

Model Selection — Bayesian Information Criterion

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Viewpoint from Homo-Bayesianis

The backdrop: We have a bunch of alternative models: \mathcal{M}_i , and each model gives a parameter space Θ and a setting for generation of data $p(\mathcal{D}|\theta_i, \mathcal{M}_i)$. For comparing fidelity of different models, under Bayesian principle, we should use $p(\mathcal{D}|\mathcal{M}_i) \propto p(\mathcal{M}_i|\mathcal{D})$, assuming the prior for different \mathcal{M}_i are equal. The previous $p(\mathcal{M}_i|\mathcal{D})$ is called *model evidence*. By Bayes' formula, the evidence is obtained by:

$$\begin{aligned} p(\mathcal{D}|\mathcal{M}_i) &= \int p(\mathcal{D}|\theta_i, \mathcal{M}_i) \pi(\theta_i|\mathcal{M}_i) d\theta \\ &=: \int f_{\mathcal{M}_i; \mathcal{D}}(\theta_i) d\theta \end{aligned}$$

So if the integration is hard to compute, it's reasonable to assume that $f_{\mathcal{M}_i; \mathcal{D}}(\theta_i)$ as a function of θ_i is close to a p.d.f. of a normal distribution [a homo-frequentist will explain it by asymptotic normality], so by Laplace approximation near the MAP point $\hat{\theta}_i$, the integration is approximately decided by the Hessian matrix of $\log f_{\mathcal{M}_i; \mathcal{D}}$ at the MAP, as follows:

$$\log p(\mathcal{D}|\mathcal{M}_i) = \log p(\mathcal{D}|\hat{\theta}_i, \mathcal{M}_i) + \log \pi(\hat{\theta}_i|\mathcal{M}_i) + \frac{k}{2} \log(2\pi) - \frac{1}{2} \log |\mathbf{A}|$$

where k represents the dimension of the i -th parameter space and \mathbf{A} is the negative Hessian matrix at $\hat{\theta}_i$:

$$\begin{aligned} \mathbf{A} &= -\nabla^2 \Big|_{\hat{\theta}_i} \log p(\mathcal{D}|\theta_i, \mathcal{M}_i) \pi(\theta_i|\mathcal{M}_i) \\ &= -\nabla^2 \Big|_{\hat{\theta}_i} \left[\log p(\mathcal{D}|\theta_i, \mathcal{M}_i) + \log \pi(\theta_i|\mathcal{M}_i) \right] \end{aligned}$$

The last 3 terms of RHS are comprised as the penalization term against complexity, called “Occam factor”, while the first term of RHS describes how well can a prediction under the model fit the given data. As N increases far larger than k : the first 2 terms of Occam factor $\log \pi(\hat{\theta}_i|\mathcal{M}_i) + \frac{k}{2} \log(2\pi)$ can be relatively ignored and, for the last term:

$$\begin{aligned} |\det(\mathbf{A})|^{\frac{1}{2}} &\sim \left| -\nabla^2 \Big|_{\hat{\theta}_i} \log p(\mathcal{D}|\theta_i, \mathcal{M}_i) \right|^{\frac{1}{2}} \\ &\sim \left| -\nabla^2 \Big|_{\hat{\theta}_i} N \cdot \log p(x_{\text{typical}}|\theta_i, \mathcal{M}_i) \right|^{\frac{1}{2}} \\ &\sim O(N^{\frac{k}{2}}) \end{aligned}$$

So we can approximately estimate $\log p(\mathcal{D}|\mathcal{M}_i)$ by

$$\log p(\mathcal{D}|\hat{\boldsymbol{\theta}}_i, \mathcal{M}_i) - \frac{k}{2} \log N$$

which we should maximize among all models.

The BIC (which we should minimize, like AIC) is formally defined as

$$BIC = K \log N - 2 \log p(\mathcal{D}|\hat{\boldsymbol{\theta}}_i, \mathcal{M}_i)$$