ENM 531: Data-driven Modeling and Probabilistic Scientific Computing

Lecture #2: Primer on Probability and Statistics



Discrete random variables

- A discrete random variable is one which may take on only a countable number of distinct values such as 0,1,2,3,4,...... Discrete random variables are usually (but not necessarily) counts. If a random variable can take only a finite number of distinct values, then it must be discrete. Examples of discrete random variables include the number of children in a family, the Friday night attendance at a cinema, the number of patients in a doctor's surgery, the number of defective light bulbs in a b
- The *probability distribution* of a discrete random variable is a list of probabilities associated with each of its possible values. It is also sometimes called the probability function or the probability mass function.

Continuous random variables

- A continuous random variable is one which takes an infinite number of
 possible values. Continuous random variables are usually measurements.
 Examples include height, weight, the amount of sugar in an orange, the time
 required to run a mile.
- A continuous random variable is not defined at specific values. Instead, it is defined over an *interval* of values, and is represented by the *area under a curve* (in advanced mathematics, this is known as an *integral*). The probability of observing any single value is equal to 0, since the number of values which may be assumed by the random variable is infinite.
 - Suppose a random variable X may take all values over an interval of real numbers. Then the probability that X is in the set of outcomes A, P(A), is defined to be the area above A and under a curve. The curve, which represents a function p(x), must satisfy the following:
 - 1: The curve has no negative values $(p(x) \ge 0 \text{ for all } x)$
 - 2: The total area under the curve is equal to 1.
 - A curve meeting these requirements is known as a *density* curve.

Probability density functions

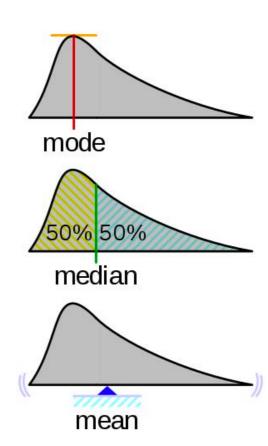
$$\Pr[a \leq X \leq b] = \int_a^b f_X(x) \, dx.$$

Hence, if F_X is the cumulative distribution function of X, then:

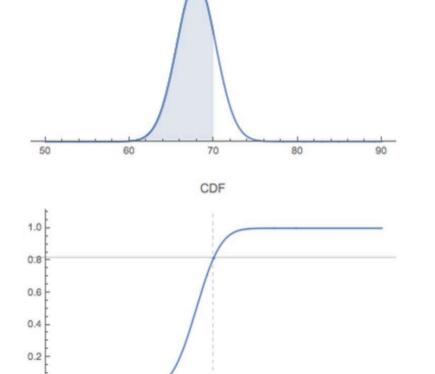
$$F_X(x) = \int_{-\infty}^x f_X(u)\,du,$$

and (if f_X is continuous at x)

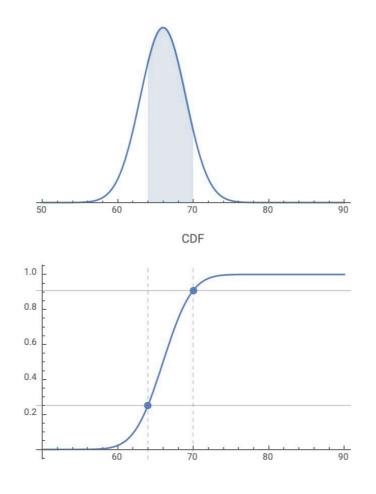
$$f_X(x)=rac{d}{dx}F_X(x).$$



Intuitively, one can think of $f_X(x) dx$ as being the probability of X falling within the infinitesimal interval [x, x + dx].



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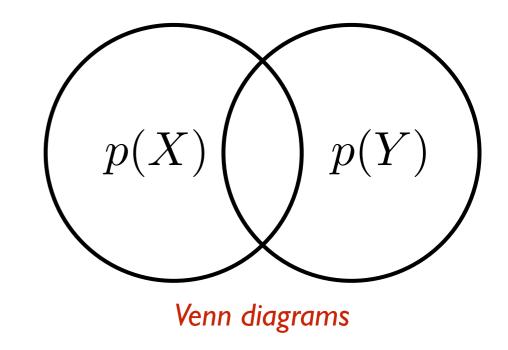


Basic rules of probability

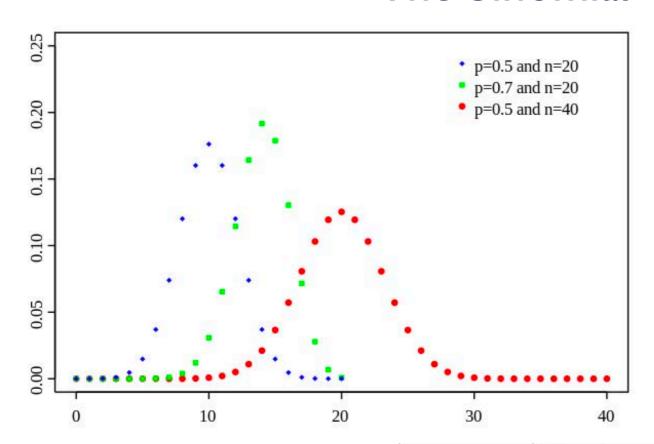
Sum rule
$$p(X) = \sum_{Y} p(X, Y)$$

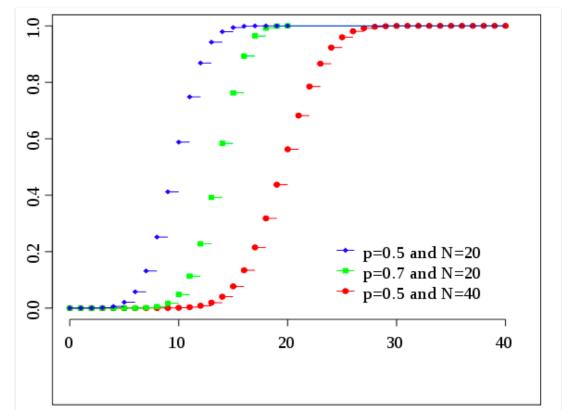
Product rule
$$p(X,Y) = p(Y|X)p(X)$$

Bayes rule
$$p(Y|X) = \frac{p(X|Y)p(Y)}{p(X)}$$



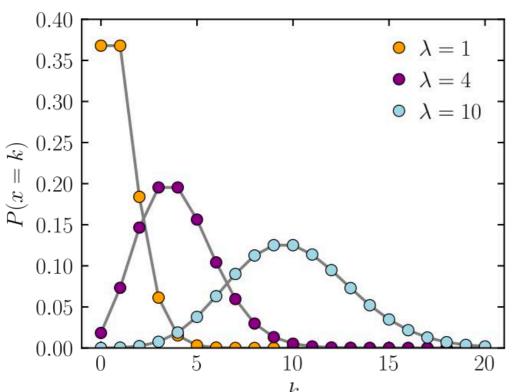
The binomial distribution

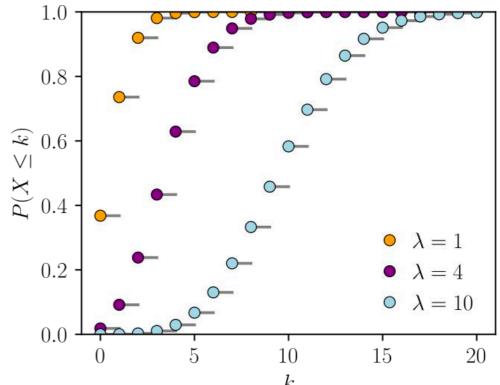




Notation	B(n,p)
Parameters	$n \in \{0,1,2,\ldots\}$ – number of trials
	$p \in [0,1]$ – success probability for
	each trial
	q=1-p
Support	$k \in \{0,1,\ldots,n\}$ – number of
	successes
PMF	$\binom{n}{k} p^k q^{n-k}$
CDF	$I_q(n-k,1+k)$
Mean	np
Median	$\lfloor np floor$ or $\lceil np ceil$
Mode	$\lfloor (n+1)p floor \lceil (n+1)p ceil -1$
Variance	npq

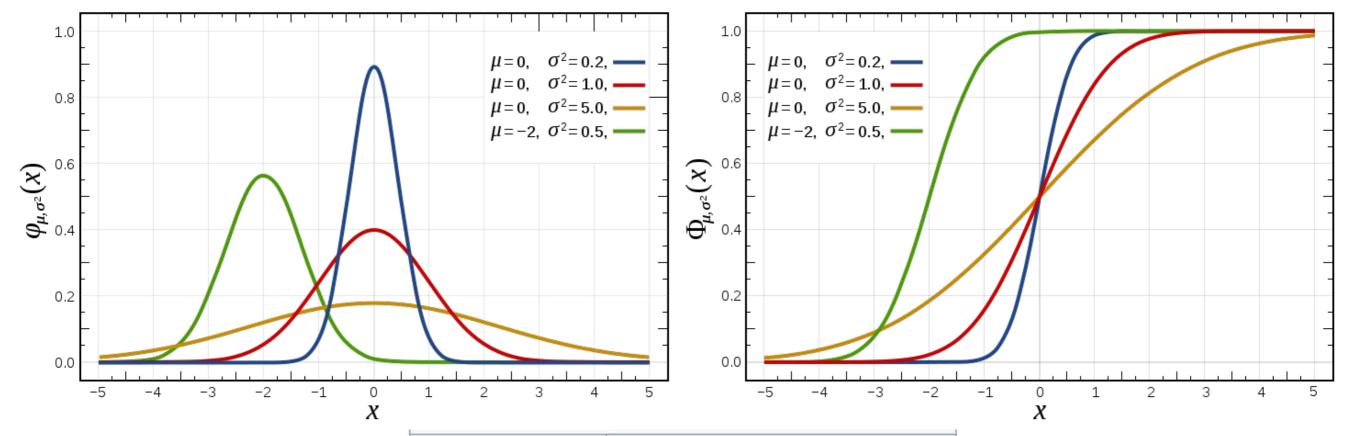
The Poisson distribution





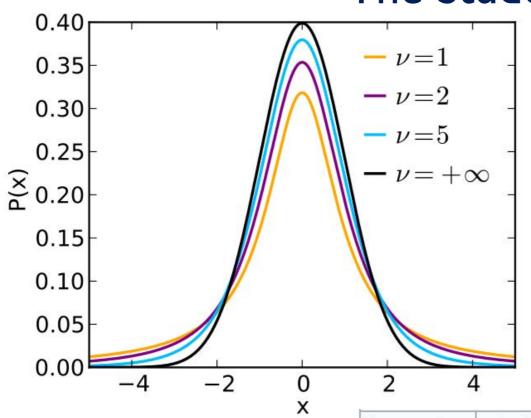
Notation	$\mathrm{Pois}(\lambda)$
Parameters	$\lambda \in (0,\infty)$ (rate)
Support	$k \in \mathbb{N}_0$ (Natural numbers starting from 0)
PMF	$\frac{\lambda^k e^{-\lambda}}{k!}$
CDF	$\frac{\Gamma(\lfloor k+1\rfloor,\lambda)}{\lfloor k\rfloor!}, \text{ or } e^{-\lambda} \sum_{i=0}^{\lfloor k\rfloor} \frac{\lambda^i}{i!} \text{ , or } Q(\lfloor k+1\rfloor,\lambda)$ (for $k\geq 0$, where $\Gamma(x,y)$ is the upper incomplete gamma function, $\lfloor k\rfloor$ is the floor function, and Q is the regularized gamma function)
Mean	λ
Median	$pprox \lfloor \lambda + 1/3 - 0.02/\lambda floor$
Mode	$\lceil \lambda ceil - 1, \lfloor \lambda floor$
Variance	λ

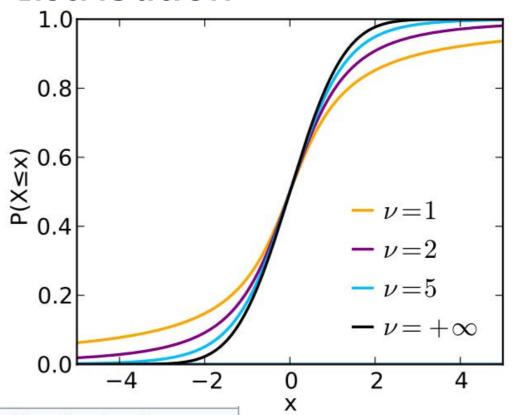
The Gaussian distribution



Notation	$\mathcal{N}(\mu,\sigma^2)$	
Parameters	$\mu \in \mathbb{R}$ = mean (location)	
	$\sigma^2>0$ = variance (squared scale)	
Support	$x\in\mathbb{R}$	
PDF	$rac{1}{\sqrt{2\pi\sigma^2}}e^{-rac{(x-\mu)^2}{2\sigma^2}}$	
CDF	$rac{1}{2}\left[1+ ext{erf}igg(rac{x-\mu}{\sigma\sqrt{2}}igg) ight]$	
Quantile	$\mu + \sigma\sqrt{2}\operatorname{erf}^{-1}(2F-1)$	
Mean	μ	
Median	μ	
Mode	μ	
Variance	σ^2	

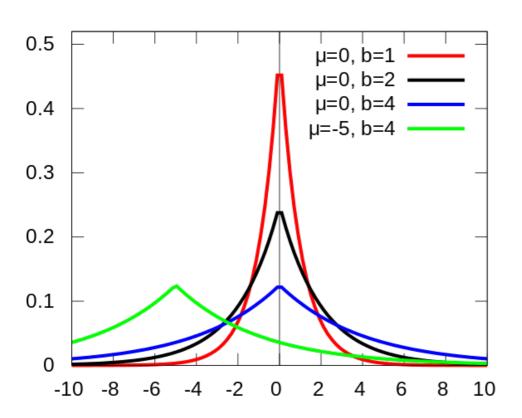
The Student-t distribution

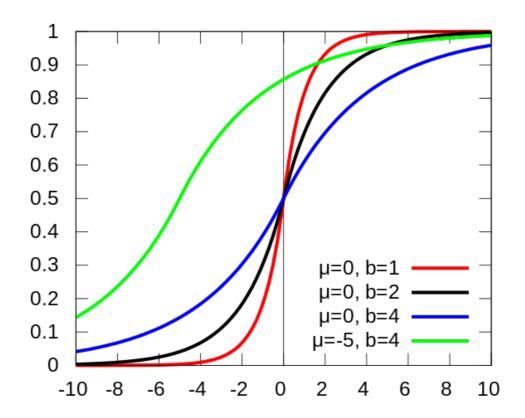




Parameters	u>0 degrees of freedom (real)
Support	$X \in (-\infty; +\infty)$
PDF	$rac{\Gamma\left(rac{ u+1}{2} ight)}{\sqrt{ u\pi}\Gamma\left(rac{ u}{2} ight)}\left(1+rac{x^2}{ u} ight)^{-rac{ u+1}{2}}$
CDF	$rac{1}{2} + x\Gamma\left(rac{ u+1}{2} ight) imes$
	$\frac{{}_2F_1\left(\frac{1}{2},\frac{\nu+1}{2};\frac{3}{2};-\frac{x^2}{\nu}\right)}{\sqrt{\pi\nu}\Gamma\!\left(\frac{\nu}{2}\right)}$
	where ₂ F ₁ is the hypergeometric function
Mean	0 for $ u > 1$, otherwise undefined
Median	0
Mode	0
Variance	$rac{ u}{ u-2}$ for $ u>2$, $ \infty$ for $ 1< u\leq 2$, otherwise undefined

The Laplace distribution





Parameters	μ location (real)	
	b>0 scale (real)	
Support	$x\in (-\infty;+\infty)$	
PDF	$\left rac{1}{2b}\exp\!\left(-rac{ x-\mu }{b} ight) ight $	
CDF	$\int \frac{1}{2} \exp\left(\frac{x-\mu}{b}\right)$	$\text{if } x < \mu$
	$\left[\left(1 - rac{1}{2} \exp\left(-rac{x-\mu}{b} ight) ight]$	$\text{if } x \geq \mu$
Mean	μ	
Median	μ	
Mode	μ	
Variance	$2b^2$	

Gaussian vs Student-t vs Laplace

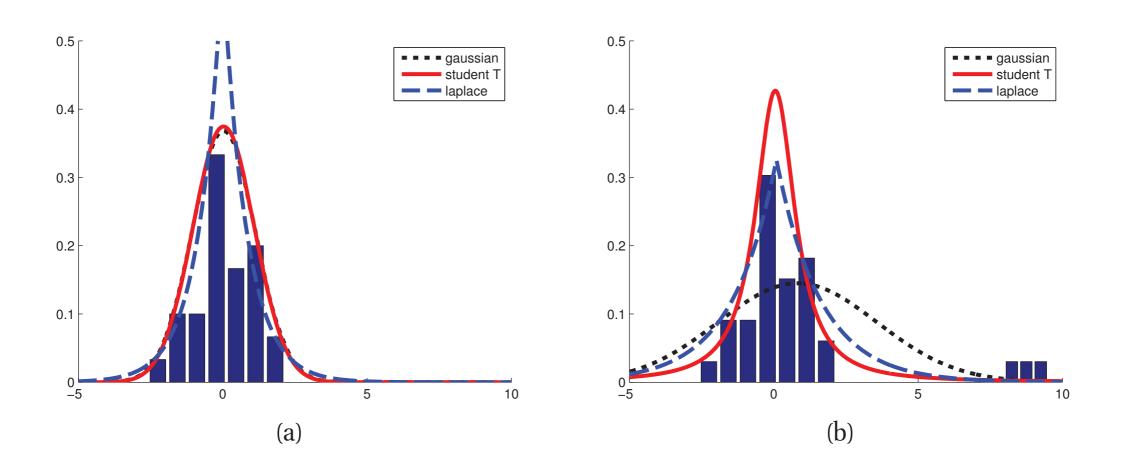
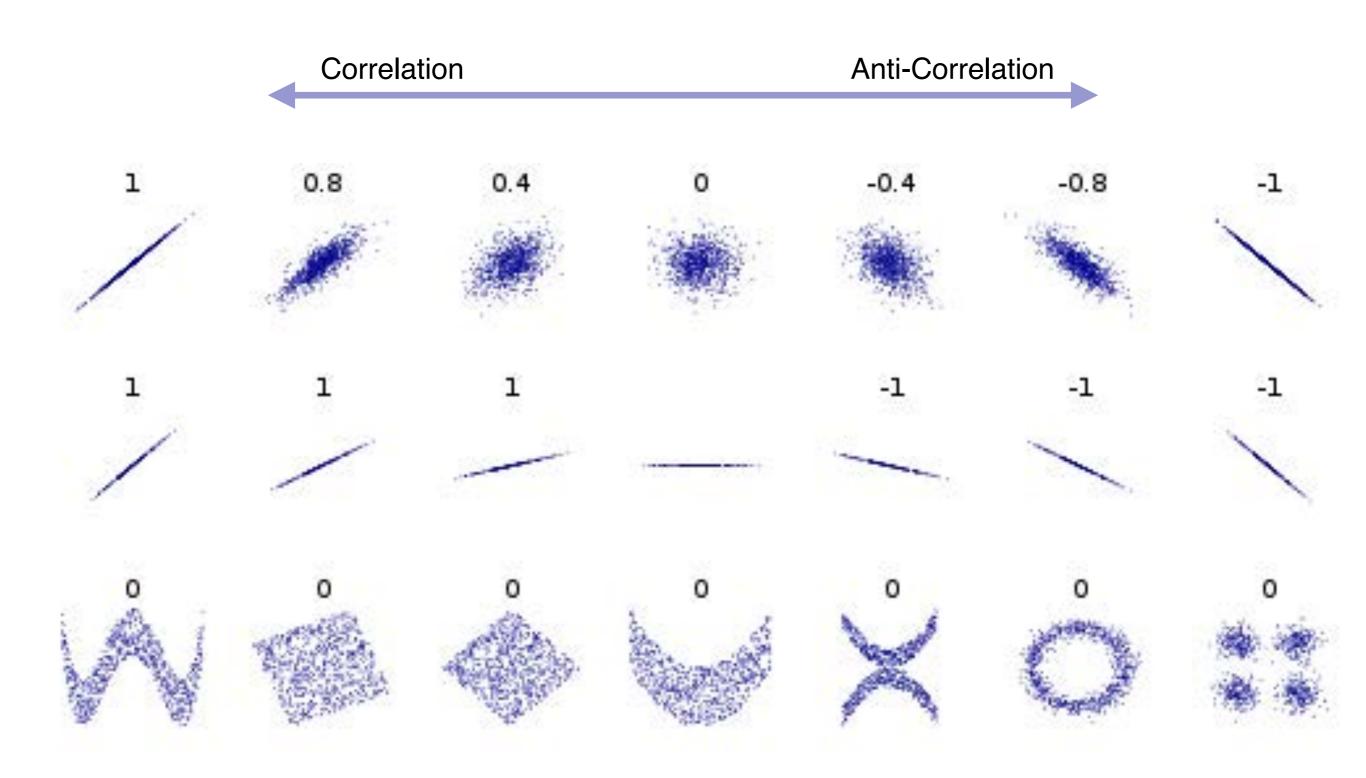
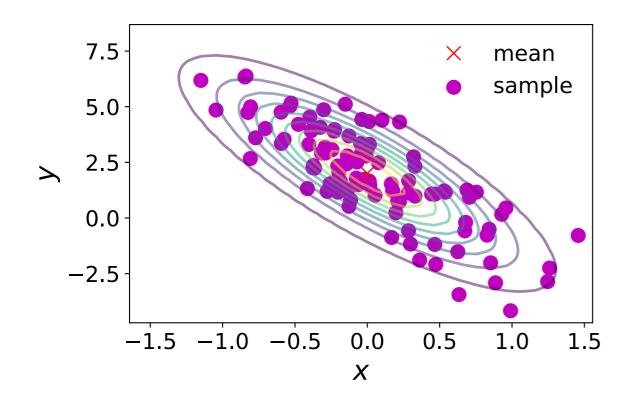


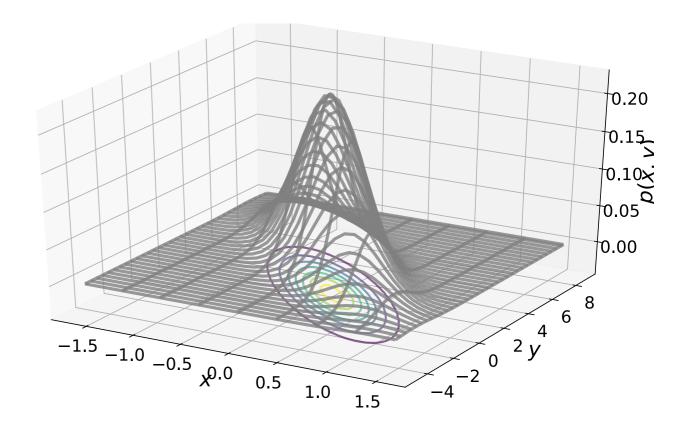
Figure 2.8 Illustration of the effect of outliers on fitting Gaussian, Student and Laplace distributions. (a) No outliers (the Gaussian and Student curves are on top of each other). (b) With outliers. We see that the Gaussian is more affected by outliers than the Student and Laplace distributions. Based on Figure 2.16 of (Bishop 2006a). Figure generated by robustDemo.

Correlation and linear dependence

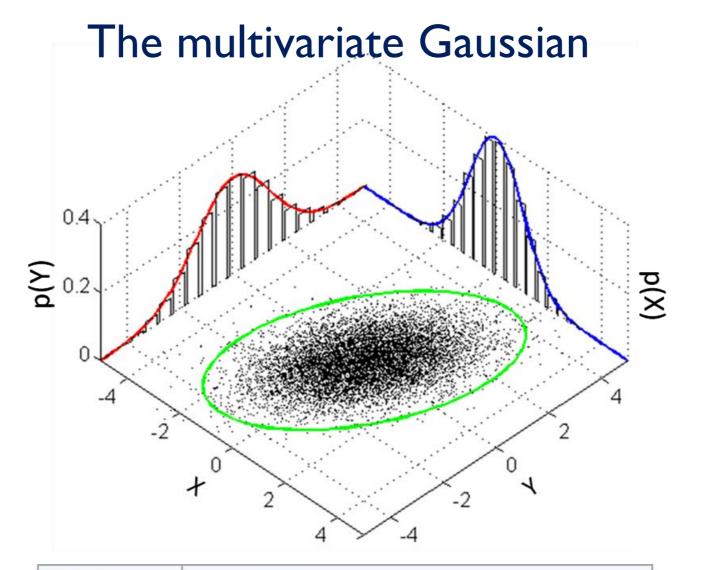


The multivariate Gaussian





$$p(\boldsymbol{x} \mid \boldsymbol{\mu}, \boldsymbol{\Sigma}) = (2\pi)^{-\frac{D}{2}} |\boldsymbol{\Sigma}|^{-\frac{1}{2}} \exp\left(-\frac{1}{2}(\boldsymbol{x} - \boldsymbol{\mu})^{\top} \boldsymbol{\Sigma}^{-1} (\boldsymbol{x} - \boldsymbol{\mu})\right)$$



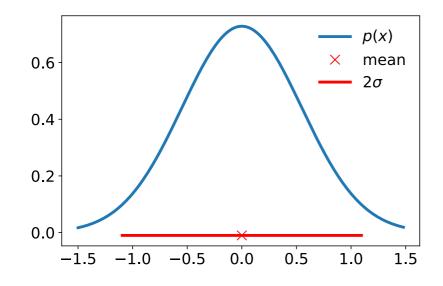
Notation	$\mathcal{N}(oldsymbol{\mu}, oldsymbol{\Sigma})$
Parameters	$\mu \in \mathbb{R}^k$ — location
	$\Sigma \in \mathbf{R}^{k \times k}$ — covariance (positive semi-
	definite matrix)
Support	$x \in \mu + \operatorname{span}(\Sigma) \subseteq \mathbf{R}^k$
PDF	$\det(2\pi\mathbf{\Sigma})^{-\frac{1}{2}}\;e^{-\frac{1}{2}(\mathbf{x}-oldsymbol{\mu})'\mathbf{\Sigma}^{-1}(\mathbf{x}-oldsymbol{\mu})},$
	exists only when Σ is positive-definite
Mean	μ
Mode	μ
Variance	Σ

Marginals and conditionals of a Gaussian

$$p(\boldsymbol{x}, \boldsymbol{y}) = \mathcal{N}\left(\begin{bmatrix}\boldsymbol{\mu}_{x}\\\boldsymbol{\mu}_{y}\end{bmatrix}, \begin{bmatrix}\boldsymbol{\Sigma}_{xx} & \boldsymbol{\Sigma}_{xy}\\\boldsymbol{\Sigma}_{yx} & \boldsymbol{\Sigma}_{yy}\end{bmatrix}\right) \xrightarrow[-2.5]{5.0}$$

Marginal distribution

$$p(\boldsymbol{x}) = \int p(\boldsymbol{x}, \boldsymbol{y}) d\boldsymbol{y} = \mathcal{N}(\boldsymbol{x} | \boldsymbol{\mu}_x, \boldsymbol{\Sigma}_{xx})$$



Conditional distribution
$$p(oldsymbol{x} \mid oldsymbol{y}) = \mathcal{N}ig(oldsymbol{\mu}_{x\mid y}, oldsymbol{\Sigma}_{x\mid y})$$
 $oldsymbol{\mu}_{x\mid y} = oldsymbol{\mu}_{x} + oldsymbol{\Sigma}_{xy}oldsymbol{\Sigma}_{yy}^{-1}(oldsymbol{y} - oldsymbol{\mu}_{y})$ $oldsymbol{\Sigma}_{x\mid y} = oldsymbol{\Sigma}_{xx} - oldsymbol{\Sigma}_{xy}oldsymbol{\Sigma}_{yy}^{-1}oldsymbol{\Sigma}_{yx}$.

These are unique properties that make the Gaussian distribution very simple and attractive to compute with! It is essentially our main building block for computing under uncertainty.