

MULTIVARIATE DISTRIBUTIONS

Covariances and correlations

Covariances of random variables

- The *covariance* between two random variables is a measure of how much the two tend to vary together.
- **Note:** Covariance is concerned only with “linear” variation.
- Definition 6: If Y_1 and Y_2 are random variables with means μ_1 and μ_2 , respectively, then the COVARIANCE of Y_1 and Y_2 is:

$$\text{Cov}(Y_1, Y_2) = E[(Y_1 - \mu_1)(Y_2 - \mu_2)]$$

$$E[Y_1] = \mu_1$$
$$E[Y_2] = \mu_2$$

Positive covariance, negative covariance, no covariance

- If $Cov(Y_1, Y_2) > 0$ then the variables are positively (linearly) related (Large values of Y_1 are associated with large values of Y_2).
- If $Cov(Y_1, Y_2) < 0$ then the variables are negatively (linearly) related (Large values of Y_1 are associated with small values of Y_2 and vice-versa).
- If $Cov(Y_1, Y_2) = 0$ then there is no (linear) association between Y_1 and Y_2 .

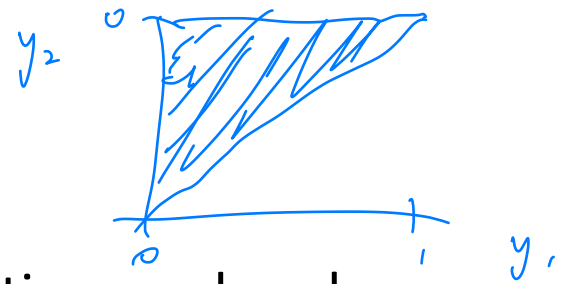
- Q: What is $Cov(Y_1, Y_1)$?

$$E[(Y_1 - \mu_1)(Y_1 - \mu_1)] = E[(Y_1 - \mu_1)^2] \\ = Var(Y_1)$$

$$Var(Y) = E[Y^2] - (E[Y])^2$$

Covariances of random variables: An alternative expression

- Theorem 10: For any random variables Y_1 and Y_2 ,
$$Cov(Y_1, Y_2) = \underbrace{E[Y_1 Y_2]} - \underbrace{E[Y_1]} \underbrace{E[Y_2]}$$



Example 7 re-revisited: Y_1 and Y_2 are both proportions and we know that $Y_1 < Y_2$. If their joint pdf is $f(y_1, y_2) = \underline{6y_1}$ for $0 < y_1 < y_2 < 1$ (and zero otherwise), find $\underline{Cov(Y_1, Y_2)}$.

$$E[Y_1] = \int_0^1 \int_0^{y_2} y_1 \cdot 6y_1 \, dy_1 \, dy_2 = \frac{1}{2}$$

$$E[Y_2] = \int_0^1 \int_0^{y_2} y_2 \cdot 6y_1 \, dy_1 \, dy_2 = \frac{3}{4}$$

$$\begin{aligned} E[Y_1 Y_2] &= \int_0^1 \int_0^{y_2} y_1 y_2 \cdot 6y_1 \, dy_1 \, dy_2 = \int_0^1 \int_0^{y_2} 6y_1^2 y_2 \, dy_1 \, dy_2 \\ &= \int_0^1 y_2 \left(2y_1^3 \Big|_0^{y_2} \right) dy_2 = \int_0^1 y_2 \cdot 2y_2^3 \, dy_2 = \int_0^1 2y_2^4 \, dy_2 = \frac{2}{5} y_2^5 \Big|_0^1 = \frac{2}{5} \end{aligned}$$

$$Cov(Y_1, Y_2) = \frac{2}{5} - \frac{1}{2} \cdot \frac{3}{4} = \frac{2}{5} - \frac{3}{8} = \frac{16 - 15}{40} = \frac{1}{40}$$

Independence and covariance

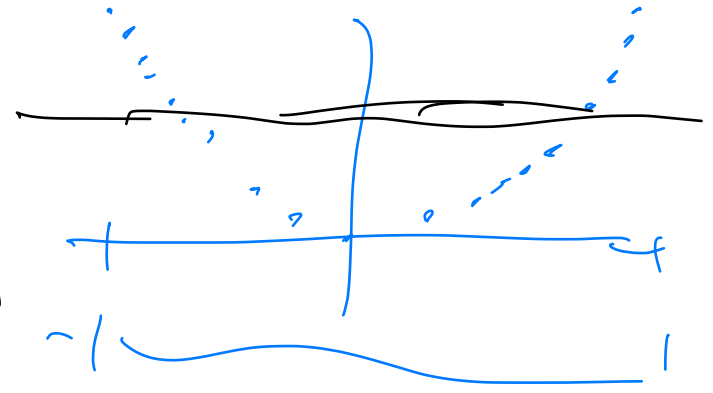
- Theorem 11 (independence and covariance): If Y_1 and Y_2 are independent random variables, then $\text{Cov}(Y_1, Y_2) = 0$.

$$\begin{aligned}\text{Cov}(Y_1, Y_2) &= \overbrace{E[Y_1 Y_2] - E[Y_1]E[Y_2]} \\ &= \overbrace{E[Y_1]E[Y_2] - E[Y_1]E[Y_2]} = 0\end{aligned}$$

because
indep.

$$Y_1 \& Y_2 \text{ are indep.} \Rightarrow \text{Cov}(Y_1, Y_2) = 0$$

$$\text{Cov}(Y_1, Y_2) = 0 \not\Rightarrow Y_1 \& Y_2 \text{ indep.}?$$



Covariance and independence

$$f_Y(y_1) = \begin{cases} \frac{1}{2} & -1 \leq y_1 \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

• **Note:** The converse (of the previous theorem) is not true!

• Example 10: $Y_1 \sim U(-1,1)$ and $Y_2 = Y_1^2$. Find $Cov(Y_1, Y_2)$ and determine whether Y_1 and Y_2 are independent.

$$Cov(Y_1, Y_2) = E[Y_1 Y_2] - E[Y_1] E[Y_2]$$

$$E[Y_1 Y_2] = E[Y_1 \cdot Y_1^2] = E[Y_1^3] = \int_{-1}^1 y_1^3 \cdot \frac{1}{2} dy_1 = \frac{1}{8} y_1^4 \Big|_{-1}^1 = \frac{1}{8} (1 - 1) = 0$$

$$E[Y_1] = \int_{-1}^1 y_1 \cdot \frac{1}{2} dy_1 = \frac{1}{4} y_1^2 \Big|_{-1}^1 = \frac{1}{4} (1 - 1) = 0$$

$$Cov(Y_1, Y_2) = 0 - 0 E[Y_2] = 0$$

Correlations of random variables

- The **covariance** will change if either Y_1 or Y_2 is expressed in different units. (Check this on your own.)
- Definition 7: For two random variables Y_1 and Y_2 , the **CORRELATION COEFFICIENT** between Y_1 and Y_2 is

$$\text{Corr}(Y_1, Y_2) = \rho_{Y_1, Y_2} = \frac{\text{Cov}(Y_1, Y_2)}{\sqrt{\text{Var}(Y_1)\text{Var}(Y_2)}}$$

- Theorem 12: For any random variables Y_1 and Y_2 ,
$$-1 \leq \text{Corr}(Y_1, Y_2) \leq 1$$
- A correlation near 1 means a strong positive (linear) relationship between Y_1 and Y_2 ; near -1 means a strong negative (linear) relationship.

“Perfect” correlations

- Theorem 13: For random variables Y_1 and Y_2 , $|Corr(Y_1, Y_2)| = 1$ if and only if there are numbers $a \neq 0$ and b such that $Y_2 = aY_1 + b$. If $a > 0$ then $Corr(Y_1, Y_2) = 1$; if $a < 0$ then $Corr(Y_1, Y_2) = -1$.

Linear functions of random variables

- In statistical analysis, we often use statistics that are linear functions of multiple random variables (data) Y_1, Y_2, \dots, Y_n , i.e., functions that can be expressed as $\sum_{i=1}^n a_i Y_i = a_1 Y_1 + a_2 Y_2 + \dots + a_n Y_n$ for constants a_1, a_2, \dots, a_n .

Linear functions of random variables

- Theorem 14: Given sequences of random variables X_1, X_2, \dots, X_m and Y_1, Y_2, \dots, Y_n with $E[X_j^2] < \infty$ for all j and $E[Y_i^2] < \infty$ for all i , define $U_1 = \sum_{i=1}^n a_i Y_i$ and $U_2 = \sum_{j=1}^m b_j X_j$ for constants a_1, a_2, \dots, a_n and b_1, b_2, \dots, b_m . Then

$$1. E[U_1] = \sum_{i=1}^n a_i E[Y_i] = a_1 E[Y_1] + a_2 E[Y_2] + \dots + a_n E[Y_n]$$

$$2. Var(U_1) = \sum_{i=1}^n a_i^2 Var(Y_i) + 2 \sum \sum_{i < j} a_i a_j Cov(Y_i, Y_j)$$

$$3. Cov(U_1, U_2) = \sum_{i=1}^n \sum_{j=1}^m a_i b_j Cov(Y_i, X_j)$$

Some "special case" corollaries:

$a_1 = 1 \quad a_2 = 1$

i. $\text{Var}(Y_1 + Y_2) = \text{Var}(Y_1) + \text{Var}(Y_2) + 2\text{Cov}(Y_1, Y_2)$

ii. $\text{Var}(Y_1 - Y_2) = \text{Var}(Y_1) + \text{Var}(Y_2) - 2\text{Cov}(Y_1, Y_2)$

iii. $\text{Var}(aY_1 + bY_2) = a^2\text{Var}(Y_1) + b^2\text{Var}(Y_2) + 2ab\text{Cov}(Y_1, Y_2)$

iv. $\text{Cov}(aW + bX, cY + dZ) = ac\text{Cov}(W, Y) + ad\text{Cov}(W, Z) + bc\text{Cov}(X, Y) + bd\text{Cov}(X, Z)$

def. of Var

Proof (ii)

$$\begin{aligned} \text{Var}(Y_1 - Y_2) &= E[(Y_1 - Y_2 - E[Y_1 - Y_2])^2] \\ &= E[(Y_1 - Y_2 - E[Y_1] + E[Y_2])^2] \end{aligned}$$

$$= E[(Y_1 - E[Y_1]) - (Y_2 - E[Y_2])]^2$$

$$= E[(Y_1 - E[Y_1])^2 + (Y_2 - E[Y_2])^2 - 2(Y_1 - E[Y_1])(Y_2 - E[Y_2])]$$

$$= E[(Y_1 - E[Y_1])^2] + E[(Y_2 - E[Y_2])^2] - 2E[(Y_1 - E[Y_1])(Y_2 - E[Y_2])]$$

$$= \text{Var}(Y_1) + \text{Var}(Y_2) - 2\text{Cov}(Y_1, Y_2) \quad \square$$

Linear functions of independent random variables

$$\Rightarrow \text{Cov}(Y_1, Y_2) = 0$$

• Corollary: If Y_1 and Y_2 are independent random variables, then

$$1. \text{Var}(Y_1 + Y_2) = \text{Var}(Y_1) + \text{Var}(Y_2)$$

$$2. \text{Var}(Y_1 - Y_2) = \text{Var}(Y_1) + \text{Var}(Y_2)$$

$$3. \text{Var}(aY_1 + bY_2) = a^2\text{Var}(Y_1) + b^2\text{Var}(Y_2)$$

Already done: $E[Y_1] = \frac{1}{2}$ $E[Y_2] = \frac{3}{4}$ $Cov(Y_1, Y_2) = \frac{1}{40}$

$$Var(Y_1) = E[Y_1^2] - (E[Y_1])^2 = \frac{1}{20}$$

Example 7 re-revisited again: Y_1 and Y_2 are both proportions and we know that $Y_1 < Y_2$. If their joint pdf is $f(y_1, y_2) = 6y_1$ for $0 < y_1 < y_2 < 1$ (and zero otherwise), Find $Var(Y_2 - Y_1)$, $Cov(Y_1, Y_2 - Y_1)$, and $Cov(Y_1 + Y_2, Y_2 - Y_1)$.

$$\textcircled{1} Var(Y_2 - Y_1) = Var(Y_2) + Var(Y_1) - 2Cov(Y_1, Y_2)$$

$$= \frac{3}{80} + \frac{1}{20} - 2 \cdot \frac{1}{40} = \frac{3}{80}$$

$$\textcircled{2} Cov(Y_1, Y_2 - Y_1) = Cov(Y_1, Y_2) - Cov(Y_1, Y_1) = \frac{1}{40} - \frac{1}{20} = -\frac{1}{40}$$

$$\textcircled{3} Cov(Y_1 + Y_2, Y_2 - Y_1) = \cancel{Cov(Y_1, Y_2)} - \overbrace{Cov(Y_1, Y_1)}^{Var(Y_1)} + \overbrace{Cov(Y_2, Y_2)}^{Var(Y_2)} - \cancel{Cov(Y_2, Y_1)}$$

$$= -\frac{1}{20} + \frac{3}{80} = -\frac{1}{80}$$

