

Lecture 7  
(26 August 2027)

Theorem. Consider a binomial RV  $X$  with parameters  $n$  and  $p$ . As  $n \rightarrow \infty$ , while keeping  $np = \lambda = \text{constant}$ , we have

$$\lim_{n \rightarrow \infty} p_X(k) = \frac{e^{-\lambda} \lambda^k}{k!}, \quad k = 0, 1, 2, \dots$$

Proof.  $p_X(k) = \binom{n}{k} p^k (1-p)^{n-k}$

$$= \frac{n!}{(n-k)! k!} \left(\frac{\lambda}{n}\right)^k \left(1 - \frac{\lambda}{n}\right)^{n-k}$$

$$= \frac{\lambda^k}{k!} \frac{n!}{(n-k)!} \frac{1}{n^k} \left(1 - \frac{\lambda}{n}\right)^{n-k}$$

$$= \frac{\lambda^k}{k!} \underbrace{\frac{n(n-1)\dots(n-k+1)}{n \cdot n \cdot \dots \cdot n}}_{\rightarrow 1} \underbrace{\left(1 - \frac{\lambda}{n}\right)^{-k}}_{\rightarrow 1} \underbrace{\left(1 - \frac{\lambda}{n}\right)^n}_{\rightarrow e^{-\lambda}}$$

$$\therefore \lim_{n \rightarrow \infty} \binom{n}{k} p^k (1-p)^{n-k} = \frac{e^{-\lambda} \lambda^k}{k!}$$

## Multiple Random Variables

Two random variables  $x$  and  $y$  on the probability space  $(\Omega, \mathcal{F}, P)$  are called jointly discrete if  $(x, y)$  takes values in some countable subset of  $\mathbb{R}^2$ .

$$x: \Omega \rightarrow \mathbb{R}$$

$$y: \Omega \rightarrow \mathbb{R}$$

$$\text{Let } \text{Range}(x) = \underline{x}, \text{Range}(y) = \underline{y}.$$

The associated joint pmf is given by

$$p_{x,y}(\underline{x}, \underline{y}) = P(\{\omega: x(\omega) = \underline{x}, y(\omega) = \underline{y}\})$$

Example, Roll a pair of dice.

$$\Omega = \{(\omega_1, \omega_2) : \omega_1, \omega_2 \in [1:6]\}.$$

$$x(\omega) = \omega_1 + \omega_2, \quad y(\omega) = \frac{\omega_1}{\omega_2}.$$

The pmfs of  $x$  and  $y$  can be obtained from joint pmf using the formulas;

$$P_x(x) = \sum_{y \in \underline{y}} p_{x,y}(\underline{x}, \underline{y}) \quad P_y(y) = \sum_{x \in \underline{x}} p_{x,y}(\underline{x}, \underline{y}).$$

Reason:  $P_x(x) = P(X=x)$

$$= P(X=x \cap \bigcup_{y \in Y} \{Y=y\})$$

$$= P(\bigcup_{y \in Y} \{X=x\} \cap \{Y=y\})$$

$$= P(\bigcup_{y \in Y} \{X=x, Y=y\})$$

$$= \sum_{y \in Y} P(X=x, Y=y)$$

$$= \sum_{y \in Y} p_{x,y}(x,y).$$

## Functions of Multiple Random Variables

Consider two jointly discrete random variables  $X$  and  $Y$ .

Let  $Z = g(X, Y)$ , i.e.,

$$Z(\omega) = g(X(\omega), Y(\omega)).$$

Analogous to the way we argued  $g(X)$  is a random variable,  $Z = g(X, Y)$  is also a random variable.

$$P_Z(z) = \sum_{(x,y): g(x,y)=z} p_{x,y}(x,y).$$

### Exercise

Prove that

$$E[g(x,y)] = \sum_{x,y} g(x,y) p_{x,y}(x,y).$$

### Independence

Two discrete random variables  $X$  and  $Y$  are said to be independent if

$$p_{x,y}(x,y) = p_X(x) p_Y(y) \quad \forall x,y, \text{ i.e., the events } \{X=x\} \text{ and } \{Y=y\} \text{ are independent for all } x,y.$$

Example. Two random variables  $X$  and  $Y$  take values in  $\{0,1\}$  and  $p_{x,y}$  is their joint pmf. Suppose  $p_{x,y}(1,1) = p_X(1) p_Y(1)$ .

Are  $X$  and  $Y$  independent?

$$\begin{aligned} p_{x,y}(1,0) &= p_X(1) - p_{x,y}(1,1) \\ &= p_X(1) - p_X(1) p_Y(1) \end{aligned}$$

$$= p_x(1) (1 - p_y(1))$$

$$= p_x(1) p_y(0),$$

Similarly  $p_{x,y}(x,y) = p_x(x) p_y(y) \forall x,y,$

yes  $x$  and  $y$  are independent.

Exercise. Prove that the indicator random variables  $1_A$  and  $1_B$  are independent if and only if the events  $A$  and  $B$  are independent.

Theorem. If  $x$  and  $y$  are independent discrete random variables then  $E[xy] = E[x] E[y]$ .

Proof.  $E[xy] = \sum_{x,y} xy p_{x,y}(x,y)$

$$= \sum_{x,y} xy p_x(x) p_y(y)$$

$$= \sum_x x p_x(x) \sum_y y p_y(y)$$

$$= E[x] E[y],$$

Exercise.  $X$  and  $Y$  are independent implies that  $g(X)$  and  $h(Y)$  are independent.

### Independence of several Rvs

$n$  random variables  $X_1, X_2, \dots, X_n$  are independent if

$$p_{X_1, X_2, \dots, X_n}(x_1, x_2, \dots, x_n) = \prod_{i=1}^n p_{X_i}(x_i) \\ \forall x_1, x_2, \dots, x_n.$$

Exercise.  $X_1, X_2, \dots, X_n$  are independent implies

$$p_{X_I}(x_I) = \prod_{i \in I} p_{X_i}(x_i) \quad \forall I \subseteq [1:n].$$

$$[x_I = (x_i : i \in I)]$$

## Computation of Mean and Variance for Examples

$X \sim \text{Bernoulli}(p)$ :

$$E[X] = p$$

$$E[X^2] = p \Rightarrow \text{Var}(X) = p - p^2 = p(1-p)$$

$X \sim \text{Binomial}(n, p)$ :  $p_X(k) = \binom{n}{k} p^k (1-p)^{n-k}$

$$E[X] = \sum_{k=0}^n k p_X(k)$$

$$= \sum_{k=0}^n k \binom{n}{k} p^k (1-p)^{n-k}$$

$$= \sum_{k=0}^n k \frac{n!}{(n-k)! k!} p^k (1-p)^{n-k}$$

$$= np \sum_{k=1}^n \frac{(n-1)!}{(n-k)! (k-1)!} p^{k-1} (1-p)^{n-k}$$

$$= np \sum_{k'=0}^{n-1} \frac{(n-1)!}{(n-k')! k'!} p^{k'} (1-p)^{n-1-k'}$$

$$= np \sum_{k=0}^{n-1} \binom{n-1}{k} p^k (1-p)^{n-1-k} = np.$$

$$E[X^2] = \sum_{k=0}^n k^2 p_X(k)$$

$$= \sum_{k=0}^n k^2 \binom{n}{k} p^k (1-p)^{n-k}$$

$$= np \sum_{k=1}^n k \binom{n-1}{k-1} p^{k-1} (1-p)^{n-k}$$

$$= np \sum_{k'=0}^{n-1} (k'+1) \binom{n-1}{k'} p^{k'} (1-p)^{n-1-k'}$$

$$= np \left[ 1 + \sum_{k=0}^{n-1} k \binom{n-1}{k} p^k (1-p)^{n-1-k} \right]$$

$$= np [1 + (n-1)p],$$

$$\text{Var}(X) = E[X^2] - E[X]^2$$

$$= np + n^2 p^2 - np^2 - n^2 p^2$$

$$= np(1-p).$$

$$X \sim \text{Geometric}(p): p_X(k) = (1-p)^{k-1} p, \quad k = 1, 2, \dots$$

$$E[X] = \sum_{k=1}^{\infty} k p_X(k)$$

$$= \sum_{k=1}^{\infty} k (1-p)^{k-1} p$$

$$\begin{cases} 1 + x + x^2 + \dots = \frac{1}{1-x} \\ 1 + 2x + 3x^2 + \dots = \frac{1}{(1-x)^2} \end{cases}$$



$$= p \sum_{k=1}^{\infty} k (1-p)^{k-1}$$

$$= p \cdot \frac{1}{(1-(1-p))^2} = \frac{1}{p}$$

$$E[x^2] = \sum_{k=1}^{\infty} k^2 p_x(k)$$

$$= \sum_{k=1}^{\infty} k^2 (1-p)^{k-1} p$$

$$\sum_{k=1}^{\infty} k^2 (1-p)^{k-1} = \sum_{k=1}^{\infty} (k^2 - k) (1-p)^{k-1} + \sum_{k=1}^{\infty} k (1-p)^{k-1}$$

$$= \frac{2(1-p)}{p^3} + \frac{1}{p^2} = \frac{2-p}{p^3}$$

$$\Rightarrow E[x^2] = p \cdot \frac{2-p}{p^3} = \frac{2-p}{p^2}$$

$$\text{Var}(x) = E[x^2] - E[x]^2$$

$$= \frac{2-p}{p^2} - \frac{1}{p^2} = \frac{1-p}{p^2}$$

$$\sum_{k=0}^{\infty} x^k = \frac{1}{1-x}$$

$$\sum_{k=1}^{\infty} k x^{k-1} = \frac{1}{(1-x)^2}$$

$$\sum_{k=2}^{\infty} k(k-1) x^{k-2} = \frac{2}{(1-x)^3}$$

$$\sum_{k=2}^{\infty} (k^2 - k) x^{k-2} = \frac{2}{(1-x)^3}$$

$$X \sim \text{Poisson}(\lambda): p_x(k) = \frac{e^{-\lambda} \lambda^k}{k!}, \quad k = 0, 1, 2, \dots$$

$$\begin{aligned}
E[x] &= \sum_{k=0}^{\infty} k p_x(k) \\
&= \sum_{k=0}^{\infty} k e^{-\lambda} \frac{\lambda^k}{k!} \\
&= \sum_{k=1}^{\infty} e^{-\lambda} \frac{\lambda^k}{(k-1)!} \\
&= \lambda \sum_{k=1}^{\infty} e^{-\lambda} \frac{\lambda^{k-1}}{(k-1)!} = \lambda.
\end{aligned}$$

$$\begin{aligned}
E[x^2] &= \sum_{k=0}^{\infty} k^2 p_x(k) \\
&= \sum_{k=0}^{\infty} k^2 \cdot e^{-\lambda} \frac{\lambda^k}{k!} \\
&= \lambda \sum_{k=1}^{\infty} k e^{-\lambda} \frac{\lambda^{k-1}}{(k-1)!} \\
&= \lambda \sum_{k'=0}^{\infty} (k'+1) e^{-\lambda} \frac{\lambda^{k'}}{k'!} \\
&= \lambda [\lambda + 1]
\end{aligned}$$

$$\begin{aligned}
\Rightarrow \text{Var}(x) &= E[x^2] - E[x]^2 \\
&= \cancel{\lambda} + 1 - \cancel{\lambda^2} \\
&= 1
\end{aligned}$$