Homework #2

Professor: Pat Kline

Students: Christina Brown, Sam Leone, Peter McCrory, Preston Mui

2. Logit MLE

a) Derive the score of the logit-likelihood: First, noting that

$$\frac{\partial \Lambda(X_i'\beta)}{\partial \beta} = \frac{\exp(X_i'\beta)X_i'(1+X_i'\beta) - \exp(X_i'\beta)\exp(X_i'\beta)X_i'}{(1+\exp(X_i'\beta))^2} = \frac{\exp(X_i'\beta)X_i'}{(1+\exp(X_i'\beta))^2}$$

The score of the logit log-likelihood is therefore

$$s(\beta) = \sum_{i} \frac{Y_{i} \exp(X'_{i}\beta)X'_{i}}{\Lambda(X'_{i}\beta)(1 + \exp(X'_{i}\beta))^{2}} - \frac{(1 - Y_{i}) \exp(X'_{i}\beta)X'_{i}}{(1 - \Lambda(X'_{i}\beta))(1 + \exp(X'_{i}\beta))^{2}}$$

$$= \sum_{i} \left(Y_{i} - (1 - Y_{i}) \exp(X'_{i}\beta)\right) \frac{X'_{i}}{1 + \exp(X'_{i}\beta)}$$

$$= \sum_{i} \left(Y_{i}(1 + \exp(X'_{i}\beta)) - \exp(X'_{i}\beta)\right) \frac{X'_{i}}{1 + \exp(X'_{i}\beta)}$$

$$= \sum_{i} \left(Y_{i} - \frac{\exp(X'_{i}\beta)}{1 + \exp(X'_{i}\beta)}\right)X'_{i}$$

b) The moment condition identifying β_{ML} is

$$E\left[\left(Y_i - \frac{\exp(X_i'\beta)}{1 + \exp(X_i'\beta)}\right)X_i'\right] = 0$$

$$E\left[E[Y_i - \frac{\exp(X_i'\beta)}{1 + \exp(X_i'\beta)}|X_i]X_i'\right] = 0$$

c) The moment conditions identifying β_{NLLS} are

$$0 = E \left[-2(Y_i - \Lambda(X_i'\beta)) \frac{\exp(X_i'\beta)}{(1 + \exp(X_i'\beta))^2} X_i' \right]$$
$$= E \left[E[(Y - \Lambda(X_i'\beta)) | X_i] \frac{\exp(X_i'\beta)}{(1 + \exp(X_i'\beta))^2} X_i' \right]$$

d) Under which conditions does β_{NLLS} coincide with β_{ML} ? The moment conditions will coincide when $E[Y_i - \Lambda(X_i'\beta)|X_i'] = 0$ for all X_i ; that is, they coincide when the logit model is correctly specified.

e)

3. Matlab Probit DGP

- a) (Matlab program attached)
- b) The ML point estimates of $\hat{\beta}^{con} = (\hat{\alpha}, \hat{b})$ are (-0.0175, 0.9159). The standard errors are (0.0918, 0.1301).
- c) Score test: Regressing the restricted model's generalized residuals, given by

$$gres_i = \frac{Y_i \phi(X_i' \hat{\beta}^{con})}{\Phi(X_i' \hat{\beta}^{con})} - \frac{(1 - Y_i) \phi(X_i' \hat{\beta}^{con})}{1 - \Phi(X_i' \hat{\beta}^{con})}$$

on X_i^2 yields a coefficient of 0.0039 with a p-value of effectively 0, rejecting the null of 0.

d) The unrestricted model yields point estimates of $\hat{\beta}^{unc} = (\hat{\alpha}, \hat{b}, \hat{b_2}) = (-0.0435, 0.9175, 0.0428)$. The wald test, which tests $g(\beta) = 0$, where

$$g(\beta) \equiv (0, 0, 1) \cdot \beta = b_2$$
$$G(\beta) \equiv \frac{\partial g(\beta)}{\partial \beta} = (0, 0, 1)$$

and the Wald statistic is given by

$$N \cdot \hat{b_2} \cdot \left(G \cdot H^{-1} \cdot G'\right)^{-1} \cdot \hat{b_2}$$

where H is the average Hessian at the ML estimate. Here, the Wald evaluates to 0.3336, which has a p-value of 0.5636 under the χ^2 distribution with d.f. 1, much higher than the 0 obtained in part b.