

# ESTIMATION

Point estimation

# Goal of statistical inference

events

- The goal of statistical inference is to use information from observed sample data to make statements about some population of interest.
- We will choose some probability distribution to be a model for the data, and then use the data to estimate the unknown parameters of the distribution.

# Modeling examples

What are the models and parameters in each of these situations?

1. To estimate the effectiveness of a drug we might sample  $n$  patients, administer the drug to each one, and count the number of patients who respond to the treatment.
2. To assess the risk of suicide, we might record the weekly number of attempts among some population.
3. To study the blood pressure of low birth weight infants, we might sample some number of such infants and measure the systolic blood pressure of each.

## Two kinds of estimates

7.14  
 ~~$\mu$~~   $\hat{p}$

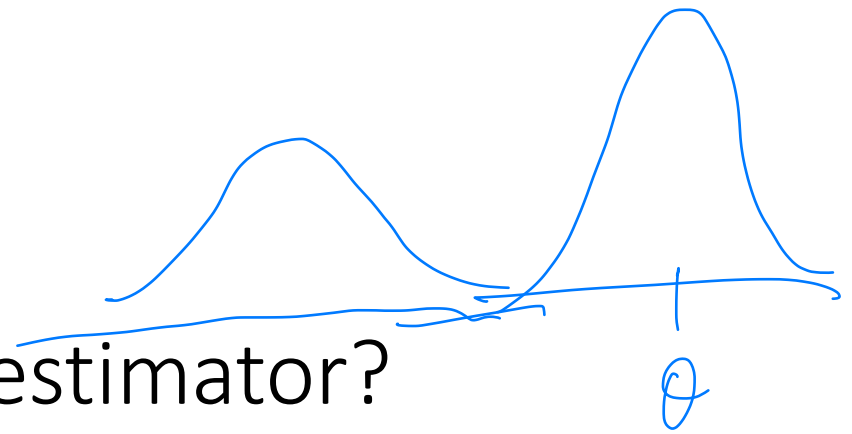
- Definition 1: A POINT ESTIMATE is a single number. It may be thought to represent the “best guess” of the unknown parameter.
- Definition 2: An INTERVAL ESTIMATE consists of a range of numbers, thought to be the most “plausible” values of the unknown parameter.

# Estimator vs. estimate

$\bar{y}$ ,  $s^2$

- An estimator is a random variable, a formula, a function of the data. We will talk about distributions of estimators, expected values, variances, etc. We will focus on the *process* of creating estimates.
- An estimate is a number, the result of applying an estimator to a specific set of observed data.

7.91  
0.4306



## What makes for a good estimator?

- A good estimator is one that will result in estimates that are close to the target parameter with high probability.
- Thinking of an estimator as being a random variable, we want its distribution to be close to the target parameter.
- An estimator is unbiased if it doesn't tend to either overestimate or underestimate the target parameter.
- All other things being equal, we would prefer an estimator with good precision.
- We will next formalize these concepts.

# Bias

$\mu, \sigma^2, \mu, P$

- Let's use  $\theta$  to denote our target parameter.
- Let's use  $\hat{\theta}$  to denote an estimator of  $\theta$ .
- Since  $\hat{\theta}$  is a random variable, we can speak of  $E[\hat{\theta}]$  and  $Var(\hat{\theta})$ .
- Definition 3: The BIAS of an estimator  $\hat{\theta}$  for  $\theta$  is  $E[\hat{\theta}] - \theta$ .
- Definition 4: An estimator  $\hat{\theta}$  is UNBIASED for  $\theta$  if  $E[\hat{\theta}] = \theta$ .
- It's good for an estimator to be unbiased, but that's not the only consideration!

# Mean squared error

- In addition to having small bias, we would also like for our estimator to have small variance.

- Definition 5: The MEAN SQUARED ERROR of an estimator  $\hat{\theta}$  for  $\theta$  is

$$MSE(\hat{\theta}) = E[(\hat{\theta} - \theta)^2]$$

- The MSE measures a combination of the bias and the variance.

of the  
estimator  
constant

$$\begin{aligned}
 & \boxed{MSE(\hat{\theta})} = E[(\hat{\theta} - \theta)^2] = E\left[\left(\underbrace{\hat{\theta} - E[\hat{\theta}]}_{\text{variance}} + \underbrace{E[\hat{\theta}] - \theta}_{\text{bias}}\right)^2\right] \\
 & = E\left[(\hat{\theta} - E[\hat{\theta}])^2 + (E[\hat{\theta}] - \theta)^2 + 2(\hat{\theta} - E[\hat{\theta}])(E[\hat{\theta}] - \theta)\right] \\
 & = E[(\hat{\theta} - E[\hat{\theta}])^2] + (E[\hat{\theta}] - \theta)^2 + 2(E[\hat{\theta}] - \theta)E[\hat{\theta} - E[\hat{\theta}]] \\
 & = \boxed{Var(\hat{\theta}) + bias^2(\hat{\theta})} + \underbrace{2 \text{bias}(\hat{\theta}) \underbrace{(E[\hat{\theta}] - E[\hat{\theta}])}_{=0}}_{=0} = 0
 \end{aligned}$$



# Mean squared error: Example

- Example 1: To estimate  $p$ , the “true probability of ‘success’” in a binomial experiment, we repeat the trial  $n$  times and observe  $Y \sim \text{Bi}(n, p)$ , the number of “successes”.

- A natural estimator is  $\hat{p} = \frac{Y}{n}$ .
- An alternative estimator is  $\hat{p}^* = \frac{Y+1}{n+1}$  ??
- How to decide which one is “better”?

$$E[\hat{p}] = E\left[\frac{Y}{n}\right] = \frac{1}{n} E[Y] = \frac{1}{n} \cdot np = p$$

$$\begin{aligned} \text{Var}(\hat{p}) &= \text{Var}\left(\frac{Y}{n}\right) = \frac{1}{n^2} \text{Var}(Y) \\ &= \frac{1}{n^2} np(1-p) = \frac{p(1-p)}{n} \end{aligned}$$

$$\text{MSE}(\hat{p}) = \text{Var}(\hat{p}) + \text{bias}^2(\hat{p}) = \frac{p(1-p)}{n} + 0 = \frac{p(1-p)}{n}$$

$$E[Y] = np$$

$$\text{Var}(Y) = np(1-p)$$

UNBIASED

$$\text{bias}(\hat{p}) = 0$$

$$E[\hat{p}^*] = E\left[\frac{Y+1}{n+1}\right] = E\left[\frac{Y}{n+1} + \frac{1}{n+1}\right] = \frac{1}{n+1} E[Y] + \frac{1}{n+1}$$

$$= \frac{np}{n+1} + \frac{1}{n+1} = \frac{np+1}{n+1} \stackrel{??}{=} p \quad \underline{\underline{No}} \quad \underline{\underline{BIASED !!}}$$

$$\text{bias}(\hat{p}^*) = \frac{np+1}{n+1} - p = \frac{\cancel{np}+1 - \cancel{np}-p}{n+1} = \frac{1-p}{n+1}$$

Bias is small  
if  $n+1$  large

$$\text{Var}(\hat{p}^*) = \text{Var}\left(\frac{Y}{n+1} + \frac{1}{n+1}\right) = \text{Var}\left(\frac{Y}{n+1}\right)$$

$$= \frac{1}{(n+1)^2} \text{Var}(Y) = \frac{np(1-p)}{(n+1)^2}$$

or if  $p$  is close to 1.

$$\text{MSE}(\hat{p}^*) = \frac{np(1-p)}{(n+1)^2} + \frac{(1-p)^2}{(n+1)^2} = \frac{(1-p)(np+1-p)}{(n+1)^2} = \frac{(1-p)(1+(n-1)p)}{(n+1)^2}$$

Which estimator has smaller MSE?

May depend on  $n, p$

~~$p = \frac{1}{2} \quad n = 10$~~

$$\text{MSE}(\hat{p}) = \frac{\frac{1}{2} \cdot \frac{1}{2}}{10} = \frac{1}{40} = 0.025$$

$$\text{MSE}(\hat{p}^*) = \frac{\frac{1}{2}(1+9(\frac{1}{2}))}{11^2} = 0.023 \quad \leftarrow$$

# Some common unbiased point estimators

- Situation 1:  $Y_1, Y_2, \dots, Y_n \sim \text{iid } N(\mu, \sigma^2)$ . How to estimate  $\mu$ ? model population mean

$$\bar{Y} = \frac{1}{n} \sum_{i=1}^n Y_i$$

$$E[\bar{Y}] = E\left[\frac{1}{n} \sum_{i=1}^n Y_i\right] = \frac{1}{n} \sum_{i=1}^n E[Y_i]$$

$$= \frac{1}{n} (\mu + \mu + \dots + \mu) = \frac{n\mu}{n} = \mu$$

$$V_2(Y) = \frac{\sigma^2}{n}$$

- Situation 2:  $X_1, X_2, \dots, X_n \sim \text{iid Bernoulli}(p)$  and  $Y = \sum_{i=1}^n X_i$ . How to estimate  $p$ ? Binomial( $n, p$ )

$$Y \sim \text{binomial}(n, p)$$

$$E[Y] = np$$

$$\hat{p} = \frac{Y}{n}$$

$$E[\hat{p}] = \frac{np}{n} = p$$

unbiased for  $p$

ditto - 0 or 1 - respond to trt.

## Another common unbiased point estimator

- Situation 3:  $Y_{1,1}, Y_{1,2}, \dots, Y_{1,n_1} \sim iid N(\mu_1, \sigma_1^2)$ ,  
 $Y_{2,1}, Y_{2,2}, \dots, Y_{2,n_2} \sim iid N(\mu_2, \sigma_2^2)$ , mutually independent. How to estimate  $\mu_1 - \mu_2$ ?

$$\bar{Y}_1 = \frac{1}{n_1} \sum_{i=1}^{n_1} Y_{1,i}$$

$$\bar{Y}_2 = \frac{1}{n_2} \sum_{i=1}^{n_2} Y_{2,i}$$

$$E[\bar{Y}_1] = \mu_1$$

$$E[\bar{Y}_2] = \mu_2$$

$$E[\bar{Y}_1 - \bar{Y}_2] = \mu_1 - \mu_2$$

UNBIASED!

$$\widehat{\mu_1 - \mu_2} = \hat{\mu}_1 - \hat{\mu}_2 = \bar{Y}_1 - \bar{Y}_2$$

$$E[\widehat{\mu_1 - \mu_2}] = E[\bar{Y}_1 - \bar{Y}_2] = E[\bar{Y}_1] - E[\bar{Y}_2] = \mu_1 - \mu_2$$

## And another common unbiased point estimator

- Situation 4:  $X_{1,1}, X_{1,2}, \dots, X_{1,n_1} \sim \text{iid Bernoulli}(p_1)$ ,  $X_{2,1}, X_{2,2}, \dots, X_{2,n_2} \sim \text{iid Bernoulli}(p_2)$ , mutually independent. How to estimate  $p_1 - p_2$ ?  
*placebo* *new drug*

$$Y_1 = \sum X_{1i} \quad Y_2 = \sum X_{2i} \quad Y_1 \sim \text{Binomial}(n_1, p_1) \\ Y_2 \sim \text{Binomial}(n_2, p_2)$$

$$\hat{p}_1 = \frac{Y_1}{n_1} \text{ unbiased for } p_1$$

$$\hat{p}_2 = \frac{Y_2}{n_2} \text{ " " } p_2$$

$$E[\hat{p}_1 - \hat{p}_2] = E[\hat{p}_1] - E[\hat{p}_2] = p_1 - p_2 \quad \text{UNBIASED!}$$

$$Var(\hat{p}_1 - \hat{p}_2) =$$

# Yet another common unbiased point estimator

properties of  $\chi^2$  dist.

- Back to Situation 1:  $Y_1, Y_2, \dots, Y_n \sim iid N(\mu, \sigma^2)$ , but how to estimate  $\sigma^2$ ?

Claim: The sample variance  $\hat{\sigma}^2 = s^2 = \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2$  is an unbiased estimator of  $\sigma^2$ .

We saw before  $\frac{(n-1)s^2}{\sigma^2} \sim \chi^2_{n-1}$  so  $E\left[\frac{(n-1)s^2}{\sigma^2}\right] = n-1$

What about  $\hat{\sigma}^{2*} = \frac{1}{n} \sum_{i=1}^n (Y_i - \bar{Y})^2$ ?

$$\hat{\sigma}^{2*} = \frac{n-1}{n} s^2 \quad E[\hat{\sigma}^{2*}] = \frac{n-1}{n} E[s^2]$$

And what about  $\hat{\sigma} = s$  as an estimator for  $\sigma$ ?

$$= \frac{n-1}{n} \sigma^2 \quad \text{BIASED}$$

$$\frac{n-1}{\sigma^2} E[s^2] = n-1$$

$$E[s^2] = \sigma^2$$

UNBIASED!

Yeah, sure, but  $E[s] \neq \sigma$  in general

BIASED

!! there is no estimator of  $\sigma$  that is unbiased for all  $\sigma$  !!

# Point estimates: What are they good for?

- A point estimate provides a “best guess” as to the unknown parameter value.
- Joe gets a sample and his estimate of  $\mu$  is 41.9.
- Dustin gets another sample and his estimate of  $\mu$  (the same one) is 48.5.
- Faye gets her own sample, and her estimate of  $\mu$  (again, the same one) is 62.5.
- Who is right? Who is wrong? Whom to believe?

A point estimate: It is what it is.

- Joe sampled 50 patients; Dustin sampled 65. Faye only sampled 15.
- Now whom do we believe (most)?