega_1) + P(\omega_2)}
\end{aligned}$$

同理可得:

$$\begin{aligned}
\lim_{x \rightarrow -\infty} P(\omega_1|x) &= \frac{P(\omega_1)}{P(\omega_1) + P(\omega_2)} \\
\lim_{x \rightarrow +\infty} P(\omega_2|x) &= \frac{P(\omega_2)}{P(\omega_1) + P(\omega_2)} \\
\lim_{x \rightarrow -\infty} P(\omega_2|x) &= \frac{P(\omega_2)}{P(\omega_1) + P(\omega_2)}
\end{aligned}$$

(3)

不失一般性地假设 $a_1 \leq a_2$, 由(1)中可以知道, 最小误差判决边界为峰值的中点。

$$\begin{aligned}
P(e) &= \int_{-\infty}^{(a_1+a_2)/2} p(x|\omega_2)p(\omega_2)dx + \int_{(a_1+a_2)/2}^{+\infty} p(x|\omega_1)p(\omega_1)dx \\
&= \frac{1}{\pi b} \int_{(a_2-a_1)/(2b)}^{\infty} \frac{1}{1+x^2} dx \\
&= \frac{1}{\pi} \arctan(x) \Big|_{\frac{a_2-a_1}{2b}}^{\infty} \\
&= \frac{1}{\pi} (\pi/2 - \arctan(\frac{a_2-a_1}{2b})) \\
&= \frac{1}{2} - \frac{1}{\pi} \arctan(\frac{a_2-a_1}{2b})
\end{aligned}$$

针对于 $a_1 > a_2$ 的情况同上可证明, 所以最小概率误差为:

$$P(e) = \frac{1}{2} - \frac{1}{\pi} \arctan \left| \frac{a_2 - a_1}{2b} \right|$$

(4)

容易发现，当 $a_1 = a_2$ 时候，有 $P(e) = 1/2$ 为最大值，因为当 $a_1 = a_2$ 时候，两个类别并没有区别，只是在两类当中随机选择一类进行猜测，所以错误率为 $1/2$ 。

3.

(1)

$$\hat{\mu} = \frac{5.0 + 7.0 + 9.0 + 11.0 + 13.0}{5} = 9.0$$

$$\hat{\sigma}^2 = \frac{(5.0 - 9.0)^2 + (7.0 - 9.0)^2 + (9.0 - 9.0)^2 + (11.0 - 9.0)^2 + (13.0 - 9.0)^2}{5} = 8.0$$

(2)

b越小似然函数越大，但是b不能小于8，这里参数b应当估计为8

(3)

由于先验概率相等只需要对比两类下样本的pdf。

$$p(x|\omega_1) = \frac{1}{\sqrt{2\pi * 8}} e^{-\frac{(6-9)^2}{2*8}} \approx 0.08$$

$$p(x|\omega_2) = 1/8 = 0.125$$

应当判定为第二类。

4.

(1)

$$\sum_{i=1}^n \ln p(x_i|\theta, \sigma) = \sum_{i=1}^n -\frac{(\ln x_i - \theta)^2}{2\sigma^2} - \ln \sigma x_i \sqrt{2\pi}$$

θ 的最大似然估计：

$$\begin{aligned}\frac{d \sum_{i=1}^n \ln p(x_i|\theta, \sigma)}{d\theta} &= 0 \\ \sum_{i=1}^n (\ln x_i - \theta) &= 0 \\ \hat{\theta}_{ML} &= \frac{\sum_{i=1}^n \ln x_i}{n}\end{aligned}$$

(2)

σ 的最大似然估计:

$$\begin{aligned}\frac{d \sum_{i=1}^n \ln p(x_i|\theta, \sigma)}{d\sigma} &= 0 \\ \sum_{i=1}^n \frac{(\ln x_i - \theta)^2}{\sigma^3} - \frac{x_i \sqrt{2\pi}}{\sigma} &= 0 \\ \hat{\sigma}^2 &= \frac{\sum_{i=1}^n (\ln x_i - \hat{\theta})^2}{\sqrt{2\pi} \sum_{i=1}^n x_i}\end{aligned}$$

(3)

θ 的MAP估计:

$$\begin{aligned}& \sum_{i=1}^n \ln p(x|\theta) + \ln p(\theta) \\ &= \sum_{i=1}^n \left(-\frac{(\ln x_i - \theta)^2}{2\sigma^2} - \ln \sigma x_i \sqrt{2\pi} \right) - \frac{1}{2} \ln (2\pi \sigma_0^2) - \frac{1}{2} \left(\frac{\theta - \theta_0}{\sigma_0} \right)^2\end{aligned}$$

于是:

$$\begin{aligned}\frac{d (\sum_{i=1}^n \ln p(x|\theta) + \ln p(\theta))}{d\theta} &= 0 \\ \sum_{i=1}^n \frac{\ln x_i - \theta}{\sigma^2} - \frac{\theta - \theta_0}{\sigma_0^2} &= 0 \\ \sigma_0^2 \sum_{i=1}^n \ln x_i - n\sigma_0^2 \theta - \sigma^2 \theta + \sigma^2 \theta_0 &= 0 \\ \hat{\theta}_{MAP} &= \frac{\sigma_0^2 \sum_{i=1}^n \ln x_i + \sigma^2 \theta_0}{n\sigma_0^2 + \sigma^2}\end{aligned}$$