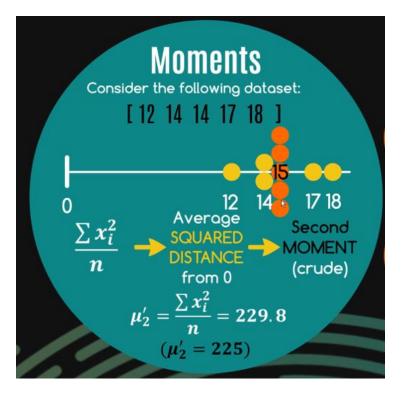
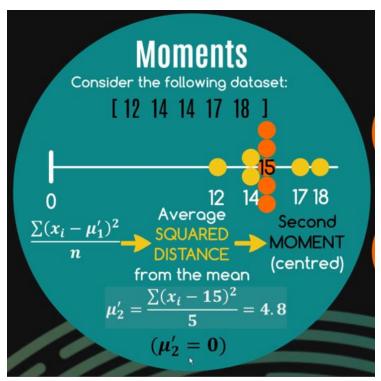


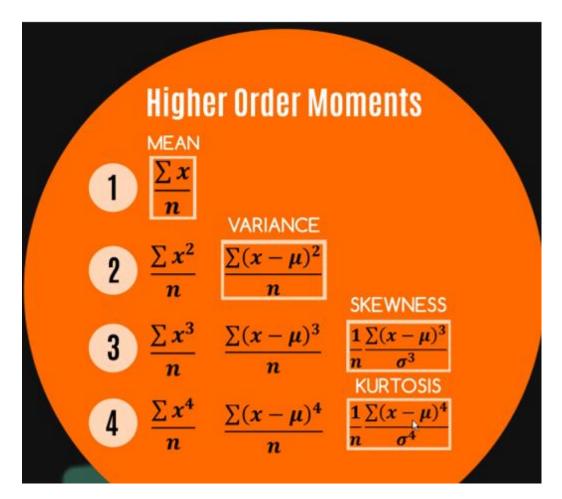
Both the average(1^{st} moment) is 15, then what is the difference between 1^{st} and 2^{nd} moment???



Greater the spread, then greater the average squared distances from 0. But the difference between two moments is less and not specifying that much.



Now it clearly tells the difference between two date-sets. It is not necessary to start from zero, the point can be anywhere, it can be mean also.



The concept of variance requires mean.

The concept of skewness requires variance.

The concept of kurtosis doesn't requires skewness.

Moments

Moments [Discrete case]

Let X be discrete R.V. taking the values $x_1, x_2, ... x_n$ with probability mass function $p_1, p_2, ... p_n$ respectively then the r^{th} moment about the origin is

origin is
$$\mu_{\mathbf{r}'} \text{ (about the origin)} = \sum_{i=1}^{n} x_{i}^{\mathbf{r}} p_{i} \qquad \dots (1)$$
and
$$\mu_{\mathbf{r}'} \text{ (about any point } x = A) = \sum_{i=1}^{n} (x_{i} - A)^{\mathbf{r}} p_{i} \qquad \dots (2)$$
and
$$\mu_{\mathbf{r}} \text{ (about mean)} = \sum_{i=1}^{n} (x_{i} - Mean)^{\mathbf{r}} p_{i} \qquad \dots (3)$$
In particular from (1)

$$\mu_{1'} = \sum_{i=1}^{n} x_{i} p_{i} = \text{Mean } (\bar{x})$$

$$\mu_{2'} = \sum_{i=1}^{n} x_{i}^{2} p_{i} = \text{Mean square value.}$$

$$\mu_{2} = \sum_{i=1}^{n} (x_{i} - \text{mean})^{2} p_{i} = \text{variance}$$

$$= \mu_{2'} - (\mu_{1'})^{2} \quad [\because \bar{x} = \mu_{1'}]$$

$$\mu_{3} = \mu_{3'} - 3\mu_{2'} \mu_{1'} + 2\mu_{1'}^{3}$$

$$\mu_{4} = \mu_{4'} - 4\mu_{3'} \mu_{1'} + 6\mu_{2'} \mu_{1'}^{2} - 3\mu_{1'}^{4}$$

Moments [Continuous case]

If X is a continuous R.V. with probability density function f(x) defined in the interval (a, b) then

$$\mu_{\mathbf{r}'} = \int_{\mathbf{a}}^{\mathbf{b}} x^{\mathbf{r}} f(x) dx$$

$$\mu_{\mathbf{r}'} \text{ (about a point A)} = \int_{\mathbf{a}}^{\mathbf{b}} (x - \mathbf{A})^{\mathbf{r}} f(x) dx$$

$$\mu_{\mathbf{r}} \text{ (about the mean)} = \int_{\mathbf{a}}^{\mathbf{b}} (x - \overline{x})^{\mathbf{r}} f(x) dx$$

Moments

Let X be a discrete random variable and X takes the values

 $x_1, x_2, x_3, \dots, x_n$ with probabilities $p_1, p_2, p_3, \dots, p_n$ Then r^{th} moment is

- $\mu'_r(about origin) = \sum_{i=1}^n x_i^r p_r$
- μ'_r (about any point x=A) = $\sum_{i=1}^n (x_i A)^r p_i$
- $\mu'_r(\text{about mean } \bar{x}) = \sum_{i=1}^n (x_i \bar{x})^r p_i$

Similarly, for continuous case on the interval (a, b), the r^{th} moment is

•
$$\mu'_r(about origin) = \int_a^b x^r f(x) dx$$

•
$$\mu'_r(\text{about any point } x=A) = \int_a^b (x-A)^r f(x) dx$$

•
$$\mu'_r(\text{about mean }\bar{x}) = \int_a^b (x - \bar{x})^r f(x) dx$$

$$\mathsf{mean} = E(X) = \mu_1' \; \mathsf{and} \;$$

Variance =
$$\sigma^2 = E(X^2) - (E(X))^2 = \mu_2' - (\mu_1')^2$$
.

If X is discrete, then
$$M_X(t) = E(e^{tx}) = \sum_{i=1}^n e^{tx_i} p(x_i)$$
,

ger Kurting Monorting where X takes the values $x_1, x_2, x_3, \dots, x_n$ with probabilities

$$p(x_1), p(x_2), p(x_3), \dots p(x_n)$$

If X is continuous, then
$$M_X(t) = E(e^{tx}) = \int_{-\infty}^{\infty} e^{tx} f(x) dx$$

$$\mathsf{mean} = \mathsf{E}(\mathsf{X}) = \mu_1'$$
 and

Variance =
$$\sigma^2 = E(X^2) - (E(X))^2 = \mu_2' - (\mu_1')^2$$
.

Moments

If X is discrete, then $M_X(t) = E(e^{tx}) = \sum_{i=1}^n e^{tx_i} p(x_i)$, where X takes the values $x_1, x_2, x_3, \ldots, x_n$ with probabilities $p(x_1), p(x_2), p(x_3), \ldots p(x_n)$

If X is continuous, then $M_X(t) = E(e^{tx}) = \int_{-\infty}^{\infty} e^{tx} f(x) dx$.

Properties of MGF

• [The r^{th} moment μ'_r is the coefficient of $\frac{t^r}{r!}$ in the expansion of $M_X(t)$ in series of powers of t.

$$M_X(t) = E(e^{tx}) = E\left[1 + \frac{tx}{1!} + \frac{(tx)^2}{2!} + \dots + \frac{(tx)^r}{r!} + \dots\right]$$

$$= 1 + tE(x) + \frac{t^2}{2!}E(x^2) + \dots + \frac{t^r}{r!}E(x^r) + \dots$$

$$= 1 + t\mu_1' + \frac{t^2}{2!}\mu_2' + \dots + \frac{t^r}{r!}\mu_r' + \dots$$

Therefore, $\mu_r^{'}$ is the coefficient of $\frac{t^r}{r!}$ in the expansion of $M_X(t)$ in series of powers of t.

$$\mu_r^{'} = E(X^r) = \left[\frac{d^r}{dt^r}(M_X(t))\right]_{t=0}.$$

In particular, if r=1 then $\mu_1'=\left[\frac{d(M_X(t))}{dt}\right]_{t=0}$ is the mean of X.

- $M_{cX}(t) = M_X(ct)$
- $M_{aX+b}(t) = e^{bt} M_X(at)$
- $M_{X+Y}(t) = M_X(t)M_Y(t)$

Characteristic function

Let X be a discrete random variable and X takes the values x_1, x_2, x_3, \ldots with probabilities p_1, p_2, p_3, \ldots Then the characteristic function is defined as $\phi_X(t) = E(e^{itx}) = \sum_r e^{itx_r} p(x_r)$.

Similarly, for continuous case $\phi_X(t) = E(e^{itx}) = \int_{-\infty}^{\infty} e^{itx} f(x) dx$.

Properties of Characteristic function

• $\mu'_r = E(X^r)$ =the coefficient of $\frac{i^r t^r}{r!}$ in the expansion of $\phi_X(t)$ in series of ascending powers of it.

$$\mu'_r = E(X^r) = \frac{1}{i^r} \left[\frac{d^r}{dt^r} (\phi_X(t)) \right]_{t=0}.$$

In particular, if r=1 then $\mu_1'=\frac{1}{i}\left[\frac{d(\phi_X(t))}{dt}\right]_{t=0}$ is the mean of X.

- $\phi_{aX+b}(t) = e^{ibt}\phi(at)$
- $\phi_{X+Y}(t) = \phi_X(t)\phi_Y(t)$

Moment generating functions and their properties

Moments Generating Function: (M.G.F)

An important device that can be used to calculate the higher moments is the moment generating function.

Moment generating function of a random variable X about the origin is defined as

$$M_X(t) = E[e^{tX}] = \sum_x e^{tx} p(x)$$
, if X is discrete
$$\int_{-\infty}^{\infty} e^{tx} f_X(x) dx$$
, if X is continuous

Find the moment generating function of the random variable X whose probability mass function $P(X=x)=\frac{1}{2^x}, x=1,2,3,\ldots$, Deduce the mean and variance from moment generating function.

$$\begin{aligned}
M_{X}(t) &= E[e^{tX}] \\
&= \sum_{x=1}^{\infty} e^{tx} p(x) \\
&= \sum_{x=1}^{\infty} e^{tx} \frac{1}{2^{x}} = \sum_{x=1}^{\infty} \left(\frac{e^{t}}{2}\right)^{x} \\
&= \left[\frac{d}{dt} \left[e^{t} \left(2 - e^{t}\right)^{-1}\right]\right]_{t=0} \\
&= \left[e^{t} \left(-1\right) \left(2 - e^{t}\right)^{-2} \left(-e^{t}\right) + \left(2 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
&= \left[e^{t} \left(1 - e^{t}\right)^{2} + \dots\right] \\
&= \left[e^{2t} \left(2 - e^{t}\right)^{-2} + \left(2 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
&= \left[e^{2t} \left(2 - e^{t}\right)^{-2} + \left(2 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
&= \left[e^{2t} \left(2 - e^{t}\right)^{-2} + \left(2 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
&= \left[e^{2t} \left(2 - e^{t}\right)^{-2} + \left(2 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
&= \left[e^{2t} \left(1 - e^{t}\right)^{2} + \left(2 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
&= \left[e^{2t} \left(2 - e^{t}\right)^{-2} + \left(2 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
&= \left[e^{2t} \left(1 - e^{t}\right)^{2} + \left(2 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
&= \left[e^{2t} \left(1 - e^{t}\right)^{-2} + \left(2 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
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&= \left[e^{2t} \left(1 - e^{t}\right)^{-2} + \left(2 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
&= \left[e^{2t} \left(1 - e^{t}\right)^{-2} + \left(2 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
&= \left[e^{2t} \left(1 - e^{t}\right)^{-2} + \left(2 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
&= \left[e^{2t} \left(1 - e^{t}\right)^{-2} + \left(1 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
&= \left[e^{2t} \left(1 - e^{t}\right)^{-2} + \left(1 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
&= \left[e^{2t} \left(1 - e^{t}\right)^{-2} + \left(1 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
&= \left[e^{2t} \left(1 - e^{t}\right)^{-2} + \left(1 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
&= \left[e^{2t} \left(1 - e^{t}\right)^{-2} + \left(1 - e^{t}\right)^{-2} + \left(1 - e^{t}\right)^{-1} e^{t}\right]_{t=0} \\
&= \left[e^{2t} \left(1 - e^{t}\right)^{-2} + \left$$

If X represents the outcome, when a fair die is tossed, find the m_{0ment} generating function (MGF) of X and hence find E(X) and $Var_{(X)}$.

Homest gardenating function:

$$H_{2}(t) \Rightarrow E[e^{tx}]$$

$$\Rightarrow \underbrace{E}_{2}(e^{tx}\varphi(x))$$

$$1=0$$

$$\Rightarrow e'(//_{6}) + e'(//_{6}) + e'(//_{6}) + \cdots + e^{t}(//_{6})$$

$$H_{2}(t) \Rightarrow (//_{6}) = e^{t} + e^{2t} + e^{2t} + e^{4t}$$

$$Hean \Rightarrow \underbrace{d}_{dt} [H_{x}(t)]_{t=0}$$

$$\Rightarrow \underbrace{d}_{dt} ['/_{6}(e^{t} + e^{t} + e^{2t} + e^{2t} + e^{2t})]_{t=0}$$

$$\frac{1}{6} \left[e^{\frac{t}{2}} + e^{\frac{2t}{2}} \right] + e^{\frac{3t}{3}} + e^{\frac{6t}{60}} \right]_{t=0}^{t=0}$$

$$\frac{1}{6} \left[e^{\frac{t}{2}} + e^{\frac{2t}{2}} \right] + e^{\frac{6t}{60}} + e^{\frac{3t}{20}}$$

$$\frac{1}{6} \left[e^{\frac{t}{2}} + e^{\frac{2t}{20}} \right] + e^{\frac{6t}{60}}$$

$$\frac{1}{6} \left[e^{\frac{t}{2}} + e^{\frac{3t}{20}} \right] + e^{\frac{3t}{60}}$$

$$\frac{1}{6} \left[e^{\frac{t}{2}} + e^{\frac{3t}{20}} \right] + e^{\frac{3t}{20}}$$

$$\frac{1}{6} \left[e^{\frac{3t}{2}} + e^{\frac{3t}{20}} \right] + e^{\frac{3t}{20}}$$

$$\frac{1}{6} \left[$$

```
Balanced win
```

Ly tossed 3 times

(HHH); (HHH); (HHH); (THH)
$$\frac{1}{2}$$
 8 $\frac{1}{2}$ 9 \frac

Bolaned win

Ly tossed 4 times (2" > 16 page ible sutrames)

- *) Exactly I head > 4
- +) Exactly 2 heads > 6
- *) Exactly 3 loads > 4
- *) Exactly 4 heads =)

$$\left[4c_{1} \Rightarrow \frac{4!}{1!(3!)} \Rightarrow \frac{4*3!}{3!} \Rightarrow 4 \right]$$

$$\begin{bmatrix} 4C_2 \Rightarrow 41 \\ 21 + 21 \end{bmatrix} \Rightarrow 4 + 3 + 21 \Rightarrow 6$$

Find the probability distribution of the total number of heads obtained in four tosses of a balanced coin. Hence obtain the MGF of X, mean of X and variance of X.

[AU A/M 2008]

Solution:

X :	Numbe	r of heads	obtained in	4 tosses of	a coin
<i>x</i> :	0	1	2	3	4
p (x):	1/16	$\frac{4}{16}$	$\frac{6}{16}$	<u>4</u> 16	1/16

7: 0 | 2 | 3 | 4

P(2):
$$\frac{1}{16}$$
 | $\frac{4}{16}$ | $\frac{6}{16}$ | $\frac{4}{16}$ | $\frac{1}{16}$ | $\frac{1}{$

$$| \frac{1}{3} \frac{1}{16} | \frac{1}{16} |$$

For the triangular distribution

friangular distribution
$$f(x) = \begin{cases} x, & 0 < x \le 1 \\ 2 - x, & 1 \le x < 2 \\ 0, & \text{otherwise} \end{cases}$$
[A.U. M/J 2006, N/D 2013]

find the mean, variance and the moment generating function (MGF) [A.U CBT M/J 2010, CBT N/D 2011] also find cdf of F(x).

[A.U.N/D 2013] Scheraling function of the Random variable X

Solution: Given:
$$f(x) = \begin{cases} x, & 0 < x \le 1 \\ 2 - x, & 1 \le x < 2 \\ 0, & \text{otherwise} \end{cases}$$

Mean =
$$E(X) = \int_{-\infty}^{\infty} x f(x) dx$$
 ... (1)
= $\int_{0}^{1} (x) (x) dx + \int_{1}^{2} (x) (2-x) dx$

$$= \int_{0}^{1} x^{2} dx + \int_{1}^{2} (2x - x^{2}) dx = \left[\frac{x^{3}}{3} \right]_{0}^{1} + \left[\frac{2x^{2}}{2} - \frac{x^{3}}{3} \right]_{1}^{2}$$

$$= \left[\frac{1}{3} - 0\right] + \left[x^2 - \frac{x^3}{3}\right]_1^2 = \frac{1}{3} + \left[\left(4 - \frac{8}{3}\right) - \left(1 - \frac{1}{3}\right)\right]$$
$$= \frac{1}{3} + \frac{4}{3} - \frac{2}{3} = \frac{4 + 1 - 2}{3} = 1$$

Variance,
$$V(X) = E(X^2) - [E(X)]^2$$
 ... (2)

Variance,
$$V(X) = E(X)$$
 [10]
$$E[X^2] = \int_{-\infty}^{\infty} x^2 f(x) dx = \int_{0}^{1} x^2 x dx + \int_{1}^{2} x^2 (2-x) dx$$

$$= \int_{0}^{1} x^3 dx + \int_{1}^{2} (2x^2 - x^3) dx = \left[\frac{x^4}{4}\right]_{0}^{1} + \left[2\frac{x^3}{3} - \frac{x^4}{4}\right]_{1}^{2}$$

$$= \left(\frac{1}{4} - 0\right) + \left(\frac{16}{3} - \frac{16}{4}\right) - \left(\frac{2}{3} - \frac{1}{4}\right) = \frac{1}{4} + \frac{16}{3} - \frac{16}{4} - \frac{2}{3} + \frac{1}{4}$$

$$= -\frac{14}{4} + \frac{14}{3} = \frac{-42 + 56}{12} = \frac{14}{12} = \frac{7}{6}$$

$$\therefore (2) \Rightarrow \text{Var } (X) = E[X^2] - [E(X)]^2$$

(NGF) also from the first
$$\frac{7}{6} = (1)^2 = \frac{7}{6} = 1 = \frac{1}{6}$$
 and the first size from (NGF)

Someth gonerating
$$\int_{0}^{\infty} H_{1}(t) \Rightarrow E(e^{tx}) \Rightarrow \int_{0}^{\infty} (e^{tx}f(x)) dx$$

$$\Rightarrow \int_{0}^{\infty} e^{tx}(x) dx + \int_{0}^{\infty} e^{tx}(2-x) dx$$

$$\Rightarrow \int_{0}^{\infty} (e^{tx}) dx + 2 \int_{0}^{\infty} e^{tx} dx - \int_{0}^{\infty} (e^{tx}) dx$$

$$\Rightarrow \left[x \frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{0}^{t} + 2 \left[\frac{e^{tx}}{t} \right]_{1}^{2}$$

$$- \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t} + 2 \left[\frac{e^{tx}}{t} \right]_{1}^{2}$$

$$- \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t} + 2 \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t}$$

$$- \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t} + 2 \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t}$$

$$\Rightarrow \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t} + \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t}$$

$$\Rightarrow \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t} + \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t}$$

$$\Rightarrow \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t} + \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t}$$

$$\Rightarrow \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t} + \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t}$$

$$\Rightarrow \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t} + \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t}$$

$$\Rightarrow \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t} + \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t}$$

$$\Rightarrow \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t} + \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t}$$

$$\Rightarrow \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t} + \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t}$$

$$\Rightarrow \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t}$$

$$\Rightarrow \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t}$$

$$\Rightarrow \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t}$$

$$\Rightarrow \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} - \frac{e^{tx}}{t^{2}} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t}$$

$$\Rightarrow \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} - \frac{e^{tx}}{t^{2}} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t}$$

$$\Rightarrow \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} - \frac{e^{tx}}{t^{2}} - \frac{e^{tx}}{t^{2}} - \frac{e^{tx}}{t^{2}} \right]_{1}^{t}$$

$$\Rightarrow \left[\frac{e^{tx}}{t} - \frac{e^{tx}}{t} - \frac{e^{tx}}{t^{2}} - \frac$$

To find the cdf of F(x)

$$F(x) = P[X \le x] = \int_{0}^{x} f(x) dx$$

- (i) If $x \le 0$, then F(x) = 0
- (ii) If $0 < x \le 1$, then

$$F(x) = \int_{0}^{x} x dx = \left[\frac{x^{2}}{2}\right]_{0}^{x} = \frac{x^{2}}{2}$$

(iii) If $1 \le x < 2$, then

$$F[x] = \int_{0}^{1} x \, dx + \int_{1}^{x} (2 - x) \, dx$$

$$= \left[\frac{x^{2}}{2} \right]_{0}^{1} + \left[2x - \frac{x^{2}}{2} \right]_{1}^{x} = \frac{1}{2} + \left(2x - \frac{x^{2}}{2} \right) - \left(2 - \frac{1}{2} \right)$$

$$= \frac{1}{2} + 2x - \frac{x^{2}}{2} - 2 + \frac{1}{2} = 2x - \frac{x^{2}}{2} - 1$$

(iv) If x > 2, then

$$F(x) = \int_{-\infty}^{x} f(x) dx$$

$$= \int_{0}^{1} x dx + \int_{1}^{2} (2 - x) dx + \int_{2}^{x} 0 dx$$

$$= \left[\frac{x^{2}}{2} \right]_{0}^{1} + \left[2x - \frac{x^{2}}{2} \right]_{1}^{2}$$

$$= \frac{1}{2} + (4 - 2) - \left(2 - \frac{1}{2} \right)$$

$$= \frac{1}{2} + 2 - 2 + \frac{1}{2} = 1$$

Let the random variable X have the p.d.f

$$f(x) = \begin{cases} \frac{1}{2} e^{-x/2}, & x > 0 \\ 0, & \text{otherwise} \end{cases}$$

Find the moment generating function, mean and variance of X.

[A.U. A/M. 2005, N/D 2012]

Solution: The m.g.f is given by

$$M_{x}(t) = E[e^{tx}] = \int_{-\infty}^{\infty} e^{tx} f(x) dx = \int_{0}^{\infty} e^{tx} \frac{1}{2} e^{-x/2} dx$$

$$= \frac{1}{2} \int_{0}^{\infty} e^{(t - \frac{1}{2})x} dx = \frac{1}{2} \int_{0}^{\infty} e^{-(\frac{1}{2} - t)x} dx$$

$$= \frac{1}{2} \left[\frac{e^{-(\frac{1}{2} - t)x}}{-(\frac{1}{2} - t)} \right]_{0}^{\infty} = -\frac{1}{2} \left[\frac{e^{-(\frac{1}{2} - t)x}}{\frac{1}{2} - t} \right]_{0}^{\infty}$$

$$= -\frac{1}{2} \left[0 - \frac{1}{\frac{1}{2} - t} \right] = -\frac{1}{2} \left[-\frac{1}{\frac{1 - 2t}} \right]$$

$$= \frac{1}{2} \left[\frac{2}{1 - 2t} \right] = \frac{1}{1 - 2t}$$

$$E(X) = \text{Mean} = M_{X}'(0) = \frac{d}{dt} \left[\frac{1}{1 - 2t} \right]_{t=0}$$

$$= \left[\frac{-1}{(1 - 2t)^{2}} (-2) \right]_{t=0} = 2$$

$$E(X^{2}) = M_{X}''(0) = \frac{d}{dt} \left[M_{X}'(t) \right]_{t=0}$$

$$= \frac{d}{dt} \left[\frac{2}{(1 - 2t)^{2}} \right]_{t=0} = \left[\frac{-4}{(1 - 2t)^{3}} (-2) \right]_{t=0}$$

$$= \left[\frac{8}{(1 - 2t)^{3}} \right]_{t=0} = 8$$

$$\text{Varience} = E(X^{2}) - (E(X))^{2}$$

$$= 8 - (2)^{2} = 8 - 4 = 4$$

The density function of a random variable x is given by f(x) = Kx(2-x), $0 \le x \le 2$. Find K, mean, variance and r^{th} moment.

[A.U. N/D 2006] [A.U. M/J 2007] [A.U Trichy A/M 2010]

Given: $f(x) = Kx(2-x), 0 \le x \le 2$ is a p.d.f.

We know that, if f(x) is a p.d.f then,

$$\int_{-\infty}^{\infty} f(x) dx = 1 \quad \Rightarrow \quad \int_{0}^{2} Kx (2-x) dx = 1$$

$$K \int_{0}^{2} (2x - x^{2}) dx = 1 \Rightarrow K \left[\frac{2x^{2}}{2} - \frac{x^{3}}{3} \right]_{0}^{2} = 1$$

$$K\left[\left(4-\frac{8}{3}\right)-(0-0)\right] = 1 \quad \Rightarrow \quad K\left[\frac{4}{3}\right] = 1 \Rightarrow K = \frac{3}{4}$$

Mean = E(X) =
$$\int_{-\infty}^{\infty} xf(x) dx = \int_{0}^{2} x Kx (2-x) dx$$

= $\int_{0}^{2} \frac{3}{4} (2x^{2} - x^{3}) dx$ [: $K = \frac{3}{4}$]
= $\frac{3}{4} \left[\frac{2x^{3}}{3} - \frac{x^{4}}{4} \right]_{0}^{2} = \frac{3}{4} \left[\left(\frac{16}{3} - \frac{16}{4} \right) - (0 - 0) \right]$
= $\frac{3}{4} (16) \left[\frac{1}{3} - \frac{1}{4} \right] = 12 \left[\frac{1}{12} \right] = 1$
E[X²] = $\int_{-\infty}^{\infty} x^{2} f(x) dx = \int_{0}^{2} x^{2} Kx (2-x) dx$
= $\frac{3}{4} \int_{0}^{2} (2x^{3} - x^{4}) dx$ [: $K = \frac{3}{4}$]
= $\frac{3}{4} \left[\frac{2x^{4}}{4} - \frac{x^{5}}{5} \right]_{0}^{2} = \frac{3}{4} \left[\left(8 - \frac{32}{5} \right) - (0 - 0) \right]$
= $\frac{3}{4} \left[\frac{40 - 32}{5} \right] = \frac{3}{4} \left[\frac{8}{5} \right] = \frac{6}{5}$
Var (X) = E(X²) - [E(X)]² = $\frac{6}{5}$ - 1 = $\frac{1}{5}$
 μ_{Γ} '= E[X^T] = $\int x^{T} f(x) dx = \int_{0}^{2} x^{T} kx (2 - x) dx$
= $\frac{3}{4} \int_{0}^{2} (2x^{T+1} - x^{T+2}) dx$
= $\frac{3}{4} \left[\left(2\left(\frac{2^{T+2}}{r+2} - \frac{x^{T+3}}{r+3} \right) - (0 - 0) \right]$

$$= \frac{3}{4} \left[\frac{2^{r+3}}{r+2} - \frac{2^{r+3}}{r+3} \right]$$

$$= \frac{(3)(2^{r+3})}{4} \left[\frac{1}{r+2} - \frac{1}{r+3} \right]$$

$$= \frac{(3)(2^{r+3})}{4} \left[\frac{r+3-r-2}{(r+2)(r+3)} \right]$$

$$= \frac{(3)(2^{r+1})}{(r+2)(r+3)}$$

A continuous R.V. X has the p.d.f f(x) given by $f(x) = c e^{-|x|}$, $-\infty < x < \infty$. Find the value of c and moment generating function of X. [A.U. M/J 2007]

Solution: Given :
$$f(x) = c e^{-|x|}$$

Given $f(x)$ is a p.d.f.

$$\therefore \int_{-\infty}^{\infty} f(x) dx = 1 \Rightarrow \int_{-\infty}^{\infty} c e^{-|x|} dx = 1$$

$$\Rightarrow 2 \int_{0}^{\infty} c e^{-x} dx = 1 \Rightarrow 2c \left[\frac{e^{-x}}{-1} \right]_{0}^{\infty} = 1$$

$$\Rightarrow -2c \left[e^{-x} \right]_{0}^{\infty} = 1 \Rightarrow -2c \left[0 - 1 \right] = 1$$

$$\Rightarrow 2c = 1 \Rightarrow c = \frac{1}{2}$$

$$\therefore f(x) = \frac{1}{2} e^{-|x|}$$

$$M_{X}(t) = E[e^{tx}] = \int_{-\infty}^{\infty} e^{tx} f(x) dx = \int_{-\infty}^{\infty} e^{tx} \frac{1}{2} e^{-|x|} dx$$
$$= \frac{1}{2} \int_{-\infty}^{\infty} e^{tx} e^{-|x|} dx = \frac{1}{2} 2 \int_{0}^{\infty} e^{tx} e^{-x} dx$$