

Markov Chain Monte Carlo :

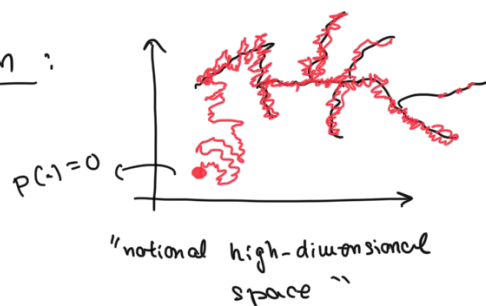
Goal: Sample from $p(x)$, $x \in \mathbb{R}^d$ and compute :

$$\mathbb{E}_{x \sim p(x)} [f(x)] \approx \frac{1}{n} \sum_{i=1}^n f(x_i), \quad x_i \stackrel{i.i.d}{\sim} p(x)$$

Challenge: $p(x)$, $f(x)$ may be complicated.

$$p(x_0) = 0$$

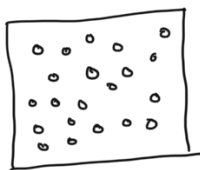
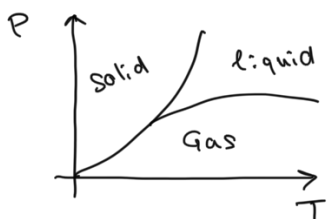
Intuition :



Start at some x_0 , find some region of high-probability and then explore the space by moving randomly yet staying close to regions of high-probability.

Metropolis 1953 :

"Hard disks in a box"



$$x \in \mathbb{R}^{2n}$$

$$\mathbb{E}_{x \sim p(x)} [f(x)]$$

$$p(x) = \frac{1}{Z_p} e^{-\frac{E(x)}{kT}}$$

Goal: Start from x_0 and simulate a sequence of state x_1, \dots, x_n that have high probability under $p(x)$.

Bayesian inference using the Metropolis algorithm :

Given some model with parameters $\theta \in \mathbb{R}^d$ and some data \mathcal{D} :

$$\text{simulate } \theta \sim p(\theta | \mathcal{D})$$

$$p(\theta | \mathcal{D}) = \frac{p(\mathcal{D} | \theta) p(\theta)}{Z}$$

Goal: sample $\theta \sim p(\theta | \mathcal{D})$

$$\propto p(\mathcal{D} | \theta) p(\theta)$$

Approach: Construct a Markov Chain with a stationary distribution $\pi(\theta | \mathcal{D}) \xrightarrow{i \rightarrow \infty} p(\theta | \mathcal{D})$.

- Proposal distribution / proposal matrix: $Q_{ab} = p(x=b | x=a)$
stochastic matrix: $\sum_b Q_{ab} = 1$, $Q_{ab} = Q_{ba}$, $a, b >$
- $p(\theta | \mathcal{D}) = \frac{\tilde{p}(\theta | \mathcal{D})}{p(\mathcal{D})} := \frac{p(\mathcal{D} | \theta) p(\theta)}{p(\mathcal{D})}$, $p(\mathcal{D}) = \int p(\mathcal{D} | \theta) p(\theta) d\theta$

Algorithm:

1. Choose a symmetric proposal matrix Q (Metropolis-Hastings does not require a symmetric Q)

2. Initialize a state $\theta_0 \in \mathbb{R}^d$

3. For $i = 0, 1, \dots, n-1$:

i.) Sample θ from $Q = p(\theta | \theta_i)$ (e.g. $\theta = \theta_i + \epsilon$, $\epsilon \sim \mathcal{N}(0, I)$)

ii.) Sample $u \sim \mathcal{U}(0, 1)$

iii.) Accept or reject θ according to the random Metropolis rule:

$$\text{If } u < \frac{\tilde{p}(\theta | \mathcal{D})}{\tilde{p}(\theta_i | \mathcal{D})} \begin{cases} \text{True, } \theta_{i+1} = \theta \\ \text{False, } \theta_{i+1} = \theta_i \end{cases}$$

$$u < \frac{p(\theta | \mathcal{D}) p(\theta)}{p(\theta_i | \mathcal{D}) p(\theta_i)}$$

4.) Output: $\theta_1, \dots, \theta_n \sim \text{M.C.}$

$$\mathbb{E}_{\theta \sim p(\theta|D)} [p(f(x^*)|\theta|D)] \approx \frac{1}{m} \sum_{i=1}^m p(f(x^*)|\theta_i, D)$$

$\theta_i \sim \text{M.C.}$

Theorem: If $(\theta_1, \dots, \theta_n)$ is an irreducible (time-homogeneous) discrete Markov Chain with a stationary distribution

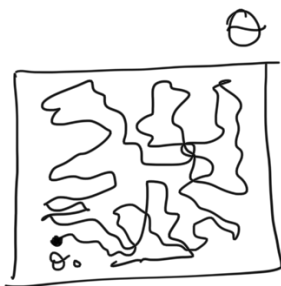
$\pi(\theta|D)$, then:

i.) $\frac{1}{n} \sum_{i=1}^n f(\theta_i) \xrightarrow[n \rightarrow \infty]{\text{a.s.}} \mathbb{E}_{\theta \sim p(\theta|D)} [f(\theta)]$ \forall any bounded function $f: \Theta \rightarrow \mathbb{R}$

If further the chain is aperiodic, then:

$$p(\theta_n | \theta_0, D) \xrightarrow[n \rightarrow \infty]{} p(\theta_n | D)$$

i.e. θ_n is a good sample from the target posterior $p(\theta|D)$



A dynamical system is ergodic if it can reach all possible states in a finite time, regardless of the initial condition θ_0 .

What is a Markov Chain?

$$\underbrace{(\theta_0) \rightarrow (\theta_1) \rightarrow \dots \rightarrow (\theta_n)}_{i=0, 1, \dots, n}$$

$$\Theta := \{ \underset{\uparrow}{\theta_0}, \underset{\uparrow}{\theta_1}, \dots, \underset{\uparrow}{\theta_n} \}$$

}

Markov property

$$p(\theta_i | \theta_0, \theta_1, \dots, \theta_{i-1}) = p(\theta_i | \theta_{i-1})$$

$$p(\theta_0, \theta_1, \dots, \theta_n) = p(\theta_0) p(\theta_1 | \theta_0) \dots p(\theta_n | \theta_{n-1})$$

If a discrete Markov Chain is irreducible, has a stationary distribution, and is aperiodic, then it is an Ergodic Markov Chain.

Definitions :

1.) A Markov - Chain is time-homogeneous / discrete-time if :

$$p(\theta_{i+1} = b \mid \theta_i = a) = Q_{ab}, \quad \forall i, \forall a, b \in \Theta$$

i.e. the transition probabilities do not depend on time (or the iteration index i)

2.) A pmf π on Θ is a stationary / invariant distribution with respect to Q_{ab} if : $\pi Q = \pi$, i.e. $\sum_{a \in \Theta} \pi_a Q_{ab} = \pi_b, \forall b$

* π is called a left-eigenvector and has eigenvalue 1.

3.) A M.C. is irreducible if $\forall a, b \in \Theta, \exists t \geq 0$ s.t.:

$$p(\theta_t = b \mid \theta_0 = a) > 0, \quad \forall a, b \in \Theta$$

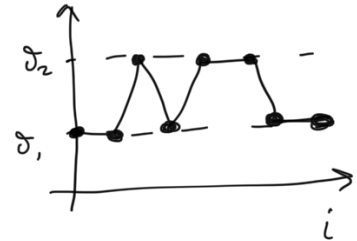
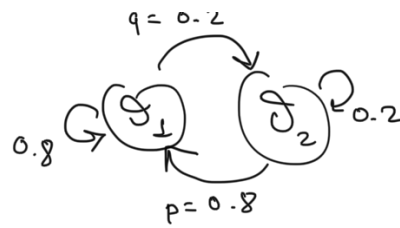
4.) An irreducible M.C. is called aperiodic if :

$$\forall a \in \Theta, \gcd \{ t : p(\theta_t = a \mid \theta_0 = a) > 0 \} = 1$$

↳ set of times for which if we start at $a \in \Theta$, we get back to a after some time

e.g. $\Theta = \{1, 2\}$

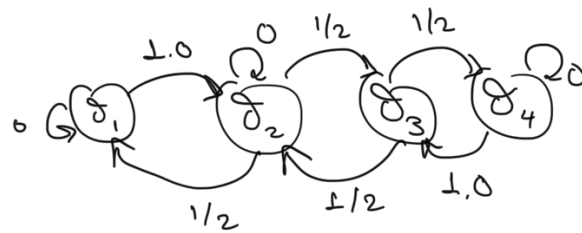
$$Q = \begin{bmatrix} p & q \\ p & q \end{bmatrix}$$



- irreducible ✓
- aperiodic ✓

$\Theta = \{1, 2, 3, 4\}$

$$Q = \begin{bmatrix} 0 & 1 & 0 & 0 \\ 1/2 & 0 & 1/2 & 0 \\ 0 & 1/2 & 0 & 1/2 \\ 0 & 0 & 1 & 0 \end{bmatrix}$$



- irreducible ✓
- aperiodic X

