

1 Limited predictability of amino acid substitutions in seasonal influenza viruses

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14 Seasonal influenza viruses repeatedly infect humans in part because they rapidly change
15 their antigenic properties and evade host immune responses, necessitating frequent
16 updates of the vaccine composition. Accurate predictions of strains circulating in the
17 future could therefore improve the vaccine match. Here, we studied the predictability of
18 frequency dynamics and fixation of amino acid substitutions. Current frequency was the
19 strongest predictor of eventual fixation, as expected in neutral evolution. Other properties,
20 such as occurrence in previously characterized epitopes or high *Local Branching Index*
21 (LBI) had little predictive power. Parallel evolution was found to be moderately predictive
22 of fixation. While the LBI had little power to predict frequency dynamics, it was still
23 successful at picking strains representative of future populations. The latter is due to a
24 tendency of the LBI to be high for consensus-like sequences that are closer to the future
25 than the average sequence. Simulations of models of adapting populations, in contrast,
26 show clear signals of predictability. This indicates that the evolution of influenza HA and
27 NA, while driven by strong selection pressure to change, is poorly described by common
28 models of directional selection such as travelling fitness waves.

29 INTRODUCTION

30 Seasonal influenza A viruses (IAV) infect about 10% of 53
31 the global population every year, resulting in hundreds 54
32 of thousands of deaths (Organization, 2018; Petrova and 55
33 Russell, 2017). Vaccination is the primary measure to 56
34 reduce influenza morbidity. However, the surface proteins 57
35 hemagglutinin (HA) and neuraminidase (NA) continu- 58
36 ously accumulate mutations at a high rate, leading to fre- 59
37 quent antigenic changes (Petrova and Russell, 2017; Shih 60
38 et al., 2007; Bhatt et al., 2011; Koel et al., 2013). While 61
39 a vaccine targeting a particular strain may be efficient for 62
40 some time, antigenic drift will sooner or later render it ob- 63
41 solete. The World Health Organization (WHO) regularly 64
42 updates influenza vaccine recommendations to best match 65
43 the circulating strains. Since developing, manufacturing, 66
44 and distributing the vaccine takes many months, forecast- 67
45 ing the evolution of influenza is of essential interest to 68
46 public health (Morris et al., 2018; Klingen et al., 2018a). 69

47 The number of available high quality HA and 70
48 NA sequences has increased rapidly over the last 20 71
49 years (Bogner et al., 2006; Shu and McCauley, 2017) 72
50 and virus evolution and dynamics can be now be tracked 73

51 at high temporal and spatial resolution (Rambaut et al.,
52 2008). This wealth of data has given rise to an active
53 field of predicting influenza virus evolution (Morris et al.,
54 2018; Klingen et al., 2018a). These models predict the
55 future population of influenza viruses by estimating strain
56 fitness or proxies of fitness. Luksza and Lässig (2014),
57 for example, train a fitness model to capture antigenic
58 drift and protein stability on patterns of epitope and non-
59 epitope mutations. Other approaches by Steinbrück et al.
60 (2014); Neher et al. (2016) predict fitness by using hemag-
61 glutination inhibition (HI) data to determine possible
62 antigenic drift of clades in the genealogy of the HA pro-
63 tein. Finally, Neher et al. (2014) use branching patterns
64 of HA phylogenies as a proxy for fitness. These branching
65 patterns are summarized by the Local Branching Index
66 (LBI), which was shown to be a proxy of relative fitness
67 in mathematical models of rapidly adapting populations
68 (Neher et al., 2014).

69 The underlying assumption of all these methods is that
70 (i) differences in growth rate between strains can be es-
71 timated from sequence or antigenic data and (ii) that
72 these growth rate differences persist for long enough to be
73 predictive of future success. Specific positions in surface
74 proteins are of particular interest in this context. The
75 surface proteins are under a strong positive selection and
76 change their amino acid sequence much more rapidly than
77 other IAV proteins or than expected under neutral evo-
78 lution (Bhatt et al., 2011; Strelkowa and Lässig, 2012).

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Epitope positions, i.e., positions targeted by human antibodies, are expected to change particularly often since viruses with altered epitopes can evade existing immune responses (Shih et al., 2007; Koel et al., 2013; Wolf et al., 2006). It therefore seems plausible that mutations at these positions have a tendency to increase fitness and a higher probability of fixation (Strelkowa and Lässig, 2012). But one has to be careful to account for the fact that these positions are often ascertained post-hoc (Shih et al., 2007) and human immune responses are diverse with substantial inter-individual variation (Lee et al., 2019).

In this work, we use HA and NA sequences of A/H3N2 and A/H1N1pdm influenza from year 2000 to 2019 to perform a retrospective analysis of frequency trajectories of amino acid mutations. We quantify how rapidly mutations at different frequencies are lost or fixed and how rapidly they spread through the population. We further investigate whether any properties or statistics are predictive of whether a particular mutation fixes or not. To our surprise, we find that the predictability of these trajectories is very limited: The probability that a mutation fixes differs little from its current frequency, as would be expected if fixation happened purely by chance. This observation holds for many different categories of mutations, including mutations at epitope positions. This weak predictability is not attributable solely to clonal interference and genetic linkage, as simulation of models including even strong interference retain clear signatures of predictability. Consistent with these observations, we show that a simple predictor uninformed by fitness, the consensus sequence, performs as well as the Local Branching Index (LBI), the growth measure based on the genealogy used in (Neher et al., 2014). This suggests that although LBI has predictive power, the reason for its success may not be related to it approximating fitness of strains.

qualitative question, is to ask whether X will fix in the population, will disappear, or whether the site will stay polymorphic.

We use amino-acid sequences of the HA and NA genes of A/H3N2 (since the year 2000) and A/H1N1pdm (since the year 2009) influenza available in GISAID (Shu and McCauley, 2017) (see supplementary materials for an acknowledgment of all data contributors). This amounts to 44 976 HA and 36 300 NA sequences for A/H3N2 and 45 350 HA and 40 412 NA sequences for A/H1N1, with a minimum of 100 per year. These sequences are binned in non-overlapping intervals of one month. Each single-month time bin and the sequences that it contains represent a (noisy) snapshot of the influenza population at a given date. The number of sequences per time bin varies strongly both with year and according to the season, with earlier time bins containing around 10 sequences while more recent bins contain several hundreds (see figures S9 and S10 in SM for details).

The central quantities that we derived from this data are *frequency trajectories* of amino acids at each position in the sequences. If an amino acid X_i is found at position i at a frequency between 5% and 95% in the population of a given time bin t , then the population is considered polymorphic at position i and at time t . This polymorphism is characterized by the frequency of X_i , $f_{X_i}(t)$, and also by frequencies of other amino acids at i . The series of values $f_{X_i}(t)$ for contiguous time bins constitutes the frequency trajectory of X_i . A trajectory is terminated if the corresponding frequency is measured above 95% (resp. below 5%) for two time bins in a row, in which case amino acid X_i is considered as *fixed* (resp. *absent*) in the population. Otherwise, the trajectory is considered *active*. Examples of trajectories can be seen in figure S11 of the Supplement.

Since segments in influenza do not recombine, trajectories defined as above are sometimes heavily correlated with each other. Indeed, the trajectories corresponding to two mutations that partly appear on the same background will tend to show a similar behavior. In the extreme case where the two mutations always appear on the same strains, their two trajectories will be identical. This “nested” characterized or frequency trajectories may introduce biases when performing statistical analysis. In section .3 of the SM, we attempt to correct for nested trajectories by clustering them based on the strains that compose them. However, this does not significantly change the results presented below.

In the rest of this work, we will focus on frequency trajectories that are starting at a zero (low) frequency, *i.e.* $f(t=0)=0$. These represent new amino acid variants which were absent in the population at the time bin when the trajectory started and are currently rising in the population (see Methods). Such distinction in novel and ancestral variants is necessary to meaningfully interrogate predictability. Each rising trajectory of a new mutation

RESULTS

The main underlying question asked in this work is the following: given a mutation X in the genome of influenza that we observe at a frequency f in the population at a given date, what can we say about the future of X ? The trajectory of a mutation will depend on its own effect on fitness, the contribution of the genetic background on the same segment, and the effect of the remaining seven segments. Here, we investigate properties of broad categories of mutations effectively averaging over different genetic backgrounds to isolate the effects intrinsic to the mutation.

First, we ask whether we can quantitatively predict the frequency of X at future times $f(t)$. In other words, having observed a mutation at frequencies (f_1, f_2, \dots, f_n) at dates (t_1, t_2, \dots, t_n) , what can we say about its frequency at future dates $(t_{n+1}, t_{n+2}, \dots)$? A simpler, more

implies the existence of another decreasing one at the²⁴¹
 same position, since frequencies of all amino acids at a²⁴²
 given position must sum to one. If novel variants arise by²⁴³
 selection, we expect to see a stronger signal of selection²⁴⁴
 after conditioning on these novel variants. In classic mod-²⁴⁵
 els of population genetics, strongly advantageous variants²⁴⁶
 undergo rapid selective *sweeps*, i.e., the rapid rise and²⁴⁷
 fixation. The sweep of a mutation can be due to its own²⁴⁸
 fitness effect, to the genetic background or to the effect of²⁴⁹
 the seven other segments. By considering the ensemble of²⁵⁰
 novel variants that are rising in frequency, we effectively²⁵¹
 average over backgrounds, obtaining a set of mutations²⁵²
 that we expect to be beneficial on average. If such sweeps²⁵³
 are common in the evolution of HA and NA, the restric-²⁵⁴
 tion to trajectories that start at low frequency should²⁵⁵
 thus enrich for mutations that are positively selected and²⁵⁶
 on their way to fixation.²⁵⁷

205 Predicting future frequencies

Having observed the frequency trajectory $f(t)$ of a²⁶²
 mutation until a given date t_0 , how much can we say²⁶³
 about the future values of f after t_0 ? We consider the²⁶⁴
 idealized case sketched in panel **A** of figure 1: given²⁶⁵
 the trajectory of a *new* mutation, i.e. that started at a²⁶⁶
 frequency of 0, and that we observe at frequency f_0 at²⁶⁷
 time t_0 , what is the probability $P_{\Delta t}(f)$ of observing it at²⁶⁸
 a value f at time $t_0 + \Delta t$?²⁶⁹

To answer this question retrospectively, we use all fre-²⁷⁰
 quency trajectories extracted from HA and NA sequences²⁷¹
 that satisfy these conditions for a given f_0 . The num-²⁷²
 ber of trajectories is limited and the frequency estimates²⁷³
 themselves are based on a finite sample and are hence im-²⁷⁴
 precise. Therefore, we consider trajectories in an interval²⁷⁵
 $[f_0 - \delta f, f_0 + \delta f]$ with $\delta f = 0.05$.²⁷⁶

For $f_0 = 0.3$, we found 120 such trajectories in the²⁷⁷
 case of A/H3N2 influenza, represented on the panel **B**²⁷⁸
 of figure 1, where time is shifted such that $t_0 = 0$. The²⁷⁹
 same analysis was performed for A/H1N1pdm, with the²⁸⁰
 89 found trajectories displayed in figure S13. Some trajec-²⁸¹
 tories fall in the frequency bin around f_0 while decreasing²⁸²
 even though they crossed that bin at an earlier time. This²⁸³
 is due to the fact that some trajectories “skipped” the²⁸⁴
 interval f_0 in question on their initial rise due to sparse²⁸⁵
 sampling. These trajectories are nevertheless rising in
 the sense that they start at frequency 0 for $t \rightarrow -\infty$ ²⁸⁶
 Removing them does not change results significantly.²⁸⁷

Since rapid sequence evolution of influenza HA and²⁸⁸
 NA mediates immune evasion, one could expect that²⁸⁹
 a significant fraction of new amino acid mutations on²⁹⁰
 rising trajectories in figure 1 are *adaptive*. We could thus²⁹¹
 expect that most of these trajectories continue to rise after²⁹²
 reaching frequency f_0 , at least for some time. A fraction²⁹³
 of those would then sweep through the population and²⁹⁴
 fix.²⁹⁵

To quantify the extent to which this preconception of sweeping adaptive mutations is true, we estimated the probability distribution $P_{\Delta t}(f|f_0)$ of finding a trajectory at frequency f after a time Δt given that it was observed at f_0 at time 0. The results for different Δt are shown in figure 1C. Initially, i.e. at time $t_0 = 0$, this distribution is by construction peaked around f_0 . If a large fraction of the trajectories keep increasing after this time, we should see the “mass” of $P_{\Delta t}(f|f_0)$ move to the right towards higher frequencies as time progresses.

However, future distributions for $\Delta t > 0$ do not seem to follow a pattern compatible with selective sweeps. The thick black line in Figure 1B shows the average frequency of all trajectories. This average makes a sharp turn at $t = 0$ and is essentially flat for $t > 0$ in the case of A/H3N2, and slightly increasing for A/H1N1pdm (see supplement). Hence, the fact that this average rose for $t < 0$ gives little information for $t > 0$, and is due to the conditions by which these trajectories were selected. This shows that sweep-like trajectories rising steadily from frequency 0 to 1 are not common enough to dominate the average trajectory. **Pierre:** This result is consistent with (Illingworth and Mustonen, 2012). They show that sweep-like trajectories are very bad fits to actual trajectories.

Consistent with the average, the frequency distribution of the selected trajectories broadens in time without a significant shift of the mean as time passes. After 60 days, the distribution is rather symmetrical around the initial $f_0 = 0.3$ value, suggesting that the knowledge that the trajectories were rising is lost after two months. On a timescale of 60 to 120 days, the only possible prediction is that trajectories are likely to be found in a broad interval around the initial frequency f_0 . After one year the distribution becomes almost flat (excluding mutations that have disappeared or fixed), and the initial peak at f_0 is not visible anymore. The only information remaining from the initial frequency is the fraction that fixed or was lost (see below). This behavior is expected in neutral models of evolution (Kimura, 1964) but incompatible with a dynamic dominated by sweeps taking over the population.

While this observation does not rule out that signatures exist that predict future frequency dynamics, past dynamics alone is weakly informative.

Prediction of fixation or loss

Instead of predicting future frequency, let’s consider the long-term goal of predicting the probability that a mutation fixes in the population. We first estimate the fraction of frequency trajectories that either fix in the population or are lost, as well as the time it takes for one or the other to happen. Panels **A** and **B** of figure 2 shows the fraction of frequency trajectories in HA and NA that either have fixed, were lost or remained active as

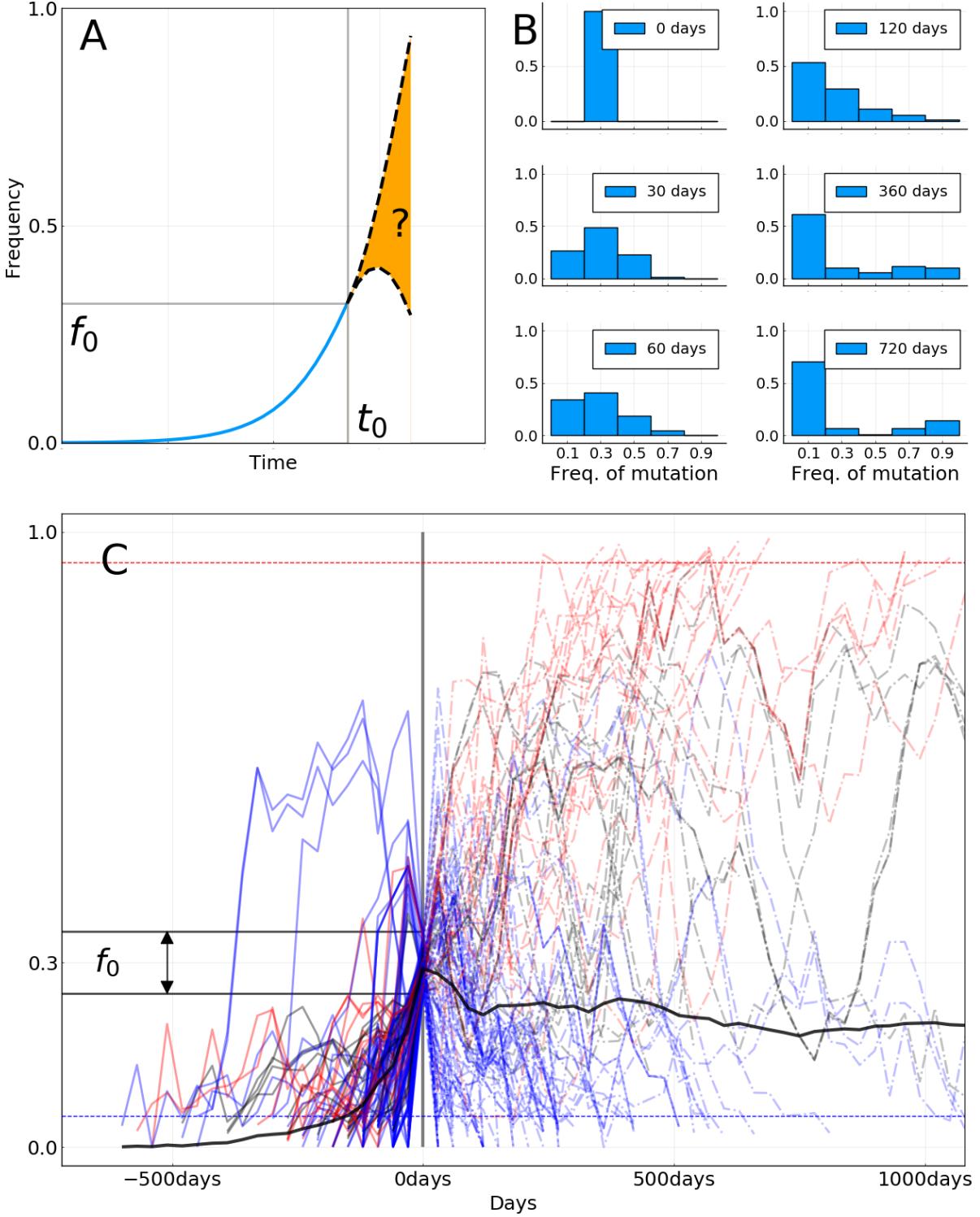


FIG. 1 **A:** Sketch of the idea behind the short term prediction of frequency trajectories. Given a mutation that we have seen increasing in frequency and that we “catch” at frequency f_0 at time t_0 , what can we say about the distribution of future frequencies $P_{\Delta t}(f|f_0)$? **B:** All frequency trajectories of amino acid mutations in the A/H3N2 HA and NA genes that were absent in the past, are seen around $f_0 = 30\%$ frequency at time $t_0 = 0$, and are based on more than 10 sequences at each time point. Red curves represent mutations that will ultimately fix, blue the ones that will be lost, and black the ones for which we do not know the final status. Dashed horizontal lines (blue and red) represent loss and fixation thresholds. The thick black line is the average of all trajectories, counting those that fix (resp. disappear) as being at frequency 1 (resp. 0). Figure S12 shows equivalent figures for other values of f_0 . **C:** Distribution of future frequencies $P_{\Delta t}(f|f_0)$ for the trajectories shown in panel **B** and for specific values of Δt .

a function of the time elapsed since they were first seen³⁵⁰
 above 25% frequency. Most mutations are either lost or³⁵¹
 become fixed after 2-3 years, with very few trajectories³⁵²
 remaining active after 5 years. This time scale of 2-3 years³⁵³
 is consistent with the typical coalescence time observed in³⁵⁴
 phylogenetic trees of A/H3N2 influenza (Rambaut et al.,³⁵⁵
 2008; Yan et al., 2019). We also note that the fraction³⁵⁶
 of lost trajectories increases sharply at small times with³⁵⁷
 40% of mutations observed above 25% frequency being³⁵⁸
 lost within one year for A/H3N2, while it takes longer to³⁵⁹
 fix a mutation in the whole population.³⁶⁰

We then examined the probability of mutations to fix³⁶¹
 in the population as a function of the frequency at which³⁶²
 they are seen. For different values of frequency f , we³⁶³
 consider all trajectories that started at a null frequency³⁶⁴
 and are seen in the interval $[f - 7.5\%, f + 7.5\%]$ at any³⁶⁵
 given time. The probability of a mutation fixing given³⁶⁶
 that it is seen at frequency f , $P_{fix}(f)$, is then estimated³⁶⁷
 by the fraction of those trajectories which terminate at³⁶⁸
 a frequency larger than 95%, *i.e.* our fixation threshold.³⁶⁹
 Panels **C** and **D** of figure 2 show $P_{fix}(f)$ as a function³⁷⁰
 of f for NA and HA. For both proteins, the probability³⁷¹
 of fixation of a new mutation at frequency f is close to³⁷²
 f itself, that is $P_{fix}(f) \simeq f$. This result is exactly what³⁷³
 is expected in a population evolving in the *absence of*³⁷⁴
 selection. A mutation or trait appearing at frequency f ³⁷⁵
 is shared by $f \cdot N$ individuals, and the probability for³⁷⁶
 one of them to become the ancestor of all the future³⁷⁷
 population is $f \cdot N/N = f$. Thus, the probability of this³⁷⁸
 mutation or trait to fix in the population is equal to its³⁷⁹
 current frequency, a case which we will refer to as the³⁸⁰
 neutral expectation. Panel **C** of figure 2 indicates that³⁸¹
 mutations in the surface proteins of A/H3N2 influenza³⁸²
 are in good agreement with the neutral expectation, while³⁸³
 those in A/H1N1pdm show only small deviations from³⁸⁴
 it. In both cases, the probability of fixation seems to be³⁸⁵
 mainly dictated by the current frequency f at which the³⁸⁶
 mutation is observed.³⁸⁷

This dynamics is in apparent contradiction with evidence³⁸⁸
 that influenza surface proteins are under strong selective³⁸⁹
 pressure to evade human immune responses (Bhatt³⁹⁰
 et al., 2011). If strong selection was present, we would³⁹¹
 expect rising amino acid mutations to fix at a distinctively³⁹²
 higher frequency than the one at which they are measured.³⁹³
 In an extreme case where most trajectories³⁹⁴
 would be clean sweeps, $P_{fix}(f)$ should be close to 1 for³⁹⁵
 all but very small values of f .³⁹⁶

Next, we searched for features of mutations that allow³⁹⁷
 low prediction of fixation beyond frequency by dividing³⁹⁸
 frequencies into categories that deviate from the diagonal³⁹⁹
 in panels **C** and **D** of figure 2. We first turn to the Local⁴⁰⁰
Branching Index (LBI), a quantity calculated for each⁴⁰¹
 node in a phylogenetic tree that indicates how dense⁴⁰²
 the branching of the tree is around that node. LBI has⁴⁰³
 previously been successfully used as a predictor of the future⁴⁰⁴
 population of influenza (Neher et al., 2014), and⁴⁰⁵

was shown to be a proxy for fitness of leaves or ancestral nodes in mathematical models of evolution. Here, we define the LBI of a mutation at date t as the average LBI of strains that carry this mutation and that were sampled in the time bin corresponding to t . Panel **A** of figure 3 shows fixation probability for HA mutations with LBI in the top or bottom half of the distribution. Both groups have identical probability of fixation, suggesting that LBI carries very little information on the probability of fixation of a mutation.

Next, we focused on previously reported antigenic sites in the A/H3N2 HA protein, referred to as *epitope* positions. Mutations at these position might mediate immune escape and are therefore likely under strong selection and show sweep-like behavior. We used four lists of relevant epitope positions from different sources comprising from 7 to 129 positions in the sequence of the HA1 protein (Shih et al., 2007; Koel et al., 2013; Luksza and Lässig, 2014; Wolf et al., 2006). Panel Fig. 3**B** shows fixation probability as a function of frequency for the four lists of epitopes. Only mutations at the 7 epitope sites reported in (Koel et al., 2013) have higher chances of fixation than expected by chance. No clear difference is found for the lists by Luksza and Lässig (2014); Wolf et al. (2006), while positions from Shih et al. (2007) show lower chances of fixation. One should also note that many of these positions were determined post-hoc and might be enriched for positions that experienced rapid substitutions before the publication of the respective studies.

Two ways of categorising mutations, however, suggest some power to predict fixation. In panel Fig. 3**C**, we split trajectories into those occurring at binary positions where only two amino acid variants co-circulate and non-binary positions with more than two variants. Novel variants at non-binary positions, *i.e.* ones for which competition between three amino acids or more has occurred at least once, have a higher chance of fixation. In panel **D**, we separated mutations that appear more than once or only once in the reconstructed tree (see methods), and found that the former fix more often. Panels **C** and **D** show that it is possible to gain some information on the chance of fixation of a particular mutation, as was done in panel **B**. However, the predictive power remains small, with the “top” curves in panels **C&D** being very close to the diagonal.

We conduct the same analysis on A/H1N1pdm influenza, with results shown in figure S15. Results are qualitatively similar to those obtained for A/H3N2, with LBI giving little information and mutations at non-binary positions having a higher chance of fixation. Panel **D** differs between figures 3 and S15, with convergent evolution giving less information on fixation in the latter case. However, this could be due to the shorter time period over which A/H1N1pdm evolved, resulting in a shorter tree and less possibilities of convergent evolution. Indeed, error bars for mutations appearing multiple times

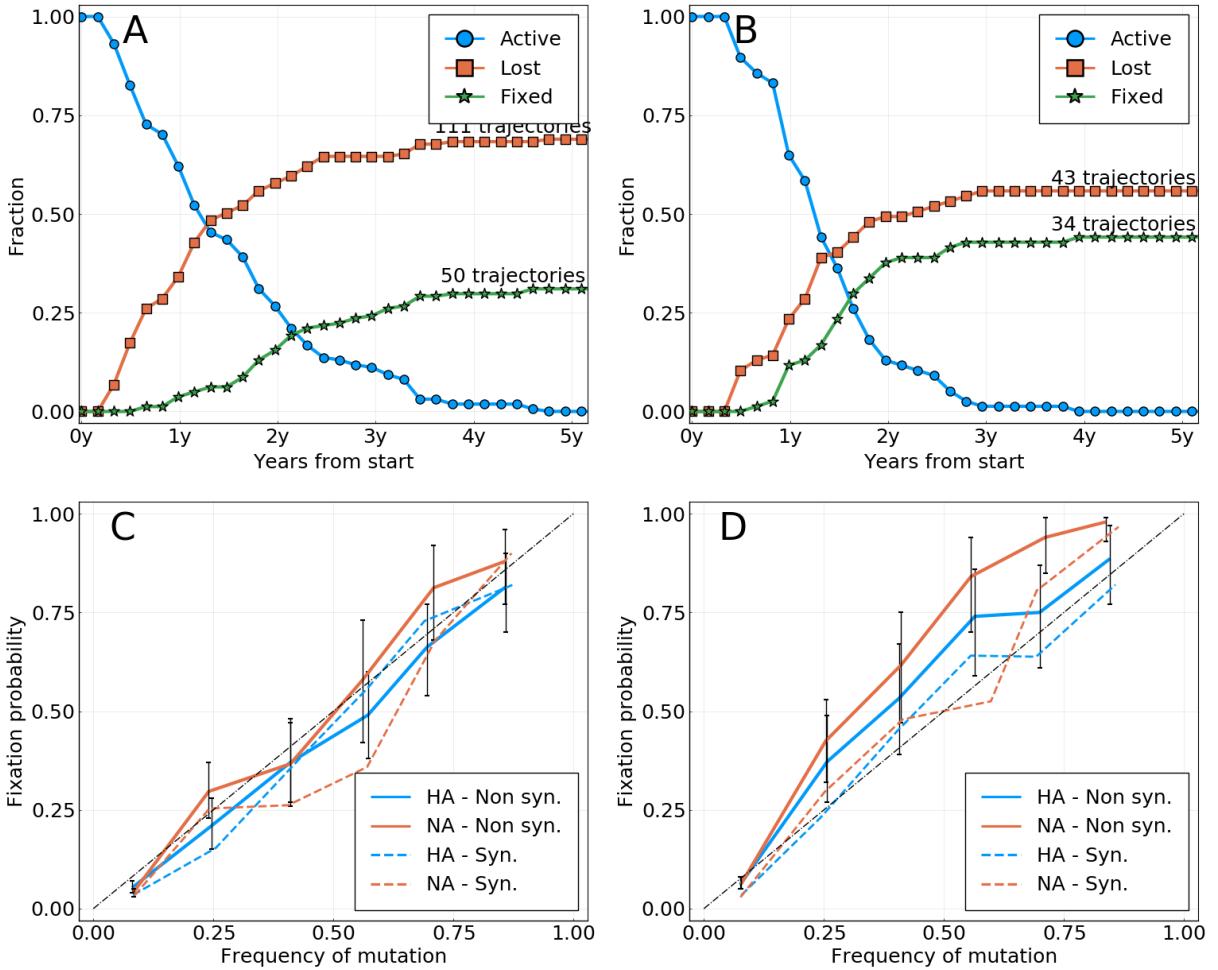


FIG. 2 **A:** Activity of all rising frequency trajectories seen above 25% frequency for A/H3N2 HA and NA. **B:** Same as **A** for A/H1N1. **C:** Probability of fixation of a mutation (amino acid or synonymous) $P_{fix}(f)$ as a function of the frequency f at which it is measured, for A/H3N2 HA and NA. Only new mutations are considered, *i.e.* mutations that were absent in the past. The diagonal dashed line is the expectation from a neutrally evolving population. Colored dashed lines represent synonymous mutations. Colored solid lines represent amino acid mutations. Error bars represent a 95% confidence interval. **D:** Same as **C** for A/H1N1.

in **D** of figure S15 are relatively large, indicating a lower amount of trajectories.

ones, but the effect is small. These two scores are in fact correlated, with older trajectories representing mutations that are more geographically spread, as can be seen in figure S18 of SM. However, it is important to note that sampling biases and heterogeneity across time and space (see supplementary figures S9 and S10) make answering such specific hypothesis challenging. Frequency of mutations might thus be amplified through different sampling biases, making the connection between geographic spread, seasonality and mutation frequency non-trivial to measure.

Simulations of models of adaptation

The results shown in figures 2 and 3 are difficult to reconcile with the idea that seasonal influenza virus evo-

Since influenza is seasonal in temperate regions, geographic spread and persistence might be predictive of the success of mutations. We quantify geographic spread of a mutation by the entropy of its frequency distribution across regions (see methods) and its persistence by the age of the trajectory by the time it reaches frequency f . Figures S16 and S17 show the fixation probabilities as a function of observed frequency for mutations classified according to these scores. The two scores also allow a quantitatively moderate distinction between mutations: for a given frequency f , mutations found in many regions or those that are older (in the sense that they have taken more time to reach frequency f) tend to fix more often than geographically localized mutations or more recent

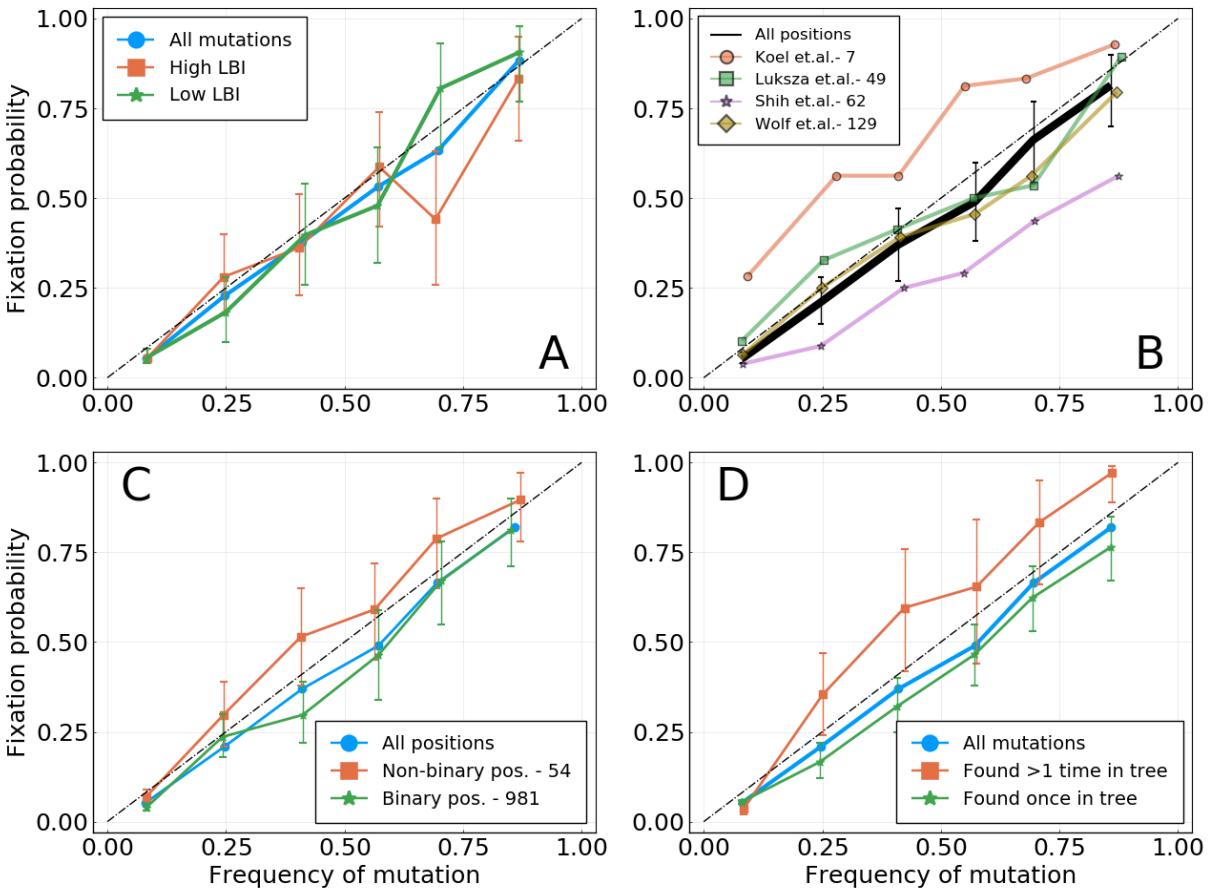


FIG. 3 Fixation probability $P_{fix}(f)$ as a function of frequency, for A/H3N2 influenza. Figure S15 shows the same analysis for A/H1N1. **A:** HA mutations with higher or lower LBI values, based on their position with respect to the median LBI value. **B:** Different lists of epitope positions in the HA protein. The authors and the number of positions is indicated in the legend. **C:** HA and NA mutations for binary positions, *i.e.* positions for which we never see more than two amino acids in the same time bin. **D:** HA and NA mutations that appear once or more than once in the tree for a given time bin.

437 lution is driven by rapid directed positive selection. One⁴⁵⁷
 438 possible explanation for the weakly predictable behaviour⁴⁵⁸
 439 of mutations (beyond their current frequency) might be⁴⁵⁹
 440 tight genetic linkage inside each segment and strong com-⁴⁶⁰
 441 petition between different adaptive mutations (Strelkowa⁴⁶¹
 442 and Lässig, 2012; Neher and Shraiman, 2011). We de-⁴⁶²
 443 sign a simple model of population evolution based on⁴⁶³
 444 the `ffpopsim` simulation software to test this hypothe-⁴⁶⁴
 445 sis (Zanini and Neher, 2012). The model represents a⁴⁶⁴
 446 population of binary genomes of length $L = 200$ evolving⁴⁶⁵
 447 in a fitness landscape that changes through time.⁴⁶⁶

448 First, we use an additive fitness function, with sequence⁴⁶⁸
 449 ($x_1 \dots x_L$) having a fitness $\sum_i h_i x_i$. This implies that for⁴⁶⁹
 450 a given genome position i , the trait $x_i = 1$ is favored if⁴⁷⁰
 451 $h_i > 0$ whereas $x_i = -1$ is favored if $h_i < 0$. All h_i 's⁴⁷¹
 452 have the same magnitude, and only their signs matter⁴⁷²
 453 Every Δt generations, we randomly choose a position i_{473}
 454 and flip the sign of h_i , effectively changing the fitness⁴⁷⁴
 455 landscape. Individuals in the population now have the⁴⁷⁵
 456 opportunity to make an adaptive mutation at site i giving⁴⁷⁶

them a fitness advantage $2|h_i|$. A “flip” at position i of the fitness landscape will decrease fitness of all individuals that carried the adapted variant at position i and increases the fitness of those that happened to carry a deleterious variant.

To increase competition between genomes, we designed a second model that includes epistasis. Once again, the baseline fitness of a genome is an additive function, this time with values of h_i that do not change through time. In addition, we added a component that mimics immune selection. Every Δt generation, we now introduce “antibodies” that target a specific sub-sequence of length $l = 5$, noted $(x_{i_1}^{ab}, \dots, x_{i_l}^{ab})$. The positions $(i_1 \dots i_l)$ are chosen at random, while the targeted sub-sequence is the dominant state at each position. Genomes that include the *exact* sub-sequence targeted by the antibody suffer a strong fitness penalty. However, a single mutation away from that sub-sequence removes this penalty completely, resulting in a fitness landscape with very strong epistasis. This has the effect of triggering a strong competition

477 between adaptive mutations: for a given antibody, $l = 5_{530}$
 478 possible mutations are now adaptive, but combinations 5_{531}
 479 of these mutations do not bring any fitness advantage. 5_{532}

480 Having simulated populations in these two fitness land 5_{533}
 481 scapes, we perform the same analysis of frequency trajec 5_{534}
 482 tories as for the real influenza data. Figure S20 of the SM 5_{535}
 483 shows the $P_{fix}(f)$ as a function of f for the two models 5_{536}
 484 and for different values of the inverse rate of change Δt 5_{537}
 485 of the fitness landscape. For all models, this curve devi 5_{538}
 486 ates significantly from the diagonal. This is most evident 5_{539}
 487 for the case of a simple additive fitness landscape that 5_{540}
 488 changes rarely $\Delta t = 1000$: rising mutations almost always 5_{541}
 489 fix in the population, with $P_{fix}(f) \simeq 1$ for any f larger 5_{542}
 490 than a few percent. This is corroborated by visual inspec 5_{543}
 491 tion of the trajectories, which shows that evolution in this 5_{544}
 492 regime is driven by regular selective sweeps that take a 5_{545}
 493 typical time of ~ 400 generations. In other regimes, with 5_{546}
 494 smaller Δt or with strong epistatic competition, $P_{fix}(f)$ 5_{547}
 495 is reduced and closer to the diagonal. However, it takes 5_{548}
 496 an extremely fast changing fitness landscape to push P_{fix} 5_{549}
 497 close to the diagonal: with $\Delta t = 10$, that is about 40 5_{550}
 498 changes to the fitness landscape in the time it would take 5_{551}
 499 a selective sweep to go from 0% to fixation, $P_{fix}(f)$ differs 5_{552}
 500 from f in a way that is comparable to what is observed 5_{553}
 501 in A/H1N1pdm influenza. 5_{554}

502 These models are not meant to be accurate models of in 5_{555}
 503 influenza viruses evolution. But figure S20 does show is that 5_{556}
 504 the patterns observed in influenza virus evolution are only 5_{557}
 505 reproduced by models of adapting populations when push 5_{558}
 506 clonal competition to extreme values. We conclude 5_{559}
 507 that the pattern in figure 2 may not be a straightforward 5_{560}
 508 manifestation of genetic linkage and clonal interference 5_{561}
 509 but that some more intricate interplay of epidemiology 5_{562}
 510 seasonality, human immunity and chance gives rise to 5_{563}
 511 the weakly predictable yet strongly selected evolutionary 5_{564}
 512 dynamics of IAVs. 5_{565}

513 Why do predictions work?

514 The statistics of frequency trajectories seem to be in 5_{570}
 515 conflict with the notion that influenza evolution is pre 5_{571}
 516 dictable. Likewise, the LBI, a quantity that correlates 5_{572}
 517 with fitness in mathematical models and is used to predict 5_{573}
 518 future influenza populations (Neher et al., 2014), does not 5_{574}
 519 seem to contain any information on whether a specific 5_{575}
 520 mutation is going to fix or not, see figure 3. To resolve 5_{576}
 521 this conundrum, we first note that the criterion by which 5_{577}
 522 predictive power for influenza was measured in (Neher 5_{578}
 523 et al., 2014) was the distance between the strain with the 5_{579}
 524 highest LBI and the future population, not the ability of 5_{580}
 525 the LBI to predict dynamics. The distance was compared 5_{581}
 526 to the average distance between the present and future 5_{582}
 527 population, as well as the post-hoc optimal representative 5_{583}
 528 and the future. 5_{584}

529 To quantify the ability of the LBI and other measures 5_{585}

to pick good representatives of the future, we construct a large tree of HA sequences with 100 sequences in non-overlapping time bins of 4 months from year 2003 to 2019 (a total of 4402 as some 4 month intervals contain less than 100 sequences). Each time bin is considered as a snapshot of the A/H3N2 influenza population and we will refer to sequences in time bin t as the population of the *present*. From this present population, we predict *future* populations in time bin $t + \Delta t$, using only sequences in time bin t and before.

To assess the ability of the LBI to pick a close representative of the future, we compute the LBI of each node of one time bin in the tree using only the leaves that belong to that time bin. The top panel in figure 4 shows the hamming distance of the strain with the highest LBI to future populations at different Δt along with the same distance for a randomly chosen strain. The figure shows the distance averaged over all possible values of t for Δt between 0 and 32 months, giving us an average efficiency of a predictor over 16 years of influenza evolution.

The strain with the highest LBI is consistently closer to the future than the average strain by about 1-2 amino acids, while the overall distance increases linearly due to the continuous evolution of the population. We hence reproduce previous results showing that the LBI picks closer than average representatives (Neher et al., 2014). To investigate whether this apparent success is due to the ability of the LBI to predict fitness or not, we explored a different predictor: the amino acid consensus sequence of the present population (see Methods for a definition of the consensus sequence). The choice is motivated by the fact that it can be shown to be the best possible long term predictor for a neutrally evolving population in terms of Hamming distance (see SM section .1). Figure 4 shows that the consensus sequence is in fact a equally good or even slightly better representative of the future than the sequence with highest LBI (note that the consensus sequence does *not* necessarily exist in the population).

This near equivalence of the consensus and the strain with highest LBI can be explained as follows: The LBI tends to be high for nodes in a tree that are close to the root of a dense and large clade. A typical sample of influenza HA sequences fall into a small number of recognizable clades, and the strains with maximal LBI will often be close to the root of the largest of those clades. This root of the largest clade will often be close to the consensus of the whole population, explaining the similar distance patterns. To test that hypothesis, we measure the hamming distance from the sequence of the top LBI strain to the consensus sequence for populations of all time bins. Panel **B** of figure 4 shows these distances, scaled with respect to an average strain (details in caption). It clearly shows that the top-LBI strain and the consensus sequence are indeed quite similar: out of 48 time bins, only once is the sequence of the top-LBI strain farther away from the consensus than the average sequence is. Moreover,

586 the sequence of the top-LBI strain *exactly* matches the⁶³⁹
 587 consensus in 19 cases.⁶⁴⁰

588 DISCUSSION

589 Predicting the trajectory of a mutation requires (i)⁶⁴⁵
 590 significant fitness difference between genomes carrying⁶⁴⁶
 591 different variants at the site and (ii) a selection pressure⁶⁴⁷
 592 that changes slowly over time. Under such conditions, it is⁶⁴⁸
 593 expected that frequency trajectories will show a persistent⁶⁴⁹
 594 behavior which would make them predictable for some⁶⁵⁰
 595 time. However, we could find only limited evidence for⁶⁵¹
 596 such persistent behavior in the past 19 years of IAV⁶⁵²
 597 evolution. This lead us to conclude that (i) influenza⁶⁵³
 598 virus evolution is qualitatively different from models of⁶⁵⁴
 599 rapidly adapting population (despite clear evidence for⁶⁵⁵
 600 frequent positive selection), and (ii) previous methods to⁶⁵⁶
 601 predict influenza evolution work primarily because they⁶⁵⁷
 602 pick strains that represent the future well, not because⁶⁵⁸
 603 they predict future dynamics.⁶⁵⁹

604 The primary focus in this work was the investigation⁶⁶⁰
 605 of frequency trajectories of new amino acid mutations. In⁶⁶¹
 606 the short term, we found that on average the direction of⁶⁶²
 607 trajectory does not persist for longer than a few months⁶⁶³
 608 Indeed, the average trajectory in figure 1 takes a sharp⁶⁶⁴
 609 turn when going from $t < 0$ to $t > 0$, instead of showing⁶⁶⁵
 610 “inertia”. This suggests that selective sweeps are not⁶⁶⁶
 611 representative of typical trajectories. **Pierre: We could⁶⁶⁷**
612 cite the Illingworth paper here. They do show that as⁶⁶⁸
613 well, in a different way. The fact many trajectories are not⁶⁶⁹
614 sweeps is expected when there is strong linkage. What⁶⁷⁰
615 cannot be easily explained by linkage only is the complete⁶⁷¹
616 lack of inertia on average (at least I think?), and the⁶⁷²
617 $P_{fix}(f) = f$.⁶⁷³

618 On a longer timescale, we investigated the probability⁶⁷⁴
 619 that a novel mutation observed at frequency f fixes. In⁶⁷⁵
 620 neutral models of evolution this probability equals f , while⁶⁷⁶
 621 it should be higher or lower than f for mutations with⁶⁷⁷
 622 a beneficial or deleterious effect on fitness, respectively⁶⁷⁸
 623 However, in the case of influenza, this probability differs⁶⁷⁹
 624 little from f , making current frequency the best predictor⁶⁸⁰
 625 for fixation. In figure 3, we split trajectories into groups⁶⁸¹
 626 for which we expected P_{fix} to deviate from f . Many of⁶⁸²
 627 these splits, such as high/low LBI or epitope/non-epitope⁶⁸³
 628 positions, did not result in an increased predictability,⁶⁸⁴
 629 while others gave limited information on fixation. Despite⁶⁸⁵
 630 the lack of predictability of mutation frequency trajec-to⁶⁸⁶
 631 ries, influenza surface proteins show strong signatures of⁶⁸⁷
 632 selection (Bhatt et al., 2011; Strelkowa and Lässig, 2012)⁶⁸⁸

633 Methods for predicting the future evolution of influenza⁶⁸⁹
 634 either construct explicit fitness models (Luksza and Lässig,⁶⁹⁰
 635 2014; Huddleston et al., 2020), use historical patterns of⁶⁹¹
 636 evolution (Luksza and Lässig, 2014; Bush et al., 1999),⁶⁹²
 637 phenotypic assays (Neher et al., 2016; Steinbrück and⁶⁹³
 638 McHardy, 2012), or dynamic or phylogenetic patterns⁶⁹⁴

(Neher et al., 2014; Klingen et al., 2018b). The goal of these methods is to pick strains that are good representatives of future populations and could serve as vaccine candidates (Morris et al., 2018).

The low power to predict frequency dynamics or fixation naturally triggers the question why the above methods have been found to work. Picking representatives of the future and predicting frequency dynamics are distinct objectives and success at the former (as compared to random picks) is not necessarily inconsistent with a lack of predictable dynamics. In fact, (Huddleston et al., 2020) reports that the rate at which the frequency of a strain changes is often a poor predictor – consistent with our observations here. But despite the fact that future frequencies are not predicted by the LBI, the strain with the highest LBI in the population is a better predictor of the future population than a randomly picked one. While the LBI was shown to be a correlate of relative fitness and be predictive of fixation in mathematical models of evolution (Neher et al., 2014), it does not seem to be predictive of influenza evolution because it measures fitness from genealogical structure. Instead, we believe it picks closer than average strains simply because it has the tendency to be maximal at the base of large and dense clades. These basal genotypes are closer to the future populations than the current tips of the tree and hence a better predictor on average. The consensus sequence of all present strains performs slightly but consistently better than picking the strain with the highest LBI. The consensus sequence is the best possible predictor for a neutrally evolving population, and does not attempt to model fitness in any way.

At the same time, influenza virus phylogenies show clear deviations from those expected from the neutral Kingman coalescent, similar to those expected under Bolthausen-Sznitman coalescent (BSC) processes that are generated by traveling wave models of rapid evolution (Neher and Hallatschek, 2013; Desai et al., 2013). The correspondence between the BSC and traveling wave models comes from transient exponential amplification of fit strains before these fitness differences are wiped out by further mutation. This exponential amplification generates long-tailed effective offspring distributions which in turn can lead to genealogies described by the BSC (Neher and Hallatschek, 2013; Schweinsberg, 2003). Many processes other than selection, including seasonality and spatio-temporal heterogeneity, can generate effective long tailed offspring distributions even in absence of bona-fide fitness differences, which might explain ladder-like non-Kingman phylogenetic trees.

A recent preprint proposed that influenza virus evolution is primarily limited by an asynchrony between population level selection and generation of new variants within infected hosts (Morris et al., 2020). Along these lines, it is possible that the A/H3N2 population readily responds once population level selection is high enough by

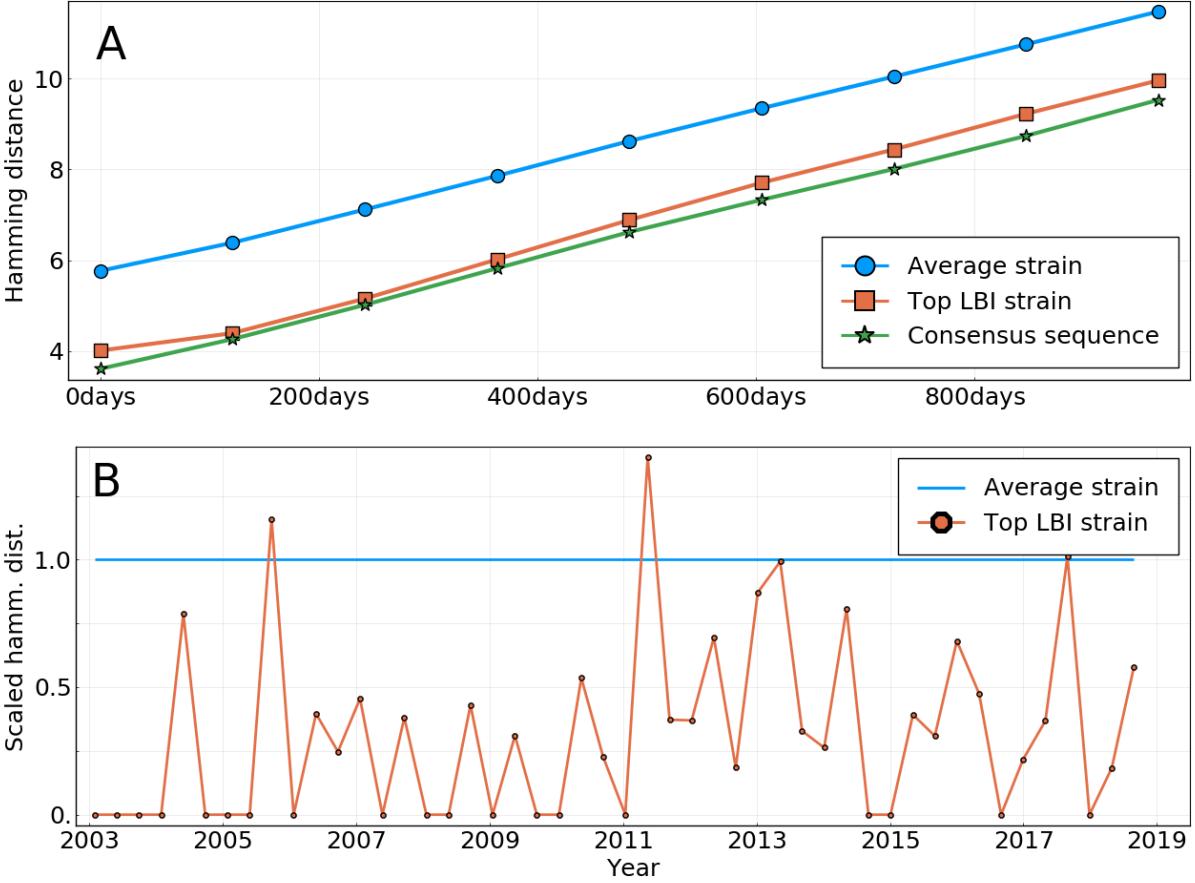


FIG. 4 **A:** Average Hamming distance of the sequences of different predictors to HA sequences of future influenza populations, themselves averaged over all “present” populations from year 2003 to 2019. Predictors are: a randomly picked sequence in the present population; the sequence of the strain with the highest LBI in the present population; the consensus sequence of the present population. **B:** Scaled Hamming distance between the sequence of the top LBI strain and the consensus sequence for populations at different dates. The scaling is such that for each date, the Hamming distance between a strain from the population and the consensus is on average 1. The strain with the highest LBI is almost always closer to the consensus sequence than the average strain.

giving rise to essentially equivalent variants. Furthermore, selection might cause the rapid rise of a novel variant to macroscopic frequencies (observable in a global sample) but its benefit rapidly “expires” because competing variants catch up and/or it mediates immune escape only to a small fraction of the population. These considerations might explain the disconnect between models of rapid adaptation and the frequency dynamics observed in influenza virus populations.

METHODS

Data and code availability

The sequences used are obtained from the GISAID database (Shu and McCauley, 2017). Strain names and accession numbers are given as tables in two supplementary files. Outliers strains listed at <https://github.com/PierreBarrat/FluPredictability>.

[com/PierreBarrat/FluPredictability/src/config](https://github.com/PierreBarrat/FluPredictability/src/config) were removed.

The code used to generate the figures presented here is available at <https://github.com/PierreBarrat/FluPredictability>.

Frequency trajectories

For a set of sequences in a given time bin, we compute frequencies of amino acids at each position by simple counting. We make the choice of not applying any smoothing method in an attempt to be as close to the data and “model-less” as possible. This is especially important for the short term prediction of frequency trajectories, as estimations of the “persistence time” of a trajectory might be biased by a smoothing method.

We compute frequency trajectories based on the frequencies of amino acids. A trajectory begins at time t if an

726 amino acid is seen under the lower frequency threshold⁷⁷⁶
 727 of 5% (resp. above the higher threshold of 95%) for the
 728 two time bins preceding t , and above this lower threshold
 729 (resp. below the higher threshold) for time bin t . It ends
 730 in the reciprocal situation, that is when the frequency
 731 is measured below the lower threshold (resp. above the
 732 higher threshold) for two time bins in a row.

733 In order to avoid estimates of frequencies that are too
 734 noisy, we only keep trajectories that are based on a pop-
 735 ulation of at least 10 sequences for *each* time bin. As
 736 said in the Results section, we also restrict the analysis to
 737 trajectories that begin at a 0 frequency, in part to avoid
 738 double counting. We find a total of 460 such trajectories.
 739 However, only 106 reach a frequency of 20%, on which
 740 figure 2 is based for instance.

741 Note that the fact that we use samples of relatively small
 742 sizes – at least for some time bins – leads to biases in
 743 the estimation of frequencies. We show in Supplementary
 744 Material that these biases are generally small and do not
 745 induce any qualitative changes to results presented here.⁷⁸⁰

746 Local Branching Index

747 LBI was introduced in (Neher et al., 2014) as an approx-⁷⁸⁵
 748 imation of fitness in populations evolving under persistent
 749 selective pressure that is fully based on a phylogenetic tree.
 750 It relies on the intuition that the tree below high-fitness
 751 individuals will show dense branching events, whereas
 752 absence of branching is a sign of low-fitness individuals.
 753 Quantitatively, the LBI $\lambda_i(\tau)$ of a node i is the integral
 754 of all of the tree’s branch length around i , with an expo-
 755 nentially decreasing weight $e^{-t/\tau}$ with t being the branch
 756 length. When considering a time binned population, the
 757 LBI is computed once for each time bin by considering
 758 only the leaves of the tree that belong to the time bin.
 759 This means that only branches that ultimately lead to
 760 a leaf that belongs to the time bin are considered in the
 761 integration.

762 τ is the time scale for which the tree is informative of
 763 the fitness of a particular node. Here, we use a value of
 764 τ equal to a tenth of $T_C \simeq 6$ years, the coalescence time
 765 for influenza A/H3N2 strains, converted to units of tree
 766 branch length through the average nucleotide substitution⁷⁸⁶
 767 rate ($\simeq 4 \cdot 10^{-3}$ substitutions per site per year for HA),⁷⁸⁷
 768 We have observed that given our method to predict the⁷⁸⁸
 769 future from present populations corresponding to time
 770 bins of 4 months, changing the value of τ has little effect
 771 on the pick of the top LBI strain. By retrospectively⁷⁸⁹
 772 optimizing its value, it is possible to reduce the average
 773 distance to the population 2 years ahead by ~ 0.25 amino
 774 acids on average, making the LBI method almost as good
 775 as the consensus on figure 4.

Measuring the geographical spread of a mutation

For a mutation X we define its regional distribution using the numbers $n_r(X)$ that represent the number of sequences sampled in region r that carry X . Regional weights are then defined as

$$w_r(X) = \frac{n_r(X)}{\sum_r n_r(X)}.$$

We can then measure the geographical spread $G(X)$ of X by using the Shannon entropy of the probability distribution $w_r(X)$:

$$G(X) = \sum_r w_r(X) \log(w_r(X)).$$

$G(X)$ is a positive quantity that is larger when X is equally present in many regions, and equal to zero when X is concentrated in only one region.

Region used are the ones defined in the **Nextstrain** tool (Hadfield et al., 2018). Those are North America, South America, Europe, China, Oceania, Southeast Asia, Japan & Korea, South Asia, West Asia, and Africa.

784 Assigning a fitness to trajectories

Consensus sequence

Given a set of N sequences $(\sigma^1, \dots, \sigma^N)$ based on an alphabet \mathcal{A} (e.g. \mathcal{A} has 20 elements for amino acids, 4 for nucleotides), we can define a *profile* distribution $p_i(a)$ by the following expression:

$$p_i(a) = \sum_{n=1}^N \delta_{\sigma_i^n, a}$$

where i is a position in the sequence, σ_i^n the character appearing at position i in sequence σ^n , a a character of the alphabet and δ the Kronecker delta. The profile $p_i(a)$ simply represents the fraction of sequences which have character a at position i .

We then simply define the consensus sequence σ^{cons} such that

$$\sigma_i^{cons} = \operatorname{argmax}_a p_i(a).$$

In other words, the consensus sequence is the one that has the dominant character of the initial set of sequences at each position.

Earth Mover’s Distance

In order to measure the distance of several predictor sequences to the future population, we rely on the *Earth Mover’s Distance* (EMD), a metric commonly applied

in machine learning to compare collections of pixels or words (Rubner et al., 1998; Kusner et al., 2015). Here, we apply it to compute the distance between the sequences of two populations, noted as $\mathcal{X} = \{(x^n, p^n)\}$ and $\mathcal{Y} = \{(y^m, q^m)\}$ with $n \in \{1 \dots N\}$ and $m \in \{1 \dots M\}$. In this notation, x^n and y^m are sequences, and p^n and q^m are the frequencies at which these sequences are found in their respective populations. For convenience, we also define $d_{mn} = H(x^n, y^m)$ as the Hamming distance between pairs of sequences in the two populations.

We now introduce the following functional

$$F(\mathbf{w}) = \sum_{n,m} d_{nm} w_{nm},$$

with $\mathbf{w} = \{w_{nm}\}$ being a matrix of positive weights. The EMD between the two populations \mathcal{X} and \mathcal{Y} is now defined as the minimum value of function F under the conditions

$$\sum_{n=1}^N w_{nm} = q^m, \quad \sum_{m=1}^M w_{nm} = p^n, \text{ and } w_{nm} \geq 0$$

Intuitively, the weight w_{nm} tells us how much of sequence x^n is “moved” to sequence y^m . The functional F sums all of these moves and attributes them a cost equal to the Hamming distance d_{nm} . The conditions on weights in \mathbf{w} ensure that all the weight p^n of x^n is “moved” to elements in \mathcal{Y} and vice versa.

The minimization is easily performed by standard linear optimization libraries. Here, we use the Julia library JuMP (Dunning et al., 2017).

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1007 **SUPPLEMENTARY MATERIAL**1008 **1. Consensus sequence as a predictor for neutrally evolving populations**

1009 We consider the case of a neutrally evolving and structure-less population, such as the one in the Wright-Fisher
 1010 model of evolution (Sigwart, 2005). At an initial time $t = 0$, the population consists of N individuals with genomes
 1011 $(\sigma^1 \dots \sigma^N)$ of length L (not necessarily distinct).

1012 We make two hypotheses about this population. We first suppose that *no* mutations occur during the evolution of this
 1013 population. This may seem surprising and is of course not true in the case of influenza. This assumption is however in
 1014 line with the fact that the object of this work is to predict the outcome of *already existing* mutations in the influenza
 1015 population. The prediction of mutations that we have not yet seen is not in its scope. Thus, assuming that no new
 1016 mutations take place can be seen as a simple way to model the fact that we have no information about such events.
 1017 The second assumption is that the population evolves in a completely neutral way, meaning that the average number
 1018 of descendants of each genome σ^n is the same. Let us now consider the population after it has evolved for a long time
 1019 $t \gg T$ where T is the typical coalescence time (for the Wright-Fisher model, $T = 2N$). At this point, all individuals in
 1020 the future population will descend from a unique individual n_0 in the $t = 0$ population. Our two hypotheses now allow
 1021 us to make two statements. First, since no new mutations are allowed, the population at $t \gg T$ will be clonal, with all
 1022 individuals having genome σ^{n_0} . Second, since the evolution is neutral and does not favour any genome in particular,
 1023 the probability that σ^{n_0} is equal to a given genome σ is $1/N$. In other words, the probability that a genome at $t = 0$
 1024 ultimately becomes the ancestor of all the future population is equal to its frequency in the $t = 0$ population.

1025

1026 We now try to find the genome σ that best predicts the future population on the long run, that is for $t \gg T$. Here,
 1027 we take best to mean that the predictor minimizes $H(\sigma, \sigma^{n_0})$ where H is the Hamming distance defined by

$$H(\sigma^a, \sigma^b) = \sum_{i=1}^L (1 - \delta_{\sigma_i^a, \sigma_i^b}), \quad (1)$$

1028 with σ_i being the character appearing at position i of genome σ and δ the Kronecker delta. Since we do not know n_0 ,
 1029 we have to average over all its possible values. σ must thus minimize the following quantity:

$$\begin{aligned} \langle H(\sigma, \sigma^{n_0}) \rangle_{n_0} &= \sum_{n=1}^N H(\sigma, \sigma^n) \\ &= \sum_{i=1}^L \sum_{n=1}^N (1 - \delta_{\sigma_i, \sigma_i^n}) \end{aligned} \quad (2)$$

by using the definition of the Hamming distance. We now assume that characters at each positions of the genomes can
 be indexed by an integer a running from 1 to q . For instance, if these were amino acid sequences, we could index the
 20 amino acids by a running from 1 to $q = 20$. We rewrite the Kronecker delta in the previous expression using this
 indexation:

$$\delta_{\sigma_i, \sigma_i^n} = \sum_{a=1}^q \delta_{\sigma_i, a} \delta_{\sigma_i^n, a}.$$

1030 We also introduce the *profile* frequencies $p_i(a)$ of the population at time $t = 0$:

$$p_i(a) = \sum_{n=1}^N \delta_{\sigma_i^n, a}. \quad (3)$$

1031 $p_i(a)$ represents the frequency at which character a appears at position i in genomes of the initial population.

1032 Equation 2 now becomes

$$\begin{aligned}
 \langle H(\sigma, \sigma^{n_0}) \rangle_{n_0} &= \sum_{i=1}^L \sum_{n=1}^N \left(1 - \sum_{a=1}^q \delta_{\sigma_i, a} \delta_{\sigma_i^n, a} \right) \\
 &= \sum_{i=1}^N \left(1 - \sum_{a=1}^q \delta_{\sigma_i, a} p_i(a) \right) \\
 &= \sum_{i=1}^L (1 - p_i(\sigma_i))
 \end{aligned} \tag{4}$$

1033 This means that the genome $\sigma = (\sigma_1 \dots \sigma_L)$ which best predicts the future population according to our definition is
 1034 the one that minimizes the quantity $(1 - p_i(\sigma_i))$ for all positions i . This obviously implies that each σ_i must be chosen
 1035 as to maximize $p_i(a)$, that is σ_i must be the character that appears the most frequently at position i . Thus, σ must be
 1036 the *consensus* sequence of the initial population.

1037 2. Predictor based on the local LBI maxima

In figure 19, we use several sequences as a predictor of the future population. Distance between two sets of sequences, *i.e.* the predictor sequences and the ones of the future population, is defined as the Earth Mover's Distance (EMD). Here, we show that for a population evolving under the same hypotheses as in section .1, the best *multiple* sequence long term predictor is again the consensus sequence with weight 1.

Let the predictor be a set of weighted sequences $\{(s^\alpha, q_\alpha)\}$. We again use the fact that in the long term, a unique sequence σ^{n_0} from the present will be the ancestor of the entire population. We want to compute the EMD from the predictor to σ^{n_0} , that is the EMD between the sets $\mathcal{X} = \{(s^\alpha, q_\alpha)\}$ and $\mathcal{Y} = \{\sigma^{n_0}, 1\}$. Applying the definition of the Methods section, it follows that the weights \mathbf{w} are in this case equal to the q_α s. By averaging over all values of n_0 , we now obtain

$$\langle \text{EMD}(\{(s^\alpha, q_\alpha)\}) \rangle_{n_0} = \sum_{n=1}^N \sum_{\alpha} H(s^\alpha, \sigma^n) \cdot q_\alpha.$$

By the same calculation procedure as in the previous section, this expression simplifies to

$$\langle \text{EMD}(\{(s^\alpha, q_\alpha)\}) \rangle_{n_0} = \sum_{i=1}^L \left(1 - \sum_{a=1}^q p_i(a) q_i(a) \right),$$

where the profile of the present population $p_i(a)$ has already been defined, and $q_i(a)$ stands for the profile of the predictor, that is

$$q_i(a) = \sum_{\alpha} \delta_{s_i^\alpha, a} q_\alpha.$$

1038 To minimize this distance, we find a profile $q_i(a)$ that maximizes the quantity $\sum_{\alpha} \delta_{s_i^\alpha, a} q_\alpha$ for each position i . It is
 1039 clear that this is done by assigning a value $q_i(a) = 1$ if a maximizes $p_i(a)$, and $q_i(a) = 0$ otherwise. Thus, the profile of
 1040 the predictor must be that of the consensus sequence, which is only possible if the predictor becomes $\{\sigma^{cons}, 1\}$.

1041 3. Correcting for nested trajectories

1042 The analysis of the main text computes probabilities of fixation assuming that all trajectories are independent.
 1043 However, it is well-known that mutations in influenza viruses are nested: they appear on backgrounds that already
 1044 carry other mutations. Since mutations appearing on the same genomes will jointly fix or disappear, many frequency
 1045 trajectories are not independent but correlated. In order to compensate for potential biases due to this effect, we
 1046 attempted to cluster trajectories based on similarity in their strain composition. Our aim is that two trajectories
 1047 corresponding to mutations appearing mostly on the same genomes will be grouped in the same cluster. We then
 1048 conduct the same analysis as in the main text on a set of *effectively independent* trajectories constructed by taking one
 1049 trajectory from each cluster.

In order to perform clustering, we define a distance between trajectories. A frequency trajectory X is characterized by a series of frequency values and a time interval T . $f(t)$ for $t \in T$ corresponds to the frequency at which a given mutation x appears in the population at date t , and we define $S(t)$ as the strains that carry mutation x at date t . With this notation, $f(t)$ is the ratio of the number of elements in $S(t)$ to the total number of strains at date t . Let us now consider two frequency trajectories X_1 and X_2 .

We define the distance $d(X_1, X_2)$ between these two trajectories based on the average similarity of the strains S_1 and S_2 that compose them:

$$d(X_1, X_2) = \frac{1}{|T_1 \cap T_2|} \sum_{t \in T_1 \cap T_2} \frac{|S_1(t) \cap S_2(t)|}{|S_1(t) \cup S_2(t)|},$$

where $T_1 \cap T_2$ is the time interval where both trajectories are active, and $|\cdot|$ denotes the number of elements of a set. The quantity summed corresponds to the Jaccard index between strains composing X_1 and X_2 at a given date. It is 1 if the two trajectories share exactly the same strains for this date, and 0 if they share no strain at all. This leads to the two following properties of d :

- if $d(X_1, X_2) = 0$, then X_1 and X_2 represent the same frequency trajectory. The mutations x_1 and x_2 that they correspond to always appear on the same strains and are totally linked.
- if $d(X_1, X_2) = 1$, then X_1 and X_2 can be considered completely independent. This can be the case if the two trajectories do not occur at the same dates, *i.e.* $|T_1 \cap T_2| = 0$, or if their respective mutations are never present on the same genomes.

we attempt to reduce the potential statistical bias due to the nesting of trajectories by grouping them based on the above defined distance. Given a set of trajectories $\{X\}$, we perform a decomposition of $\{X\}$ into disjoint clusters $C_1(d^*) \cup \dots \cup C_n(d^*) = \{X\}$ where d^* is an arbitrary threshold distance. Clusters are built in such a way that given two trajectories X_i and X_j

$$d(X_i, X_j) \leq d^* \Rightarrow \exists k : X_i, X_j \in C_k$$

and

$$X_i \in C_k \Rightarrow \exists j \in C_k : d(X_i, X_j) \leq d^*.$$

These condition imply that the clusters formed are the minimal ones that guarantee that any two trajectories closer than the threshold distance d^* belong to the same cluster. The number of clusters n depends on the chosen value for d^* .

We compute clusters for different values of d^* for the case of the HA gene in A/H3N2. The top panel of figure S1 shows the number cluster n as a function of d^* . In the $d^* = 0$ case, only trajectories that are exactly identical in terms of strain composition are clustered together. In this case, our clustering amounts to counting mutations that appear on exactly the same strains as one, reducing the number of effective trajectories from 800 to slightly less than 700. For higher values of d^* , the number of clusters steadily goes down until it reaches 1 for $d^* = 1$, which is the maximum value of the distance $d(X_1, X_2)$. The sharp drop in n for $d^* = 0.5$ is explained by the high number of very short (typically one time point) and low frequency trajectories that share one out of two strains.

Since the choice of d^* is arbitrary and since no particular value can be chosen based on the number of clusters $n(d^*)$, we decide to test our clustering strategy for five values, namely $d^* \in \{0, 0.05, 0.1, 0.2, 0.49\}$. The bottom panel of figure S1 shows examples of a cluster for the four non-zero values of d^* . The cluster displayed in each case is the one containing the mutation HA1:33R. As d^* increases, more and more unlike trajectories are grouped together. In the case $d^* = 0.49$, the cluster consists of 13 trajectories, 5 of which end up dying while the rest fix. Since such a high value of d^* results in grouping trajectories that do not have the same fate (fixation or death), we decide to exclude it from the rest of the analysis, resulting in four remaining values $d^* \in \{0, 0.05, 0.1, 0.2\}$.

Once clustering is performed, we re-conduct the analysis of the main text on a set of effective trajectories. This set is constructed by taking one trajectory at random from each cluster. Effective trajectories are then considered independent from each other. The left panel of figure S2 shows the fixation probability of trajectories as a function of their frequency for different values of d^* , for the HA gene of A/H3N2. The result obtained in panel A of figure 2 of the main text is also showed as a reference. For the three lower values of d^* , results do not differ from the one obtained in the main text, even though the number of trajectories in each frequency bin has dropped as can be seen in the right

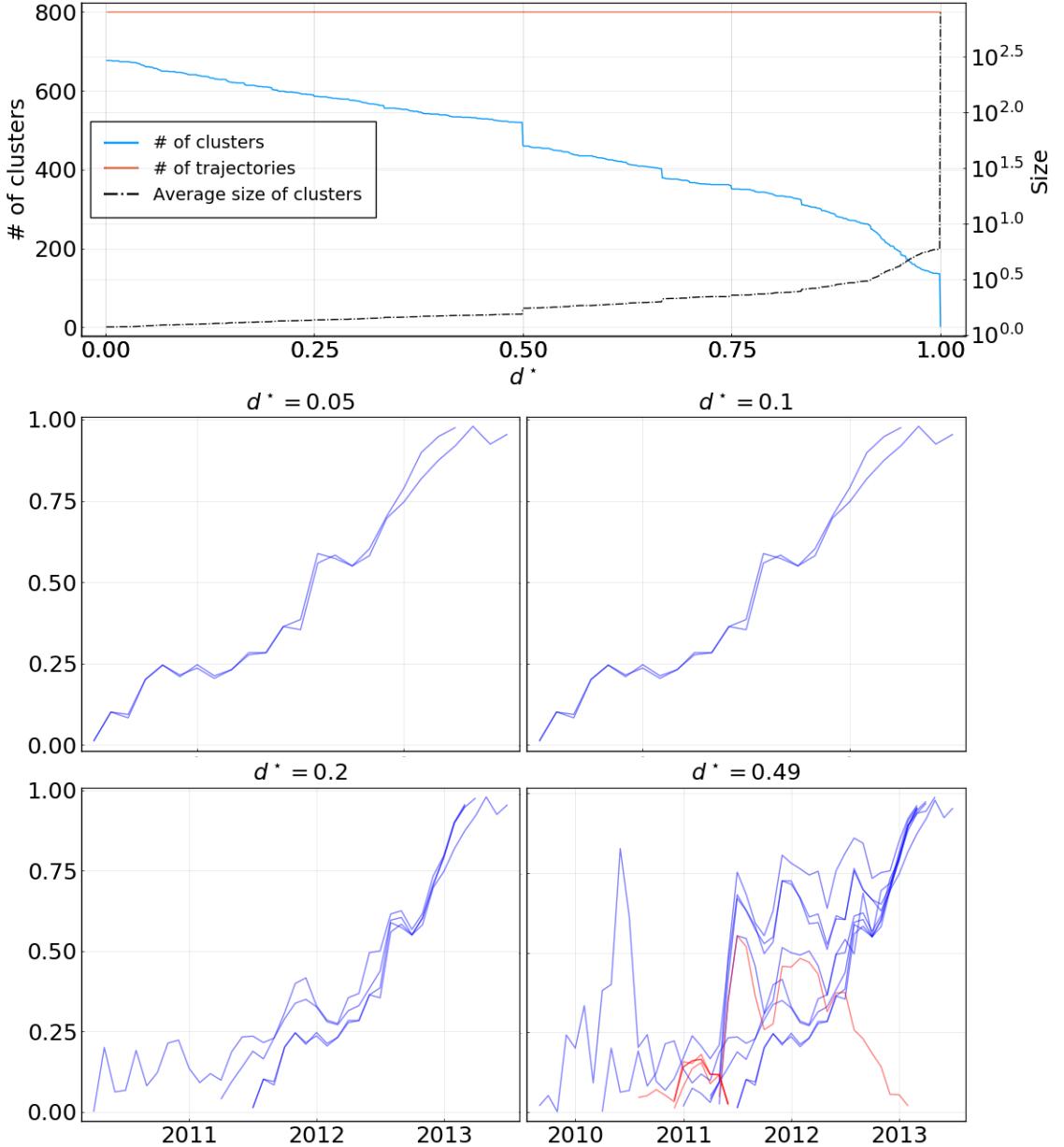


Figure S 1 **Top:** Left-axis: number n of clusters as a function of the threshold distance d^* . The total number of trajectories is shown as a flat orange line. Right-axis: average size of clusters as a function of d^* . **Bottom:** Examples of clusters for four values of d^* . The four clusters displayed are the ones to which mutation HA1:33R belongs to.

panel of figure S2. This indicates that grouping together trajectories that share most of their strains, and are thus very correlated, does not modify the computed fixation probability in any way. For the higher value $d^* = 0.2$, fixation probability drops slightly across all frequency bins, suggesting that fixating trajectories tend to be grouped together more frequently. However, this drop remains of limited amplitude.

Overall, this analysis leads us to think that even though mutations in influenza may be nested, considering trajectories as independent does not result in strong statistical biases. Indeed, clustering similar trajectories together does significantly modify results presented in the main text.

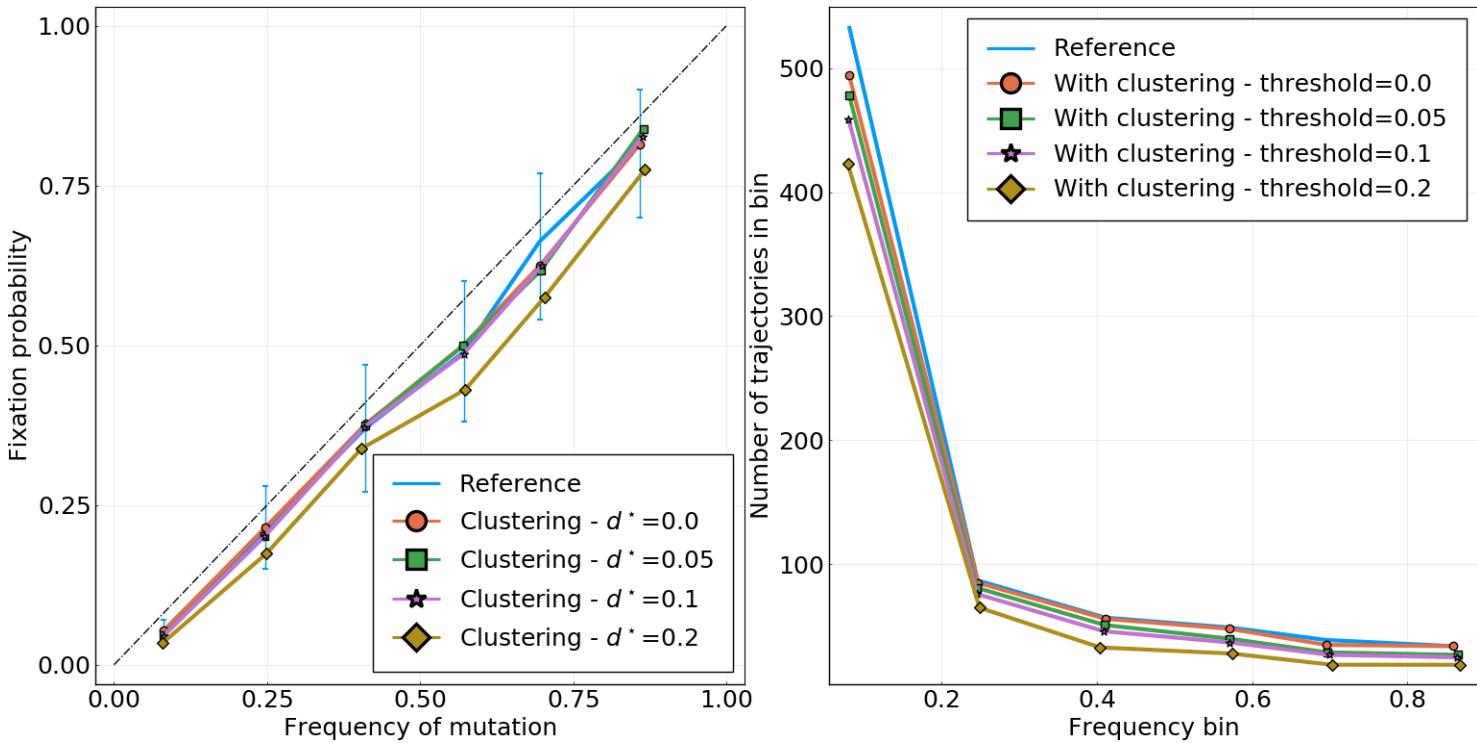


Figure S 2 **Left:** Fixation probability of trajectories as a function of probability for four values of d^* . The reference curve is the same as in panel A of figure 2 for the HA gene. For readability, error bars are only displayed for the reference curve. **Bottom:** Number of trajectories in each frequency bin corresponding to the left panel.

1090 4. Ability of fitness models to predict fixation

1091 In a companion article (Huddleston et al., 2020), authors developed several models to predict the future HA
 1092 population for A/H3N2 influenza. The premise of these models is to assign a fitness score to the HA gene of each
 1093 A/H3N2 strain in a given year, and to then apply deterministic evolutionary equations to obtain the future strain
 1094 composition. The fitness scores are derived from several quantities that have been shown to be connected with viral
 1095 fitness.

1096 Here, we assess the ability of four fitness scores computed in (Huddleston et al., 2020) to predict fixation: mutational
 1097 load and epitope antigenic novelty introduced in (Luksza and Lässig, 2014), hemagglutination inhibition (HI) antigenic
 1098 novelty (Neher et al., 2016) and the previously mentioned Local Branching Index (LBI) (Neher et al., 2014). Figure S3
 1099 shows fixation probability for HA mutations with fitness scores in the top or bottom half of the fitness distribution for
 1100 the four fitness measures. **Pierre: Maybe use delta-frequency instead of epitope mutations? They behave weirdly (high**
1101 fitness is worse), and delta-frequency uses more than the last time point in the trajectory, which may partly answer
1102 one point of Rev1.

1103 Next, we use two fitness scores based on linear combinations of the following quantities: mutational load + LBI and
 1104 mutational load + HI antigenic novelty. These two combinations were found in (Huddleston et al., 2020) to be the
 1105 best predictors of future population. The coefficients of the linear combinations were trained each year using the five
 1106 previous years of data, and here we use their average value over a 25 years training period.

1107 5. Biases in frequency estimations

1108 The frequency of mutations in a given time-bin is simply performed by computing their frequency in sequences
 1109 sampled in that time bin. This leads to potential biases in estimating frequencies, that arise for two reasons:

- 1110 (i) A mutation present at frequency p in the population might be observed at another frequency $f \neq p$ if f is
 1111 estimated using a sub-sample of the population.

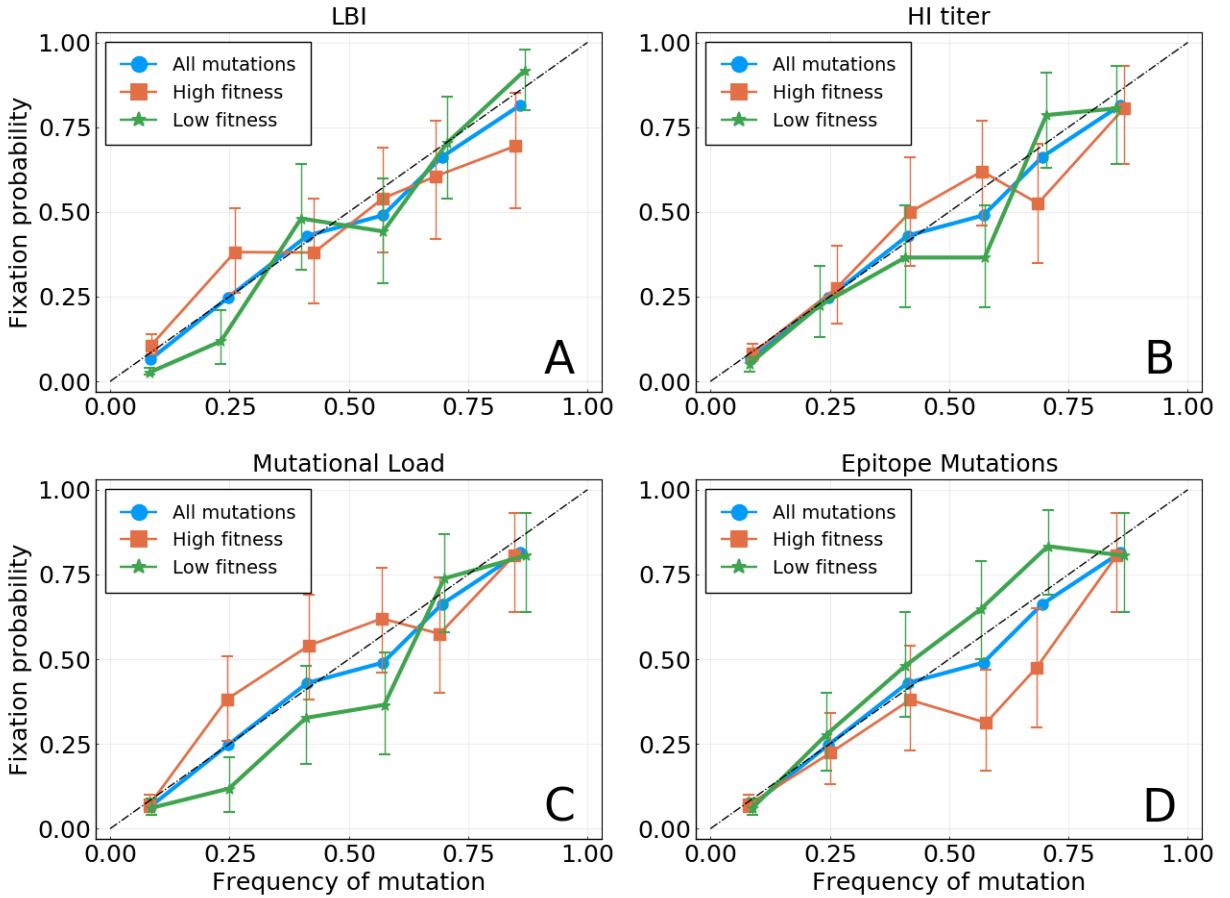


Figure S 3 Ability of fitness scores based on **A** LBI, **B** HI titer, **C** mutational load and **D** epitope antigenic novelty to predict fixation of mutation frequency trajectories in the HA gene of A/H3N2. Similar to figure 3 of the main text.

1112 (ii) For a neutrally evolving population, the distribution of frequencies of alleles is of the form $P(p) \propto 1/p$. This
1113 means that the amount of alleles at frequency p is lower when p is higher.

1114 To illustrate (i), let us compute the probability that a mutation present at “real” frequency p in the population is
1115 found to be in a given frequency bin $[f_1, f_2]$ when p is estimated from a sample of size n . The sample consists of n
1116 observations $\{x_i\}$ with $1 \leq i \leq n$, with $x_i = 1$ if sequence sequence i of the sample bears the mutation, and $x_i = 0$ if
1117 not. If n is small with regard to the total population size, we can consider the x_i as random variables with a binomial
1118 distribution, meaning that $P(x_i = 1) = p$ and $P(x_i = 0) = 1 - p$. The empirical frequency f is then estimated by
1119 taking the average of the x_i variables, that is $f = (x_1 + \dots + x_n)/n$. If those are independently sampled and n is large
1120 enough, the probability of measuring value f is given by the Central Limit Theorem:

$$P_{n,p}(f) \propto e^{(f-p)^2/2\sigma^2}, \text{ where } \sigma^2 = \frac{p(1-p)}{n}. \quad (5)$$

1121 To compute the probability that this mutation is found in a given frequency bin $[f_1, f_2]$, we integrate this distribution:

$$P_{f_1, f_2}(p, n) = \int_{f_1}^{f_2} dx P_{n,p}(x). \quad (6)$$

1122 Function $P_{f_1, f_2}(p, n)$ is shown as a function of p for a fixed interval and for different values of n in the first panel of
1123 figure S5. Note the asymmetry of it: the variance of a binomial distribution of parameter p is small when p is close to
1124 0 or 1, and goes through a maximum at $p = 0.5$. For this reason, mutations present at frequency p close to 0.5 have a
1125 higher probability of being observed in other frequency bins. On the contrary, this is unlikely for very rare or very
1126 frequent mutations.
1127

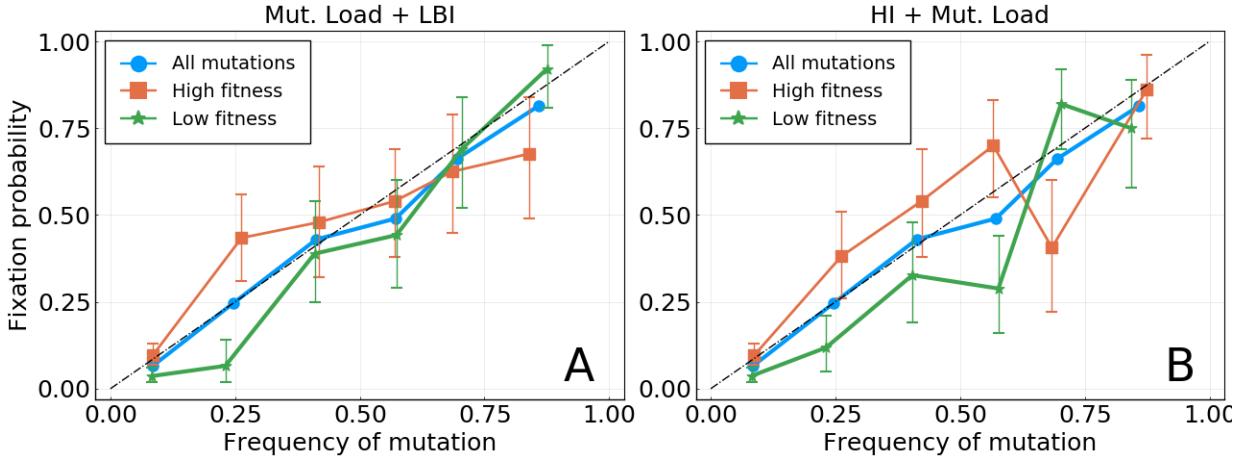


Figure S 4 Ability of fitness scores based on two combinations of the four fitness scores mentioned above: **A** mutational load + LBI; **B** mutational load + HI titer.

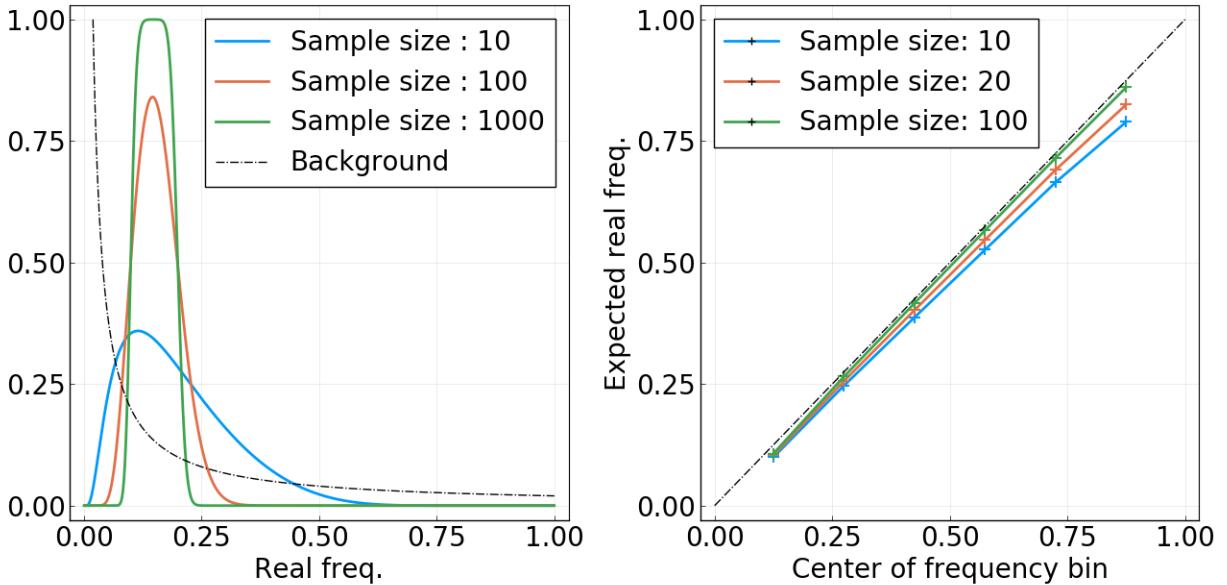


Figure S 5 **Left:** For a mutation present at frequency p in the population, probability of being observed in the frequency bin $[0.1, 0.2]$ as a function of p and for different sample sizes n . The dashed black line sketches the (non-normalized) background distribution $P_b(p)$. **Right:** Expected “real” average frequency of mutations found in frequency bin $[f_1, f_2]$ as a function of the centre of the bin $(f_1 + f_2)/2$, for different sample sizes.

We now try to estimate biases in frequency estimation due to this phenomenon. Given a set of mutations that have been measured in frequency bin $[f_1, f_2]$, what is the average *real* frequency of these mutations? To compute this, we need to sum $P_{f_1, f_2}(p, n)$ over all possible real frequencies p , giving us the amount of mutations that are observed in interval $[f_1, f_2]$, and weigh this sum by the frequency value p as well as by the background distribution of frequencies $P_b(p) \propto 1/p$. This last quantity represents the expected amount of mutations that are present at frequency p in the population. Note that there is no divergence problem as the smallest non zero frequency is $1/N$, where N is the population size. This leads us to the following expression for the average of “real” frequencies:

$$\langle p \rangle(f_1, f_2, n) = \int_{1/N}^{1-1/N} dp P_{f_1, f_2}(p, n) P_b(p) p = \int_{1/N}^{1-1/N} dp P_{f_1, f_2}(p, n). \quad (7)$$

1136 We have not made normalization explicit in these equations. It is simply achieved by dividing the above expression by
 1137 $\int dp P_{f_1, f_2}(p, n) P_b(p)$.
 1138 In the second panel of figure S5, $\langle p \rangle(f_1, f_2, n)$ is plotted as a function of the centre of the interval $[f_1, f_2]$ and for
 1139 different values of n . For sample sizes $n > 100$, the biases due to this effect are almost non existent. For smaller
 1140 samples, for instance $n = 10$, they are small but non negligible. However, we argue that this is not a significant
 1141 problem with respect to the main results presented in this article. First, figure S10 shows that sample sizes of the order
 1142 of $n = 10$ are only the case for a few months in the period going from year 2000 to 2018. From 2010 and onwards,
 1143 more than a hundred sequences are available per month for most months. Secondly, even if most samples were in the
 1144 $n = 10$ case, deviations shown in figure S5 are small enough that results shown in figures 2 and 3 would be *qualitatively*
 1145 unchanged.
 1146 Note that using the centre of the interval as a reference in figure S5, *i.e.* $(f_1 + f_2)/2$, would be correct in the case of
 1147 a very large n and a flat background distribution $P_b(p)$. For figures 2 and 3 of the main text however, the average
 1148 frequency of mutations found in an interval $[f_1, f_2]$ is computed by taking the average of the observed frequencies, and
 1149 not the centre of the interval. This partially takes into account biases considered here, as the background distribution
 1150 $P_b(p)$ is then accounted for, even though it is equivalent to assuming infinite sample sizes.

1151 6. Cutting off the HA1 159S branch

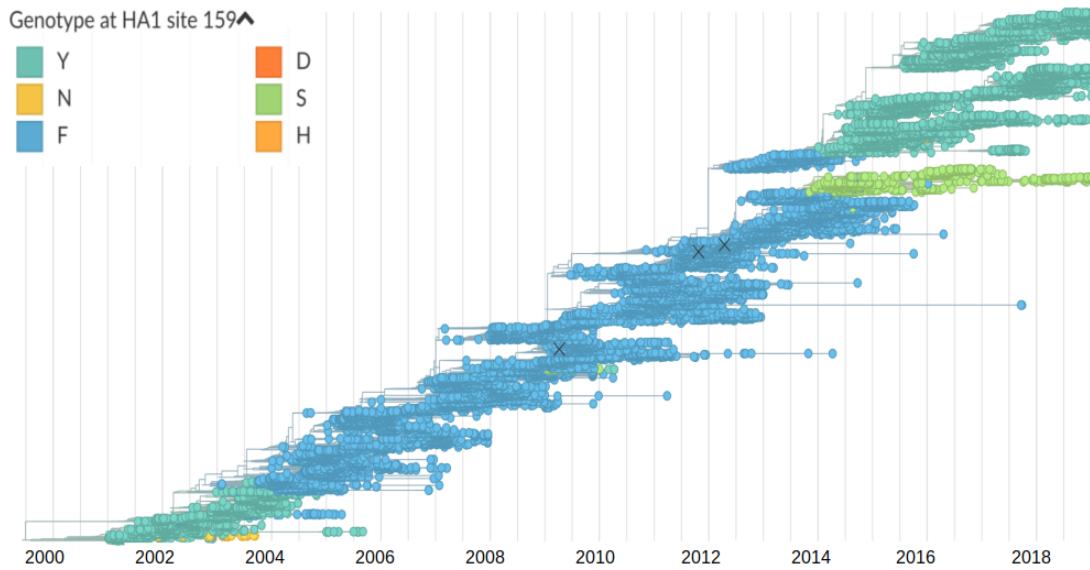


Figure S 6 Tree used for this study, based on a random selection of 100 strains per month from year 2002 to 2018. Nodes and branches are colored according to the amino acid found at position HA1:159. The HA1 159S mutation is visible as a thin but long light-greened color branch, coalescing with the “trunk” around 2013.

1152 The analysis of the main text is in a large part based on the probability of fixation of mutations. The motivation
 1153 underlying this choice is the relatively short coalescence time of the A/H3N2 influenza population, typically around
 1154 three years. This can be seen in figure 2 of the main text, which shows the typical lifetime of frequency trajectories,
 1155 ending in fixation or loss after at most 3 years in most cases. The tree in figure S6 is another illustration of this: for
 1156 the most part of it, a “trunk” is clearly identifiable, and lineages that depart from it have a relatively short lifetime.
 1157 This is no longer the case since the year ~ 2013: two clades have been competing since then, with no definite way to
 1158 identify a trunk in the tree. The clade defined by the HA1 159S mutation, colored in light green on figure S6, is one of
 1159 these two competing lineages. Because of this particular situation, the number of mutations fixating in the population
 1160 is strongly reduced, as a mutation must appear in both clades to reach a frequency of 1. This is a potential flaw in our
 1161 analysis, which concentrates on mutations fixating.

1162 For this reason, we decided to re-run our analysis after having cut off the HA1 159S clade. In other words, we remove
 1163 from the set of sequences those that carry the HA1 159S mutation. Results are shown in figures , equivalent to figures 2
 1164 and 3 of the main text. It is clear that qualitative results are left unchanged when this competing clade is removed.
 1165 This can be surprising, as almost no complete fixation of an amino acid mutation has occurred since 2013. Cutting off
 1166 the HA1 159S branch should thus result in many new fixations, changing the analysis. The reason for the similarity
 1167 of results can be explained: fixation (resp. loss) of a mutation are defined here as the frequency of this mutation
 1168 being measured above 95% (resp. 5%) frequency for two months in a row. As the HA1 159S clade is rather sparsely
 1169 populated, it reaches frequencies lower than 5% two times (in 2015 and 2017), allowing mutations in the competing
 1170 clade to “fix” as defined here. Thus, removing strains carrying HA1 159S does not introduce a significant amount of
 1171 “new” fixation events.

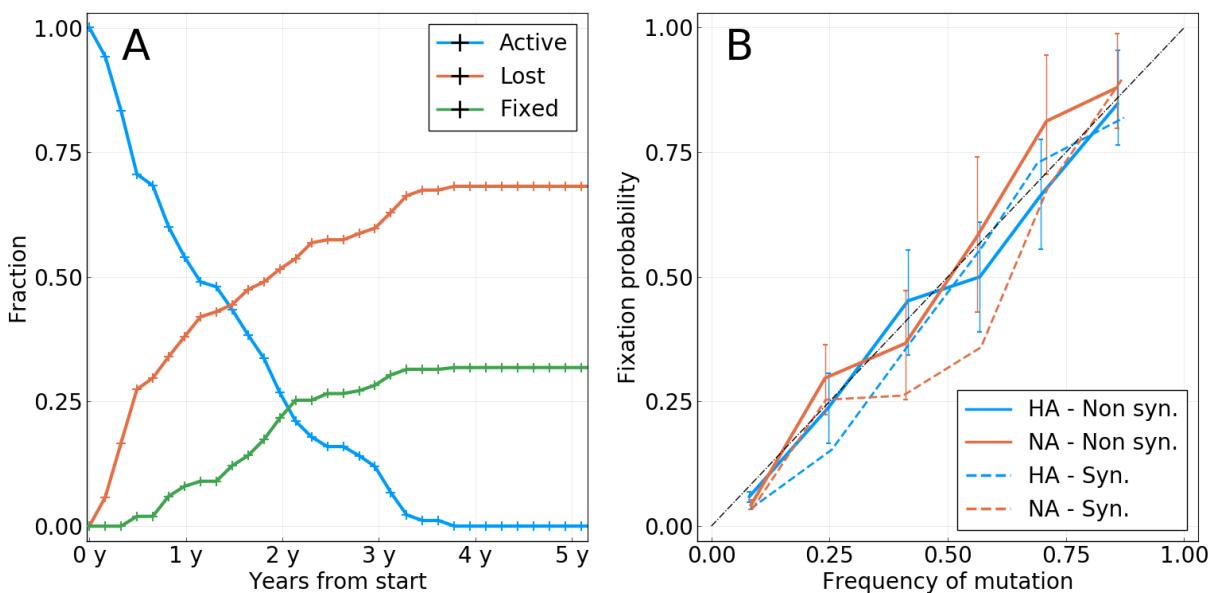


Figure S 7 Equivalent to figure 2 of the main text, but with strains carrying the HA1 159S mutation removed.

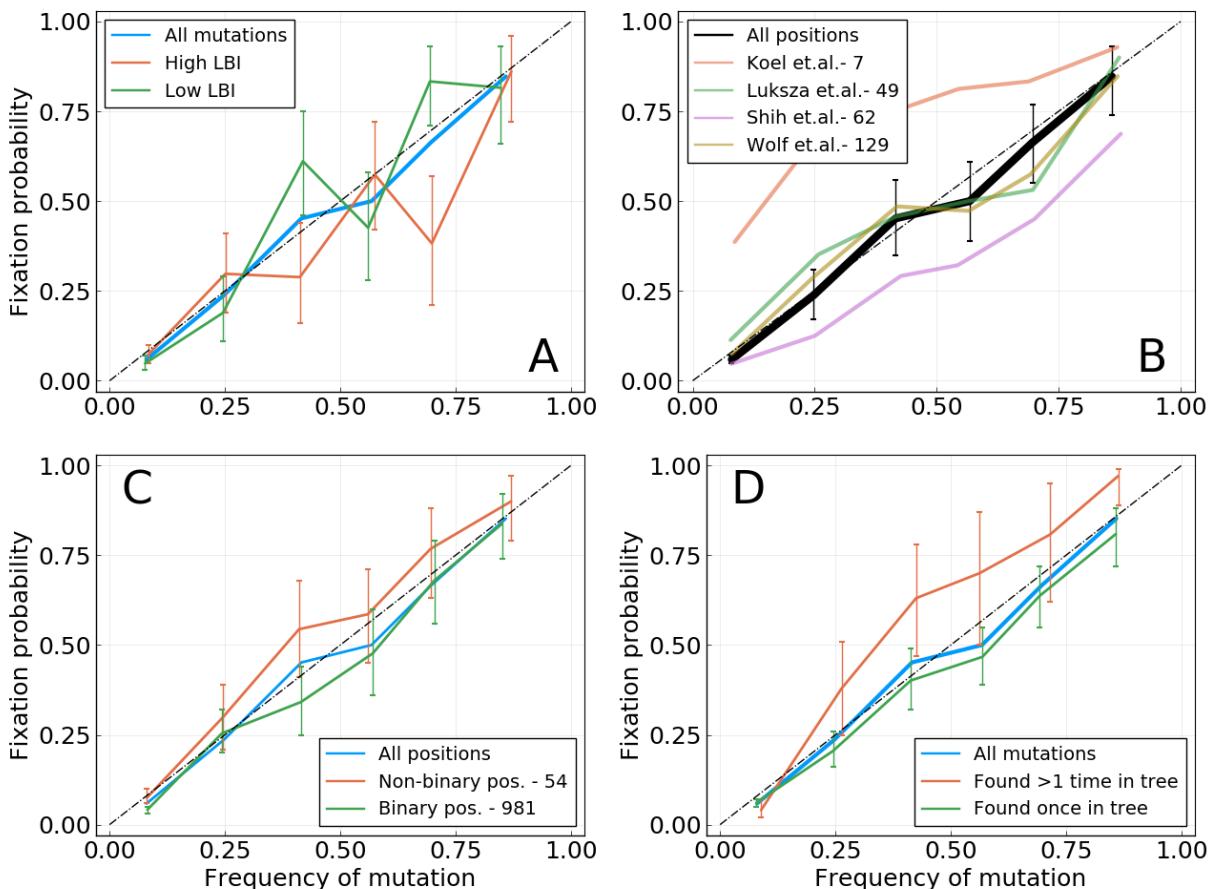


Figure S 8 Equivalent to figure 3 of the main text, but with strains carrying the HA1 159S mutation removed.

1172 **7. Probability of fixation in single locus model of evolution**

1173 In (Kimura, 1964), Kimura investigates a simple model of evolution with a single locus and a population of size
 1174 N . In this framework, a mutation at this locus with fitness effect s and observed at frequency f has the following
 1175 probability of fixation:

$$P_{fix}(f|s, N) = \frac{1 - e^{-sNf}}{1 - e^{-sN}}. \quad (8)$$

1176 Expanding this formula for $sN \ll 1$, that is in the weak selection regime, yields at the first order

$$P_{fix}(f|s, N) = f + f(1 - f) \frac{sN}{2}. \quad (9)$$

1177 Equation 9 tells us two things. First, when the mutation is neutral, that is $s = 0$, we have $P_{fix}(f) = f$. This naturally
 1178 confirms the result obtained for a neutral model of evolution. Seconds, when $sN \neq 0$, we can expect deviations from
 1179 the diagonal in a P_{fix} against f plot. The sign of these deviations is determined by the sign of s , with beneficial
 1180 mutations being found above diagonal while deleterious one are found below. The amplitude of these deviations
 1181 depends on the strength of selection sN , as well as on the frequency through the $f(1 - f)$ term, making them larger
 1182 for $f \sim 0.5$.

1183 8. Mutation tables

Gene	Position	AA	Start date	End date	Shih	Luksza	Koel	Tree counts
HA1	144	D	2001-06-09	2002-02-04	true	true	false	0
HA1	189	N	2003-07-29	2004-05-24	false	true	true	2
HA1	159	F	2003-08-28	2004-05-24	false	true	true	2
HA1	226	I	2003-09-27	2004-09-21	true	true	false	3
HA1	145	N	2003-12-26	2004-11-20	false	true	true	2
HA1	227	P	2003-05-30	2005-04-19	false	true	false	2
HA2	32	I	2004-06-23	2005-07-18	false	false	false	1
HA1	193	F	2004-12-20	2006-03-15	false	true	true	1
HA2	46	D	2006-06-13	2007-05-09	false	false	false	2
HA2	121	K	2006-06-13	2007-06-08	false	false	false	1
HA1	50	E	2006-09-11	2007-06-08	false	true	false	2
HA1	140	I	2006-11-10	2007-11-05	true	false	false	1
HA1	173	Q	2007-07-08	2009-01-28	true	true	false	2
HA2	32	R	2007-07-08	2009-01-28	false	false	false	1
HA1	158	N	2009-01-28	2009-07-27	true	true	true	2
HA1	189	K	2009-01-28	2009-07-27	false	true	true	2
HA1	212	A	2009-03-29	2011-01-18	false	false	false	2
HA1	45	N	2010-03-24	2013-02-06	false	false	false	3
HA1	223	I	2010-12-19	2013-02-06	false	false	false	2
HA1	48	I	2011-03-19	2013-02-06	false	false	false	1
HA1	198	S	2011-03-19	2013-02-06	false	false	false	1
HA1	312	S	2009-08-26	2013-03-08	false	false	false	3
HA1	278	K	2011-06-17	2013-03-08	false	true	false	1
HA1	145	S	2011-04-18	2013-04-07	false	true	true	4
HA1	33	R	2011-06-17	2013-06-06	false	false	false	2
HA2	160	N	2012-07-11	2015-09-24	false	false	false	3
HA1	225	D	2013-08-05	2015-09-24	false	false	false	3
HA1	3	I	2013-08-05	2016-11-17	false	false	false	2
HA1	159	Y	2014-02-01	2016-11-17	false	true	true	2
HA1	160	T	2014-01-02	2017-07-15	false	true	false	2

Table S I The 30 trajectories that took place between year 2000 and year 2018 and resulted in fixation. Columns **Shih**, **Luksza** and **Koel** respectively indicate whether the position is found in the epitopes lists in (respectively) (Shih et al., 2007), (Luksza and Lässig, 2014) and (Koel et al., 2013). The **Tree counts** column indicates the number of times the mutation corresponding to the trajectory can be found in the phylogenetic tree. Note that a trajectory is only shown in the table if the sequenced population counts more than 10 strains at its time of fixation. This explains that only 30 trajectories are displayed, whereas more mutations did fix in this period of time.

Gene	Position	AA	Start date	End date	Fixation	Max. freq.
HA1	106	A	2001-02-09	2002-02-04	lost	1.0
HA1	144	D	2001-06-09	2002-02-04	fixed	1.0
HA1	105	H	2003-04-30	2003-10-27	lost	1.0
HA1	126	D	2003-04-30	2004-05-24	lost	1.0
HA1	140	Q	2004-01-25	2004-06-23	lost	0.31
HA1	226	I	2003-09-27	2004-09-21	fixed	1.0
HA1	173	E	2004-12-20	2006-03-15	lost	0.63
HA1	142	G	2006-06-13	2007-05-09	lost	0.71
HA1	144	D	2006-07-13	2007-05-09	lost	0.67
HA1	128	A	2006-09-11	2007-05-09	lost	0.25
HA1	157	S	2006-09-11	2007-05-09	lost	0.59
HA1	140	I	2006-11-10	2007-11-05	fixed	1.0
HA1	173	N	2007-12-05	2008-07-02	lost	0.3
HA1	157	S	2007-12-05	2008-09-30	lost	0.31
HA1	173	E	2006-06-13	2008-12-29	lost	0.67
HA1	173	Q	2007-07-08	2009-01-28	fixed	0.96
HA1	158	N	2009-01-28	2009-07-27	fixed	0.96
HA1	62	K	2009-01-28	2011-05-18	lost	0.73
HA1	144	K	2009-01-28	2011-05-18	lost	0.75
HA1	62	V	2011-04-18	2011-09-15	lost	0.34
HA1	157	S	2013-05-07	2015-09-24	lost	0.35
HA1	128	A	2012-08-10	2016-11-17	lost	0.81
HA1	197	K	2015-11-23	2016-11-17	lost	0.27
HA1	142	R	2018-05-11	2018-10-08	lost	0.38
HA1	142	G	2012-03-13		poly	0.86
HA1	144	S	2013-12-03		poly	0.96
HA1	121	K	2015-12-23		poly	0.82
HA1	142	K	2016-05-21		poly	0.77
HA1	62	G	2017-03-17		poly	0.75
HA1	128	A	2018-01-11		poly	0.56

Table S II Trajectories of mutations at epitope positions in (Shih et al., 2007) (*Shih et. al.*) that have been observed at least once above frequency 0.25. The **Fixation** column indicates whether the mutation has fixed, disappeared, or is still polymorphic as of October 2018. The **Max.freq.** column indicates the maximum frequency reached by the trajectory. A maximum frequency of 1 for mutations that finally disappear is explained by trajectories reaching frequency 1 for one time bin and going back to lower values for following ones (a frequency above 0.95 for two time bins in a row defines fixation).

Gene	Position	AA	Start date	End date	Fixation	Max. freq.
HA1	50	G	2001-02-09	2002-02-04	lost	1.0
HA1	144	D	2001-06-09	2002-02-04	fixed	1.0
HA1	126	D	2003-04-30	2004-05-24	lost	1.0
HA1	189	N	2003-07-29	2004-05-24	fixed	1.0
HA1	159	F	2003-08-28	2004-05-24	fixed	1.0
HA1	226	I	2003-09-27	2004-09-21	fixed	1.0
HA1	145	N	2003-12-26	2004-11-20	fixed	1.0
HA1	188	N	2004-07-23	2005-02-18	lost	0.36
HA1	227	P	2003-05-30	2005-04-19	fixed	1.0
HA1	173	E	2004-12-20	2006-03-15	lost	0.63
HA1	193	F	2004-12-20	2006-03-15	fixed	0.97
HA1	142	G	2006-06-13	2007-05-09	lost	0.71
HA1	144	D	2006-07-13	2007-05-09	lost	0.67
HA1	157	S	2006-09-11	2007-05-09	lost	0.59
HA1	50	E	2006-09-11	2007-06-08	fixed	0.95
HA1	173	N	2007-12-05	2008-07-02	lost	0.3
HA1	157	S	2007-12-05	2008-09-30	lost	0.31
HA1	173	E	2006-06-13	2008-12-29	lost	0.67
HA1	173	Q	2007-07-08	2009-01-28	fixed	0.96
HA1	158	N	2009-01-28	2009-07-27	fixed	0.96
HA1	189	K	2009-01-28	2009-07-27	fixed	0.96
HA1	213	A	2009-01-28	2010-02-22	lost	0.68
HA1	144	K	2009-01-28	2011-05-18	lost	0.75
HA1	53	N	2009-11-24	2013-02-06	lost	0.72
HA1	278	K	2011-06-17	2013-03-08	fixed	0.98
HA1	145	S	2011-04-18	2013-04-07	fixed	0.99
HA1	159	S	2013-11-03	2015-08-25	lost	0.46
HA1	157	S	2013-05-07	2015-09-24	lost	0.35
HA1	159	Y	2014-02-01	2016-11-17	fixed	0.97
HA1	159	S	2015-10-24	2016-11-17	lost	0.4
HA1	197	K	2015-11-23	2016-11-17	lost	0.27
HA1	160	T	2014-01-02	2017-07-15	fixed	0.96
HA1	142	R	2018-05-11	2018-10-08	lost	0.38
HA1	135	N	2018-06-10	2018-10-08	lost	0.38
HA1	142	G	2012-03-13		poly	0.86
HA1	144	S	2013-12-03		poly	0.96
HA1	121	K	2015-12-23		poly	0.82
HA1	142	K	2016-05-21		poly	0.77
HA1	131	K	2016-09-18		poly	0.77
HA1	135	K	2016-11-17		poly	0.47

Table S III Same as table SII, for (Luksza and Lässig, 2014) (*Luksza et. al.*).

Gene	Position	AA	Start date	End date	Fixation	Max. freq.
HA1	189	N	2003-07-29	2004-05-24	fixed	1.0
HA1	159	F	2003-08-28	2004-05-24	fixed	1.0
HA1	145	N	2003-12-26	2004-11-20	fixed	1.0
HA1	193	F	2004-12-20	2006-03-15	fixed	0.97
HA1	158	N	2009-01-28	2009-07-27	fixed	0.96
HA1	189	K	2009-01-28	2009-07-27	fixed	0.96
HA1	145	S	2011-04-18	2013-04-07	fixed	0.99
HA1	159	S	2013-11-03	2015-08-25	lost	0.46
HA1	159	Y	2014-02-01	2016-11-17	fixed	0.97
HA1	159	S	2015-10-24	2016-11-17	lost	0.4

Table S IV Same as table SII, for (Koel et al., 2013) (*Koel et. al.*).

1184 9. Supplementary figures

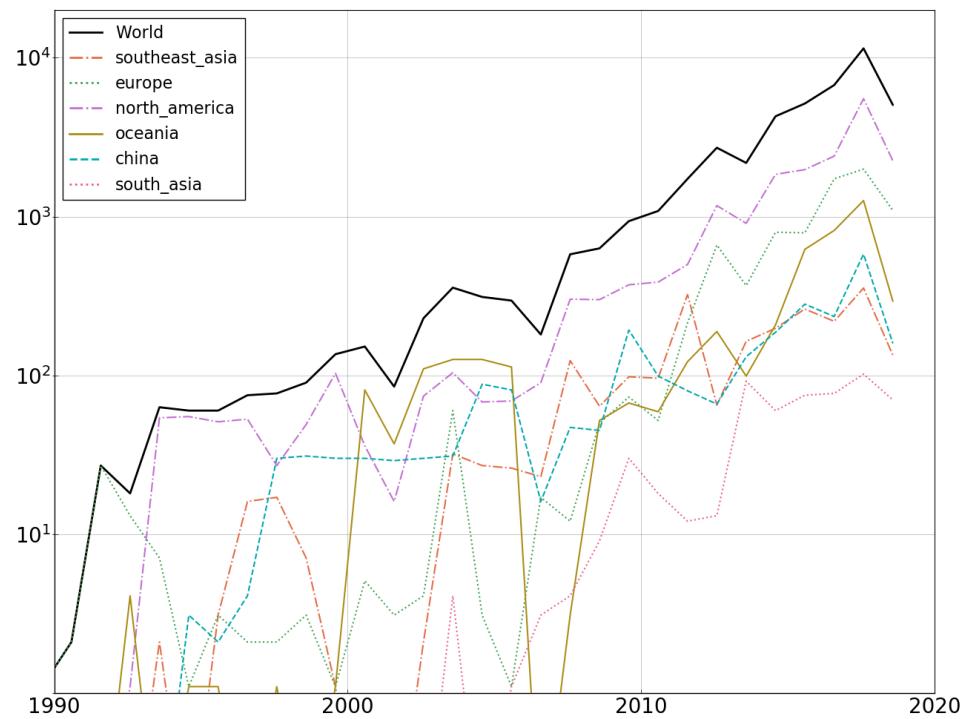


Figure S 9 Number of A/H3N2 HA sequences per year from year 1990.

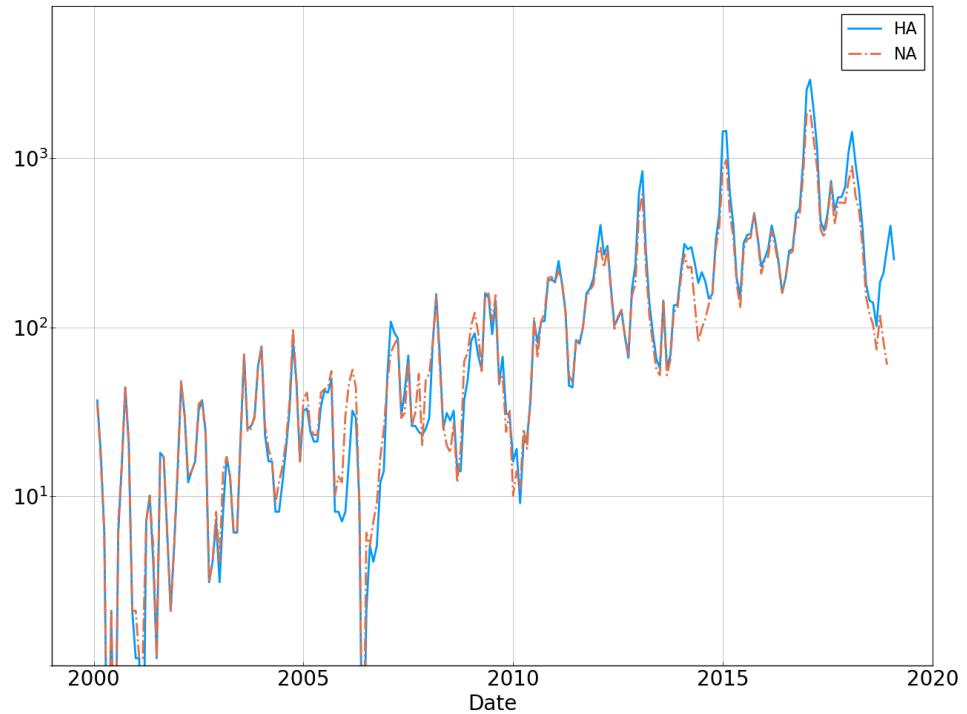


Figure S 10 Number of H3N2 HA and NA sequences per month from year 2000.

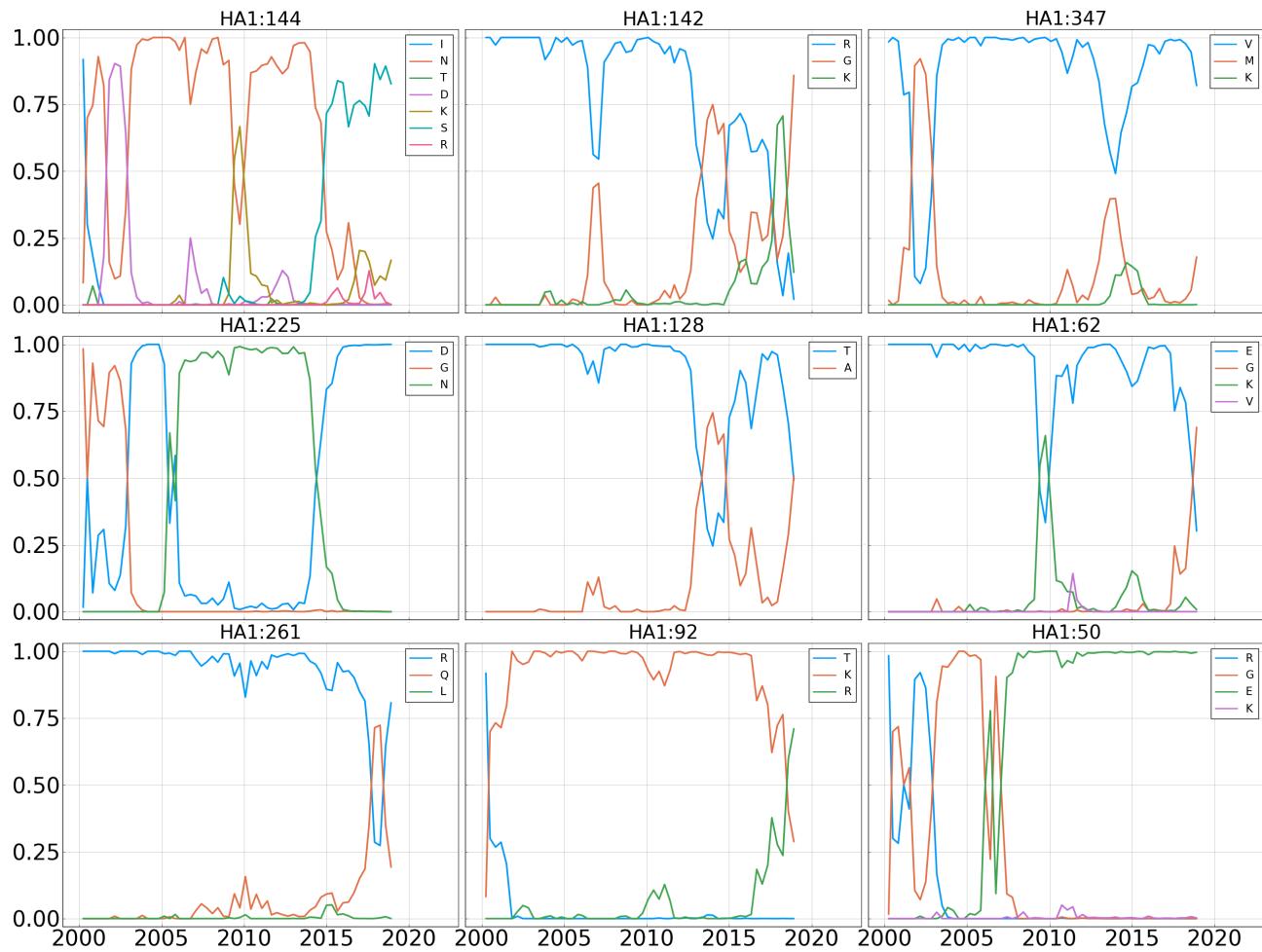


Figure S 11 Frequency trajectories for the 9 most entropic positions in the A/H3N2 HA protein.

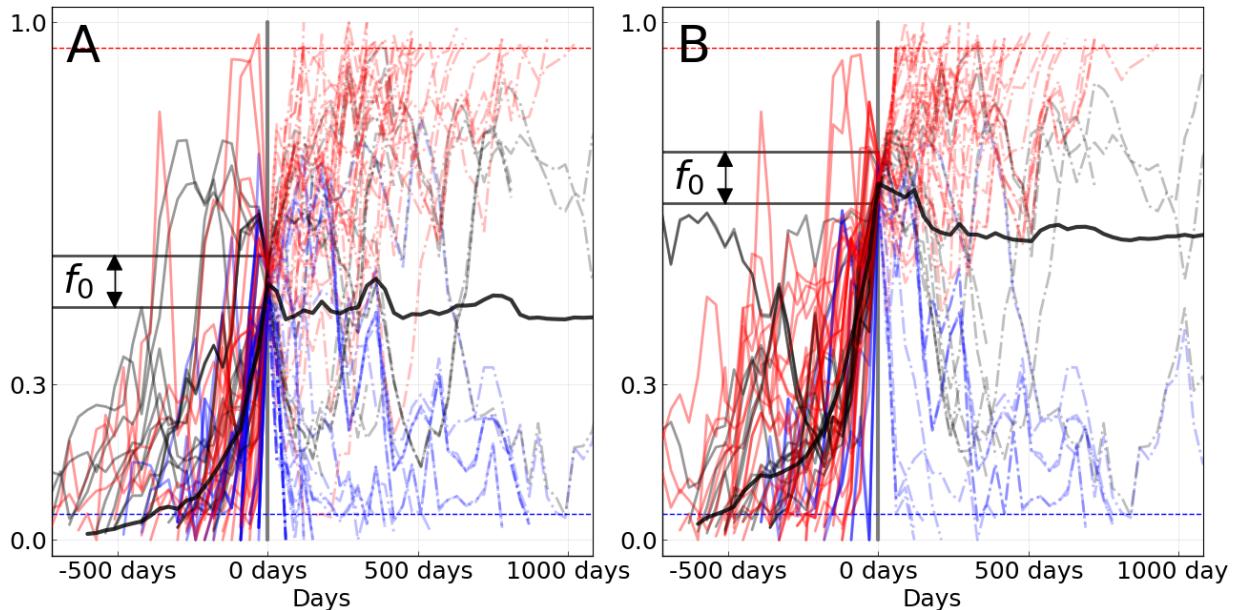


Figure S 12 Equivalent to panel B of figure 1 of the main text for A/H3N2, with f_0 equal to 0.5 in A (76 trajectories), and 0.7 in B (63 trajectories).

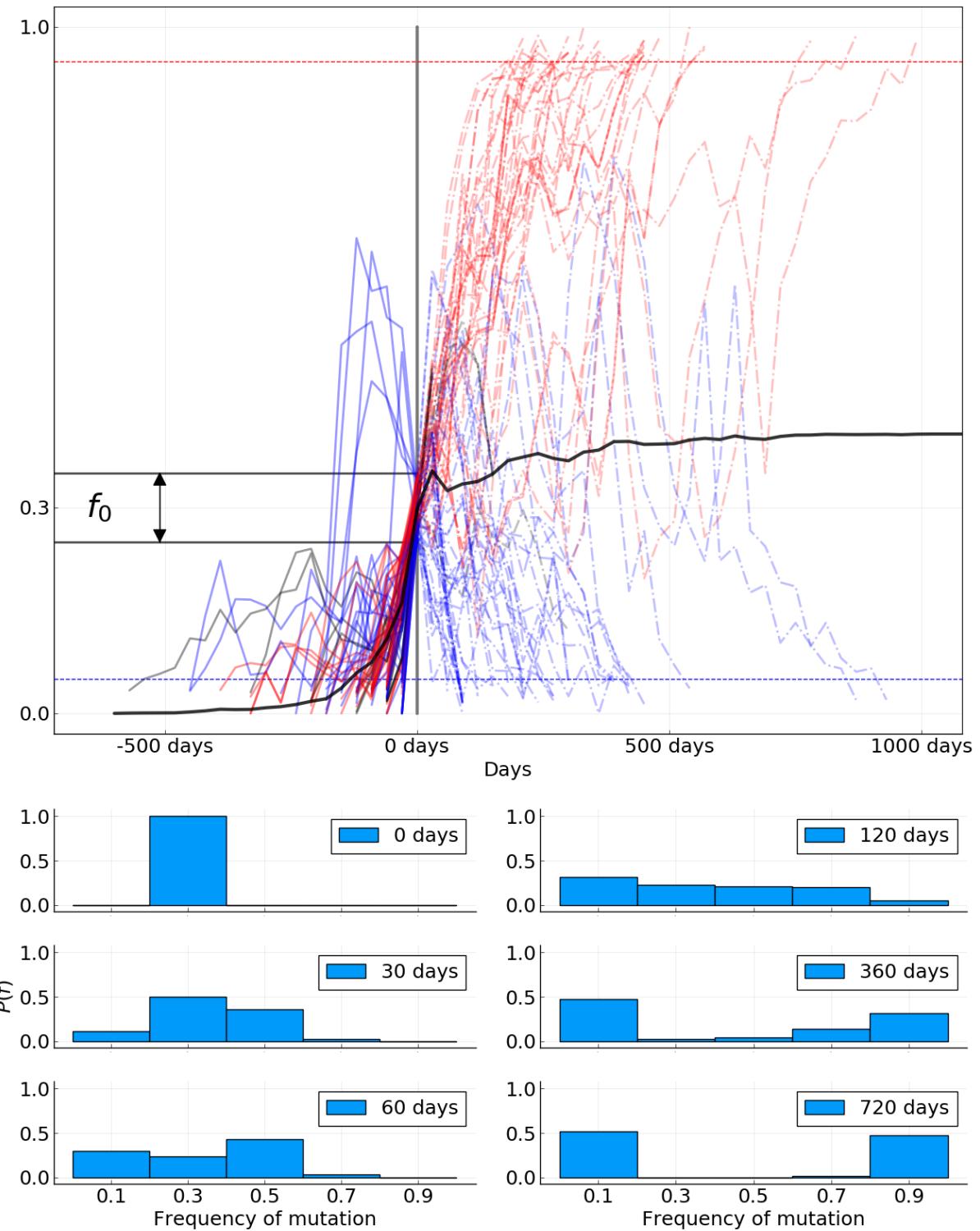


Figure S 13 Equivalent to panels **B** and **C** of figure 1 of the main text for A/H1N1pdm influenza. 89 trajectories are shown and participate to the mean (thick black line).

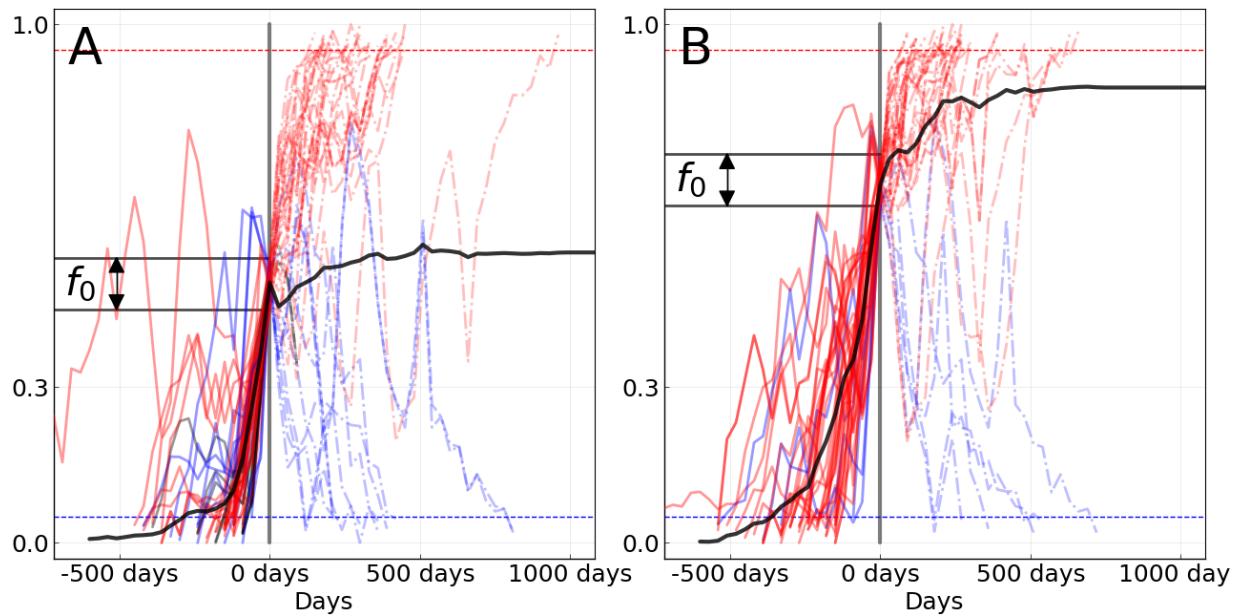


Figure S 14 Equivalent to panel **B** of figure 1 of the main text for A/H1N1pdm, with f_0 equal 0.5 in **A** (50 trajectories), and 0.7 in **B** (41 trajectories).

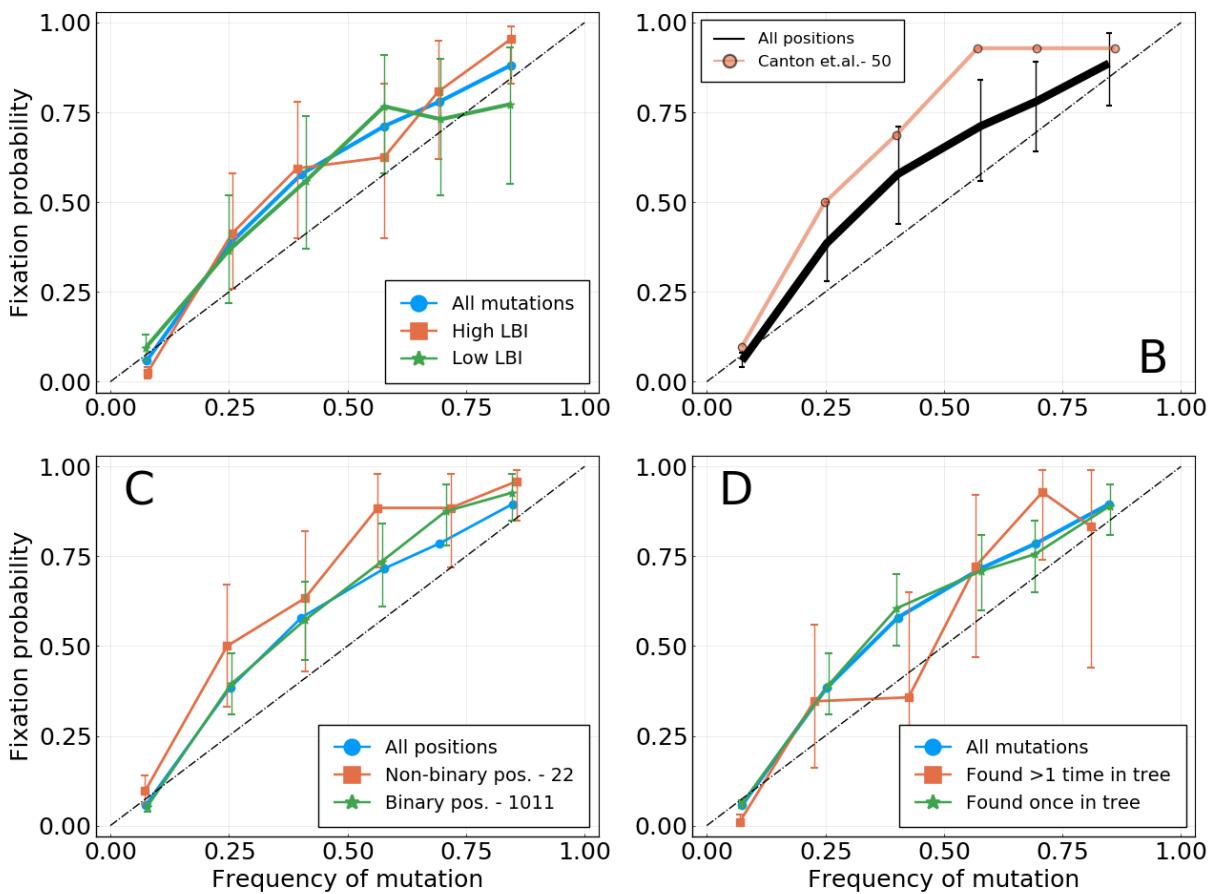


Figure S 15 Equivalent of figure 3 of the main text for the HA gene of A/H1N1pdm influenza. Fixation probability $P_{fix}(f)$ as a function of frequency. **A:** Mutation with higher or lower LBI values, based on their position with respect to the median LBI value. **B:** Different lists of epitope positions in the HA protein. The authors and the number of positions is indicated in the legend. **C:** Mutations for binary positions, *i.e.* positions for which we never see more than two amino acids in the same time bin. **D:** Mutations that appear once or more than once in the tree for a given time bin.

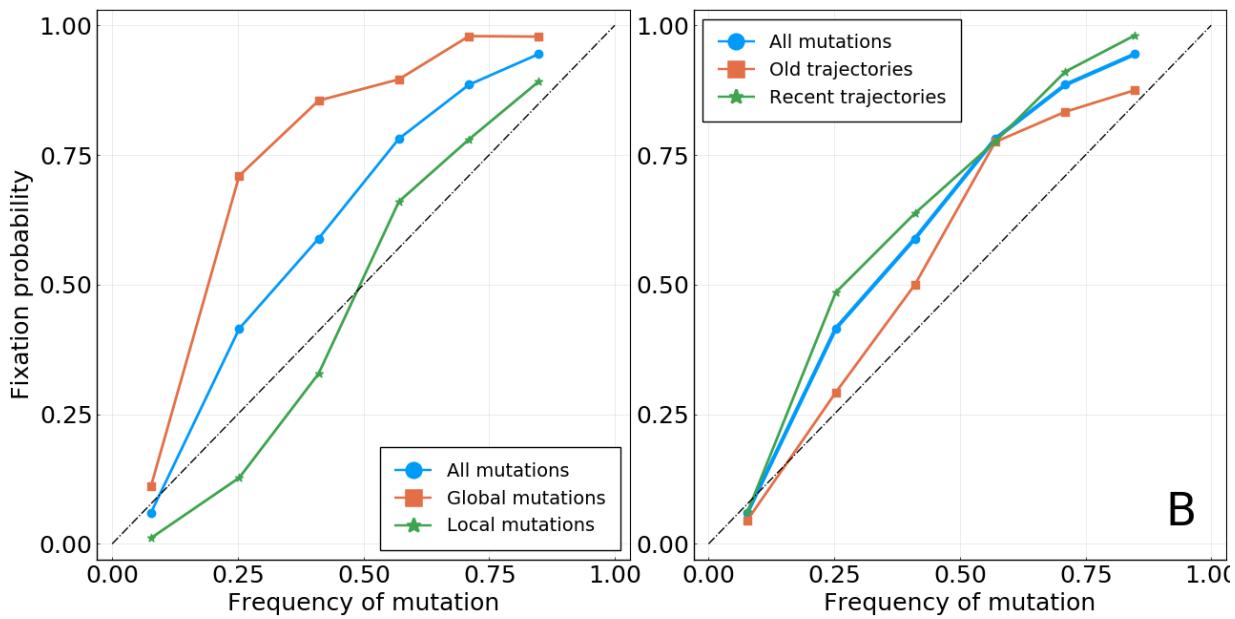
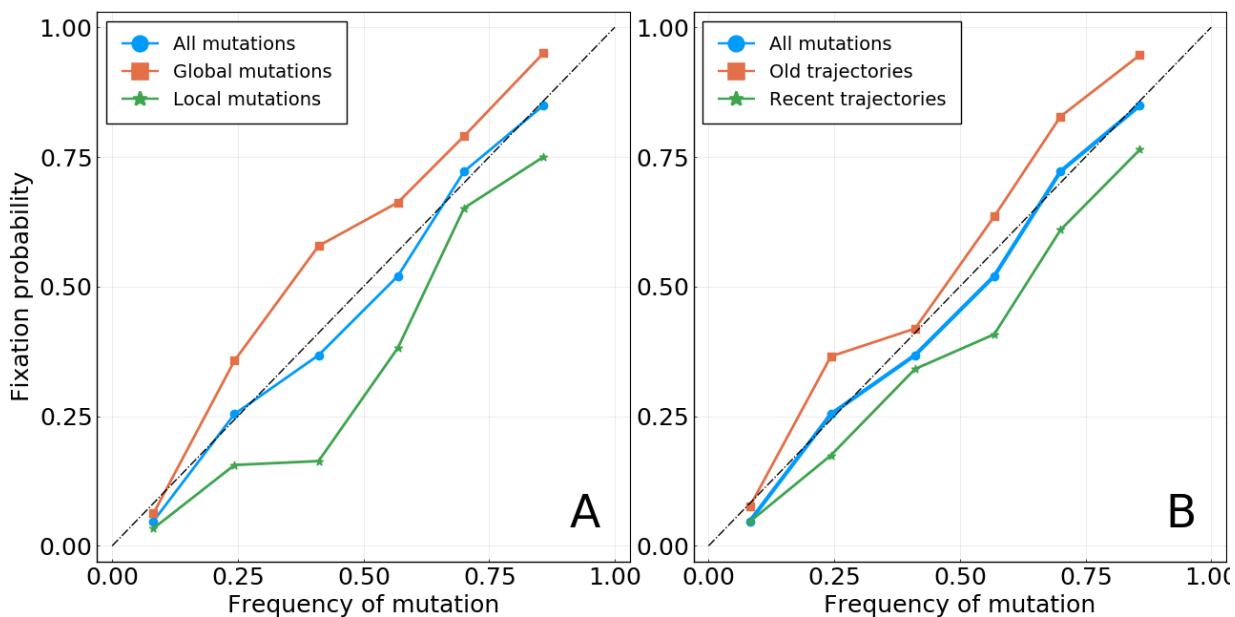


Figure S 17 Based on A/H1N1pdm HA and NA. **A:** Mutations with a higher or lower geographical spread, based on the median value of the score used (see Methods). Note: the words *local* and *global* only reflect the position of the geographic spread of the mutation relative to the median value computed for all mutations found at this frequency. As this median value may change with the considered frequency bin, so does the definition of local and global mutations. **B:** Mutations whose trajectories are older or more recent, based on the median age of trajectories when reaching the considered frequency f .

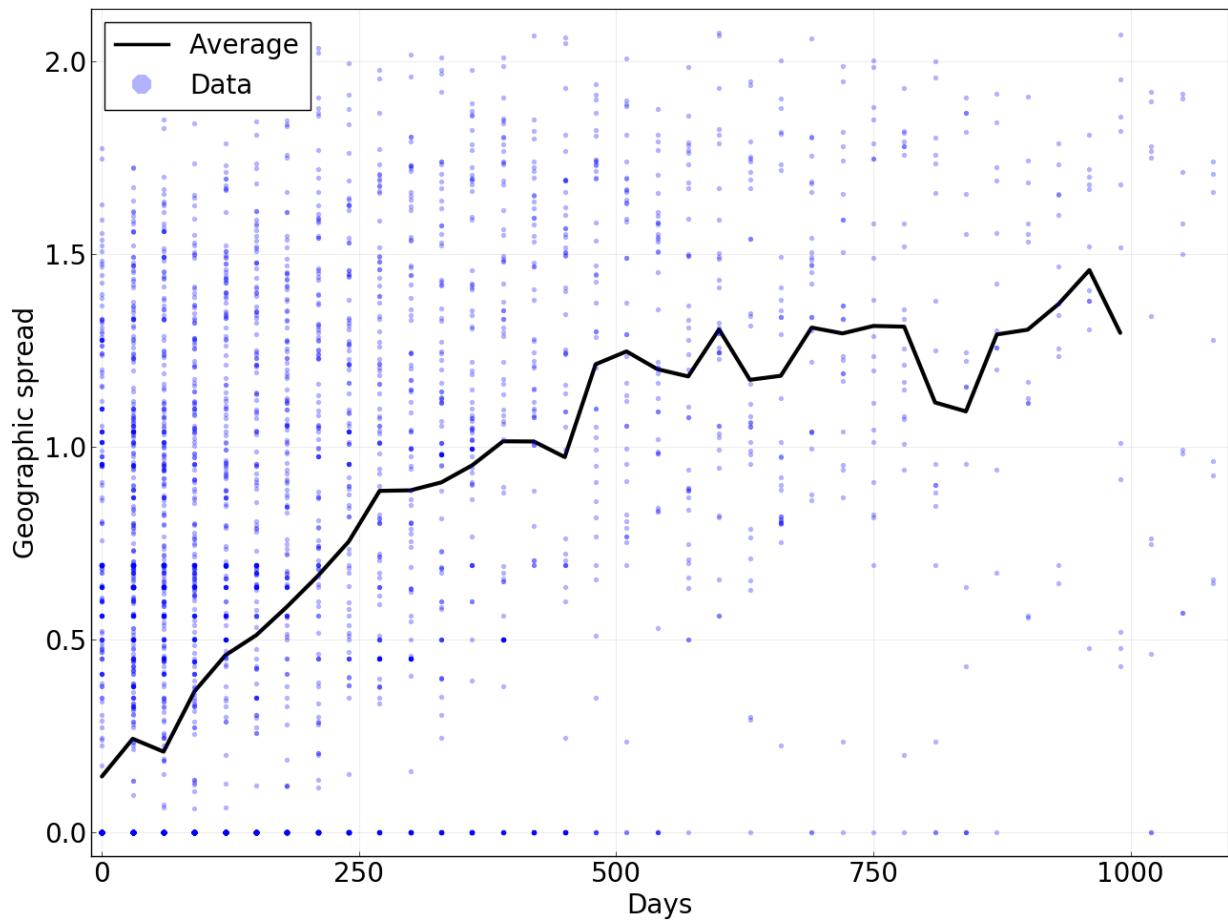


Figure S 18 Geographic spread of mutations as a function of the time for which they have been present in the population above a frequency of 5%. Points represent individual mutations and for a population in a given time bin. The line is the average of dots for a given value on the x -axis. Based on data for A/H3N2 HA.

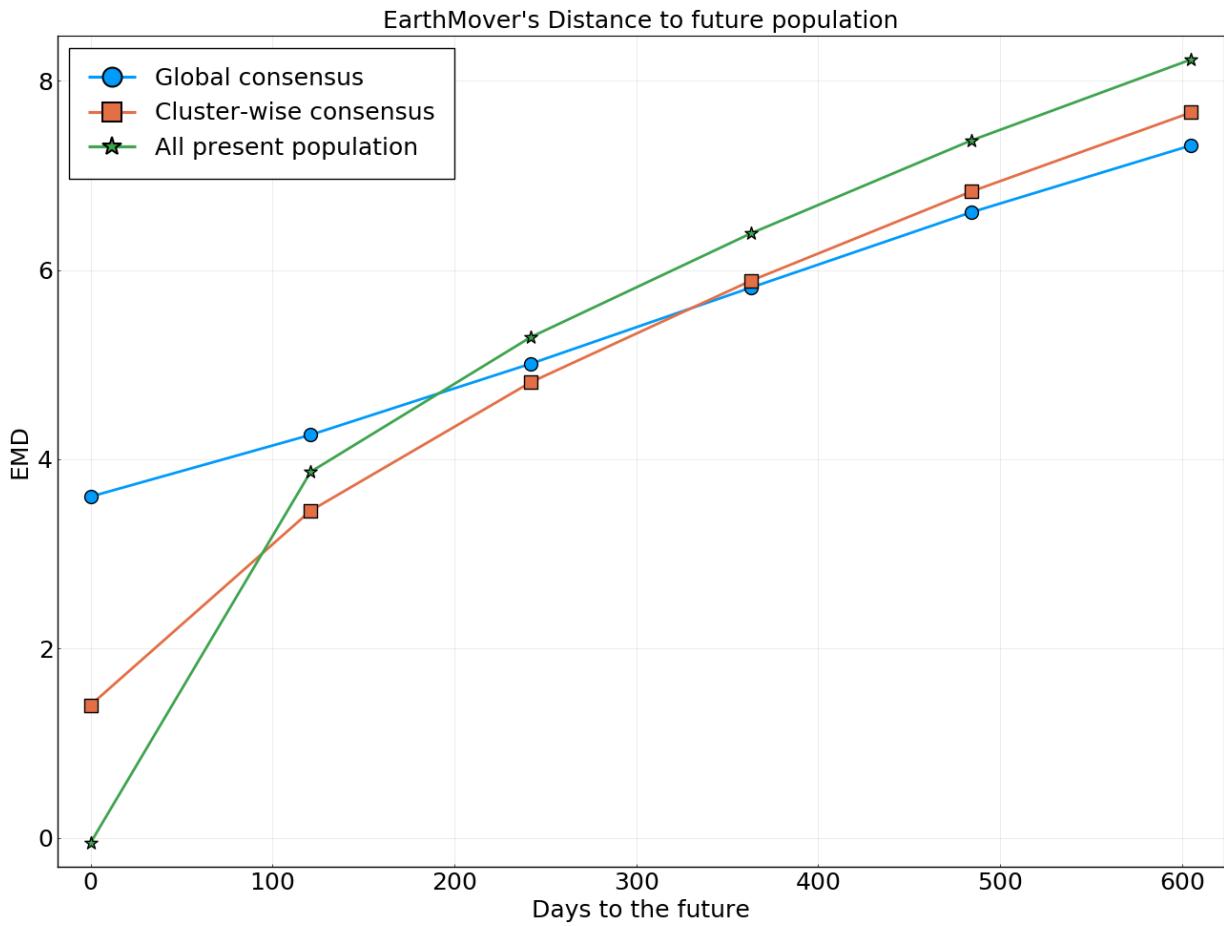


Figure S 19 Earth mover's distance to the future population for different predictors. A present population consists of all A/H3N2 HA sequences sampled in a 4 months time window. Quantities are averaged over all possible "present" populations from the year 2002. Predictors are: **Global consensus**: Consensus sequence of the present population. Best long-term predictor for a structure-less neutrally evolving population. **All present population**: All sequences in the present population. Perfect predictor if the population does not change at all through time. **Cluster-wise consensus**: Consensus sequence for each cluster in the present population. Clusters are based on local maxima of the LBI. Sequences are assigned to a given cluster based on their tree branch-length distance to the corresponding local maximum.

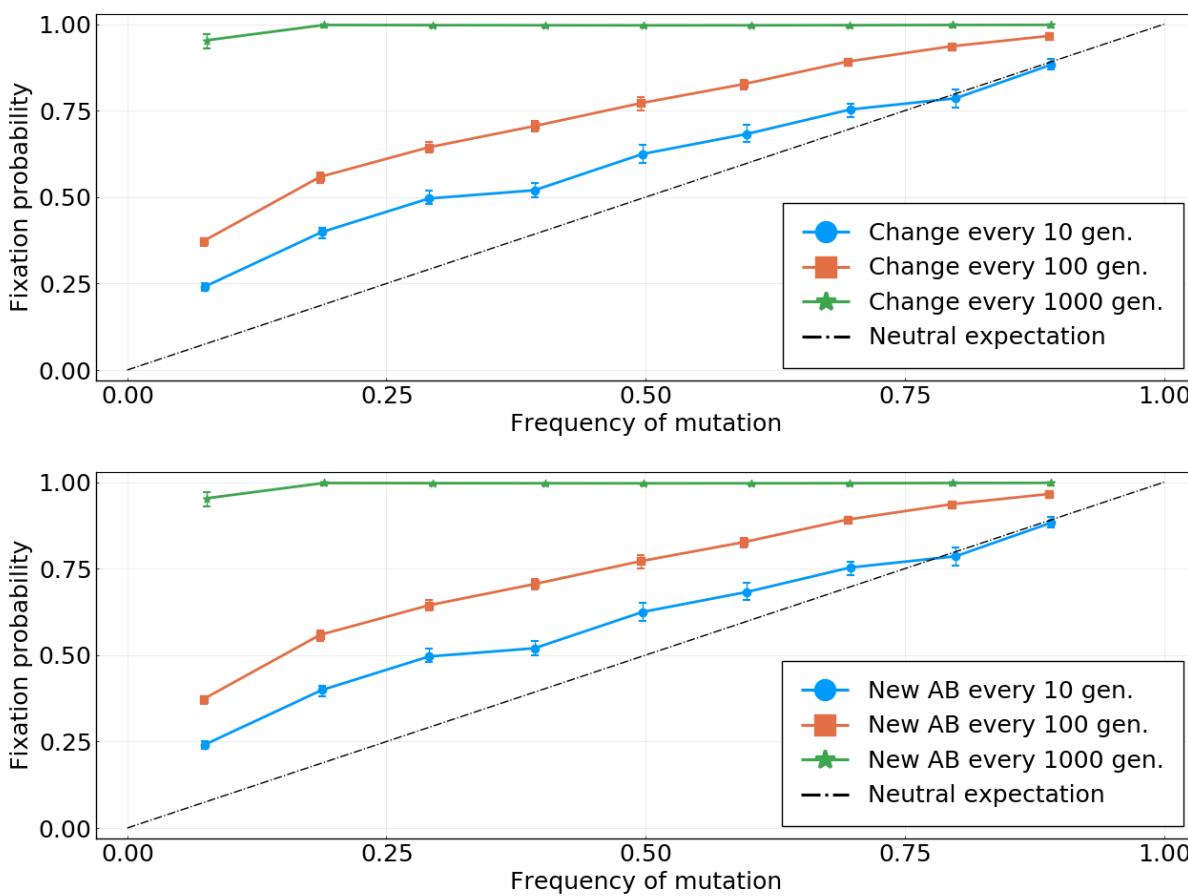


Figure S 20 Fixation probability as a function of frequency for the simulations discussed in the main text. **Top:** Simulation without antibodies. The three colored curves reflect different rate of change for the fitness landscape. Visual inspection of the frequency trajectories indicates a typical sweep time of ~ 400 generations. **Bottom:** Simulation with antibodies. The different colored curves indicate the rate at which antibodies are introduced.

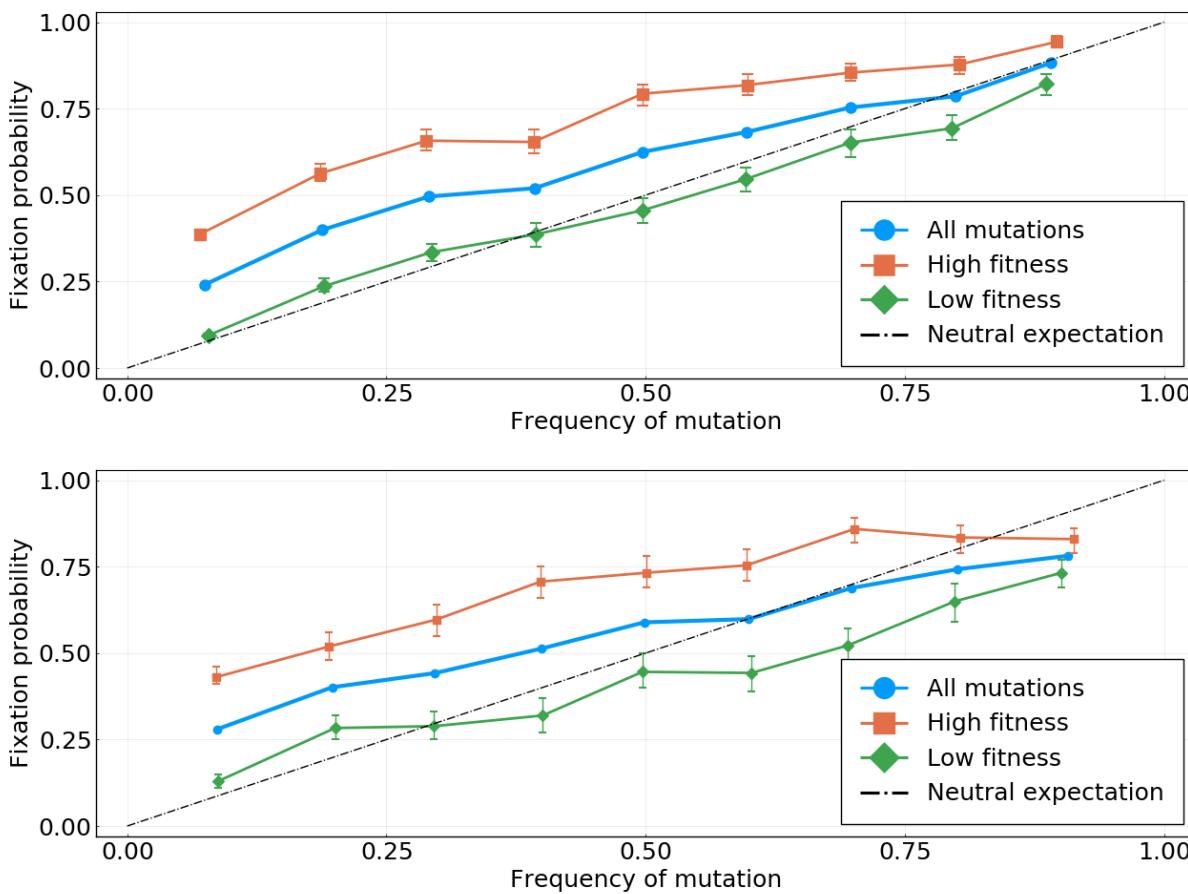


Figure S 21 Fixation probability as a function of frequency for the simulations discussed in the main text, with trajectories stratified according to real fitness values. “High” and “low” fitness classes are defined with respect to the median value. **Top:** Simulation without antibodies and with changes to the fitness landscape every $dt = 10$ generations. **Bottom:** Simulation with antibodies, with a new antibody every $dt = 10$ generations.