

Homework 4 STAT 632

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2022-03-12

Exercise 1

MLR: $Y = X\beta + e$ $\text{Var}(e) = \sigma^2 I$ $I = n \times n$ identity matrix

$$\hat{Y} = X\hat{\beta} = X(X'X)^{-1}X'Y = HY \quad ; \quad H = X(X'X)^{-1}X'$$

(a) $HH' = HH = H$

$\therefore H$ is idempotent & symmetric so $HH = H$ & $H = H'$

$\therefore HH' = HH$

\therefore some prove $HH' = H$

$$\begin{aligned} & X(X'X)^{-1}X' \cdot X(X'X)^{-1}X' \\ &= X(X'X)^{-1}X' \quad \therefore \text{identity matrix } I \\ &\therefore \text{which is just equal to } H. \end{aligned}$$

(b) $E(\hat{Y}) = X\beta$

$$\begin{aligned} E(HY) &= X(X'X)^{-1}X' \cdot E(Y) \quad \therefore E(Y) = E(X\beta + e) \\ &= \therefore X X' (X'X)^{-1} = I \quad = E(X\beta) + E(e) \\ &= I \cdot X\beta \quad = X\beta + 0 \\ &= X\beta \quad = X\beta \end{aligned}$$

(c) $\text{Var}(\hat{Y}) = \sigma^2 H$

$$\begin{aligned} \text{Var}(HY) &= H \text{Var}(Y) H' \\ \text{Var}(\hat{Y}) &= \sigma^2 H \end{aligned}$$

$\therefore \hat{Y} = HY$

\therefore by linearity of Expectation: $\text{Var}(AX) = A \text{Var}(X) A'$

$\therefore H$ is idempotent & symmetric so $H \cdot H' = H$

$$\begin{aligned} \therefore \text{Var}(Y) &= \text{Var}(X\beta + e) = \text{Var}(e) + \text{Var}(X\beta) \\ &= \text{Var}(e) \\ &= \sigma^2 I_n \end{aligned}$$

$\therefore X\beta = \text{fixed}$

Exercise 2

Exercise 2.

$$\text{Var}(\hat{\beta}) = \sigma^2 (X'X)^{-1}$$

$$X = \begin{bmatrix} 1 & x_1 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix} \quad X' = \begin{bmatrix} 1 & 1 & \dots & 1 \\ x_1 & x_2 & \dots & x_n \end{bmatrix}$$

$$\text{So, } X'X = \begin{bmatrix} 1 & 1 & \dots & 1 \\ x_1 & x_2 & \dots & x_n \end{bmatrix} \begin{bmatrix} 1 & x_1 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix} = \begin{bmatrix} n & \sum x_i \\ \sum x_i & \sum x_i^2 \end{bmatrix}$$

$2 \times n \qquad \qquad \qquad n \times 2 \qquad \qquad \qquad 2 \times 2$

$$\text{Var}(\hat{\beta}) = \sigma^2 (X'X)^{-1}$$

$$\therefore \text{ if } A = \begin{bmatrix} a & b \\ c & d \end{bmatrix} \text{ then, } A^{-1} = \frac{1}{ad-bc} \begin{bmatrix} d & -b \\ -c & a \end{bmatrix}$$

$$\hookrightarrow \frac{\sigma^2}{n \sum x_i^2 - (\sum x_i)^2} \begin{bmatrix} \sum x_i^2 & -\sum x_i \\ -\sum x_i & n \end{bmatrix}$$

$$\therefore \text{ compare it w/ } \text{Var}(\hat{\beta}) : \quad \text{Var}(\hat{\beta}) = \begin{bmatrix} \text{Var}(\hat{\beta}_0) & \text{Cov}(\hat{\beta}_0, \hat{\beta}_1) \\ \text{Cov}(\hat{\beta}_1, \hat{\beta}_0) & \text{Var}(\hat{\beta}_1) \end{bmatrix}$$

$$\begin{aligned}
 \bullet \text{Var}(\hat{\beta}_0) &= \frac{\sigma^2 \sum x_i^2}{n \sum x_i^2 - (\sum x_i)^2} = \frac{\sigma^2 \sum x_i^2}{\sum x_i^2 - \frac{(\sum x_i)^2}{n}} \\
 &= \frac{\sigma^2 \sum x_i^2}{\sum x_i^2 - n \bar{x}^2} \\
 &= \frac{\sigma^2 \sum x_i^2}{n \cdot SXX} \\
 &= \sigma^2 \left[\frac{\sum x_i^2 - n \bar{x}^2 + n \bar{x}^2}{n \cdot SXX} \right] \\
 &= \sigma^2 \left[\frac{SXX + n \bar{x}^2}{n SXX} \right] \\
 \therefore \sigma^2 \left[\frac{1}{n} + \frac{\bar{x}^2}{SXX} \right] &\quad \therefore \text{Var}(\hat{\beta}_0) = \sigma^2 \left(\frac{1}{n} + \frac{\bar{x}^2}{SXX} \right) \\
 \therefore SXX &= \sum_{i=1}^n (x_i - \bar{x})^2 \\
 &\hookrightarrow \sum x_i^2 - \sum \bar{x}^2 \\
 \bullet \text{Var}(\hat{\beta}_1) &= \frac{\sigma^2 n}{n \sum x_i^2 - (\sum x_i)^2} \quad \therefore \bar{x} = \frac{\sum x_i}{n} \\
 &= \frac{\sigma^2}{\sum x_i^2 - \frac{n^2 \bar{x}^2}{n}} \\
 &= \frac{\sigma^2}{\sum x_i^2 - n \bar{x}^2} \\
 &= \frac{\sigma^2}{SXX} \quad \therefore \text{Var}(\hat{\beta}_1) = \frac{\sigma^2}{SXX}
 \end{aligned}$$

Exercise 3

```
library(MASS)
head(Boston)
```

```
##      crim zn indus chas   nox   rm age   dis rad tax ptratio  black lstat
## 1 0.00632 18  2.31    0 0.538 6.575 65.2 4.0900   1 296    15.3 396.90  4.98
## 2 0.02731  0  7.07    0 0.469 6.421 78.9 4.9671   2 242    17.8 396.90  9.14
## 3 0.02729  0  7.07    0 0.469 7.185 61.1 4.9671   2 242    17.8 392.83  4.03
## 4 0.03237  0  2.18    0 0.458 6.998 45.8 6.0622   3 222    18.7 394.63  2.94
## 5 0.06905  0  2.18    0 0.458 7.147 54.2 6.0622   3 222    18.7 396.90  5.33
## 6 0.02985  0  2.18    0 0.458 6.430 58.7 6.0622   3 222    18.7 394.12  5.21
##   medv
## 1 24.0
## 2 21.6
## 3 34.7
## 4 33.4
## 5 36.2
## 6 28.7
```

(a)

- The results are the same as the parameter estimates provided by the `lm()` function.

```
#response vector
Y <-matrix(Boston$medv, ncol=1)

#design matrix
X <-cbind(Intercept =1, Boston[, c('dis', 'rm', 'tax', 'chas')])
X <-as.matrix(X)

#manually calculate least squares estimate
betahat <-solve(t(X) %*% X) %*% t(X) %*% Y
betahat
```

```
##              [,1]
## Intercept -20.16720221
## dis       -0.10656777
## rm         7.88589232
## tax       -0.01647039
## chas       3.87901205
```

```
#compare with lm()
lm1 <-lm(medv ~ dis + rm + tax + chas, data=Boston)
coef(lm1)
```

```
## (Intercept)      dis      rm      tax      chas
## -20.16720221 -0.10656777  7.88589232 -0.01647039  3.87901205
```

(b)

- The square root of the diagonal entries of the computed var-cov matrix is the same as the standard errors provided by the `lm()` function.

```
#manually calculate standard errors for least squares estimate
n <- nrow(Boston)
p <- 4
resid <- as.numeric(Y - X %*% betahat)
sigmahat2 <- sum(resid^2) / (n-p-1)
covbetahat <- sigmahat2 * solve(t(X) %*% X)
covbetahat
```

```
##           Intercept          dis          rm          tax          chas
## Intercept  8.510688175 -0.1236974635 -1.0632321031 -3.155777e-03  0.014868410
## dis       -0.123697464  0.0233402861 -0.0045240274  1.516224e-04  0.023912819
## rm        -1.063232103 -0.0045240274  0.1616761190  1.644565e-04 -0.040647554
## tax       -0.003155777  0.0001516224  0.0001644565  3.759877e-06  0.000171935
## chas       0.014868410  0.0239128186 -0.0406475538  1.719350e-04  1.151456124
```

```
#compare with lm()
vcov(lm1)
```

```
##           (Intercept)          dis          rm          tax          chas
## (Intercept)  8.510688175 -0.1236974635 -1.0632321031 -3.155777e-03  0.014868410
## dis         -0.123697464  0.0233402861 -0.0045240274  1.516224e-04  0.023912819
## rm          -1.063232103 -0.0045240274  0.1616761190  1.644565e-04 -0.040647554
## tax         -0.003155777  0.0001516224  0.0001644565  3.759877e-06  0.000171935
## chas         0.014868410  0.0239128186 -0.0406475538  1.719350e-04  1.151456124
```

```
sebetahat <- sqrt(diag(covbetahat))
sebetahat
```

```
## Intercept          dis          rm          tax          chas
## 2.91730838 0.15277528 0.40208969 0.00193904 1.07305924
```

```
#compare with lm()

summary(lm1)$coef
```

```
##           Estimate Std. Error    t value    Pr(>|t|)
## (Intercept) -20.16720221  2.91730838 -6.9129484 1.450596e-11
## dis         -0.10656777  0.15277528 -0.6975459 4.857848e-01
## rm           7.88589232  0.40208969 19.6122719 5.663118e-64
## tax         -0.01647039  0.00193904 -8.4940923 2.286843e-16
## chas         3.87901205  1.07305924  3.6149095 3.308341e-04
```