

## STATS 406F15 Lab 05

### 1 Review of simulating Discrete random variables via inversion

Discrete random variables can be generated by slicing up the interval  $(0, 1)$  into subintervals which define a partition of  $(0, 1)$ :

$$(0, F(x_1)), (F(x_1), F(x_2)), (F(x_2), F(x_3)), \dots, (F(x_{k-1}), 1)$$

and generating  $U = \text{Uniform}(0, 1)$  random variables, and seeing which subinterval  $U$  falls into.

Write a function to generate  $n$  random numbers from binomial distribution with  $m$  trials and  $p$  using inversion sampling. The mass function of binomial distribution is

$$P(X = k) = \binom{m}{k} p^k (1 - p)^{m-k}$$

where  $k = 0, 1, \dots, n$ . We set  $n = 1000$ ,  $p = 0.2$ . Plot the histograms with  $m = 5$  and  $m = 50$ , respectively.

#### Solution:

```
## Function to implement the inversion sampling
## for binomial distribution
binomial <- function(n, num_trials, prop)
{
  ## Generate n uninform numbers
  z <- runif(n)
  ## Obtain P(X=0), P(X=1), ..., P(X=num_trials)
  p <- choose(num_trials, seq(0, num_trials))
  * prop^seq(0, num_trials) * (1-prop)^(num_trials-seq(0, num_trials))
  ## Intialize the binomial numbers
  x <- rep(0, n)
  for (i in seq(1, n))
  {
    s <- 0 ## Initialize the sum
```

```

    k <- -1 ## Initialize
    ## While loop
    ## Finding the smallest k such that p[1]+...p[k] > z[i]
    while (s < z[i])
    {
        k <- k + 1
        s <- s + p[k+1]
    }
    x[i] <- k
}

return(x)
}

## Call binomial generator
x <- binomial(n=1000, num_trials=50, prop=0.2)
hist(x)

```

## 2 Basic Monte Carlo Integration

### 2.1 Moments of functions of random variables

Example: Let  $U$  and  $V$  be independent Uniform  $(-1,1)$ . Compute  $P(|U + V| \leq 1)$  using Monte Carlo method. Report the monte carlo error.

Solution:  $P(|U + V| \leq 1) = E(1(|U + V| \leq 1)) = \int_{-1}^1 \int_{-1}^1 1(|u + v| \leq 1) du dv$ . Estimate this by generating  $U_1, U_2, \dots, U_N$  and  $V_1, V_2, \dots, V_N$  from the uniform distribution on  $(-1,1)$  and calculating

$$\frac{1}{N} \sum_{i=1}^N 1(|U_i + V_i| \leq 1).$$

```
rm(list=ls(all=TRUE)) ## clear out old variables from the memory
```

```
## The input parameter is #samples and level of significance
```

```

integral <- function(n, alpha)
{
    u1 <- runif(n,min=-1, max =1)
    u2 <- runif(n,min=-1, max =1)
    h<- (abs(u1 + u2) <= 1)
    integral <- mean(h)
    mc_error <- sqrt(var(h) / n)
    z_alphaby2 <- qnorm(1- alpha/2)

```

```

ci <- c( integral - z_alphaby2*mc_error, integral + z_alphaby2*mc_error)
  output <- list(integral=integral, mc_error=mc_error, ci = ci )
  return(output)
}

print(integral(10000,.05))

```

## 2.2 Calculating arbitrary integrals

The integral of an arbitrary function  $h$  over the interval  $(a, b)$  is :

$$\int_a^b h(x)dx = (b-a) \int_a^b h(x) \frac{1}{b-a} dx = (b-a)E(h(V)),$$

where  $V$  follows Uniform( $a, b$ ).

Example: Compute  $\int_0^{2\pi} \sin(x \cos(x))dx$ .

Solution: This is equal to  $2\pi E(\sin(V \cos(V)))$  where  $V \sim U(0, 2\pi)$

```

rm(list=ls(all=TRUE))

## The input parameter is #samples
integral1 <- function(n)
{
  x <- runif(n=n, min=0, max=2*pi)
  integral <- mean(2*pi*sin(x*cos(x)))
  mc_error <- sqrt(var(2*pi*sin(x*cos(x))) / n)

  output <- list(integral=integral, mc_error=mc_error)
  return(output)
}

```

Example: Consider again  $h(x) = \sin(x \cos(x))$ . Compute the integral

$$\int_{-\infty}^{\infty} h(x)dx.$$

Solution:

```

#### Computation of the second integral
## The input parameter is #samples
integral2 <- function(n)
{
  x <- rnorm(n)

```

```

integral <- mean(sin(x*cos(x)) / dnorm(x))
mc_error <- sqrt(var(sin(x*cos(x)) / dnorm(x)) / n)

output <- list(integral=integral, mc_error=mc_error)
return(output)
}

print(integral2(10000))

```

### 3 Rejection Sampling

Write a function to generate  $n$  random numbers from standard normal distribution using rejection sampling. The density function of standard normal is:

$$f(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right)$$

Using Cauchy distribution as the trial (reference) distribution:

$$g(x) = \frac{1}{\pi(1+x^2)}$$

#### Solution:

Consider the ratio:

$$h(x) = \frac{f(x)}{g(x)} = \sqrt{\pi/2}(1+x^2) \exp\left(-\frac{x^2}{2}\right)$$

Setting  $h'(x) = 0$ , we can obtain  $x = \{-1, 0, +1\}$ , where  $x = -1, +1$  correspond to the two peaks. Then we set  $M = h(1) = \sqrt{2\pi} \exp(-1/2)$ .

```

## Density function of standard normal distribution
f <- function(x)
{
  y <- 1/sqrt(2*pi) * exp(-x^2/2)
  return(y)
}

## Density function of Cauchy distribution
g <- function(x)
{
  y <- 1 / (pi * (1 + x^2))
  return(y)
}

```

```

## Rejection sampling for standard normal distribution
rejection_sampling_standard_normal <- function(n)
{
  ## Boundary M
  M <- f(1)/g(1)

  ## Initialize the output
  x <- rep(0, n)

  k <- 0 ## Number of accepted samples
  ## While loop until there are n accepted samples
  while (k <= n)
  {
    y <- rcauchy(1)
    u <- runif(1)
    if (u*M*g(y) <= f(y))
    {
      k <- k + 1
      x[k] <- y
    }
  }

  return(x)
}

## Call function
x <- rejection_sampling_standard_normal(1000)
hist(x, breaks=20)

```