# Sensitivity and Effective Samples

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## Plan for today

- First homework
  - Some issues with identification / estimation
- Second homework
  - Some basic asymptotics
- Effective Samples in regression
  - Aronow & Samii (2013)
- Confounding & Sensitivity Analysis
  - Imbens (2003)
  - And generalizations (which are easy)

# Identification / Estimation

- The "Four questions" may have been a little confusing.
- Let me re-break down the last two questions as "Identification" and "Estimation"
- "A regression is causal when the CEF it approximates is causal" MHE
- Identification consists (in this context) as the set of things you need to believe in order to believe the CEF is causal.
- Thus, when we think about identification, we should think about assumptions.
- Estimation is the process you use to estimate the CEF.

### In practice

- The Conditional Independence Assumption would fall in identification.
- I'd also throw in assumptions of additivity / linearity in this pot.
- Estimation would include the populations/samples of interest, and all statistical inference.
- But there's a close connection between the two.
- Problems in estimation can lead to changes in identifying assumptions.
- So they aren't completely separable.

### Basic Identities and Definitions

• Convergence in probability:

$$X_n \stackrel{p}{\to} X$$
 if  $\forall \epsilon > 0$   $\Pr(|X_n - X| > \epsilon) \to 0$  as  $n \to \infty$ 
• A very basic decomposition of variance:

- $var(X) = E[X^2] E[X]^2$
- Something we do NOT know:  $E[XY] = E[X] \times E[Y]$
- Pretty much everybody assumed this to be true.

### Problem Set 2

- We want to find the probability limit of  $\frac{D_d'D_d}{N}$
- Pretty much everyone correctly observed that this was:

$$\begin{pmatrix} \frac{1}{N} \sum_{i} (X_{i} - \bar{X})^{2} & \frac{1}{N} \sum_{i} (X_{i} - \bar{X})(Z_{i} - \bar{Z}) \\ \frac{1}{N} \sum_{i} (X_{i} - \bar{X})(Z_{i} - \bar{Z}) & \frac{1}{N} \sum_{i} (Z_{i} - \bar{Z})^{2} \end{pmatrix}$$

• We can't simply say that this is equal to the following, though:

$$\begin{pmatrix} \frac{1}{N} \sum_{i} (X_i - p)^2 & \frac{1}{N} \sum_{i} (X_i - p)(Z_i - q) \\ \frac{1}{N} \sum_{i} (X_i - p)(Z_i - q) & \frac{1}{N} \sum_{i} (Z_i - q)^2 \end{pmatrix}$$

This is a fact which you should all be able to recognize as false IMMEDI-ATELY.

#### Continued

- The sample variance is BIASED when divided by N rather than N-1.
- It is still consistent, though, which is what is relevant to us.
- The simple proof is as follows:

$$\frac{\sum_i (X_i - \bar{X})^2}{N} = \frac{1}{N} \left( \sum_i X_i^2 - 2 \sum_i \bar{X} X_i + \sum_i \bar{X}^2 \right)$$

- We can pull out constants from the sums:  $\frac{1}{N} \left( \sum_{i} X_i^2 - 2\bar{X} \sum_{i} X_i + \bar{X}^2 \sum_{i} 1 \right)$
- And simplify some things:  $\frac{\sum_i X_i^2}{N} 2\bar{X}^2 + \bar{X}^2 = \frac{\sum_i X_i^2}{N} \frac{\sum_i \bar{X}^2}{N}$  This is the point at which we can use the LLN. Averages converge in probability to their expectations.  $E[X^2] - E[\bar{X}^2]$
- Use this identity:  $E[Z^2] = E[Z]^2 + var(Z)$
- Which gives us:  $E[X]^2 + var(X)$  and  $E[\bar{X}]^2 + var(\bar{X})$
- Substituting in these identities:  $E[X]^2 + var(X) - E[\bar{X}]^2 - var(\bar{X})$
- Which gives us:  $p^2+p(1-p)-p^2-\frac{p(1-p)}{N}=\frac{N-1}{N}p(1-p)$  Which converges in probability to p(1-p) (i.e. it is consistent)

## Finishing it up:

- After that there weren't many issues
- We know off-diagonals are zero because X and Z are independent.
- This means we have a diagonal matrix, which is easy to invert (just take the inverse of the diagonal elements)
- We then use a similar method to find  $D'_dY$ , and multiply them together, which completes the first part.
- For the next part, observe that we know  $var(X_i) = p(1-p)$ , so we only need to think about the expectation of the numerator.
- We end up with the following: cov(X,Y) = E[XY] - E[X]E[Y] = E[Y|X = 1]p(X = 1) - pE[Y]
- We then use the law of total probability on E[Y]: pE[Y|X=1] - p(pE[Y|X=1] + (1-p)E[Y|X=0])
- Distribute and collect terms:
  - p(1-p) (E[Y|X=1] E[Y|X=0])
- By random assignment of X, we have:  $p(1-p)(E[Y_1-Y_0])$
- So dividing by p(1-p) will leave us with  $\rho$

## Effective Samples

- We're going to be investigating how to check the properties of your effective sample in regression.
- The key result is:  $\hat{\rho}_{reg} \xrightarrow{p} \frac{E[w_i \rho_i]}{E[w_i]}$  where  $w_i = \text{var}(D_i | X_i)$

- We estimate these weights with:  $\hat{w}_i = \hat{D}_i^2$  where  $D_i^2$  is the *i*th squared residual.
- Because these estimates are "bad" for each unit, using them to reweight the sample is a bad idea.
- Instead, we just use them to get a sense for what the effective sample is by examining the weight allocated to particular strata.
- We will now explore how to do this.

### Example paper

- We will be looking at Egan and Mullin (2012)
- This paper explores the effect of local weather variations on belief in global warming.
- Very cool paper! With an interesting randomized treatment.
- But what is the effective sample?
- In other words, where is weather (conditional on covariates) most variable?
- That's what we'll explore.

#### Load in data

• Get the data from Pat's replication materials here.

```
require(foreign)
d <- read.dta("gwdataset.dta")</pre>
## Warning: value labels ('q2') for 'jan07_q2' are missing
zips <- read.dta("zipcodetostate.dta")</pre>
zips <- unique(zips[, c("statenum", "statefromzipfile")])</pre>
pops <- read.csv("population_ests_2013.csv")</pre>
pops$state <- tolower(pops$NAME)</pre>
d$getwarmord <- as.double(d$getwarmord)</pre>
# And estimate primary model of interest:
out <- lm(getwarmord ~ ddt_week + educ_hsless + educ_coll + educ_postgrad +
    educ_dk + party_rep + party_leanrep + party_leandem + party_dem + male +
    raceeth_black + raceeth_hisp + raceeth_notwbh + raceeth_dkref + age_1824 +
    age_2534 + age_3544 + age_5564 + age_65plus + age_dk + ideo_vcons + ideo_conservative +
    ideo_liberal + ideo_vlib + ideo_dk + attend_1 + attend_2 + attend_3 + attend_5 +
    attend_6 + attend_9 + as.factor(doi) + as.factor(statenum) + as.factor(wbnid_num),
# And partial regression:
```

### Base Model

• We won't worry about standard errors yet.

. .

```
summary(out)$coefficients[1:10, ]
```

```
##
                Estimate Std. Error t value Pr(>|t|)
                1.945740 0.771479 2.5221 1.169e-02
## (Intercept)
## ddt_week
                ## educ_hsless
                         0.024484 2.3583 1.839e-02
                0.057740
## educ_coll
                0.021457
                         0.027095 0.7919 4.284e-01
## educ_postgrad 0.049387
                         0.030734 1.6069 1.081e-01
## educ dk
                         0.250495 0.4109 6.811e-01
                0.102936
## party rep
               -0.243900
                          0.036627 -6.6591 2.992e-11
## party_leanrep -0.092809
                         0.041812 -2.2197 2.648e-02
## party leandem 0.147409
                          0.039024 3.7774 1.600e-04
## party_dem
                          0.035267 4.9743 6.725e-07
                0.175427
```

#### Estimate D^2

• We can simply square the residuals of a partial regression to get  $D^2$ :

. . .

```
outD <- lm(ddt_week ~ educ_hsless + educ_coll + educ_postgrad + educ_dk + party_rep +
    party_leanrep + party_leandem + party_dem + male + raceeth_black + raceeth_hisp +
    raceeth_notwbh + raceeth_dkref + age_1824 + age_2534 + age_3544 + age_5564 +
    age_65plus + age_dk + ideo_vcons + ideo_conservative + ideo_liberal + ideo_vlib +
    ideo_dk + attend_1 + attend_2 + attend_3 + attend_5 + attend_6 + attend_9 +
    as.factor(doi) + as.factor(statenum) + as.factor(wbnid_num), d)</pre>
D2 <- residuals(outD)^2
```

## **Effective Sample Statistics**

• We can use these estimated weights for examining the sample.

```
compare_samples <- d[, c("wave", "ddt_week", "ddt_twoweeks", "ddt_threeweeks",
   "party_rep", "attend_1", "ideo_conservative", "age_1824", "educ_hsless")]
compare_samples <- apply(compare_samples, 2, function(x) c(mean(x), sd(x), weighted.mean(x,
   D2), sqrt(weighted.mean((x - weighted.mean(x, D2))^2, D2))))
compare_samples <- t(compare_samples)</pre>
colnames(compare_samples) <- c("Nominal Mean", "Nominal SD", "Effective Mean",</pre>
   "Effective SD")
compare_samples
                 Nominal Mean Nominal SD Effective Mean Effective SD
##
## wave
                   3.09694 1.4253 3.20788 1.5609
                                          5.11579
                    3.83549 5.9047
                                                     10.8980
## ddt_week
                                         5.00137
5.10859
                    3.85506
                             5.4572
## ddt_twoweeks
                                                      9.2263
                 3.96720
                             4.7690
## ddt_threeweeks
                                                      8.4348
## party_rep
                    0.29527 0.4562
                                          0.28978
                                                      0.4537
## attend_1
                    0.11433 0.3182
                                          0.12343
                                                      0.3289
0.29325
                                                      0.4553
                    0.07196 0.2584
## age_1824
                                          0.06881
                                                      0.2531
## educ_hsless 0.34151 0.4743 0.31220
                                                      0.4634
```

## Effective Sample Maps

- But one of the most interesting things is to see this visually.
- Where in the US does the effective sample emphasize?
- To get at this, we'll use some tools in R that make this incredibly easy.
- In particular, we'll do this in ggplot2.

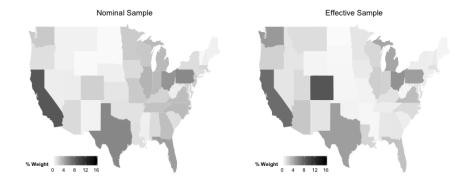
```
# Effective sample by state
wt.by.state <- tapply(D2, d$statenum, sum)
wt.by.state <- wt.by.state/sum(wt.by.state) * 100
wt.by.state <- cbind(D2 = wt.by.state, statenum = names(wt.by.state))
data_for_map <- merge(wt.by.state, zips, by = "statenum")
# Nominal Sample by state
wt.by.state <- tapply(rep(1, 6726), d$statenum, sum)
wt.by.state <- wt.by.state/sum(wt.by.state) * 100
wt.by.state <- cbind(Nom = wt.by.state, statenum = names(wt.by.state))
data_for_map <- merge(data_for_map, wt.by.state, by = "statenum")
# Get correct state names
require(maps, quietly = TRUE)
data(state.fips)
data_for_map <- merge(state.fips, data_for_map, by.x = "abb", by.y = "statefromzipfile")</pre>
```

### More setup

```
plotEff <- ggplot(data_for_map, aes(map_id = state))</pre>
plotEff <- plotEff + geom map(aes(fill = D2), map = state map)</pre>
plotEff <- plotEff + expand_limits(x = state_map$long, y = state_map$lat)</pre>
plotEff <- plotEff + scale_fill_continuous("% Weight", limits = c(0, 16), low = "white",</pre>
    high = "black")
plotEff <- plotEff + labs(title = "Effective Sample")</pre>
plotEff <- plotEff + theme(legend.position = c(0.2, 0.1), legend.direction = "horizontal",</pre>
    axis.line = element_blank(), axis.text = element_blank(), axis.ticks = element_blank(),
    axis.title = element_blank(), panel.background = element_blank(), plot.background = element_blank()
    panel.border = element_blank(), panel.grid = element_blank())
plotNom <- ggplot(data_for_map, aes(map_id = state))</pre>
plotNom <- plotNom + geom_map(aes(fill = Nom), map = state_map)</pre>
plotNom <- plotNom + expand_limits(x = state_map$long, y = state_map$lat)</pre>
plotNom <- plotNom + scale_fill_continuous("% Weight", limits = c(0, 16), low = "white",</pre>
    high = "black")
plotNom <- plotNom + labs(title = "Nominal Sample")</pre>
plotNom <- plotNom + theme(legend.position = c(0.2, 0.1), legend.direction = "horizontal",
    axis.line = element_blank(), axis.text = element_blank(), axis.ticks = element_blank(),
    axis.title = element_blank(), panel.background = element_blank(), plot.background = element_blank()
    panel.border = element_blank(), panel.grid = element_blank())
```

# And the maps

```
require(gridExtra, quietly = TRUE)
grid.arrange(plotNom, plotEff, ncol = 2)
```



### Setup Comparison Plot

```
plotDiff <- ggplot(data_for_map, aes(map_id = state))
plotDiff <- plotDiff + geom_map(aes(fill = Diff), map = state_map)
plotDiff <- plotDiff + expand_limits(x = state_map$long, y = state_map$lat)
plotDiff <- plotDiff + scale_fill_gradient2("% Weight", low = "red", mid = "white",
    high = "black")
plotDiff <- plotDiff + labs(title = "Effective Weight Minus Nominal Weight")
plotDiff <- plotDiff + theme(legend.position = c(0.2, 0.1), legend.direction = "horizontal"
    axis.line = element_blank(), axis.text = element_blank(), axis.ticks = element_blank(),
    axis.title = element_blank(), panel.background = element_blank(), plot.background = element_blank())</pre>
```

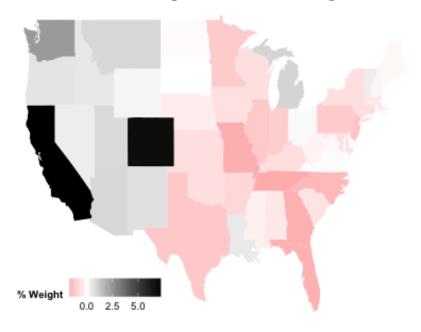
## Difference in Weights

plotDiff

## **Population Comparison**

```
plotEff <- ggplot(data_for_map, aes(map_id = state))
plotEff <- plotEff + geom_map(aes(fill = PopDiffEff), map = state_map)
plotEff <- plotEff + expand_limits(x = state_map$long, y = state_map$lat)
plotEff <- plotEff + scale_fill_gradient2("% Weight", limits = c(-2, 6), low = "red",
    mid = "white", high = "black")
plotEff <- plotEff + labs(title = "Effective Sample")
plotEff <- plotEff + theme(legend.position = c(0.2, 0.1), legend.direction = "horizontal",
    axis.line = element_blank(), axis.text = element_blank(), axis.ticks = element_blank(),
    axis.title = element_blank(), panel.background = element_blank(), plot.background = element_blank(),</pre>
```

# Effective Weight Minus Nominal Weight



```
panel.border = element_blank(), panel.grid = element_blank())

plotNom <- ggplot(data_for_map, aes(map_id = state))
plotNom <- plotNom + geom_map(aes(fill = PopDiffNom), map = state_map)
plotNom <- plotNom + expand_limits(x = state_map$long, y = state_map$lat)
plotNom <- plotNom + scale_fill_gradient2("% Weight", limits = c(-2, 6), low = "red",
    mid = "white", high = "black")
plotNom <- plotNom + labs(title = "Nominal Sample")
plotNom <- plotNom + theme(legend.position = c(0.2, 0.1), legend.direction = "horizontal",
    axis.line = element_blank(), axis.text = element_blank(), axis.ticks = element_blank(),
    axis.title = element_blank(), panel.background = element_blank(), plot.background = element_blank())</pre>
```

### **Population Comparison Plots**

grid.arrange(plotNom, plotEff, ncol = 2)



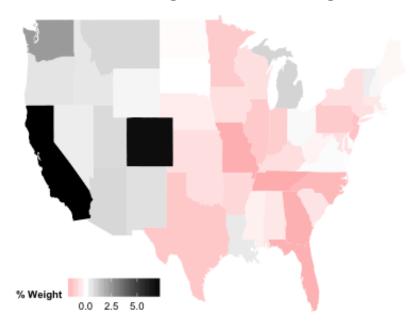
# Setup New Comparison Plot

```
plotDiff <- ggplot(data_for_map, aes(map_id = state))
plotDiff <- plotDiff + geom_map(aes(fill = PopDiff), map = state_map)
plotDiff <- plotDiff + expand_limits(x = state_map$long, y = state_map$lat)
plotDiff <- plotDiff + scale_fill_gradient2("% Weight", low = "red", mid = "white",
    high = "black")
plotDiff <- plotDiff + labs(title = "Effective Weight Minus Nominal Weight")
plotDiff <- plotDiff + theme(legend.position = c(0.2, 0.1), legend.direction = "horizontal"
    axis.line = element_blank(), axis.text = element_blank(), axis.ticks = element_blank(),
    axis.title = element_blank(), panel.background = element_blank(), plot.background = element_blank())</pre>
```

## Plot Difference

plotDiff

## Effective Weight Minus Nominal Weight



## Sensitivity Analysis

- The homework mentions a couple places to find pre existing code.
- I'm going to walk you through how to do a generalized version of the Imbens (2003) method.
- It may be easier to use one of the canned routines for your homework, though.
- We're going to keep working with Pat's data, since we already have it handy.
- Imbens process:
  - Simulate (or imagine simulating) an unobserved confounder like the following:

```
Y_d|X, U \sim \mathcal{N}(\tau d + \boldsymbol{\beta}'X + \delta U, \sigma^2)

D|X, U \sim f(\boldsymbol{\gamma}'X + \alpha U) \text{ (with } f \text{ known)}

- That is, Y_1, Y_0 \perp D|X, U
```

- So we want to generate an additively linear confounder with both D and Y.

## Generate a confounder

- For our analysis, Y is belief in global warming and D is local variation in temperature.
- We want to standardize these variables first.

. . .

```
d$getwarmord <- scale(d$getwarmord)
d$ddt_week <- scale(d$ddt_week)
genConfound <- function(alpha, delta) {
    e <- rnorm(nrow(d), 0, 1)
    U <- alpha * d$ddt_week + delta * d$getwarmord + e
    return(U)
}</pre>
```

 So we can vary partial correlations with D and Y by varying alpha and delta.

```
X <- "educ_hsless+educ_coll+educ_postgrad+educ_dk+party_rep+party_leanrep+party_leandem+party_c(D = coef(lm(paste0("ddt_week~U1+", X), d))["U1"], Y = coef(lm(paste0("getwarmord~U1+", X), d))["U1"])

## D.U1 Y.U1

## 0.005302 0.388986

c(D = coef(lm(paste0("ddt_week~U2+", X), d))["U2"], Y = coef(lm(paste0("getwarmord~U2+", X), d))["U2"])

## D.U2 Y.U2

## 0.03226 0.06704</pre>
```

### Continued

 More importantly, we can see how this changes our estimate of the treatment effect:

```
out <- lm(paste0("getwarmord~ddt_week+", X), d)
coef(out)["ddt_week"]

## ddt_week
## 0.03618

coef(lm(paste0("getwarmord~ddt_week+U1+", X), d))["ddt_week"]

## ddt_week
## 0.01399

coef(lm(paste0("getwarmord~ddt_week+U2+", X), d))["ddt_week"]

## ddt_week
## -0.9855</pre>
```

• Now we want to do this over a larger number of values of alpha and delta

```
alphas <- rnorm(1000, 0, 0.5)
deltas <- rnorm(1000, 0, 0.5)
results <- NULL
for (i in seq_len(length(alphas))) {
    U <- genConfound(alphas[i], deltas[i])
    corD <- cor(U, d$ddt_week)
    corY <- cor(U, d$getwarmord)
    estTE <- coef(lm(paste0("getwarmord~ddt_week+U+", X), d))["ddt_week"]
    names(estTE) <- NULL
    res <- c(estTE = estTE, corD = corD, corY = corY)
    results <- rbind(results, res)
}
results <- cbind(results, TEchange = (results[, "estTE"] - coef(out)["ddt_week"]))</pre>
```

#### Plot Simulation Code

#### Plot of the Results

