# FTML

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# 2 Exercice 2

We consider a supervised regression problem, Y = R. We have seen that when the loss used is the square loss l2(y, z) = (y z), then the Bayes predictor is the conditional expectation

# 2.1 Question 1

Subject : Propose a setting where the Bayes predictor is different for the square loss and for the absolute loss.

We could use the continuous distribution of Exp() with the a theta parameter

First, let's calculate a bayes estimator with the squared loss function

$$f * (x) = E[Y|X = x] = \frac{1}{\theta}$$

When we use the squared error we get the Esp() of the exponential Now if we switch to an absolute loss function we have

$$f * (x) = E[|y - z||X = x] = \frac{ln(2)}{\theta}$$

We can conclude that it's possible that the bayes estimator can change if the loss function is changed

## 2.2 Question 2

# 3 Exercice 3

We want to show that in the linear model, fixed design we have

$$E[R_x(\widehat{\theta})] = \frac{n-d}{n}\sigma^2$$

In this expression, both y and ^ are random variables, that are not independent, as theta is the OLS estimator. The expectation is over the distribution of both variables.

#### 3.1 step 1

Show that:

$$E[R_{n}(\widehat{\theta})] = E_{\epsilon} \left[ \frac{1}{n} \| I_{n} - X(X^{t}X)^{-1}X^{t}) \epsilon \|^{2} \right]$$

$$E[R_{n}(\widehat{\theta})] = E\left[ \frac{1}{n} \| y - X\widehat{\theta} \|^{2} \right]$$

$$E[R_{n}(\widehat{\theta})] = E\left[ \frac{1}{n} \| X\theta + \epsilon - X\widehat{\theta} \|^{2} \right]$$

$$E[R_{n}(\widehat{\theta})] = E\left[ \frac{1}{n} \| X\theta + \epsilon - X(X^{t}X)^{-1}X^{t}y \|^{2} \right]$$

$$E[R_{n}(\widehat{\theta})] = E\left[ \frac{1}{n} \| X\theta + \epsilon - X(X^{t}X)^{-1}X^{t}(X\theta + \epsilon) \|^{2} \right]$$

$$E[R_{n}(\widehat{\theta})] = E_{\Theta}\left[ \frac{1}{n} \| X\theta - X(X^{t}X)^{-1}X^{t}X\theta \|^{2} \right] + E_{\epsilon}\left[ \frac{1}{n} \| \epsilon - X(X^{t}X)^{-1}X^{t}\epsilon \|^{2} \right]$$

$$E[R_{n}(\widehat{\theta})] = E_{\Theta}\left[ \frac{1}{n} \| 0 \|^{2} \right] + E_{\epsilon}\left[ \frac{1}{n} \| (I_{n} - X(X^{t}X)^{-1}X^{t})\epsilon \|^{2} \right]$$

$$E[R_{n}(\widehat{\theta})] = E_{\epsilon}\left[ \frac{1}{n} \| (I_{n} - X(X^{t}X)^{-1}X^{t})\epsilon \|^{2} \right]$$

# 3.2 step 2

Let A Rn,n. Show that

$$\sum_{i=1}^{n} A_{ij}^2 = tr(A^t A)$$

To show this demonstration was can use the definition of the scalaire and the trace function

$$tr(A^t A) = \langle A, A \rangle = \sum_{i=1}^n A_{ij}^2$$

## 3.3 step 3

Show that

$$E\left[\frac{1}{n}\|A\epsilon\|^2\right] = \frac{\sigma^2}{n}tr(A^tA)$$

$$E\left[\frac{1}{n}\|A\epsilon\|^2\right] = E\left[\frac{1}{n}\sum_{i=1}^{n}\left(A_{ij}\epsilon\right)^2\right]$$

$$E[\frac{1}{n}\|A\epsilon\|^2] = \frac{1}{n}tr(A^tA)E[\epsilon^2]$$

But

$$E[\epsilon^2] = E[\epsilon^2 - E[\epsilon]^2] = Var = \sigma^2$$

Since in the subject its said that E(epsilon) is actualy equal to 0 So in the end we have

$$E\left[\frac{1}{n}\|A\epsilon\|^2\right] = \frac{\sigma^2}{n}tr(A^tA)$$

## 3.4 step 4

We note

$$A = I_n - X(X^t X)^{-1} X^t$$

Show that

$$A^t A = A$$

$$A^{t}A = (I_{n} - X(X^{t}X)^{-1}X^{t})^{t}(I_{n} - X(X^{t}X)^{-1}X^{t})$$
$$A^{t}A = A$$

## 3.5 step 5

We can now show that our proposition in the start is true

$$E[R_x(\widehat{\theta})] = E[[E[R_n(\widehat{\theta})]]$$

we have

$$[E[R_n(\widehat{\theta})] = \frac{\sigma^2}{n} tr(I_n - X(X^t X)^{-1} X^t)$$

we now use trace proprieties tr(a+b) = tr(a) + tr(b):

$$[E[R_n(\widehat{\theta})] = \frac{\sigma^2}{n} tr(I_n) - tr(X(X^t X)^{-1} X^t)$$

we now use trace proprieties tr(abc) = tr(cab):

$$[E[R_n(\widehat{\theta})] = \frac{\sigma^2}{n} tr(I_n) - tr(X^t X (X^t X)^{-1})$$

$$[E[R_n(\widehat{\theta})] = \frac{\sigma^2}{n} tr(I_n) - tr(I_d)$$
$$[E[R_n(\widehat{\theta})] = \frac{\sigma^2}{n} (n - d)$$

if we take back our first equation:

$$E[R_x(\widehat{\theta})] = E[[E[R_n(\widehat{\theta})]] = \frac{\sigma^2}{n}(n-d)$$