# ST 516: Foundations of Data Analytics

t-based procedures for the mean

Module 5 Lecture 1 1/ 12

### Recall: estimating a population mean

We have a random sample,  $X_1,...,X_n$ , of size n from the population. We calculate the sample mean,  $\overline{X}$  and wish to perform inferences about the population mean,  $\mu$ .

Last module we constructed confidence intervals for  $\mu$  of the form,

$$\overline{X} \pm z_{\alpha/2} \times SE_{\overline{X}}$$

where  $SE_{\overline{X}} = \frac{s}{\sqrt{n}}$ . We performed hypothesis tests for  $\mu = \mu_0$  by comparing

$$\frac{\overline{X} - \mu_0}{SE_{\overline{Y}}}$$

to a standard Normal distribution.

Module 5 Lecture 1 2/ 12

#### Motivation

To construct our tests and confidence intervals, we relied on the result that the sampling distribution of  $\overline{X}$  was approximately Normal with mean  $\mu$  and standard deviation  $\frac{s}{\sqrt{n}}$ .

The standard deviation should really be  $\frac{\sigma}{\sqrt{n}}$  but we don't know  $\sigma$ . Instead we used s as an estimate for  $\sigma$ .

Using an estimate for  $\sigma$  introduces extra uncertainty. This lecture, we are going to investigate how ignoring this extra uncertainty can be problematic in small samples and introduce an approach to remedy it.

 $\sigma$ : the population standard deviation

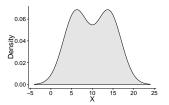
s: the sample standard deviation

Module 5 Lecture 1 3/ 12

## An example

Let's consider an example where our sample size is relatively small, n = 15.

Consider the following population distribution (with mean,  $\mu = 10$ ):



Let's look at the sampling distribution for the statistic  $t = \frac{\overline{X} - \mu}{s/\sqrt{n}}$ .

Our intervals and tests from last week, relied on the sampling distribution for this quantity being a Normal with zero mean and standard deviation 1.

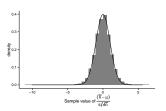
Module 5 Lecture 1 4/ 12

### An example

We repeat many times by simulation:

- take a sample of size 15 from the population
- use the sample to calculate t.

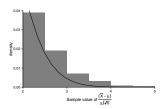
This histogram of the resulting statistics is shown below. I've overlaid a Normal curve on our result.



Doesn't look too bad, but let's take a closer look at the extremes

Module 5 Lecture 1 5/ 12

### An example



The Normal curve is under estimating how likely it is to see t values that are large (and similarly extreme and negative).

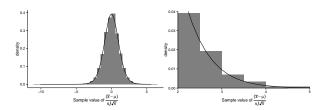
The additional uncertainty due to the estimation of  $\sigma$  results in a sampling distribution with more probability of extreme values than a Normal curve would suggest.

If you consider 95% confidence intervals created the way we did last week ( $\overline{X} \pm 1.96 \times SE_{\overline{X}}$ ), about 92.7% cover the true parameter, not the 95% we would expect.

Module 5 Lecture 1 6/ 12

#### Enter the t

Turns out there is a better distribution to describe the sampling distribution of  $\frac{\overline{X} - \mu}{s/\sqrt{n}}$ , it's called Student's t-distribution, or the t-distribution for short.



When we base our confidence intervals on our statistic having this distribution they cover 95.2% of the time

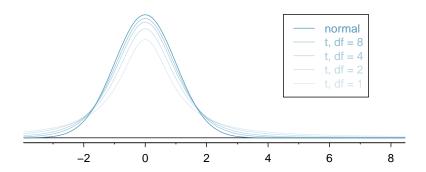
Module 5 Lecture 1 7/ 12

#### A few facts about the t-distribution

The t-distribution is centered around zero.

It has a parameter called the degrees of freedom, which controls the shape of the distribution.

As degrees of freedom increases, the shape of the t-distribution gets closer and closer to the shape of the Normal distribution.



### In practice

To use the t-based methods for inference on a population mean, the calculations are mostly the same, but the reference distribution changes.

For hypothesis tests we calculate the statistic the same way,

$$t = \frac{\overline{X} - \mu}{s / \sqrt{n}}$$

but to find a p-value (or critical value) we now compare our statistic to the **t-distribution with** n-1 **degrees of freedom**.

For a t-based confidence interval, the form is the same

$$\overline{X}$$
 ± critical value  $\times \frac{s}{\sqrt{n}}$ 

but the critical value is now the corresponding quantile from a **t-distribution with** n-1 **degrees of freedom**.

Module 5 Lecture 1 9/ 12

### Returning to our example

To construct a 95% t based confidence intervals, we need the 0.975 quantile from a t-distribution with 14 d.f.

## [1] 2.144787

If, in our simulation, we calculate our confidence interval with

$$\overline{X} \pm 2.14 \times \frac{s}{\sqrt{n}}$$

the coverage of our intervals is 95.2%. (Much closer to our claim of 95%)

Module 5 Lecture 1 10/ 12

### Approximate versus exact

For small samples using the t-distribution as a model for the sampling distribution of the t-statistic is better than using the Normal distribution because it takes into account the added uncertainty due to estimating  $\sigma$ .

However, the t-distribution still might not be quite the right distribution.

If the sampling distribution of the t-statistic is exactly t-distributed, we say our procedures are **exact**.

This means our 95% confidence intervals will have exactly 95% coverage, and our level 0.05 t-test will reject the null exactly 5% of the time when the null is true.

When procedures aren't exact, but there is some reason to believe they are close to exact, we might label them as **approximate**.

Module 5 Lecture 1 11/ 12

#### CLT to the rescue

If the population distribution is Normal, then the t-statistic is exactly t-distributed, regardless of the sample size.

However, it's pretty rare (impossible?) to know your population distribution is exactly Normal.

We do know for large samples  $\overline{X}$  is approximately Normal, regardless of the population distribution (CLT) and we know that the t-distribution is approximately Normal for reasonably large sample sizes.

So, our t-based procedures should be approximately correct for large samples.

It turns out that *large* here is surprisingly small for populations that aren't too non-Normal. So the t-based methods work very well for small samples even when the population distribution isn't Normal.

We'll demonstrate this in the next module.

Module 5 Lecture 1 12/ 12