Optimistic Estimation of Convergence in Markov Chains with the Average Mixing Time

Pierre Alquier



SciCADE, July 15, 2024, Singapore

Talk is based on the preprint :



Wolfer, G. and Alquier, P. (2024). Optimistic Estimation of Convergence in Markov Chains with the Average Mixing Time. *Preprint arXiv*: 2402.10506.



Geoffrey WOLFER
Waseda University, Tokyo

In this talk, $(X_1, X_2, X_3...)$ is an ergodic Markov chain on a finite or countable space \mathcal{X} , with transition kernel P and invariant probability π .

$$d(t) = \sup_{x \in \mathcal{X}} \|\delta_x P^t - \pi\|_{\text{TV}}.$$

Mixing time

$$t_{\text{mix}}(\xi) = \min\{t : d(t) \le \xi\} \in \{1, 2, \dots\} \cup \{+\infty\}.$$

Multiplicative property

$$t_{\text{mix}}(\xi) \leq \lceil \log_2(1/\xi) \rceil t_{\text{mix}}(1/4).$$

$$t_{\text{mix}} := t_{\text{mix}}(1/4).$$



Hsu, D., Kontorovich, A., Levin, D. A., Peres, Y., Szepesvári, C. & Wolfer, G. (2019). Mixing time estimation in reversible Markov chains from a single sample path. *Annals of Applied Probability*.



Wolfer, G. & Kontorovich, A. (2024). Improved estimation of relaxation time in nonreversible Markov chains. *Annals of Applied Probability*.

Based on a trajectory (X_1, \ldots, X_n) on a finite \mathcal{X} we can

• estimate the spectral gap γ of P : with probability $1-\delta$,

$$|\hat{\gamma} - \gamma| \le C \sqrt{\frac{\log\left(\frac{\operatorname{card}(\mathcal{X})}{\delta}\right)\log\left(\frac{n}{\pi_*\delta}\right)}{\pi_*\gamma n}}$$

where $\pi_* = \min_x \pi(x)$;

- ullet estimate the relaxation time $t_{\rm rel}=1/\gamma$:
- finally,

$$(t_{\mathrm{rel}}-1)\log 2 \leq t_{\mathrm{mix}} \leq t_{\mathrm{rel}}\log \frac{4}{\pi}$$

The mixing time is a pessimistic notion :

- $t_{\rm mix}(\xi)$ can be infinite,

This motivates:

$$d^{\sharp}(t) = \sum_{\mathbf{x} \in \mathcal{X}} \pi(\mathbf{x}) \| \delta_{\mathbf{x}} P^{t} - \pi \|_{\text{TV}}.$$

Average mixing time

$$t_{\mathrm{mix}}^{\sharp}(\xi) = \min\{t : d^{\sharp}(t) \leq \xi\}.$$

Proposition 1

$$t_{\mathrm{mix}}^{\sharp}(\xi)<+\infty.$$

Proposition 2

Assume $\operatorname{card}(\mathcal{X}) = 2$. Then, for any M > 0, there exists a transition kernel P such that

$$rac{t_{ ext{mix}}(\xi)}{t_{ ext{mix}}^{\sharp}(\xi)} > M.$$

Remarks:

- no multiplicative property,
- this talk : $t_{\text{mix}}^{\sharp}(\xi)$ is still informative and can be estimated.

A little detour to show that the average mixing time is a useful notion : let us explore its connection to mixing coefficients.

- Mixing coefficients are general tools developed to study the convergence of general stochastic processes.
- There are many notions : α -mixing, β -mixing and φ -mixing are the most popular.

Definition : β -mixing coefficients

For two σ -fields \mathcal{A} and \mathcal{B} .

$$eta(\mathcal{A},\mathcal{B}) = rac{1}{2} \sup \sum_{i=1}^{J} \sum_{i=1}^{J} |\mathbb{P}(A_i \cap B_j) - \mathbb{P}(A_i)\mathbb{P}(B_j)|,$$

sup over all partitions $(A_i)_{i=1}^I$ and $(B_j)_{j=1}^J$ of Ω by sets in \mathcal{A} and \mathcal{B} respectively;

$$\beta(s) = \sup_{t,h,\ell} \beta(\sigma(X_t,\ldots,X_{t+h}),\sigma(X_{t+h+s},\ldots,X_{t+h+s+\ell})).$$

Let f be a measurable function with $\mathbb{E}_{\pi}(f) = 0$ and $||f||_{\infty} \leq 1$. Let n, B, s be three integers with n = Bs.

Theorem (Yu, 1984)

$$\mathbb{P}\left(\frac{1}{n}\sum_{i=1}^{n}f(X_{i})>\varepsilon\right)\leq2\exp\left(-\frac{n\varepsilon^{2}}{4s}\right)+2(B-1)\beta(s).$$

Theorem (Mohri and Rostmizadeh, 2008)

$$\mathbb{P}\left(\sup_{f\in\mathcal{F}}\frac{1}{n}\sum_{i=1}^{n}f(X_{i})>\varepsilon+2\mathfrak{Rad}_{B}(\mathcal{F})\right)$$

$$\leq 2\exp\left(-\frac{n\varepsilon^{2}}{4s}\right)+2(B-1)\beta(s)$$

where $\mathfrak{Rad}_{B}(\mathcal{F})$ is the Rademacher complexity.

- Many classical processes satisfy $\beta(s) \leq \beta_0 \exp(-\beta_1 s^b)$ or $\beta(s) \leq \beta_1/s^b$.
- In theory, it is possible to estimate $\beta(s)$ based on a trajectory (X_1, \ldots, X_n) , see



Khaleghi, A. and Lugosi, G. (2023). Inferring the mixing properties of a stationary ergodic process from a single sample-path. *IEEE Transactions on Information Theory*.

• However, their procedure is hardly feasible in practice, and the convergence $\hat{\beta}(s) \xrightarrow[n \to \infty]{} \beta(s)$ can be arbitrarily slow...

Reminder

In this talk, $(X_1, X_2, X_3...)$ is an ergodic Markov chain with transition kernel P and invariant probability π .

Theorem (Davydov, 1973)

Assume the chain has initial distribution μ .

$$\beta(s) = \sup_{t} \sum_{x \in \mathcal{X}} \mu P^{t}(x) \|\delta_{x} P^{s} - \mu P^{t+s}\|_{\text{TV}},$$

$$\varphi(s) = \sup_{t} \sup_{x \in \mathcal{X}} \|\delta_{x} P^{s} - \mu P^{t+s}\|_{\text{TV}}.$$

Thus if the chain is stationary ($\mu = \pi$),

$$\beta(s) = \sum_{x} \pi(x) \|\delta_{x} P^{s} - \pi\|_{\text{TV}} = d^{\sharp}(s),$$

$$\varphi(s) = \sup_{x} \|\delta_{x} P^{s} - \pi\|_{\text{TV}} = d(s).$$

Corollary

When the chain is stationary,

$$t_{\min}^{\sharp}(\xi) = \min\{t : \beta(t) \leq \xi\}.$$

Corollary

Assume $\beta(s) \leq \beta_1/s^b$. Then

$$t_{\mathrm{mix}}^{\sharp}(\xi) \leq \left\lceil (\beta_1/\xi)^{\frac{1}{b}} \right\rceil.$$

For any $\varepsilon, \delta > 0$, we have

$$\mathbb{P}\left(\frac{1}{n}\sum_{i=1}^n f(X_i) > \varepsilon\right) \le \delta$$

as soon as

$$n \geq C \left(\frac{8}{\varepsilon^2} \log \frac{4}{\delta}\right)^{\frac{b+1}{b}} \left(2t_{\mathrm{mix}}^\sharp(\delta)\right)^{\frac{1}{b}}.$$

Estimation of $\beta(t)$ and $t_{\text{mix}}^{\sharp}(\xi)$ for stationary, erdogic Markov chains from (X_1, \ldots, X_n) .

$$\beta(s) = \sum_{x} \pi(x) \|\delta_x P^s - \pi\|_{\text{TV}}$$
$$= \frac{1}{2} \sum_{x,y} \pi(x) |P^s(x,y) - \pi(y)|$$

- Obvious empirical estimates for $\hat{\pi}(x)$ and $\widehat{P}^s(x,y)$.
- Leads to an estimator $\hat{\beta}(s)$ of $\beta(s)$ (see the paper for the exact formula).

$\mathsf{Theorem}$

For any p > 1,

$$\mathbb{E}|\hat{eta}(s)-eta(s)| \leq 4\sqrt{rac{s\mathcal{B}_p^{(s)}\mathcal{I}_p^{(s)}}{n-s-1}},$$
 where

$$\mathcal{B}_{p}^{(s)} = \sum_{t=0}^{\infty} \beta(st)^{\frac{1}{p}} \text{ and } \mathcal{I}_{p}^{(s)} = \sum_{x,y \in \mathcal{X}} [\pi(x)P^{s}(x,y)]^{1-\frac{1}{p}}.$$

- The term $\mathcal{B}_{n}^{(s)}$ is bounded explicitly and finite under the standard assumptions $\beta(s) < \beta_0 \exp(-\beta_1 s^b)$ or $\beta(s) < \beta_1/s^b$, with conditions on p and b.
- Compare $\mathcal{I}_p^{(s)}$ to $\sum_{y} \sqrt{P(x,y)}$ that already appears when bounding $\|\widehat{P} - P\|_{\infty}$:



Wolfer, G. (2024). Empirical and Instance-Dependent Estimation of Markov Chain and Mixing Time. Scandinavian Journal of Statistics.

Study of
$$\mathcal{I}_p^{(s)} = \sum_{x,y \in \mathcal{X}} [\pi(x)P^s(x,y)]^{1-\frac{1}{p}}$$
.

• Case s = 0:

$$\mathcal{I}_p^{(0)} = \sum_{x \in \mathcal{X}} [\pi(x)]^{1 - \frac{1}{p}}.$$

• When $s \to \infty$, $P^s(x, y) \to \pi(y)$. Thus, at least when $\mathcal X$ is finite, we expect :

$$\mathcal{I}_p^{(s)} \xrightarrow[s \to \infty]{} \sum_{x,y \in \mathcal{X}} [\pi(x)\pi(y)]^{1-\frac{1}{p}} \leq \sum_{x \in \mathcal{X}} [\pi(x)]^{2\left(1-\frac{1}{p}\right)}.$$

Theorem

Assume:

- \mathcal{X} is finite, then $\mathcal{I}_p^{(s)} \leq [\operatorname{card}(\mathcal{X})]^{2/p}$.
- the chain is V-geometrically ergodic, then, under suitable moment conditions on 1/V and π , then

$$\sup_{s\in\mathbb{N}}\mathcal{I}_p^{(s)}<+\infty.$$

• various weak conditions : $\mathcal{I}_p^{(s)}$ sub-linear in s.

Estimation of $\beta(t)$ and $t_{\text{mix}}^{\sharp}(\xi)$ for stationary, uniformly erdogic Markov chains from (X_1, \ldots, X_n) .

Theorem

For any $p \ge 1$,

$$|\mathbb{E}|\hat{\beta}(s) - \beta(s)| \leq 8\sqrt{\frac{\left(s + \frac{t_{\min}\rho}{\log(2)}\right)\mathcal{I}_{\rho}^{(s)}}{n-1}}.$$

In this case, we can say more...

Theorem

Let p>1 and assume $\sup_s \mathcal{I}_p^{(s)} = \mathcal{I}_p < +\infty$. Put $\tilde{\beta}(s) = \hat{\beta}(s) 1_{\{s \leq S\}}$ for some $S = S(n, \varepsilon, \delta, p, \mathcal{I}_p)$. Then, as soon as

$$n \gtrsim \frac{t_{\min} \rho \log \frac{1}{\varepsilon}}{\varepsilon^2} \max \left[\mathcal{I}_p, \log \left(\frac{t_{\min} \rho \log \frac{1}{\varepsilon}}{\varepsilon^2} \right) \right]$$

we have, with probability $1 - \delta$,

$$\sup_{s} |\tilde{\beta}(s) - \beta(s)| \le \varepsilon.$$

Put $\hat{t}^{\sharp}_{mix}(\xi) = \min\{s : \tilde{\beta}(s) \leq \xi\}$. Then, for n (explicitly) large enough,

$$\hat{t}_{ ext{mix}}^{\sharp}(\xi) \in \left[t_{ ext{mix}}^{\sharp}(\xi(1+arepsilon)), t_{ ext{mix}}^{\sharp}(\xi(1-arepsilon))
ight].$$

Thank you!