



# PROBABILITY REVIEW(?)

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# Predicting amount of rainfall



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# (Linear) Regression

- $h_{\theta}(x) = \theta_0 + \theta_1 x_1 + \theta_2 x_2 + \theta_3 x_3 + \theta_4 x_4 + \theta_5 x_5$

- $\theta$ s are the parameter (or weights)

Assume  $x_0$  is always 0

- We can rewrite

$$h_{\theta}(x) = \sum_{i=0}^n \theta_i x_i = \theta^T \mathbf{x}$$

- Notation: vectors are bolded
- Notation: vectors are column vectors

## LMS regression with gradient descent

$$\frac{\partial J}{\partial \theta_j} = -\sum_{i=1}^m (y_i - \theta^T \mathbf{x}_i) x_i^{(j)}$$

$$\theta_j \Leftarrow \theta_j + r \sum_{i=1}^m (y_i - \theta^T \mathbf{x}_i) x_i^{(j)}$$

Interpretation?

# Logistic Regression

- Pass  $\theta^T \mathbf{x}$  through the logistic function

$$h_{\theta}(x) = \frac{1}{1 + e^{-\theta^T x}}$$

# Logistic Regression update rule

$$\theta_j \Leftarrow \theta_j - r \sum_{i=1}^m (y_i - h_{\theta}(x_i)) x_i^{(j)}$$

Update rule for linear regression

$$\theta_j \Leftarrow \theta_j - r \sum_{i=1}^m (y_i - \theta^T \mathbf{x}_i) x_i^{(j)}$$

# What is Probability?

- Frequentist

Probability = rate of occurrence in a large number of trials

- Bayesian

Probability = uncertainty in your knowledge of the world



# Bayesian vs Frequentist





# Bayesian vs Frequentist


- Toss a coin
- Frequentist
  - $P(\text{head}) = \theta$ ,  $\theta = \text{\#heads}/\text{\#tosses}$
- Bayesian
  - $P(\text{head}) = \theta$ ,  $\theta \sim U(0.6, 1.0)$
  - Parameters of distributions can now have probabilities
  - Bayesian interpretation can give prior knowledge to the phenomena – subjective view of the world
  - Prior knowledge can be updated according to the observed frequency

# Bayesian statistics

- Coin with  $P(\text{head}) = p$
- Observed frequency of heads  $\hat{p} = \text{\#heads}/\text{\#n}$
- In Bayesian view, we can talk about  $P(p \mid \hat{p})$  by using Bayes's rule

$$P(p \mid \hat{p}) = \frac{P(\hat{p} \mid p) P(p)}{P(\hat{p})}$$

Prior probability



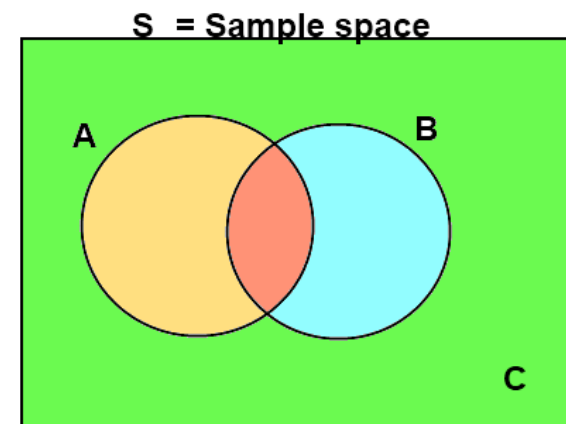
# Important concepts

- Conditional probability
  - Independence
- Bayes' Rule
- Expected Value and Variance
- CDFs
- Sum of RVs
- Gaussian Random Variable
  - Multivariate Gaussian

# Conditional probability

- $P(A|B)$  probability of A given B has occurred

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$



- A student posts a facebook status after finishing Pattern Recognition homework
- $P(\text{he is happy})$
- $P(\text{he is happy} \mid \text{the post starts with “\#\$@\#\$!@\#\$”})$

# Independence

- Two events are independent (statistically independent or stochastically independent) if the occurrence of one does not affect the probability of occurrence of the other.

$$P(A \cap B) = P(A)P(B) \Leftrightarrow P(B) = P(B | A)$$

- $P(\text{he is happy} \mid \text{His friend posted a cat picture on instgram})$

# Bayes' Rule (Bayes's theorem or Bayes' law)

$$P(A | B) = \frac{P(B | A) P(A)}{P(B)},$$

Usefulness: We can find  $P(A|B)$  from  $P(B|A)$  and vice versa

# Expected value

- Expected value

$$E[x] = \int_{-\infty}^{\infty} xp(x)dx$$

$$E[g(x)] = \int_{-\infty}^{\infty} g(x)p(x)dx$$

- Variance ( $\sigma^2$ ) (Standard Deviation =  $\sigma$ )

$$Var[x] = E[(x - E[x])^2] = \sigma^2 = \int_{-\infty}^{\infty} (x - E[x])^2 p(x)dx$$

$$E[(x - E[x])^2] = E[x^2] - (E[x])^2$$

# Expected Value notes

- It's a weighted sum.
- Something can have a high expected value but low probability of occurring
- Lottery:  $P(\text{win}) = 10^{-20}$ , winner gets  $10^{30}$
- $P(\text{loss}) = 1 - P(\text{win})$ , loser gets  $-10$
- $E(\text{Lottery earnings}) = 10^{-20}10^{30} + (1 - P(\text{win}))(-10)$
- $= 10^{10} - 10$
- Humans are not good at gauging probability at extreme points



## Expected value and Variance properties

- $E[a] = a$ ;  $a$  is a constant.
- $E[aX+b] = aE[X]+b$
- $E[X+Y] = E[X]+E[Y]$
- $\text{Var}[a] = 0$
- $\text{Var}[aX+b] = a^2\text{Var}[X]$

Conditional Expected Value

$$E[x | A] = \int_{-\infty}^{\infty} xp(x | A)dx$$

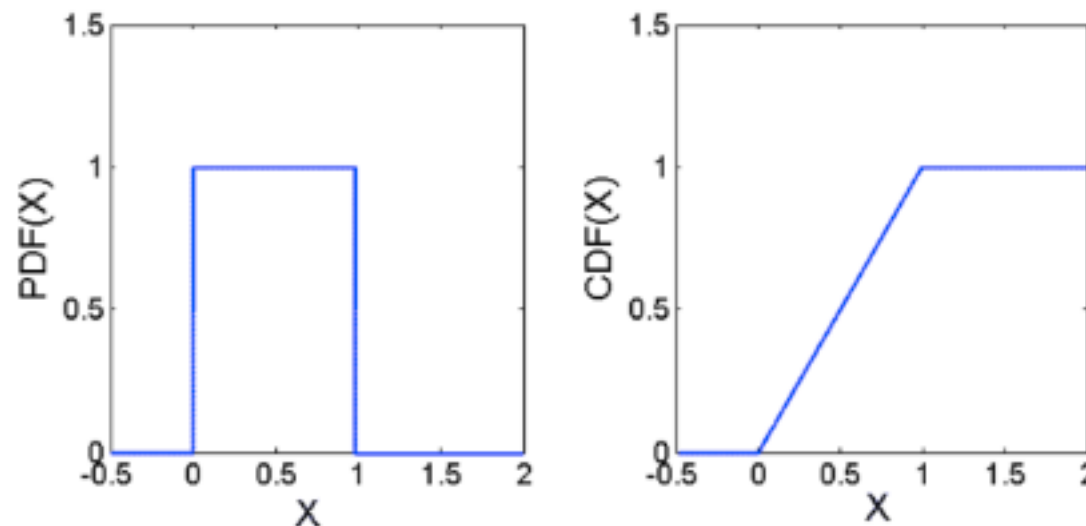
$$E[g(x) | A] = \int_{-\infty}^{\infty} g(x)p(x | A)dx$$

# Cumulative Distribution Functions CDFs

- Probability that the RV is less than a certain amount

$$F_X(x_0) = P(X \leq x_0) = \int_{-\infty}^{x_0} p(x)dx$$

- CDF is the integral of PDF. Differentiating CDF wrt x gives the PDF



# Joint distributions

- If we want to monitor how two events are jointly occurring, we consider the joint distribution  $p_{X,Y}(x,y)$
- $p_{X,Y}(x,y) = p_X(x)p_Y(y)$  if  $x$  and  $y$  are independent

$$P(A) = \int \int_A p_{XY}(x, y) dx dy$$

$$p_X(x) = \int_{-\infty}^{\infty} p_{XY}(x, y) dy$$

$$p_Y(y) = \int_{-\infty}^{\infty} p_{XY}(x, y) dx$$

# Sum of Random variables

- $Z = Y + X$
- What is the pdf of  $Z$ ? Where  $Y$  and  $X$  continuous RVs

$$p_{X+Y}(z) = (p_X * p_Y)(z) = (p_Y * p_X)(z)$$

# Central Limit Theorem (CLT)

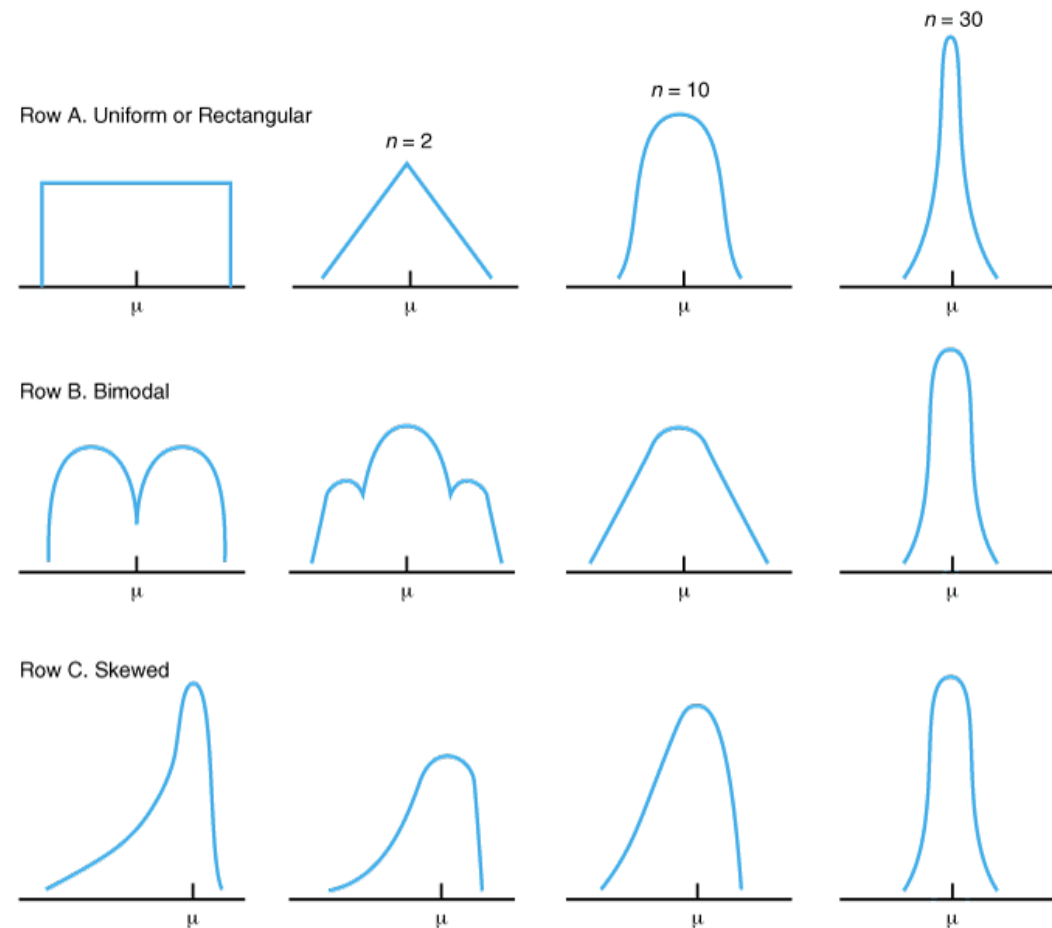
- Supposed  $X_1, X_2, \dots$  is a sequence of iid (independent and identically distributed) RVs. As  $n$  approaches infinity the sum of the sequence converge in distribution to a Normal distribution

$$\sqrt{n} \left( \left( \frac{1}{n} \sum_{i=1}^n X_i \right) - \mu \right) \xrightarrow{d} N(0, \sigma^2)$$

- Other variants of CLT exists, without the dependence or identically distributed assumption

# CLT implications

- A sum of RVs tends to become Normally distributed very quickly



# Gaussian distribution (normal distribution)

- $X$  is normal (Gaussian):  $X \sim N(\mu, \sigma^2)$

$$p(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/2\sigma^2}$$

$$E[x] = \mu$$

$$Var[x] = \sigma^2$$

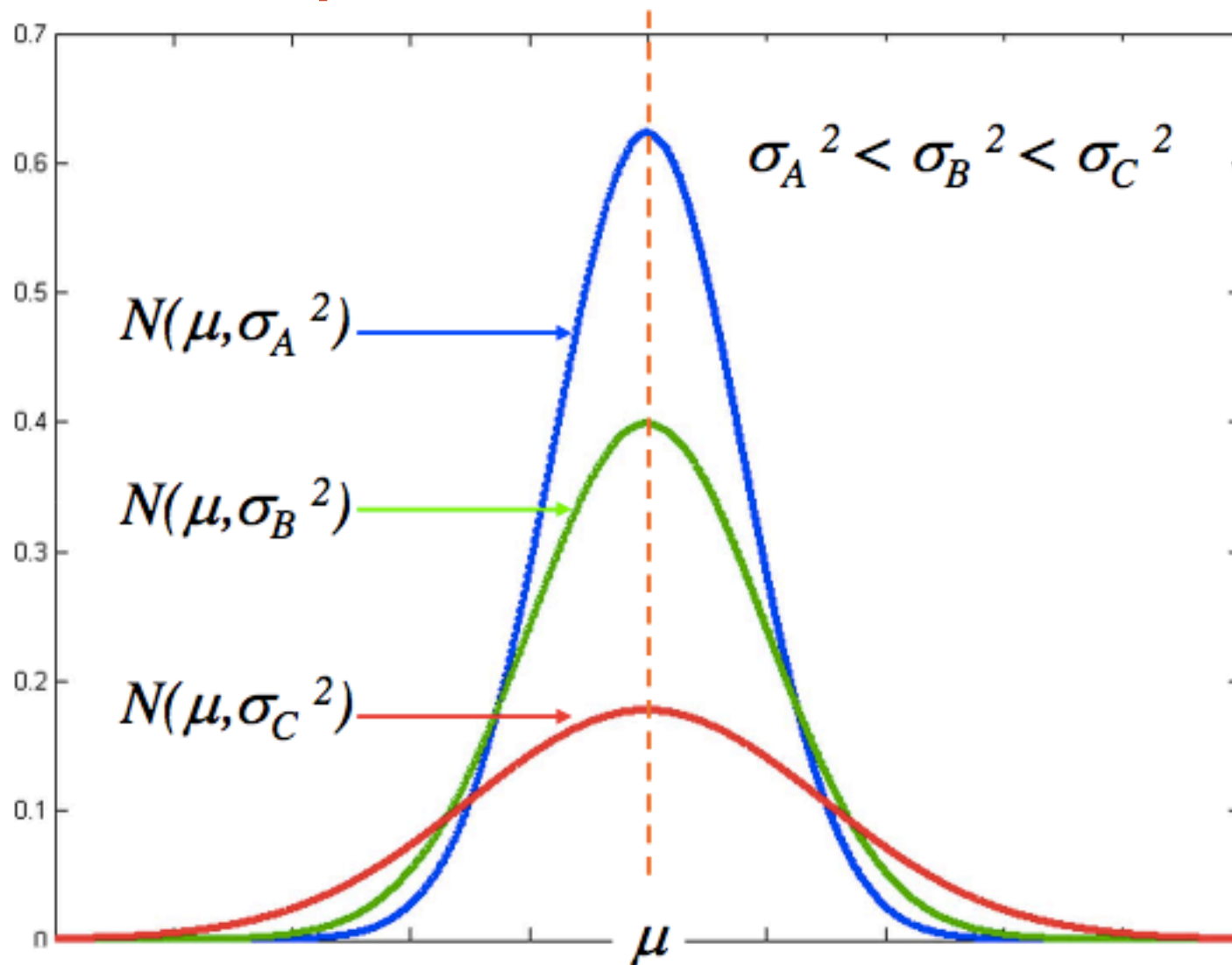
- $X$  is Standard normal (Standard Gaussian):  
 $X \sim N(0,1)$  when  $\mu=0, \sigma^2=1$

$$p(x) = \frac{1}{\sqrt{2\pi}} e^{-(x-\mu)^2/2}$$

$$E[x] = 0$$

$$Var[x] = 1$$

# Gaussian pdf





# Linear transformation of Gaussian RV

- Normality is preserved by linear transformation. Calculation involving the normal variable is usually done in terms of standard normal.
- Let  $Y=aX+b$ ,  
if  $X \sim N(\mu, \sigma^2) \rightarrow Y \sim N(a\mu+b, a^2 \sigma^2)$
- Let  $Z=(X-\mu)/\sigma$ ,  
if  $X \sim N(\mu, \sigma^2) \rightarrow Z \sim N(0,1) : \text{Standard Normal}$

Can you prove this?

# Expectation of multivariate distributions

$$E[g(X_1, X_2)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x_1, x_2) p_{X_1, X_2}(x_1, x_2) dx_1 dx_2$$

$$E[g(X_1)h(X_2)] = E[g(X_1)]E[h(X_2)]$$

If  $X_1$  and  $X_2$  independent

# Covariance of multivariate distributions

- $\text{cov}(X_1, X_2) = E[(X_1 - m_1)(X_2 - m_2)]$
- $\text{cov}(X_1, X_2) = E[(X_1)(X_2)] - m_1 m_2$
- Covariance with itself is just the Variance
- Correlation

$$\rho = \frac{\text{cov}(X_1, X_2)}{\sqrt{V(X_1)V(X_2)}}$$

# Covariance matrix

- Given a set of RVs,  $X_1 X_2 \dots X_n$
- The covariance matrix is a matrix which has the covariance of the  $i$  and  $j$  RV in position  $(i,j)$

$$\Sigma = \begin{bmatrix} E[(X_1 - \mu_1)(X_1 - \mu_1)] & E[(X_1 - \mu_1)(X_2 - \mu_2)] & \cdots & E[(X_1 - \mu_1)(X_n - \mu_n)] \\ E[(X_2 - \mu_2)(X_1 - \mu_1)] & E[(X_2 - \mu_2)(X_2 - \mu_2)] & \cdots & E[(X_2 - \mu_2)(X_n - \mu_n)] \\ \vdots & \vdots & \ddots & \vdots \\ E[(X_n - \mu_n)(X_1 - \mu_1)] & E[(X_n - \mu_n)(X_2 - \mu_2)] & \cdots & E[(X_n - \mu_n)(X_n - \mu_n)] \end{bmatrix}.$$

# Understanding the Covariance matrix

$$= \begin{matrix} \text{A} & & \text{B} \\ \left[ \begin{array}{cc} \text{Cov}(X, X) & \text{Cov}(Y, X) \\ \text{Cov}(X, Y) & \text{Cov}(Y, Y) \end{array} \right] & & \\ \text{C} & & \text{D} \end{matrix}$$

Which statements are true?

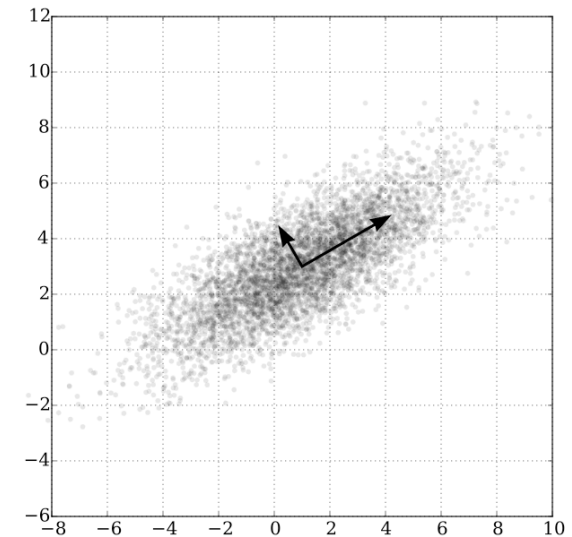
$$C = B$$

$$B < 0$$

$$D < 0$$

$$A < D$$

$$A > B$$



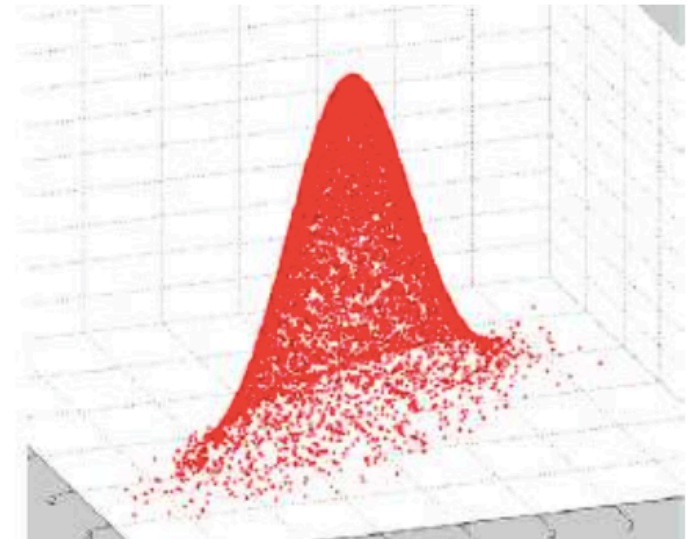
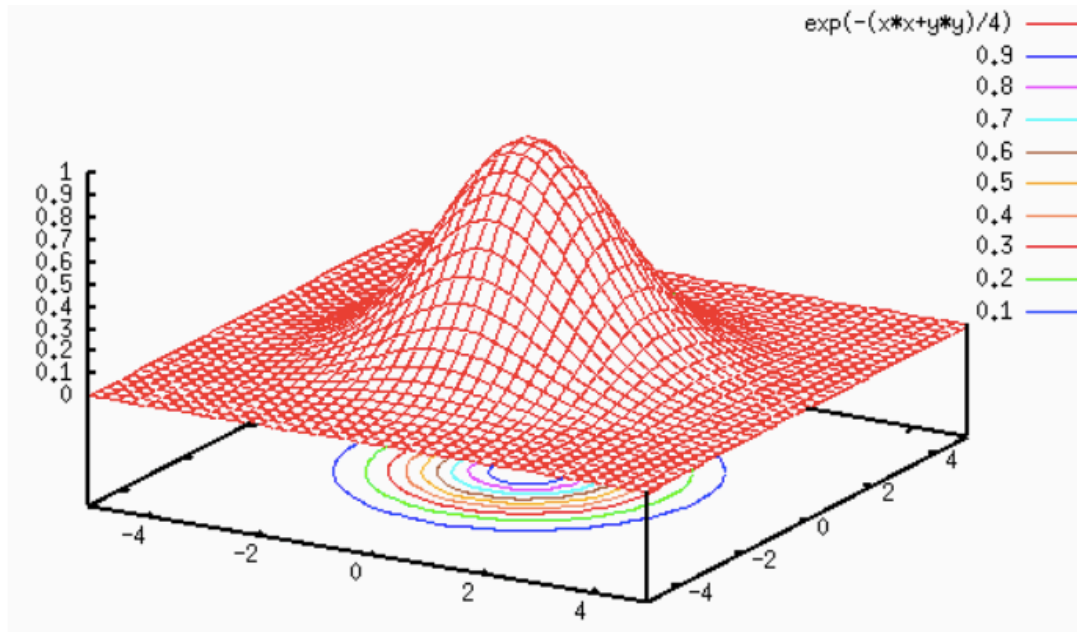
# Covariance matrix observations

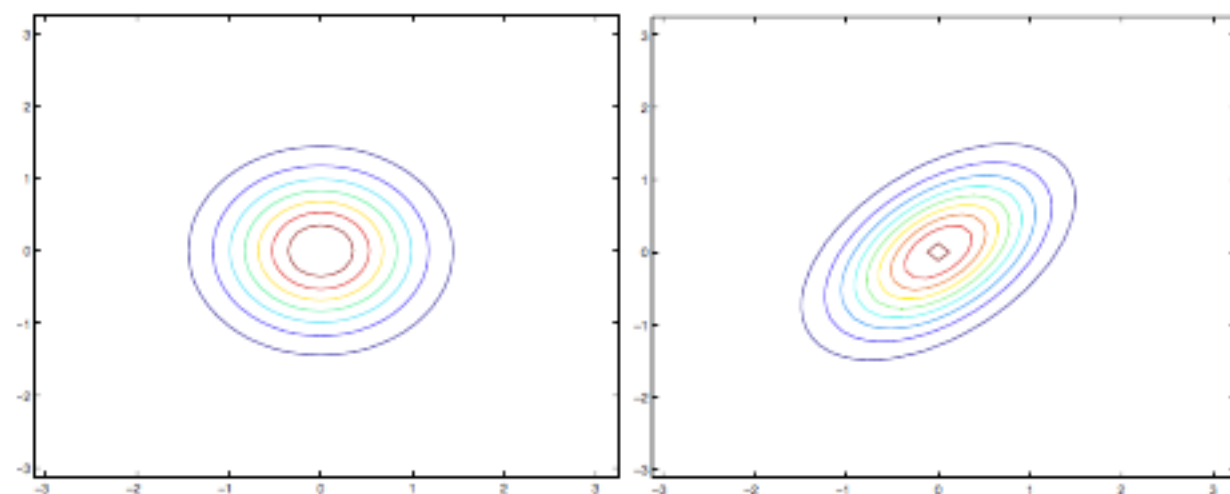
- $\Sigma = \Sigma^T$
- If the covariance matrix is diagonal, all RVs are mutually independent.
- Covariance matrix is positive-semidefinite

# Multivariate Gaussian distribution

- Put  $X_1, X_2, X_3 \dots X_n$  into a vector  $\mathbf{x}$

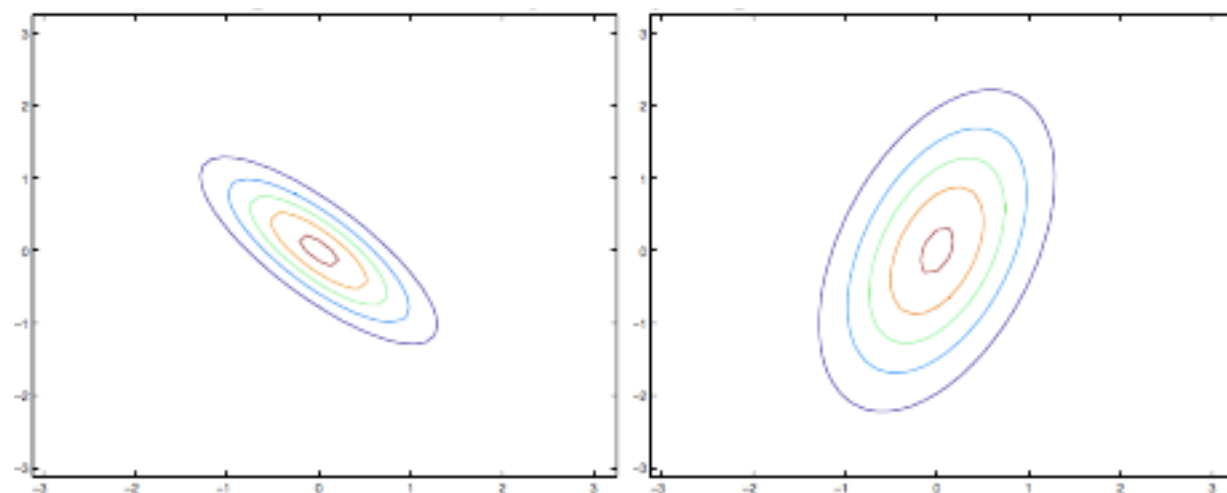
$$p(\mathbf{x}; \mu, \Sigma) = \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp\left[-\frac{1}{2} (\mathbf{x} - \mu)^t \Sigma^{-1} (\mathbf{x} - \mu)\right]$$





$$\Sigma = \begin{bmatrix} 1 & -0.8 \\ -0.8 & 1 \end{bmatrix}$$

$$\Sigma = \begin{bmatrix} 3 & 0.8 \\ 0.8 & 1 \end{bmatrix}$$



$$\Sigma = \begin{bmatrix} 1 & 0.5 \\ 0.5 & 1 \end{bmatrix}$$

$$\Sigma = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$



# Affine transformation of multivariate Gaussians

- $\mathbf{y} = \mathbf{Ax} + \mathbf{b}$  , Assuming A has full rank (invertible)

$$\mathbf{x} \sim N(\mu, \Sigma)$$

$$\mathbf{y} \sim N(A\mu + b, A\Sigma A^T)$$

# Important concepts

- Conditional probability
  - Independence
- Bayes' Rule
- Expected Value and Variance
- CDFs
- Sum of RVs
  - CLT
- Gaussian Random Variable
  - Multivariate Gaussian

# Distribution parameter estimation

- [illegible]

# Linear Regression Revisit

- $h_{\theta}(x) = \theta_0 + \theta_1 x_1 + \theta_2 x_2 + \theta_3 x_3 + \theta_4 x_4 + \theta_5 x_5$
- $\theta$ s are the parameter (or weights)
- We can rewrite

$$h_{\theta}(x) = \sum_{i=0}^n \theta_i x_i = \theta^T \mathbf{x}$$

- Notation: vectors are bolded
- Notation: vectors are column vectors

# Probabilistic Interpretation of linear regression

- Real world data is our model plus some error term
  - Noise in the data
  - Something that we do not model (features we are missing)
- Let's assume the error is normally distributed with mean zero and variance  $\sigma^2$ 
  - Why Gaussian?
  - Why saying mean is zero is a valid assumption?

$$y_i = \theta^T \mathbf{x}_i + \epsilon_i$$

# Probabilistic view of Linear regression

- Find  $\theta$
- Maximize Likelihood of seeing  $x$  and  $y$  in training
- From our assumption we know that

$$y_i = \theta^T \mathbf{x}_i + \epsilon_i$$

error

$$p(y_i | \mathbf{x}_i; \theta) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(y_i - \theta^T \mathbf{x}_i)^2}{2\sigma^2}\right)$$

Error term is normally distributed with mean 0 and variance  $\sigma^2$

# Maximizing Likelihood

What is the assumption here?  
Is it accurate?

- Max  $L(\theta) = \prod_{i=1}^m p(y_i | \mathbf{x}_i; \theta)$
- We use the log likelihood instead  $\log(L(\theta)) = l(\theta)$

From our previous lecture

$$\text{Min } J(\theta) = \frac{1}{m} \sum_{i=1}^m (y_i - \theta^T \mathbf{x}_i)^2$$

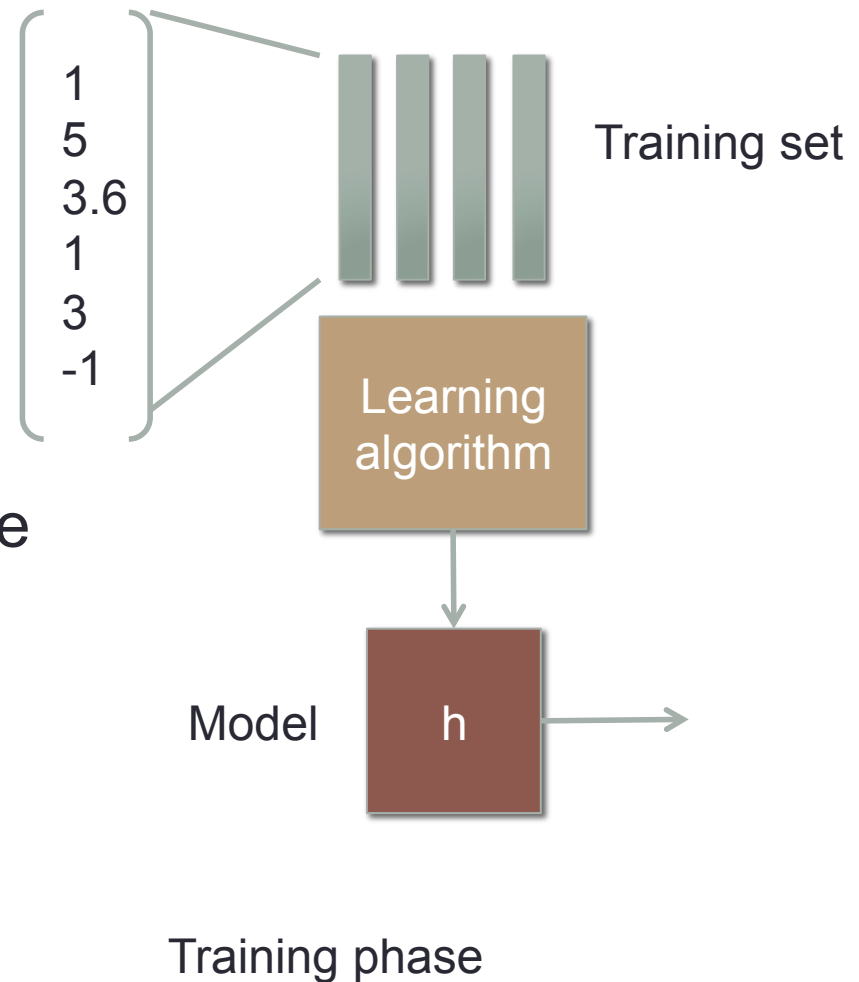
# Mean square error solution and MLE solution

- Turns out MLE and MSE gets to the same solution
  - This justifies our choice of MSE as the Loss for linear regression
  - This does not mean MSE is the best Loss for regression, but you can at least justify it with a probabilistic reasoning
- Note how our choice of variance  $\sigma^2$  falls out of the maximization, so this derivation is true regardless of which assumption for variance is.



# Flood or no flood

- What would be the output?
- $y = 0$  if not flooded
- $y = 1$  if flooded
- Anything in between is a score for how likely it is to flood

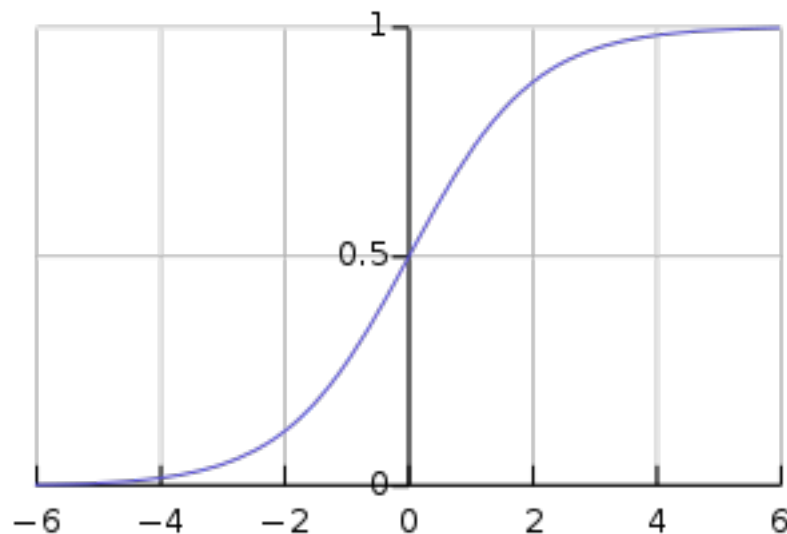


# Can we use regression?

- Yes
- $h_{\theta}(x) = \theta_0 + \theta_1 x_1 + \theta_2 x_2 + \theta_3 x_3 + \theta_4 x_4 + \theta_5 x_5$
- But
- What does it mean when  $h$  is higher than 1?
- Can  $h$  be negative? What does it mean to have a negative flood value?

# Logistic function

- Let's force h to be between 0 and 1 somehow
- Introducing the logistic function (sigmoid function)



$$\begin{aligned} f(x) &= \frac{1}{1 + e^{-x}} \\ &= \frac{e^x}{1 + e^x} \end{aligned}$$

# Logistic Regression

- Pass  $\theta^T \mathbf{x}$  through the logistic function

$$h_{\theta}(x) = \frac{1}{1 + e^{-\theta^T x}}$$

# Loss function?

- MSE error no longer a good candidate
- Let's turn to use probabilistic argument for logistic regression

# Logistic Function derivative

The derivative has a nice property by design.

This is also why many algorithm we'll learn later in class also uses the logistic function

$$\begin{aligned}g'(z) &= \frac{d}{dz} \frac{1}{1 + e^{-z}} \\&= \frac{1}{(1 + e^{-z})^2} (e^{-z}) \\&= \frac{1}{(1 + e^{-z})} \cdot \left(1 - \frac{1}{(1 + e^{-z})}\right) \\&= g(z)(1 - g(z)).\end{aligned}$$

# Probabilistic view of Logistic Regression

- Let's assume, we'll classify as 1 with probability in accordance to the output of

$$h_{\theta}(x) = \frac{1}{1 + e^{-\theta^T x}}$$

$$P(y = 1 \mid x; \theta) = h_{\theta}(x)$$

$$P(y = 0 \mid x; \theta) = 1 - h_{\theta}(x)$$

or

$$p(y \mid x; \theta) = (h_{\theta}(x))^y (1 - h_{\theta}(x))^{1-y}$$

# Maximizing log likelihood

$$p(y \mid x; \theta) = (h_{\theta}(x))^y (1 - h_{\theta}(x))^{1-y}$$

$$\begin{aligned} g'(z) &= \frac{d}{dz} \frac{1}{1 + e^{-z}} \\ &= \frac{1}{(1 + e^{-z})^2} (e^{-z}) \\ &= \frac{1}{(1 + e^{-z})} \cdot \left(1 - \frac{1}{(1 + e^{-z})}\right) \\ &= g(z)(1 - g(z)). \end{aligned}$$

$$\theta_j \Leftarrow \theta_j + r \sum_{i=1}^m (y_i - h_{\theta}(x_i)) x_i^{(j)}$$



## Logistic Regression update rule

$$\theta_j \Leftarrow \theta_j + r \sum_{i=1}^m (y_i - h_{\theta}(x_i)) x_i^{(j)}$$

Update rule for linear regression

$$\theta_j \Leftarrow \theta_j + r \sum_{i=1}^m (y_i - \theta^T \mathbf{x}_i) x_i^{(j)}$$

## Loose ends from previous lecture

$$\theta_j \Leftarrow \theta_j + r \sum_{i=1}^m (y_i - \theta^T \mathbf{x}_i) x_i^{(j)}$$

$$\theta = (X^T X)^{-1} X^T y$$

$$X = \begin{bmatrix} - & x_1^T & - \\ - & x_2^T & - \\ - & x_m^T & - \end{bmatrix}$$

X is mxn    The inverse is on nxn matrix

# Homework

- $\nabla_{Atr}ABA^TC = CAB + C^TAB^T$

Hint: Try first solving the easier equation of  $\nabla_{Atr}BAC = (CB)^T = B^TC^T$

# Homework II

Can I just use scikit-learn?

No

# One button machines

- Machine learning as a tool for non-experts
- Can a non-expert just provide the data and let the machine decide how to proceed

DataRobot

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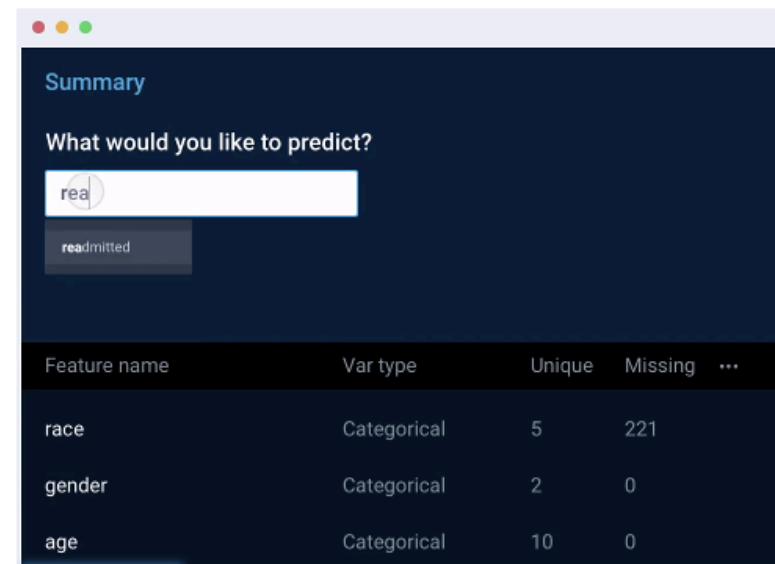
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# Reinforcement Learning for Model Selection

- Tuning a network takes time
- Let machine learning learn how to tune a network
- Matches or outperforms ML experts performance

