

Homework 10

Solution

April 30th

Exercises from the book

Exercise 2.35

(i) $T \sim \text{Exp}(\lambda)$, $10 = \mu = \mathbb{E}[T] = \frac{1}{\lambda}$ so that $\lambda = 0.1$

$$\text{Var}(T) = \mu^2 = 100.$$

(ii)

$$\mathbb{P}(T \leq 5) = 10 \int_0^5 e^{-0.1t} dt = 1 - e^{-0.5}.$$

(iii)

$$\mathbb{P}(T \leq 30 | T > 25) = 1 - \mathbb{P}(T > 30 | T > 25) = 1 - \mathbb{P}(T > 5) = \mathbb{P}(T \leq 5).$$

(iv)

$$\mathbb{P}(T > \mathbb{E}[T]) = \mathbb{P}(T > 10) = e^{-10\lambda} = e^{-1}.$$

Exercise 2.42 We compute the cumulative distribution for Y , and deduce its density by taking the derivative. It will help us to keep in mind that the inverse of the logarithm is the exponential.

$$F_Y(y) = \mathbb{P}(Y \leq y) = \mathbb{P}(\ln(X) \leq y) = \mathbb{P}(X \leq e^y) = F_X(e^y).$$

Since $F_X(x) = 1 - e^{-3x}$ for all $x > 0$, we get that

$$F_Y(y) = 1 - \exp(-3e^y)$$

and

$$p_Y(y) = 3e^y \cdot \exp(-3e^y) = 3 \exp(-3e^y + y).$$

Exercise 2.46 Again, we first compute the cumulative distribution for Y . Setting $f(x)$ such that $Y = f(X)$, we see that for $0 < y < 1$,

$$f(x) \leq y \Leftrightarrow (x \leq y \text{ and } x \leq 1) \text{ or } (1/x \leq y \text{ and } x > 1) \Leftrightarrow x \leq y \text{ or } x \geq 1/y.$$

Then for $0 < y < 1$ we have

$$F_Y(y) = \mathbb{P}(Y \leq y) = \mathbb{P}(\{X \leq y\} \cup \{X \geq 1/y\}) = \int_0^y e^{-x} dx + \int_{1/y}^{\infty} e^{-x} dx = 1 - e^{-y} + e^{-1/y}.$$

Differentiating, for $0 < y < 1$ the density function is $p_Y(y) = e^{-y} + \frac{1}{y^2}e^{-1/y}$. Since $f(x)$ has values in $[0, 1]$ for $x \geq 0$, Y has values in $[0, 1]$ and $F_y(y) = 0$ out of $(0, 1)$.

Exercise 2.48

(i)

$$\int_{-\infty}^{\infty} f(x) dx = \int_{c_n}^{\infty} \frac{c_n}{x^{n+1}} dx = \frac{1}{nc_n^{n-1}} = 1$$

since $f(x)$ is the density of a random variable. Therefore, if $n > 1$

$$c_n = \frac{1}{n^{1/(n-1)}}$$

while if $n = 1$, then c_n could be anything.

(ii)

$$\mathbb{E}[X] = \int_{c_n}^{\infty} \frac{c_n}{x^n} dx.$$

This diverges if $n = 1$ but for $n > 1$ gives

$$\frac{1}{n-1} \cdot n^{(n-2)/(n-1)} = \frac{n}{(n-1)n^{1/(n-1)}}.$$

(iii) The random variable Z_n has range $[\ln c_n, \infty)$. For z in this range we have

$$F_{Z_n}(z) = \mathbb{P}(Z_n \leq z) = \mathbb{P}(\ln(X_n) \leq z) = \mathbb{P}(X_n \leq e^z) = \int_{c_n}^{e^z} \frac{c_n}{x^{n+1}} dx = \frac{1}{nc_n^{n-1}} - \frac{c_n}{ne^{nz}}.$$

The density is then the derivative of the function above:

$$p_{Z_n}(z) = \begin{cases} c_n e^{-nz} & \text{for } z \geq \ln(c_n), \\ 0 & \text{else.} \end{cases}$$

(iv)

$$\mathbb{E}[X_n^{m+1}] = \int_{c_n}^{\infty} x^{m+1} \frac{c_n}{x^{n+1}} dx = \int_{c_n}^{\infty} \frac{c_n}{x^{n-m}} dx.$$

This is convergent only of $m < n - 1$.

Exercise 3.5 The joint probability mass function of the random vector (X, Y) is

$$\mathbb{P}(X = b, Y = 2b) = \mathbb{P}(X = 2b, X = b) = \frac{1}{2}, \quad \mathbb{P}(X = b, Y = b) = \mathbb{P}(X = 2b, Y = 2b) = 0.$$

This shows that X and Y have identical distributions

$$p_X(b) = p_Y(b) = \frac{1}{2} = p_X(2b) = p_Y(2b).$$

(i) The conclusion follows from the above equalities. Indeed, X, Y have identical distributions so

$$\mathbb{E}[X] = \mathbb{E}[Y] = \frac{1}{2}b + \frac{2b}{2} = \frac{3b}{2}.$$

(ii) Using the law of the subconscious statistician we deduce

$$\mathbb{E}\left[\frac{Y}{X}\right] = \sum_{x,y} \frac{y}{x} \mathbb{P}(X = x, Y = y) = \frac{1}{2} \frac{2b}{b} + \frac{1}{2} \frac{b}{2b} = 1 + \frac{1}{4} = \frac{5}{4}.$$

(iii) We have

$$\mathbb{E}[Z] = b\mathbb{P}(Z = b) + 2b\mathbb{P}(Z = 2b).$$

From the law of total probability we deduce

$$\mathbb{P}(Z = b) = \underbrace{\mathbb{P}(Z = b|X = b)}_{=1-p(b)} \mathbb{P}(X = b) + \underbrace{\mathbb{P}(Z = b|X = 2b)}_{=p(2b)} \mathbb{P}(X = 2b) = \frac{1}{2}(1 - p(b) + p(2b)).$$

Similarly

$$\mathbb{P}(Z = 2b) = \underbrace{\mathbb{P}(Z = 2b|X = b)}_{=p(b)} \mathbb{P}(X = b) + \underbrace{\mathbb{P}(Z = 2b|X = 2b)}_{=1-p(2b)} \mathbb{P}(X = 2b) = \frac{1}{2}(1 - p(2b) + p(b)).$$

Hence

$$\begin{aligned} \mathbb{E}[Z] &= \frac{b}{2} \left((1 - p(b) + p(2b)) + 2(1 - p(2b) + p(b)) \right) \\ &= \frac{b}{2} (3 + p(b) - p(2b)) = \frac{3b}{2} + \underbrace{\frac{b}{2} \left(\frac{1}{1+e^{2b}} - \frac{1}{1+e^{4b}} \right)}_{>0} > \frac{3b}{2} = \mathbb{E}[X]. \end{aligned}$$

Exercise 3.6 As suggested, call X_1, \dots, X_4 the variables defined by $X_i = 1$ if the i th component works, $X_i = 0$ otherwise. This means that the X_i are independent Bernoulli variables; more precisely, $X_1 \sim \text{Ber}(0.9)$, $X_2 \sim \text{Ber}(0.8)$, $X_3 \sim \text{Ber}(0.6)$, $X_4 \sim \text{Ber}(0.6)$.

(i)

$$\mathbb{E}[X] = \mathbb{E}[X_1 + \dots + X_4] = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_4] = 0.9 + 0.8 + 0.6 + 0.6 = 2.9$$

(ii) Because the variables are independent,

$$\text{Var}(X) = \text{Var}(X_1 + \dots + X_4) = \text{Var}(X_1) + \dots + \text{Var}(X_4) = 0.9 \cdot 0.1 + 0.8 \cdot 0.2 + 0.6 \cdot 0.4 + 0.6 \cdot 0.4 = 0.73$$

(iii)

$$\mathbb{P}(X > 0) = 1 - \mathbb{P}(X = 0) = 1 - \mathbb{P}(X_1 = 0) \times \dots \times \mathbb{P}(X_4 = 0) = 1 - 0.9 \cdot 0.8 \cdot 0.6 \cdot 0.6 = 0.7408.$$

(iv) As far as I know, there is no clever trick for this one as there was for the ones before.

$$\begin{aligned} \mathbb{P}(X = 1) &= \mathbb{P}(\text{The first component is the only one working}) \\ &\quad + \dots + \mathbb{P}(\text{The fourth component is the only one working}) \\ &= 0.9 \cdot 0.2 \cdot 0.4 \cdot 0.4 + 0.1 \cdot 0.8 \cdot 0.4 \cdot 0.4 + 0.1 \cdot 0.2 \cdot 0.6 \cdot 0.4 + 0.1 \cdot 0.2 \cdot 0.4 \cdot 0.6 \\ &= 0.0512 \end{aligned}$$

Exercise 1

Write Y for the maximum of X_1 to X_n . It has to be between 0 and 1, so we already know that $p_Y(y) = 0$ for $y \leq 0$ or $y \geq 1$. We choose $0 < y < 1$ and compute the cumulative distribution function.

$$\begin{aligned} F_Y(y) &= \mathbb{P}(Y \leq y) = \mathbb{P}(\max(X_1, \dots, X_n) \leq y) \\ &= \mathbb{P}(X_1 \leq y \text{ and } \dots \text{ and } X_n \leq y) = \mathbb{P}(X_1 \leq y) \times \dots \times \mathbb{P}(X_n \leq y) = F_X(y)^n \end{aligned}$$

where $F_X(x)$ is the cumulative distribution function of one of the X_i (they all have the same one). Since they are uniform between 0 and 1, $F_X(x) = x$ for $0 < x < 1$, so for $0 < y < 1$ we have

$$F_Y(y) = y^n.$$

This means

$$p_Y(y) = \begin{cases} F'_Y(y) = ny^{n-1} & \text{for } 0 < y < 1 \\ 0 & \text{else} \end{cases}$$

This is a Beta distribution $\text{Beta}(n, 1)$.

Exercise 2

1. We must have $0 \leq X < Y$.
2. Fix x and y such that $0 \leq x < y$. Then

$$\begin{aligned} \mathbb{P}(X = x, Y = y) &= \mathbb{P}(\text{We do } y - 1 \text{ throws giving } 1, 2, 3, 4 \text{ or } 5, \text{ within which } x \text{ 1's, then a 6.}) \\ &= \frac{\#\{\text{possible positions for the 1's}\} \cdot \#\{\text{possible outcomes for the remaining throws}\}}{6^y} \\ &= \frac{\binom{y-1}{x} \cdot 4^{y-1-x}}{6^y} \end{aligned}$$