

CSC 411 Fall 2018  
Machine Learning and Data Mining

Homework 4

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## Solution 1

(a)

Input:

 $W \times H \times C$ kernel size  $K$ Output Maps  $M$ 

$$\begin{aligned}\text{Units} &= \text{pixels} \times \text{output units} \\ &= WH \times M = WHM\end{aligned}$$

$$\begin{aligned}\text{Weights} &= \text{size of each piece of input} \times \text{No of Filters} \\ &= K^2 C \times M \\ &= K^2 CM\end{aligned}$$

$$\begin{aligned}\text{Connections} &= \text{values processed by a kernel} \times \text{output units} \\ &= K^2 C \times WHM \\ &= K^2 C WHM.\end{aligned}$$



Layer 1.

$M = 96$  Kernels  $\rightarrow$   $11 \times 11 \times 3$  each  
 stride = 4

$$\text{Input} = 224^w \times 224^h \times 3^c$$

$$\text{New Input: } w = h = 55 \quad c = 3$$

$$\text{Units} = 55 \times 55 \times 96 = \boxed{290,400}$$

$$\text{Weight} = K^2 c M = 11^2 \times 3 \times 96 = \boxed{34,848}$$

$$\text{Connections} = K^2 c M w h = 11^2 \times 3 \times 96 \times 55 \times 55 = \boxed{105,415,200}$$

Layer 2.

$$M = 256 \quad \text{Size} \rightarrow 5 \times 5 \times 48 = K^2 c$$

$$w = h = 55 / 2 = 27 \quad (\text{max pooling})$$

$$\text{Units} = w h M = 27^2 \times 256 = \boxed{186,624}$$

$$\text{Weight} = K^2 c M = 5^2 \times 48 \times 256 = \boxed{307,200}$$

$$\text{Connections} = K^2 c w h M = 5^2 \times 48 \times 27^2 \times 256 = \boxed{223,948,800}$$

Layer 3

$$M = 384 \quad K^2 C = 3 \times 3 \times 256$$

$$W = H = 27/2 = 13 \text{ (max pooling)}$$

$$\text{Units} = W H M = 13^2 \times 384 = \boxed{64,896}$$

$$\begin{aligned} \text{Weights} &= K^2 C M = 3 \times 3 \times 256 \times 384 \\ &= \boxed{884,736} \end{aligned}$$

$$\begin{aligned} \text{Connections} &= K^2 C W H M = 3^2 \times 256 \times 13^2 \times 384 \\ &= \boxed{1149,520,384} \end{aligned}$$

Layer 4

$$M = 384 \quad K^2 C = 3 \times 3 \times 192$$

$$\text{units} = 13^2 \times 384 = \boxed{64,896}$$

$$\text{weight} = 3^2 \times 192 \times 384 = \boxed{663,552}$$

$$\text{Connection} = 3^2 \times 192 \times 13^2 \times 384$$

$$= \boxed{112,140,288}$$



layer 5

$$M = 256 \quad K^2C = 3 \times 3 \times 192$$

$$\text{units} = 13^2 \times 256 = \boxed{43,264}$$

$$\text{weight} = 3^2 \times 192 \times 256 = \boxed{442,368}$$

$$\begin{aligned} \text{connection} &= 3^2 \times 192 \times 13^2 \times 256 \\ &= \boxed{74,760,192} \end{aligned}$$

layer 6

$$\text{units} = 4096$$

$$\text{weight} = 6^2 \times 256 \times 4096 = \boxed{37,748,736}$$

$$\text{connection} = 6^2 \times 256 \times 4096 = \boxed{37,748,736}$$

layer 7

$$\text{units} = 4096$$

$$\text{weight} = 4096 \times 4096 = \boxed{16,777,216}$$

$$\text{connection} = 4096^2 = \boxed{16,777,216}$$

layer 8

$$\text{units} = \boxed{1000}$$

$$\text{weight} = 4096 \times 1000 = \boxed{4,096,000}$$

$$\text{Connection} = 4096 \times 1000 = \boxed{4,096,000}$$

layer	#Units	#weights	#Connections
convolution Layer 1	290,400	34,848	105,415,200
convolution layer 2	186,624	307,200	223,948,800
convolution layer 3	64,896	884,736	149,520,384
convolution layer 4	64,896	663,552	112,140,288
convolution layer 5	43,264	442,368	74,760,192
Fully connected layer 1	4096	37,748,736	37,748,736
Fully connected layer 2	4096	16,777,216	16,777,216
Output Layer	1000	4,096,000	4,096,000



Solution 1.  
(b)

(i)

Fully connected layers contribute to the most number of parameters i.e the weights.

$$\text{weights} = K^2 C M$$

(hence reducing  $K$  and  $M$  would help)  
 Also, Alexnet works with same accuracy even if its parameters are reduced certain times.

(b)  
 (ii) Convolution Layers contributes to most number of connections

$$\text{connection} = K^2 C M W H$$

hence, suggestion: Reduce the No. of kernels; or reduce size of kernels. would both help.



soln Q2. Gaussian Naive Bayes

(a)  $X$  is  $(x_1, \dots, x_d)$  i.e.  $d$  feature  $X$ .

$X = (x_1, x_2, \dots, x_d)$ , where each  $x_i$  is a continuous random variable.

$Y \in \{1, 2, \dots, K\}$  but to start and for simplicity let's first assume that  $Y$  is Boolean and has parameter say  $\pi = P(Y=1)$

Given, that for each  $x_i$ ,  $P(x_i | Y=k)$  is a Gaussian Distribution of the form  $N(\mu_{ik}, \sigma_i)$

For all  $i$  and  $j \neq i$ ,  $x_i$  and  $x_j$  are conditionally independent given  $Y$ .

$\sigma_i$  varies from attribute to attribute but doesn't depend on  $Y$ .

Now,

$$P(Y=1 | X) = \frac{P(Y=1) P(X | Y=1)}{P(Y=1) P(X | Y=1) + P(Y=0) P(X | Y=0)}$$

Divide both numerator and denominator by numerator.



$$p(Y=1|X) = \frac{1}{1 + \frac{p(Y=0)p(X|Y=0)}{p(Y=1)p(X|Y=1)}}$$

or equivalently,

$$p(Y=1|X) = \frac{1}{1 + \exp\left(\ln\left(\frac{p(Y=0)p(X|Y=0)}{p(Y=1)p(X|Y=1)}\right)\right)}$$

As taken conditional independence assumption, we can say

$$p(Y=1|X) = \frac{1}{1 + \exp\left(\ln \frac{p(Y=0)}{p(Y=1)} + \sum_i \ln \frac{p(X_i|Y=0)}{p(X_i|Y=1)}\right)}$$

$$= \frac{1}{1 + \exp\left(\ln \frac{1-\pi}{\pi} + \sum_i \ln \frac{p(X_i|Y=0)}{p(X_i|Y=1)}\right)}$$

$\rightarrow \textcircled{1}$

Since it is given that  $P(X_i | Y = \alpha_k)$  is Gaussian, hence

$$\sum_i \ln \frac{P(X_i | Y=0)}{P(X_i | Y=1)} = \sum_i \ln \frac{1}{\sqrt{2\pi\sigma_i^2}} \exp\left(-\frac{(X_i - \mu_{i0})^2}{2\sigma_i^2}\right)$$

$$= \sum_i \ln \exp\left(\frac{(X_i - \mu_{i1})^2 - (X_i - \mu_{i0})^2}{2\sigma_i^2}\right)$$

$$= \sum_i \left( \frac{(X_i - \mu_{i1})^2 - (X_i - \mu_{i0})^2}{2\sigma_i^2} \right)$$

$$= \sum_i \left( \frac{(X_i^2 - 2X_i\mu_{i1} + \mu_{i1}^2) - (X_i^2 - 2X_i\mu_{i0} + \mu_{i0}^2)}{2\sigma_i^2} \right)$$

$$= \sum_i \left( \frac{2X_i(\mu_{i0} - \mu_{i1}) + \mu_{i1}^2 - \mu_{i0}^2}{2\sigma_i^2} \right)$$

$$= \sum_i \left( \frac{\mu_{i0} - \mu_{i1}}{\sigma_i^2} X_i + \frac{\mu_{i1}^2 - \mu_{i0}^2}{2\sigma_i^2} \right)$$

→ (2)



Now, this expression is a linear weighted sum of  $X_i$ 's. Substituting expression (2) in (1), we get

$$P(Y=1|X) = \frac{1}{1 + \exp\left(\ln \frac{1-\pi}{\pi} + \sum_i \left( \frac{\mu_{i0} - \mu_{i1}}{\sigma_i^2} X_i + \frac{(\mu_{i1}^2 - \mu_{i0}^2)}{2\sigma_i^2} \right)\right)}$$

$$= P(Y=1|X) = \frac{1}{1 + \exp\left(w_0 + \sum_{i=1}^d w_i X_i\right)}$$

$w_1, \dots, w_d$  are given by

$$w_i = \frac{\mu_{i0} - \mu_{i1}}{\sigma_i^2}$$

and

$$w_0 = \ln \frac{1-\pi}{\pi} + \sum_i \frac{\mu_{i1}^2 - \mu_{i0}^2}{2\sigma_i^2}$$

Now for

$$P(Y=0|X) = 1 - P(Y=1|X)$$

$$P(Y=0|X) = 1 - P(Y \neq 0|X) = \frac{\exp(w_0 + \sum_{i=1}^d w_i x_i)}{1 + \exp(w_0 + \sum_{i=1}^d w_i x_i)}$$

we considered only cases where  $Y$  was a boolean variable, but now if  $Y$  can take any of discrete class label values i.e.  $y \in \{1, \dots, K\}$  that is  $Y$  can take values  $(Y_1, \dots, Y_K)$

Then

$P(Y = \alpha_k | X)$  for  $Y = \alpha_1, Y = \alpha_2, \dots$

$Y = \alpha_{K-1}$  is:

$$P(Y = \alpha_k | X) = \frac{\exp(w_{k0} + \sum_{i=1}^d w_{ki} x_i)}{1 + \sum_{j=1}^{K-1} \exp(w_{j0} + \sum_{i=1}^d w_{ji} x_i)}$$

$$P(Y = k | X, \mu, \sigma) = \frac{\exp(w_{k0} + \sum_{i=1}^d w_{ki} x_i)}{1 + \sum_{j=1}^{K-1} \exp(w_{j0} + \sum_{i=1}^d w_{ji} x_i)}$$

Ans



Ans (2)  
~~B~~

$$l(\theta; D) = -\log P(y^{(1)}, x^{(1)}, y^{(2)}, x^{(2)}, \dots, y^{(N)}, x^{(N)} | \theta)$$

Assuming data are i.i.d

$$l(\theta; D) = -\log \left( \prod_{i=1}^N P(y^{(i)}, x^{(i)} | \theta) \right)$$

we can take  
 $\Sigma$  as i.i.d  
 data.

$$= -\sum_{i=1}^N \log P(y^{(i)}, x^{(i)} | \theta)$$

$$= -\sum_N \log P(x^{(i)}, y^{(i)} | \theta) \quad \text{--- (1)}$$

Since we can write

$$P(x, y | \theta) = \log P(y | \theta) + \log P(x | y, \theta)$$

putting it in equation (1)

$$\ell(\theta; D) = -\sum_N \log P(x^{(i)}, y^{(i)} | \theta)$$

$$\ell(\theta; D) = -\sum_N (\log P(y^{(i)} | \theta) + \log P(x^{(i)} | y^{(i)}, \theta))$$

Ans



Q2 Using the Multivariate Gaussian Distribution Equation since in the question it mentions shared variance ( $\sigma_i^2$ ).

$$N(x|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2} |\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)\right\}$$

→ ①

where

$x$  is  $D$ -dimensional vector

$\mu$  is  $D$  dimensional mean vector

$\Sigma$  is  $D \times D$  covariance matrix with determinant  $|\Sigma|$

$\Sigma$  (covariance matrix) is a matrix whose  $(i, j)$  entry is covariance i.e.

$$\begin{aligned} \Sigma_{ij} &= \text{cov}(x_i, x_j) \\ &= E[(x_i - \mu_i)(x_j - \mu_j)] \end{aligned}$$

$$= E[(x_i x_j)] - \mu_i \mu_j$$

Diagonal entries are variance of

each elements, hence using equation ① in the question.

$$\theta = [\mu, \Sigma, \alpha]$$

$$Z = \sqrt{(2\pi)^D |\Sigma|} = (2\pi)^{D/2} |\Sigma|^{1/2}$$

$$p(x|y) = \frac{1}{Z} \exp\left(-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)\right)$$

$$\log L(\theta) = \log P(x, y|\theta) = \log P(y|\theta) + \log p(x|y, \theta)$$

$$= \sum_{i=1}^N \log \alpha_{y^{(i)}} - \log Z - \frac{1}{2} (X^{(i)} - \mu_{y^{(i)}})^T \Sigma_{y^{(i)}}^{-1} (X^{(i)} - \mu_{y^{(i)}}) \quad \text{①}$$

Since we are aiming at maximum likelihood that is  $\arg \max_{\theta} \log L(\theta)$  this can happen when

$$\sum_K \alpha_K = 1$$



Taking derivative w.r.t  $\mu$

$$\frac{\partial \log L}{\partial \mu_K} = - \sum_{i=0}^N \mathbb{1}(y^{(i)} = K) \Sigma^{-1} (x^{(i)} - \mu_K) = 0$$

$$\mu_K = \frac{\sum_{i=1}^N \mathbb{1}(y^{(i)} = K) x^{(i)}}{\sum_{i=1}^N \mathbb{1}(y^{(i)} = K)}$$

Ans

Now take derivative of equation (1)  
w.r.t  $\Sigma^{-1}$

$$\frac{\partial \log L}{\partial \Sigma_K^{-1}} = - \sum_{i=0}^N \mathbb{1}(y^{(i)} = K) \left[ - \frac{\partial \log Z_K}{\partial \Sigma_K^{-1}} - \frac{1}{2} \frac{(x^{(i)} - \mu_K)(x^{(i)} - \mu_K)^T}{(x^{(i)} - \mu_K)^T} \right]$$

$$= 0$$

$\rightarrow (2)$

$$Z = (2\pi)^{D/2} |\Sigma_K|^{-1/2}$$

$$\frac{\partial \log Z_K}{\partial \Sigma_K^{-1}} = \frac{1}{Z_K} \frac{\partial Z_K}{\partial \Sigma_K^{-1}}$$

$$= (2\pi)^{-D/2} |\Sigma_K|^{-1/2} (2\pi)^{D/2} \frac{\partial (|\Sigma_K^{-1}|)^{-1/2}}{\partial \Sigma_K^{-1}}$$

$$= |\Sigma_K^{-1}|^{1/2} \left(-\frac{1}{2}\right) |\Sigma_K^{-1}|^{3/2} |\Sigma_K^{-1}| \Sigma_K^T = -\frac{1}{2} \Sigma_K$$

(using property  
that  $\Sigma = \Sigma^T$ )

②

Now substituting this back in equation ②  
we get

$$\frac{\partial \text{Log } L}{\partial \Sigma_K^{-1}} = -\sum_{i=0}^N \mathbb{1}(y^{(i)} = K) \left[ \frac{1}{2} \Sigma_K - \frac{1}{2} (x^{(i)} - \mu_K)(x^{(i)} - \mu_K)^T \right] = 0$$

$$\Sigma_K = \frac{\sum_{i=1}^N \mathbb{1}(y^{(i)} = K) (x^{(i)} - \mu_K)(x^{(i)} - \mu_K)^T}{\sum_{i=1}^N \mathbb{1}(y^{(i)} = K)}$$

$$\sum_{i=1}^N \mathbb{1}(y^{(i)} = K)$$

Ans



$$\boxed{\begin{matrix} Q(2) \\ (d) \end{matrix}}.$$

As suggested in the question using Lagrange multiplier

$$\frac{\partial L(\theta)}{\partial \alpha_K} + \lambda \frac{\partial \sum_K \alpha_K}{\partial \alpha_K} = 0$$

Note: Since we need to maximize we are adding the  $\lambda$  term.

$$\lambda = - \sum_{i=1}^N \mathbb{1}(y^{(i)} = K) \frac{1}{\alpha_K}$$

$$\alpha_K = - \sum_{i=1}^N \mathbb{1}(y^{(i)} = K)$$

$$\frac{\quad}{\lambda} \rightarrow \textcircled{1}$$

Since again we are wanting to maximize the likelihood as that is possible when

$$\sum_K \alpha_K = 1 \Rightarrow \lambda = -N$$

Substituting value of  $\lambda = -N$  in  
~~equation~~ equation (1)

$$\alpha_k = \frac{-\sum_{i=1}^N \mathbb{1}(y^{(i)} = k)}{-N}$$

$$\alpha_k = \frac{1}{N} \sum_{i=1}^N \mathbb{1}(y^{(i)} = k)$$

Ans.