

HW Problems for Assignment 2 - Part 2

Due 6:30 PM Tuesday, October 7th, 2020

1. (20 Points, 5 Points Each) Practice with ES. Explicitly compute $\text{ES}_\alpha(L)$ assuming L has the following distributions/probability distribution functions (pdf).

- (a) $L \sim \text{Exp}(1/\theta)$ is exponentially distributed with mean θ .
- (b) $L = \max\{L_1, L_2\}$ where L_1, L_2 are iid $\text{Exp}(1/\theta)$ random variables. How does your answer compare to (a) as $\alpha \rightarrow 1$?
- (c) If Y is a positive random variable with strictly positive pdf, and g is a strictly increasing function on $(0, \infty)$, show that $\text{VaR}_\alpha(g(Y)) = g(\text{VaR}_\alpha(Y))$.
- (d) Let $Y \sim \text{Exp}(1/\theta)$ and $g(y) = y^2$. What is $\text{ES}_\alpha(g(Y))$? Is $\text{ES}_\alpha(g(Y)) = g(\text{ES}_\alpha(Y))$? What about as $\alpha \rightarrow 1$?

2. (20 Points) Time aggregated risk measures for a constant weight portfolio of equities. In this exercise you will estimate various risk measures at a $K = 10$ day horizon for a portfolio of equities, assuming the portfolio weights are held constant over the K day horizon.

The portfolio consists of Boeing, McDonald's, Nike and Walmart stock. At the start (time $t = 9/1/2020$) the portfolio value is $V_t = \$1M$. The weights are determined using the time t market capitalizations from (in billions of dollars) of

Boeing : 97.39; McDonald's : 158.20; Nike : 179.01; Walmart : 417.97.

Historical prices from 8/31/2015 – 8/31/2020 are in the file “Prices.csv”. The first column is the date (in Excel numeric format) while columns 2-5 give the stock price data. Data is sorted oldest to newest.

Our goal is to write a simulation to estimate the distribution of the K day losses $L_{t+K\Delta}$, and then use the distribution to estimate the risk measures. We work in the normal log-returns framework, using full losses and EWMA to estimate the conditional mean and covariance. However, as our time horizon is longer than 1 day, there are some subtleties when writing the simulation. Thus, complete the following steps:

- (a) **(5 Points)** (pen and paper problem) As an abuse of notation, for a vector $x = (x^{(1)}, \dots, x^{(d)})$ write e^x for the vector $(e^{x^{(1)}}, \dots, e^{x^{(d)}})$. Show that for constant weights, the full loss over the K day horizon is

$$L_{t+K\Delta} = -V_t \left(\prod_{k=1}^K w^T e^{X_{t+k\Delta}} - 1 \right).$$

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- (b) **(15 Points)** Write the simulation to obtain the time-aggregated risk measures, building into your simulation that (1) weights are held constant and (2) we use EWMA *each day* to obtain a new conditional mean and covariance estimate.

As for (2), use the historical data to obtain an estimate $\mu_{t+\Delta}, \Sigma_{t+\Delta}$ as of the initial time t . Since we have 5 years of data (about 1200 points) it is OK to start the EWMA procedure off with $\mu_{t_0} = 0 = \Sigma_{t_0}$.

Write $T_k = t + k\Delta$. From the historical data we have obtained μ_{T_1}, Σ_{T_1} and can sample $X_{T_1} \sim N(\mu_{T_1}, \Sigma_{T_1})$. Next, for $k = 2, \dots, K$ write your simulation to include updating the mean and covariance estimates! Indeed, once we have sampled X_{T_1} we obtain our next estimate by setting

$$\begin{aligned}\mu_{T_2} &= \lambda\mu_{T_1} + (1 - \lambda)X_{T_1}; \\ \Sigma_{T_2} &= \lambda\Sigma_{T_1} + (1 - \theta)(X_{T_1} - \mu_{T_1})(X_{T_1} - \mu_{T_1})^T.\end{aligned}$$

We then sample $X_{T_2} \sim N(\mu_{T_2}, \Sigma_{T_2})$ and obtain μ_{T_3}, Σ_{T_3} by EWMA accordingly. We repeat this over the K day horizon.

With this methodology, write a simulation to estimate the K day VaR, ES as well as the spectral risk measure with exponential weighting function $\phi_\gamma(u) = \frac{\gamma}{e^\gamma - 1} e^{\gamma u}, 0 \leq u \leq 1$. To compare with square root of time, first do this for $K = 1$ to estimate the one day risk measures. Then do it for $K = 10$ for the ten day measures. Output the ten day measures as well as the square root of time approximations. For parameter values take $\alpha = .99$, $\gamma = 30$, $\lambda = .94$, $\theta = .97$. Have your simulations perform $N = 50,000$ trials.

For error checking purposes, I am obtaining around \$70,000, \$90,000 and \$78,000 for the VaR, ES and spectral risk measure respectively.

$$1. (a) L \sim \text{Exp}(1/\theta)$$

$$f(l) = \begin{cases} \frac{1}{\theta} e^{-l/\theta} & l \geq 0 \\ 0 & l < 0 \end{cases}$$

$$\Rightarrow F(l) = \begin{cases} 1 - e^{-l/\theta} & l \geq 0 \\ 0 & l < 0 \end{cases}$$

$$P[L \leq \tau] = F(\tau) = \int_0^{\tau} f(l) dl = 1 - e^{-\tau/\theta}$$

$$\alpha = 1 - e^{-\tau/\theta} \quad 1 - \alpha = e^{-\tau/\theta} \quad -\tau/\theta = \ln(1 - \alpha) \Rightarrow \tau = -\theta \ln(1 - \alpha)$$

$$\text{VaR}_{\alpha}(L) = -\theta \ln(1 - \alpha)$$

$$\text{ES}_{\alpha}(L) = \frac{1}{1 - \alpha} \int_{\alpha}^1 -\theta \ln(1 - u) du$$

$$= \frac{-\theta}{1 - \alpha} \int_{\alpha}^1 \ln(1 - u) du$$

$$= \frac{-\theta}{(1 - \alpha)} \int_0^{1 - \alpha} \ln x dx$$

$$= \frac{-\theta}{(1 - \alpha)} \left[x \ln x - \int x d \ln x \right] \Big|_0^{1 - \alpha}$$

$$= \frac{-\theta}{(1 - \alpha)} \left[x \ln x - x \right] \Big|_0^{1 - \alpha}$$

$$= \frac{-\theta}{(1 - \alpha)} \left[(1 - \alpha) \ln(1 - \alpha) - (1 - \alpha) \right]$$

$$= -\theta \ln(1 - \alpha) + \theta$$

(2) $L = \max\{L_1, L_2\}$ L_1, L_2 are $\text{Exp}(1/\theta)$ (iid)

$$F_L = P\{L < t\} = P\{L_1 < t, L_2 < t\} = P\{L_1 < t\} \cdot P\{L_2 < t\} = P^2\{L < t\} = F^2(t)$$

$$\Rightarrow F_L(t) = \begin{cases} (1 - e^{-t/\theta})^2 & t \geq 0 \\ 0 & t < 0 \end{cases}$$

$$\alpha = (1 - e^{-t/\theta})^2 \Rightarrow \text{VaR}_\alpha(L) = -\theta \ln(1 - \sqrt{\alpha})$$

$$ES_\alpha(L) = \frac{1}{1-\alpha} \int_\alpha^1 \text{VaR}_u(L) du = \frac{-\theta}{1-\alpha} \int_\alpha^1 \ln(1 - \sqrt{u}) du$$

$$= \frac{-\theta}{1-\alpha} \int_{1-\sqrt{\alpha}}^0 \ln x d(1-x)^2 = \frac{-\theta}{1-\alpha} \int_0^{1-\sqrt{\alpha}} \ln x \cdot 2 \cdot (1-x) dx$$

$$= \frac{-2\theta}{1-\alpha} \int_0^{1-\sqrt{\alpha}} \ln x dx + \frac{2\theta}{1-\alpha} \int_0^{1-\sqrt{\alpha}} x \ln x dx$$

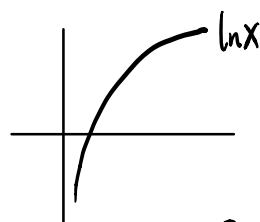
$$\textcircled{1} \int_0^{1-\sqrt{\alpha}} \ln x dx = (1-\sqrt{\alpha}) \ln(1-\sqrt{\alpha}) - (1-\sqrt{\alpha})$$

$$\textcircled{2} \int_0^{1-\sqrt{\alpha}} x \ln x dx = \left[\frac{1}{2} x^2 \ln(x) - \frac{x^2}{4} \right] \Big|_0^{1-\sqrt{\alpha}} = \frac{1}{2} \ln(1-\sqrt{\alpha}) + \frac{1}{2} \ln(1-\sqrt{\alpha}) - \frac{(1-\sqrt{\alpha})^2}{4}$$

$$\text{By } \textcircled{1} \text{ and } \textcircled{2} \Rightarrow ES_\alpha(L) = - \frac{\theta [(1-2\ln(1-\sqrt{\alpha}))\alpha + 2\sqrt{\alpha} + 2\ln(1-\sqrt{\alpha}) - 3]}{2(1-\alpha)}$$

According to (a), when $\alpha \rightarrow 1$ $ES_\alpha(L) \rightarrow \infty$

Similarly, $\alpha \rightarrow 1$ in the (b), $ES_\alpha(L)$ also go to ∞ .



$$(3) \alpha = P\{g(Y) \leq \text{VaR}_\alpha(g(Y))\} = P\{Y \leq g^{-1}(\text{VaR}_\alpha(g(Y)))\}$$

$$\Rightarrow \text{VaR}_\alpha(Y) = g^{-1}(\text{VaR}_\alpha(g(Y)))$$

$$\Rightarrow g(\text{VaR}_\alpha(Y)) = \text{VaR}_\alpha(g(Y))$$

(d) According to (c), $\text{VaR}_\alpha(g(Y)) = g(\text{VaR}_\alpha(X))$

$$\Rightarrow g(\text{VaR}_\alpha(X)) = [-\theta \ln(1-\alpha)]^2 = \theta^2 [\ln(1-\alpha)]^2$$

$$\Rightarrow \text{ES}_\alpha(g(Y)) = \frac{1}{1-\alpha} \int_\alpha^1 \theta^2 [\ln(1-u)]^2 du$$

$$= \frac{\theta^2}{1-\alpha} \cdot \int_0^{1-\alpha} [\ln x]^2 dx$$

$$= \frac{\theta^2}{1-\alpha} [x[\ln x]^2 - 2x \ln x + 2x] \Big|_0^{1-\alpha}$$

$$= \frac{\theta^2}{1-\alpha} [(1-\alpha)[\ln(1-\alpha)]^2 - 2(1-\alpha)\ln(1-\alpha) + 2(1-\alpha)]$$

$$= \theta^2 [\ln(1-\alpha)]^2 - 2\theta^2 \ln(1-\alpha) + 2\theta^2 = \theta^2 [\ln(1-\alpha)]^2 - 2\theta^2 \ln(1-\alpha) + 1 + \theta^2$$

$$g(\text{ES}_\alpha(Y))$$

$$= \theta^2 [1 - \ln(1-\alpha)]^2$$

$$= \theta^2 [1 + [\ln(1-\alpha)]^2 - 2\ln(1-\alpha)]$$

$$\Rightarrow g(\text{ES}_\alpha(Y)) = \text{ES}_\alpha(g(Y)) - \theta^2$$

So, it is not same.

when $\alpha \rightarrow 1$, $\ln(1-\alpha) \rightarrow \infty$, the $\text{ES}_\alpha(g(Y))$ and $g(\text{ES}_\alpha(Y))$ will go to infinity.

$$\begin{aligned}
2.(a) \quad L_{t+k\Delta} &= -V_t \sum_{i=1}^d w^{(i)} (e^{\sum_{k=1}^K X_{T_k}^{(i)}} - 1) \\
&= -V_t \left(\sum_{i=1}^d w^{(i)} e^{\sum_{k=1}^K X_{T_k}^{(i)}} - 1 \right) \\
&= -V_t (w_1 \cdot e^{\sum_{k=1}^K X_{T_k}^1} + w_2 e^{\sum_{k=1}^K X_{T_k}^2} + \dots + w_d e^{\sum_{k=1}^K X_{T_k}^d} - 1) \\
&= -V_t (w^T \cdot e^{\sum_{k=1}^K X_{T_k}} - 1) \quad e^{\sum_{k=1}^K X_{T_k}^1} \\
&= -V_t \left(\prod_{k=1}^K w^T e^{X_{t+k\Delta}} - 1 \right) \quad = e^{X_{t+\Delta}^1 + X_{t+2\Delta}^1 + \dots + X_{t+K\Delta}^1} \\
&\quad = \prod_{k=1}^K e^{X_{t+k\Delta}^1}
\end{aligned}$$

b) The Calculation in the Python code.

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K day VaR: 94704.97
sqrt(K) times one day VaR: 90897.60
K day ES:: 112831.38
sqrt(K) times one day ES: 105238.29
K day Spectral: 78428.17
sqrt(K) times one day Spectral: 77964.83

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