

ΕΡΓΑΣΙΑ ΚΟΥ. ΤΕΧΝΙΚΑ

B. Probability

1. (a) It holds that:

$$\int_0^{\infty} p(x; \lambda) dx = 1 \Rightarrow \int_0^{\infty} \lambda x e^{-\lambda x^2} dx = 1 \Rightarrow -\frac{\lambda}{2} \int_0^{\infty} e^{-\lambda x^2} (-2x^2)' dx = 1 \Rightarrow$$
$$-\frac{\lambda}{2} \cdot e^{-\lambda x^2} \Big|_0^{\infty} = 1 \Rightarrow -\frac{\lambda}{2} (0 - 1) = 1 \Rightarrow \frac{\lambda}{2} = 1 \Rightarrow \lambda = 2. \text{ Consequently, } p(x, \lambda) = 2\lambda x e^{-\lambda x^2}, x > 0, \lambda > 0$$

(b) It is: $E(X^{2n}) = \int_0^{\infty} x^{2n} p(x; \lambda) dx = \int_0^{\infty} x^{2n} 2\lambda x e^{-\lambda x^2} dx = - \int_0^{\infty} x^{2n} (e^{-\lambda x^2})' dx =$

$$= -x^{2n} e^{-\lambda x^2} \Big|_0^{\infty} + \int_0^{\infty} 2n x^{2n-1} e^{-\lambda x^2} dx = -0 + 0 + \frac{n}{\lambda} \int_0^{\infty} x^{2(n-1)} 2\lambda x e^{-\lambda x^2} dx = \frac{n}{\lambda} E(X^{2(n-1)})$$

It is: $E(X^0) = E(1) = 1$.

Using a recursive equation relation, it is easy to prove that:

$$E(X^{2n}) = \frac{n}{\lambda} E(X^{2(n-1)}) = \frac{n(n-1)}{\lambda^2} E(X^{2(n-2)}) = \dots = \frac{n(n-1) \cdot 2}{\lambda^n} E(X^0) = \frac{n!}{\lambda^n}, n=0,1,\dots$$

2. Usually, it is $RAND() \equiv U(0,1)$

If there is a random variable $X_i \sim U(0,1)$, then $E(X_i) = \frac{1}{2}$, $V(X_i) = \frac{1}{12}$, $i=1, \dots, 12$.

Consequently, from the Central Limit Theorem, it is: $\sum_{i=1}^{12} X_i \sim N(12 \cdot \frac{1}{2} = 6, 12 \cdot \frac{1}{12} = 1)$ (approximately)

That means: $\frac{\sum_{i=1}^{12} RAND() - 6}{1} \sim N(0,1)$

In the same way, $RAND(N) \equiv N \cdot U(0,1)$

If $X_i \sim RAND(N)$, then $E(X_i) = \frac{N}{2}$ and $V(X_i) = \frac{N^2}{12}$, $i=1, \dots, 12$.

Consequently, from the Central Limit Theorem, it is: $\sum_{i=1}^{12} X_i \sim N(12 \cdot \frac{N}{2} = 6N, 12 \cdot \frac{N^2}{12} = N^2)$

This leads to the conclusion that: $\frac{\sum_{i=1}^{12} RAND(N) - 6N}{N} \sim N(0,1)$