# Math 343 - Homework 3

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## Question 1

- **a**)
- b)
- **c**)
- d)
- **e**)
- f)

### **a**)

 $H_0$ : All means are equal.

 $H_a$ : Not all means are equal.

Performing an F test at we can see that the P-value=  $0.000 < \alpha = .05$ ,

#### **Analysis of Variance**

Source	DF	Adj SS	Adj MS	F-Value	P-Value
Technique	3	489740	163247	12.73	0.000
Error	12	153908	12826		
Total	15	643648			

Figure 1: The output of the One-way ANOVA from Minitab.

therefore we can conclude the following. There is enough statistical evidence to support the hypothesis that not all means are equal.

### b)

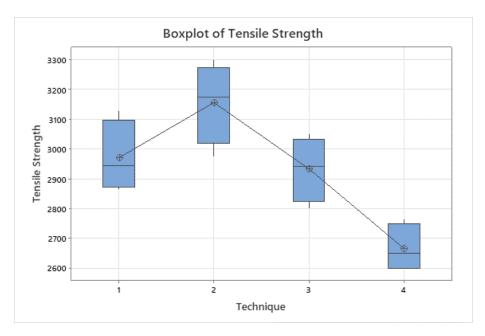


Figure 2: The Boxplot of tensile strength of the One-way ANOVA test from Minitab.

Judging by the boxplot alone, It would seem that the means are significantly different. This is consistent with the above hypothesis test. It looks as though the mean for treatment 1 and treatment 3 could be similar, in a pairwise comparison.

**c**)

The results of the Fisher's LSD test can be summarized in the graphical results below. Note that a line under the treatments indicate that they are not significantly different.

Mixing Technique (4) (3) (1) (2)

d)

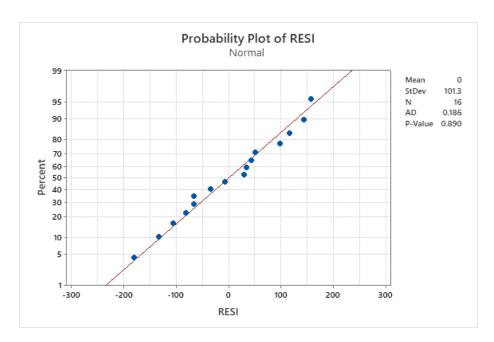


Figure 3: The normal probability plot of the residuals from Minitab.

 $H_0$ : The data are drawn from a normal disribution.

 $H_a$ : The data are not drawn from a normal disribution.

Since the P-value is very large (0.890), we can conclude the following. The evidence of the data is consistent with the data being drawn from a normal disribution.

**e**)

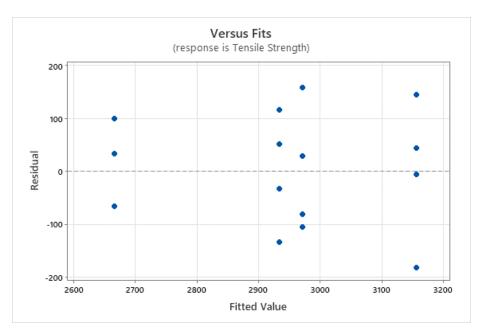


Figure 4: The residuals versus the predicted tensile strength from Minitab.

This graph indicates that there is not heterosked asticity present, that along with the conclusion that the data is drawn from a normal distribution indicates that the model assumptions are verified and the hypothesis test is valid.

**a**)

#### **a**)

 $H_0$ : All means are equal.

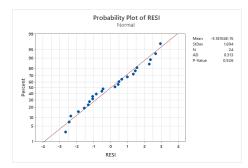
 $H_a$ : Not all means are equal.

Performing an F test at we can see that the P-value =  $.052 > \alpha = .05$ , therefore we can conclude the following. The evidence from the data is consistent with the hypothesis that all means are equal.

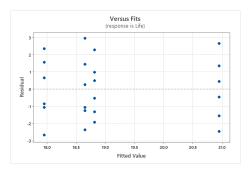
#### b)

Given that the F test concluded that all means are equal, It should not matter what fluid is selected given that the objective is long life. That being said, if forced to choose, I would pick fluid 3, as it had the highest sample mean at  $\bar{y}_{3} = 20.95$ .

#### $\mathbf{c})$



(a) Minitab output showing normal probability plot.



(b) Minitab output showing the residuals vs fitted value.

 $H_0$ : The data are drawn from a normal disribution.

 $H_a$ : The data are not drawn from a normal disribution.

Performing an Anderson Darling normality test at  $\alpha=0.05$  on the residuals gives a P-value of 0.526. Since the P-value is larger than  $\alpha$ , we can conclude the following. The evidence of the data is consistent with the data being drawn from a normal disribution.

Observing the residuals shows no indication of heterosked asticity. This, along with the normality test indicates that the basic analysis of variance assumptions are satisfied.

- d)
- **e**)

**a**)

Listing 1: R output of the summary of the aov function

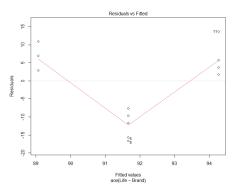
	Df	Sum Sq	Mean Sq	F	value	Pr(>F)
Brand	1	67.6	67.6		0.668	0.429
Residuals	13	1315.7	101.2			

 $H_0$ : All means are equal.

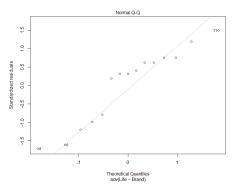
 $H_a$ : Not all means are equal.

Performing an F test at we can see that the P-value =  $0.429 > \alpha = .05$ , therefore we can conclude the following. The evidence from the data is consistent with the hypothesis that all means are equal.

b)



(a) R output showing the residuals.



(b) R output showing the normal probability plot.

 $\mathbf{c})$ 

# 95 percent confidence interval estimate on the mean life of battery brand 2.

In order to construct a 95 percent confidence interval estimate on the mean life of battery brand 2. The following equation can be used.

$$\sum_{i=1}^{a} c_i \bar{y}_i - t_{\alpha/2, N-a} \sqrt{\frac{MSE}{n} \sum_{i=1}^{a} c_i^2} \le \sum_{i=1}^{a} c_i \mu_i \le \sum_{i=1}^{a} c_i \bar{y}_i + t_{\alpha/2, N-a} \sqrt{\frac{MSE}{n} \sum_{i=1}^{a} c_i^2}$$

The contrast of intrest, where  $c_1 = 0$ ,  $c_2 = 1$ ,  $c_3 = 0$ , and a = 3 is:

$$\Gamma = \sum_{i=1}^{a} c_i \bar{y}_i.$$

$$= 0(95.2) + 1(79.4) + 0(100.4)$$

$$= 79.4$$

Using R, the following values were obtained. MSE = 1315.7, and  $t_{\alpha/2,N-a} = t_{0.025,12} = 2.178$ . Thus, the 95 percent confidence interval estimate on the mean life of battery brand 2 is:

$$79.4 - 2.178\sqrt{\frac{1315.7}{5} \cdot 1} \le \sum_{i=1}^{a} c_i \mu_i \le 79.4 + 2.178\sqrt{\frac{1315.7}{5} \cdot 1}$$
$$44.06 \le \sum_{i=1}^{a} c_i \mu_i \le 114.73$$

# 99 percent confidence interval estimate on the mean difference between the lives of battery brands 2 and 3.

The contrast of intrest, where  $c_1 = 0$ ,  $c_2 = 1$ ,  $c_3 = -1$ , and a = 3 is:

$$\Gamma = \sum_{i=1}^{a} c_i \bar{y}_i.$$

$$= 0(95.2) + 1(79.4) + -1(100.4)$$

$$= -21$$

Using R, the following values were obtained.  $t_{\alpha/2,N-a} = t_{0.005,12} = 3.054$ . Thus, the 99 percent confidence interval estimate on the mean difference between the lives of battery brands 2 and 3 is:

$$-21 - 3.054\sqrt{\frac{1315.7}{5} \cdot 2} \le \sum_{i=1}^{a} c_i \mu_i \le -21 + 3.054\sqrt{\frac{1315.7}{5} \cdot 2}$$
$$-91.06 \le \sum_{i=1}^{a} c_i \mu_i \le 49.06$$

d)

The percentage of batteries expected to fail before 85 weeks can be obtained in R using the following code.

Listing 2: Calculating the percentage of batteries expected to fail before 85 weeks.

```
# Calculate percentage of batteries expected to fail
percentage <- pnorm(85, mean = 100.4, sd = 36.27) * 100
# Print the result
percentage</pre>
```

Listing 3: Code output.

33.5566