

# Monte Carlo Methods: Monte Carlo Integration

# Monte Carlo integration

- ▶ Let  $\pi$  be a density on  $\mathbb{R}^d$ . For example if  $d = 1$ ,  $\pi$  is a density on  $\mathbb{R}$ .
- ▶ Let  $h : \mathbb{R}^d \rightarrow \mathbb{R}$  a function. Suppose we want to evaluate the integral  $\pi(h) = \int h(x)\pi(x)dx = \mathbb{E}(h(X))$ .
- ▶ For instance  $h$  could be the function  $h(x_1, \dots, x_d) = x_1$ . In which case we want to compute

$$\pi(h) = \int x_1 \pi(x) dx = \mathbb{E}(X_1),$$

the expectation of the first component of  $X$ .

# Monte Carlo integration

- ▶ Very often, such integral can be hard to evaluate analytically.
- ▶ If  $d$  is not too large, we can use numerical methods (simpson, trapezoid rules, quadrature methods etc...).
- ▶ But for  $d$  large ( $d \geq 4$ ), numerical methods become inefficient and we have to turn to Monte Carlo methods.

# Monte Carlo integration

- ▶ The basic idea of Monte Carlo integration is very simple: generate  $X_1, \dots, X_n$  i.i.d. from  $\pi$ .
- ▶ By the LLN  $n^{-1} \sum_{k=1}^n h(X_k)$  converge to  $E(h(X))$  (we will use the notation  $\pi(h)$  to denote  $E(h(X))$ ).
- ▶ In other words, we approximate the integral  $\pi(h) = E(h(X))$  by

$$\pi_n(h) = \frac{1}{n} \sum_{k=1}^n h(X_k).$$

- ▶  $\pi_n(h)$  is called the **Monte Carlo estimate** of  $\pi(h)$  or the Basic Monte Carlo estimate of  $\pi(h)$ .

# Monte Carlo integration

$\pi_n(h)$  has the following properties

## Theorem

1. *By the Law of large numbers,  $\pi_n(h)$  converges to  $\pi(h)$  as  $n \rightarrow \infty$ .*
2.  $\mathbb{E}(\pi_n(h)) = \pi(h)$  and  $\text{Var}(\pi_n(h)) = \frac{\sigma^2(h)}{n}$  where  $\sigma^2(h) = \text{Var}(h(X)) = \mathbb{E}(h^2(X)) - (\mathbb{E}(h(X)))^2$ .
3. *By the central limit theorem*

$$\pi_n(h) \approx \mathcal{N}\left(\pi(h), \frac{\sigma^2(h)}{n}\right).$$

# Monte Carlo integration

- ▶ According to the CLT, the fluctuations or precision of  $\pi_n(h)$  is given by  $\frac{\sigma^2(h)}{n}$  or its square root  $\frac{\sigma(h)}{\sqrt{n}}$ .
- ▶ We can estimate  $\sigma^2(h)$  by

$$s_n^2(h) = \frac{1}{n-1} \sum_{k=1}^n (h(X_k) - \pi_n(h))^2.$$

- ▶ The quantity  $\frac{s_n(h)}{\sqrt{n}}$  is known as the **Monte Carlo error**.
- ▶ In any Monte Carlo experiment, always report the Monte Carlo error or the Monte Carlo confidence interval

$$\pi_n(h) \pm z_{\alpha/2} \frac{s_n(h)}{\sqrt{n}}.$$

# Monte Carlo integration

## Example

- ▶ Suppose we want to calculate the integral

$$I = \int_0^1 (\cos(50x) + \sin(20x))^2 dx.$$

- ▶ It is possible to do this analytically. The exact value is 0.965.
- ▶ We can also approximate  $I$  very well using numerical integration methods (function `integrate` in R)
- ▶ To use Monte Carlo, we could see  $I$  as  $\mathbb{E}(h(U))$  where  $h(x) = (\cos(50x) + \sin(20x))^2$  and  $U \sim \mathcal{U}(0, 1)$ . Then we generate  $U_1, \dots, U_n$  i.i.d. from  $\mathcal{U}(0, 1)$  and approximate  $I$  by  $\hat{I}_n = n^{-1} \sum_{k=1}^n h(U_k)$ .

# Monte Carlo integration

```
hfun=function(x){  
  return((cos(50*x)+sin(20*x))^2)  
}  
  
nSple=10000  
uniforms=runif(nSple)  
hvalues=hfun(uniforms)  
Ihat=mean(hvalues) #Monte Carlo estimate  
se=sqrt(var(hvalues)/nSple) # Monte Carlo error  
z=qnorm(0.975)  
CI=c(Ihat-z*se,Ihat+z*se) # 95% confid. interv.  
##We report either Ihat and se, or CI.
```



# Monte Carlo integration

- ▶ An important step to apply the Monte Carlo technique to an integral  $I = \int v(x)dx$  is to rewrite  $I$  as an expectation.
- ▶ This step requires some care because it can be done in many different ways. For instance, suppose we want to approximate  $I = \int_0^1 6x^2(1-x)dx$ .
- ▶ We can write  $I = \mathbb{E}(h_1(U))$ , where  $h_1(x) = 6x^2(1-x)$  and  $U \sim \mathbf{U}(0,1)$ .
- ▶ But we could also write  $I = \mathbb{E}(h_2(V))$ , where  $h_2(x) = x$ , and  $V \sim \text{Beta}(2,2)$ , the Beta distribution with parameter  $(2,2)$ .

# Monte Carlo integration

- ▶ Another example. Consider  $I = \int_0^1 (\cos(50x) + \sin(20x))^2 dx$  seen above. We could have also used the Beta distribution. How?
- ▶ Let  $f_{\alpha,\beta}(x) = \frac{1}{\text{Beta}(\alpha,\beta)} x^{\alpha-1} (1-x)^{\beta-1}$ ,  $0 < x < 1$ , be the density of the Beta distribution. We can write.

$$I = \int_0^1 \frac{(\cos(50x) + \sin(20x))^2}{f_{\alpha,\beta}(x)} f_{\alpha,\beta}(x) dx.$$

- ▶ We have re-written the same integral as a different expectation  $\mathbb{E}(h(X))$ , where here  $X \sim \text{Beta}(\alpha, \beta)$ , and

$$h(x) = \frac{(\cos(50x) + \sin(20x))^2}{f_{\alpha,\beta}(x)}.$$

# Monte Carlo integration

- ▶ The idea of moving from one expectation representation to another, for the same integral, is known as **importance sampling**.
- ▶ We develop this in the next section.

# Importance sampling

- ▶ We have discussed the Basic Monte idea to approximate an integral  $\int h(x)\pi(x)dx$  where  $\pi$  is a density.
- ▶ We generate  $X_1, \dots, X_n$  i.i.d. from  $\pi$  and compute  $n^{-1} \sum_{k=1}^n h(X_k)$ .
- ▶ But we can write an integral in many different ways. Let  $g$  be another density such that  $g(x) > 0$  whenever  $h(x)\pi(x) > 0$ . Then write

$$\int h(x)\pi(x)dx = \int h(x)\frac{\pi(x)}{g(x)}g(x)dx = \int h(x)\omega(x)g(x)dx,$$

where  $\omega(x) = \frac{\pi(x)}{g(x)}$ . and  $g$  another density.

- ▶ Now we can apply the Basic Monte Carlo and sample from  $g$ .

# Importance sampling

## How does it work?

To evaluate  $\pi(h) = \int h(x)\pi(x)dx$  we introduce **another density**  $g$  and define the **importance ratio**  $\omega(x) = \frac{\pi(x)}{g(x)}$ :

$$\pi(h) = \int h(x)\pi(x)dx = \int h(x)\omega(x)g(x)dx.$$

- ▶ Then we generate  $Y_1, \dots, Y_n$  i.i.d. from  $g$  and approximate  $\pi(h)$  by

$$\pi_{n,IS}(h) = \frac{1}{n} \sum_{k=1}^n h(Y_k)\omega(Y_k).$$

# Importance sampling

- ▶ As in Basic Monte Carlo

$$\frac{1}{n} \sum_{k=1}^n h(Y_k) \omega(Y_k) \approx \mathcal{N} \left( \pi(h), \frac{\sigma_{IS}^2(h)}{n} \right),$$

where  $\sigma_{IS}^2(h) = \text{Var}(h(Y)\omega(Y))$ , and we can estimate  $\sigma_{IS}^2(h)$  by

$$s_{n,IS}^2 = \frac{1}{n-1} \sum_{k=1}^n (h(Y_k) \omega(Y_k) - \pi_{n,IS}(h))^2.$$

- ▶ So again we either report  $\pi_{n,IS}(h)$  and  $\sqrt{s_{n,IS}^2(h)/n}$  or the confidence interval  $\pi_{n,IS}(h) \pm z_{\alpha/2} \sqrt{s_{n,IS}^2(h)/n}$ .

# Importance sampling

How do we choose  $g$ ?

- ▶ Coverage issue: if  $\pi(x) > 0$  for some  $x$ , make sure  $g(x) > 0$ .  
For example if  $\pi(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$   $x \in \mathbb{R}$  don't use  $g(x) = e^{-x}$  for  $x > 0$  and  $g(x) = 0$  otherwise.
- ▶ Resemblance: make sure  $g$  is close to  $|h(x)|\pi(x)$ .
- ▶ Stability: make sure the function  $\omega(x)$  is bounded:  
 $\pi(x) \leq Mg(x)$  for some finite  $M > 0$ .

# Importance sampling

Exercise: Recall that  $\sigma^2(h) = \text{Var}(h(X))$  (where  $X \sim \pi$ ). Suppose that  $\sigma^2(h) < \infty$  and we choose  $g$  such that  $\pi(x) \leq Mg(x)$  for some  $M < \infty$ . Show that  $\text{Var}(h(Y)\omega(Y))$  (where  $Y \sim g$ ) is also finite.



# Importance sampling

Example: Compute  $p = \Pr(Z \geq a)$  for  $a = 2.3$ , where  $Z \sim N(0, 1)$ .

- ▶ We can write

$$p = \int_{-\infty}^{+\infty} h(x)\pi(x)dx,$$

where  $\pi$  is the density of  $N(0, 1)$ , and

$$h(x) = \begin{cases} 1 & \text{if } x \in [a, \infty), \\ 0 & \text{otherwise.} \end{cases}$$

- ▶ Basic Monte Carlo: Generate  $Z_1, \dots, Z_n \sim N(0, 1)$  and form

$$\hat{p} = \frac{1}{n} \sum_{i=1}^n h(X_i),$$

# Importance sampling

```
hfun =function(x,a) (x>=a)
##
a=2.3
n=20000
z=qnorm(0.975)
answer=1-pnorm(a)
#Estimate p by Basic Monte
Z=rnorm(n);
hX=hfun(Z,a)
p_hat=mean(hX);MC_err=sqrt(var(hX)/n)
CI=c(p_hat-z*MC_err,p_hat+z*MC_err)
```

# Importance sampling

We could also use importance sampling by generating from  $g_\alpha(x) = \alpha e^{-\alpha(x-a)}$ , for  $x \geq a$ .

$$p = \int_a^{+\infty} \pi(x) dx = \int_a^{+\infty} \frac{\pi(x)}{g_\alpha(x)} g_\alpha(x) dx = \int_a^{+\infty} \omega(x) g_\alpha(x) dx.$$

where

$$\omega(x) = \frac{1}{\alpha\sqrt{2\pi}} e^{\alpha(x-a) - \frac{1}{2}x^2}.$$

```
omfun =function(x,a,alpha){  
y=(1/sqrt(2*pi))*(1/alpha)  
      *exp(-0.5*x^2+alpha*(x-a))  
return(y)  
}
```

# Importance sampling

```
#Estimate by Importance Smpling  
alpha=a;  
Y=a + rexp(n,alpha)  
hY=hfun(Y,a)*omfun(Y,a,alpha)  
p_hat2=mean(hY);  
MC_err2=sqrt(var(hY)/n)  
CI=c(p_hat2-z*MC_err2,p_hat2+z*MC_err2)
```

# Importance sampling

- ▶ Consider again importance sampling for calculating  $\pi(h) = \int h(x)\pi(x)dx$ :

$$\pi(h) = \int h(x)\pi(x)dx = \int h(x)\omega(x)g(x)dx,$$

where  $\omega(x) = \pi(x)/g(x)$ .

- ▶ In many applications of IS the importance ratio  $\omega$  cannot be computed. The prime example where this happens is in [Bayesian inference](#).
- ▶ Often all that can be computed is  $\tilde{\omega}(x) = C\omega(x)$  for some constant  $C$ .
- ▶ In these cases we do IS as follows: generate  $Y_1, \dots, Y_n$  from  $g$  and form

$$\tilde{\pi}_{n,IS} = \frac{\sum_{k=1}^n h(Y_k)\tilde{\omega}(Y_k)}{\sum_{k=1}^n \tilde{\omega}(Y_k)}.$$

# Importance sampling

- Notice that

$$\begin{aligned}\tilde{\pi}_{n,IS} &= \frac{\sum_{k=1}^n h(Y_k) \tilde{\omega}(Y_k)}{\sum_{k=1}^n \tilde{\omega}(Y_k)} \\ &= \frac{\frac{1}{n} \sum_{k=1}^n h(Y_k) \omega(Y_k)}{\frac{1}{n} \sum_{k=1}^n \omega(Y_k)} \\ &\rightarrow \frac{\int h(x) \pi(x) dx}{1}.\end{aligned}$$

- Hence it is also correct to use this estimate.

# Importance sampling

- ▶ But the Monte Carlo error for the new estimate is more difficult to evaluate.
- ▶ Notice that the quantity  $n^{-1} \sum_{k=1}^n \tilde{\omega}(Y_k)$  converges to  $\int C\omega(x)g(x)dx = C$ .
- ▶ Thus we can assess the estimate  $\tilde{\pi}_{n,IS}$  by looking how well the estimation of  $C$  is doing. To do so we compute the sample coefficient of variation:

$$CV = \sqrt{\frac{1}{n-1} \sum_{k=1}^n \left( \frac{\tilde{\omega}(Y_k)}{\bar{\omega}} - 1 \right)^2},$$

where  $\bar{\omega} = \frac{1}{n} \sum_{k=1}^n \tilde{\omega}(Y_k)$ .

# Importance sampling

Take home method:

- Calculate the fluctuation of the weights:

$$CV = \sqrt{\frac{1}{n-1} \sum_{k=1}^n \left( \frac{\tilde{\omega}(Y_k)}{\bar{\omega}} - 1 \right)^2},$$

where  $\bar{\omega} = \frac{1}{n} \sum_{k=1}^n \tilde{\omega}(Y_k)$ .

- If  $CV$  is small (say smaller than 5) then compute  $\tilde{\pi}_{n,IS}$  and its Monte Carlo error  $\sqrt{\hat{\sigma}_{n,IS}^2/n}$ .

$$\hat{\sigma}_{n,IS}^2 = \frac{1}{n-1} \sum_{k=1}^n (Z_{n,k} - \tilde{\pi}_{n,IS})^2, \quad Z_{n,k} = \frac{\tilde{\omega}(Y_k)}{\frac{1}{n} \sum_{k=1}^n \tilde{\omega}(Y_k)} h(Y_k).$$

- Otherwise  $\tilde{\pi}_{n,IS}$  is too unstable and should not be trusted.



# Importance sampling

## Example:

- ▶ We want to sample from the distribution

$$\pi(x) = \frac{1}{C} \frac{1}{1+x^2} e^{-(x-2)^2/2} \quad (1)$$

and calculate  $\delta = \int x\pi(x)dx$ .

- ▶  $C$  is called the normalizing constant.  $C = \int \frac{1}{1+x^2} e^{-(x-2)^2/2} dx$  and is unknown.
- ▶ It is thus important to have Monte Carlo methods where the knowledge of  $C$  is not required like the **weighted average importance sampling**.

# Importance sampling

- ▶ We use importance sampling with the Cauchy distribution  $\mathcal{C}(1, 1)$ .

$$g(x) = \frac{1}{c_0(1 + (x - 1)^2)},$$

where  $c_0 = \pi$  (the irrational number).

- ▶ We can plot the functions  $\tilde{\pi}(x) = \frac{1}{1+x^2}e^{-(x-2)^2/2}$  and  $g(x) = \frac{1}{c_0(1+(x-1)^2)}$  to see how close they are.
- ▶ We need the importance function

$$\omega(x) = \frac{\pi(x)}{g(x)} = \frac{c_0}{C} \frac{1 + (x - 1)^2}{1 + x^2} e^{-\frac{1}{2}(x-2)^2}.$$

# Importance sampling

- ▶ To do importance sampling, we generate  $Y_1, \dots, Y_n$  iid from  $\mathcal{C}(1, 1)$  and if we knew  $C$  we can estimate  $\delta$  by:

$$\frac{1}{n} \sum_{i=1}^n Y_k \omega(Y_k).$$

- ▶ Clearly, the problem is that  $C$  is not known!
- ▶ This is where the weighted average version becomes useful.

# Importance sampling

- ▶ We drop  $C$  (and  $c_0$ ). The new importance function is

$$\tilde{\omega}(x) = \frac{1 + (x - 1)^2}{1 + x^2} e^{-\frac{1}{2}(x-2)^2}.$$

- ▶ The new importance sampling: We sample  $Y_1, \dots, Y_n$  iid from  $\mathcal{C}(1, 1)$  and estimate  $\delta$  by:

$$\frac{\sum_{k=1}^n Y_k \tilde{\omega}(Y_k)}{\sum_{i=1}^n \tilde{\omega}(Y_i)}.$$

# Importance sampling

To implement this method in R we need to

- ▶ write a function *omegafun* that calculates the function  $\tilde{\omega}$
- ▶ know how to generate from  $\mathcal{C}(1, 1)$  (`rcauchy(n,mu,sigma)`).
- ▶ That's it.

# Importance sampling

```
omfun=function(x){  
  v=exp(-(2-x)^2/2)*(1+(x-1)^2)/(1+x^2)  
}
```

# Importance sampling

```
n=10000
Sple=1+rcauchy(n)
Om=omfun(Sple)
est=sum(Sple*Om)/sum(Om)# estimate
CVsq=var(Om/mean(Om))
v1=var((Om/mean(Om))*Sple)
se=sqrt(v1/n)
z=qnorm(0.975)
CI=(est-z*se,est+z*se)
```

Question: Use the function `integrate` to confirm the answer we obtained.