

Optimistic Estimation of Convergence in Markov Chains with the Average Mixing Time

Pierre Alquier



SciCADE, July 15, 2024, Singapore

Talk is based on the preprint :



Wolfer, G. and Alquier, P. (2024).
Optimistic Estimation of
Convergence in Markov Chains with
the Average Mixing Time. Preprint
arXiv :2402.10506.



Geoffrey WOLFER

Waseda University, Tokyo

In this talk, $(X_1, X_2, X_3 \dots)$ is an **ergodic Markov chain** on a finite or countable space \mathcal{X} , with transition kernel P and invariant probability π .

$$d(t) = \sup_{x \in \mathcal{X}} \|\delta_x P^t - \pi\|_{\text{TV}}.$$

Mixing time

$$t_{\text{mix}}(\xi) = \min\{t : d(t) \leq \xi\} \in \{1, 2, \dots\} \cup \{+\infty\}.$$

Multiplicative property

$$t_{\text{mix}}(\xi) \leq \lceil \log_2(1/\xi) \rceil t_{\text{mix}}(1/4).$$

$$t_{\text{mix}} := t_{\text{mix}}(1/4).$$



Hsu, D., Kontorovich, A., Levin, D. A., Peres, Y., Szepesvári, C. & Wolfer, G. (2019). Mixing time estimation in reversible Markov chains from a single sample path. *Annals of Applied Probability*.



Wolfer, G. & Kontorovich, A. (2024). Improved estimation of relaxation time in nonreversible Markov chains. *Annals of Applied Probability*.

Based on a trajectory (X_1, \dots, X_n) on a **finite** \mathcal{X} we can

- estimate the spectral gap γ of P : with **probability** $1 - \delta$,

$$|\hat{\gamma} - \gamma| \leq C \sqrt{\frac{\log\left(\frac{\text{card}(\mathcal{X})}{\delta}\right) \log\left(\frac{n}{\pi_* \delta}\right)}{\pi_* \gamma n}}$$

where $\pi_* = \min_x \pi(x)$;

- estimate the relaxation time $t_{\text{rel}} = 1/\gamma$;
- finally,

$$(t_{\text{rel}} - 1) \log 2 \leq t_{\text{mix}} \leq t_{\text{rel}} \log \frac{4}{\pi_*}.$$

The mixing time is a pessimistic notion :

- $t_{\text{mix}}(\xi)$ can be infinite,
- $d(t) = \sup_{x \in \mathcal{X}} \|\delta_x P^t - \pi\|_{\text{TV}}$.

This motivates :

$$d^\#(t) = \sum_{x \in \mathcal{X}} \pi(x) \|\delta_x P^t - \pi\|_{\text{TV}}.$$

Average mixing time

$$t_{\text{mix}}^\#(\xi) = \min\{t : d^\#(t) \leq \xi\}.$$

Proposition 1

$$t_{\text{mix}}^{\#}(\xi) < +\infty.$$

Proposition 2

Assume $\text{card}(\mathcal{X}) = 2$. Then, for any $M > 0$, there exists a transition kernel P such that

$$\frac{t_{\text{mix}}(\xi)}{t_{\text{mix}}^{\#}(\xi)} > M.$$

Remarks :

- no multiplicative property,
- this talk : $t_{\text{mix}}^{\#}(\xi)$ is still informative and can be estimated.

A little detour to show that the average mixing time is a useful notion : let us explore its connection to **mixing coefficients**.

- Mixing coefficients are general tools developed to study the convergence of **general stochastic processes**.
- There are many notions : **α -mixing**, **β -mixing** and **φ -mixing** are the most popular.

Definition : β -mixing coefficients

For two σ -fields \mathcal{A} and \mathcal{B} ,

$$\beta(\mathcal{A}, \mathcal{B}) = \frac{1}{2} \sup \sum_{i=1}^I \sum_{j=1}^J |\mathbb{P}(A_i \cap B_j) - \mathbb{P}(A_i)\mathbb{P}(B_j)|,$$

sup over all partitions $(A_i)_{i=1}^I$ and $(B_j)_{j=1}^J$ of Ω by sets in \mathcal{A} and \mathcal{B} respectively ;

$$\beta(s) = \sup_{t, h, \ell} \beta(\sigma(X_t, \dots, X_{t+h}), \sigma(X_{t+h+s}, \dots, X_{t+h+s+\ell})).$$

Let f be a measurable function with $\mathbb{E}_\pi(f) = 0$ and $\|f\|_\infty \leq 1$. Let n, B, s be three integers with $n = Bs$.

Theorem (Yu, 1984)

$$\mathbb{P} \left(\frac{1}{n} \sum_{i=1}^n f(X_i) > \varepsilon \right) \leq 2 \exp \left(-\frac{n\varepsilon^2}{4s} \right) + 2(B-1)\beta(s).$$

Theorem (Mohri and Rostmizadeh, 2008)

$$\begin{aligned} \mathbb{P} \left(\sup_{f \in \mathcal{F}} \frac{1}{n} \sum_{i=1}^n f(X_i) > \varepsilon + 2\mathfrak{Rad}_B(\mathcal{F}) \right) \\ \leq 2 \exp \left(-\frac{n\varepsilon^2}{4s} \right) + 2(B-1)\beta(s) \end{aligned}$$

where $\mathfrak{Rad}_B(\mathcal{F})$ is the Rademacher complexity.

- Many classical processes satisfy $\beta(s) \leq \beta_0 \exp(-\beta_1 s^b)$ or $\beta(s) \leq \beta_1/s^b$.
- In theory, it is possible to estimate $\beta(s)$ based on a trajectory (X_1, \dots, X_n) , see



Khaleghi, A. and Lugosi, G. (2023). Inferring the mixing properties of a stationary ergodic process from a single sample-path. *IEEE Transactions on Information Theory*.

- However, their procedure is hardly feasible in practice, and the convergence $\hat{\beta}(s) \xrightarrow{n \rightarrow \infty} \beta(s)$ can be arbitrarily slow...

Reminder

In this talk, $(X_1, X_2, X_3 \dots)$ is an **ergodic Markov chain** with **transition kernel P** and **invariant probability π** .

Theorem (Davydov, 1973)

Assume the chain has initial distribution μ .

$$\beta(s) = \sup_t \sum_{x \in \mathcal{X}} \mu P^t(x) \|\delta_x P^s - \mu P^{t+s}\|_{\text{TV}},$$

$$\varphi(s) = \sup_t \sup_{x \in \mathcal{X}} \|\delta_x P^s - \mu P^{t+s}\|_{\text{TV}}.$$

Thus if the chain is stationary ($\mu = \pi$),

$$\beta(s) = \sum_x \pi(x) \|\delta_x P^s - \pi\|_{\text{TV}} = d^\#(s),$$

$$\varphi(s) = \sup_x \|\delta_x P^s - \pi\|_{\text{TV}} = d(s).$$

Corollary

When the chain is stationary,

$$t_{\text{mix}}^{\#}(\xi) = \min\{t : \beta(t) \leq \xi\}.$$

Corollary

Assume $\beta(s) \leq \beta_1/s^b$. Then

$$t_{\text{mix}}^{\#}(\xi) \leq \left\lceil (\beta_1/\xi)^{\frac{1}{b}} \right\rceil.$$

For any $\varepsilon, \delta > 0$, we have

$$\mathbb{P} \left(\frac{1}{n} \sum_{i=1}^n f(X_i) > \varepsilon \right) \leq \delta$$

as soon as

$$n \geq C \left(\frac{8}{\varepsilon^2} \log \frac{4}{\delta} \right)^{\frac{b+1}{b}} \left(2t_{\text{mix}}^{\#}(\delta) \right)^{\frac{1}{b}}.$$

Estimation of $\beta(t)$ and $t_{\text{mix}}^{\sharp}(\xi)$ for stationary, ergodic Markov chains from (X_1, \dots, X_n) .

$$\begin{aligned}\beta(s) &= \sum_x \pi(x) \|\delta_x P^s - \pi\|_{\text{TV}} \\ &= \frac{1}{2} \sum_{x,y} \pi(x) |P^s(x,y) - \pi(y)|\end{aligned}$$

- Obvious empirical estimates for $\hat{\pi}(x)$ and $\widehat{P^s}(x,y)$.
- Leads to an estimator $\hat{\beta}(s)$ of $\beta(s)$ (see the paper for the exact formula).

Theorem

For any $p \geq 1$,

$$\mathbb{E}|\hat{\beta}(s) - \beta(s)| \leq 4\sqrt{\frac{s\mathcal{B}_p^{(s)}\mathcal{I}_p^{(s)}}{n-s-1}}, \text{ where}$$

$$\mathcal{B}_p^{(s)} = \sum_{t=0}^{\infty} \beta(st)^{\frac{1}{p}} \text{ and } \mathcal{I}_p^{(s)} = \sum_{x,y \in \mathcal{X}} [\pi(x)P^s(x,y)]^{1-\frac{1}{p}}.$$

- The term $\mathcal{B}_p^{(s)}$ is bounded explicitly and finite under the standard assumptions $\beta(s) \leq \beta_0 \exp(-\beta_1 s^b)$ or $\beta(s) \leq \beta_1/s^b$, with conditions on p and b .
- Compare $\mathcal{I}_p^{(s)}$ to $\sum_y \sqrt{P(x,y)}$ that already appears when bounding $\|\hat{P} - P\|_{\infty}$:



Wolfer, G. (2024). Empirical and Instance-Dependent Estimation of Markov Chain and Mixing Time. *Scandinavian Journal of Statistics*.

Study of $\mathcal{I}_p^{(s)} = \sum_{x,y \in \mathcal{X}} [\pi(x) P^s(x, y)]^{1 - \frac{1}{p}}$.

- Case $s = 0$:

$$\mathcal{I}_p^{(0)} = \sum_{x \in \mathcal{X}} [\pi(x)]^{1 - \frac{1}{p}}.$$

- When $s \rightarrow \infty$, $P^s(x, y) \rightarrow \pi(y)$. Thus, at least when \mathcal{X} is finite, we expect :

$$\mathcal{I}_p^{(s)} \xrightarrow{s \rightarrow \infty} \sum_{x,y \in \mathcal{X}} [\pi(x)\pi(y)]^{1 - \frac{1}{p}} \leq \sum_{x \in \mathcal{X}} [\pi(x)]^{2(1 - \frac{1}{p})}.$$

Theorem

Assume :

- \mathcal{X} is finite, then $\mathcal{I}_p^{(s)} \leq [\text{card}(\mathcal{X})]^{2/p}$.
- the chain is V -geometrically ergodic, then, under suitable moment conditions on $1/V$ and π , then

$$\sup_{s \in \mathbb{N}} \mathcal{I}_p^{(s)} < +\infty.$$

- various weak conditions : $\mathcal{I}_p^{(s)}$ sub-linear in s .

Estimation of $\beta(t)$ and $t_{\text{mix}}^{\#}(\xi)$ for stationary, **uniformly** ergodic Markov chains from (X_1, \dots, X_n) .

Theorem

For any $p \geq 1$,

$$\mathbb{E}|\hat{\beta}(s) - \beta(s)| \leq 8 \sqrt{\frac{\left(s + \frac{t_{\text{mix}} p}{\log(2)}\right) \mathcal{I}_p^{(s)}}{n - 1}}.$$

In this case, we can say more...

Theorem

Let $p > 1$ and assume $\sup_s \mathcal{I}_p^{(s)} = \mathcal{I}_p < +\infty$. Put $\tilde{\beta}(s) = \hat{\beta}(s)1_{\{s \leq S\}}$ for some $S = S(n, \varepsilon, \delta, p, \mathcal{I}_p)$. Then, as soon as

$$n \gtrsim \frac{t_{\text{mix}} p \log \frac{1}{\varepsilon}}{\varepsilon^2} \max \left[\mathcal{I}_p, \log \left(\frac{t_{\text{mix}} p \log \frac{1}{\varepsilon}}{\varepsilon^2} \right) \right]$$

we have, with probability $1 - \delta$,

$$\sup_s |\tilde{\beta}(s) - \beta(s)| \leq \varepsilon.$$

Put $\hat{t}_{\text{mix}}^\#(\xi) = \min\{s : \tilde{\beta}(s) \leq \xi\}$. Then, for n (explicitly) large enough,

$$\hat{t}_{\text{mix}}^\#(\xi) \in \left[t_{\text{mix}}^\#(\xi(1 + \varepsilon)), t_{\text{mix}}^\#(\xi(1 - \varepsilon)) \right].$$

Thank you !