Using Lotteries to Encourage Saving: Experimental Evidence from Kenya

Justin Abraham, Merve Akbas, Dan Ariely, and Chaning Jang, July 21, 2016

Abstract

This paper reports a "lab-in-the-field" experiment that test a savings scheme incorporating lottery-like payoffs to traditional savings accounts. We provide a mobile savings product to 311 informal residents in Nairobi, Kenya and observe account activity over a 60-day period. We find that respondents with lottery-linked deposit accounts (LLDAs) made 30-40% more deposits on average over the project period than respondents receiving a matched incentive. Moreover, this increase in account activity is due to respondents making more deposits per day and depositing for more days. We also show that when presented with potential winnings from previous days, respondents with LLDAs increased self-reported gambling activity by 15%. Our results suggest that LLDAs are a promising tool to improve savings among the poor but that product design has considerable implications on gambling behavior.

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I. Introduction

Although access to savings accounts for the poor has improved in recent years, demand for and usage of savings accounts remains a stumbling block hindering full financial inclusion (Kendall, Mylenko, and Ponce 2010). Standard interventions have focused on lowering transaction costs, boosting incentives to save, and improving financial infrastructure in underserved areas (Ashraf, Karlan, and Yin 2006a; Schreiner 2006; Allen et al. 2014). The growth of mobile money services has been especially important in encouraging the use of savings account in Sub-Saharan Africa (Demirgüç-Kunt et al. 2015). A recent focus on behavioral barriers to saving has been fruitful. Behaviorally-informed product designs, including SMS reminders (Karlan et al. 2010), commitment devices (Ashraf, Karlan, and Yin 2006b; Dupas and Robinson 2013), default contributions (Thaler and Benartzi 2004; Chetty and Friedman 2014), and tokens (Akbas et al. 2016), have been shown to be extremely cost-effective compared to traditional economic interventions.

Our study focuses on an alternative scheme that incorporates lottery-like payoffs to traditional savings accounts. Lottery-linked deposit accounts (LLDAs) have been in use since at least the 17th century and presently exist in various forms around the globe (Murphy 2005; Kearney et al. 2010). NS&I Premium Bonds in the U.K. and First National Bank's "A-Million-A-Month" Account in South Africa are two prominent examples of this type of savings product. Lottery-linked savings accounts are unique in that they provide stochastic returns as a function of savings or deposits. Savers will forego interest for a probabilistic payoff but face no risk of losing their principal. This unique feature makes the product attractive as a tool to promote financial inclusion. Lottery expenditures as a proportion of income are highest among the poor, which suggests that they may be especially responsive to lottery-like incentive structures (Brown, Kaldenberg, and Browne 1992). Furthermore, there is some evidence that usage of lottery-linked accounts displaces regular gambling behavior, thereby redirecting gambling expenditures into savings (Cookson 2016).

How might lotteries induce savings? Extensions of von Neumann-Morgenstern expected utility can explain why risk-averse individuals might prefer stochastic over fixed returns. These include the over-weighting of small probabilities (Kahneman and Tversky 1979), preferences for skewness (Golec and Tamarkin 1998), excitement from gambling (Conlisk 1993), or consumption indivisibilities faced by poor households (McCaffrey 1994). Persistent gambling behavior in the face of losses might also stem from cognitive biases such as the gambler's fallacy and the "hot hand" fallacy (Gilovich, Vallone, and Tversky 1985). There also exists evidence that aversion to anticipated regret could drive gambling behavior (Zeelenberg and Pieters 2004).

Literature on the potential demand for LLDAs is extensive, but empirical evidence as to its effect on savings behavior is scarce. Two recent experimental studies provide support for an extant benefit of stochastic returns on saving for the future. Atalay et al.

(2014) conducted an online portfolio-choice experiment that resulted in respondents saving an additional 12 percentage points more with lottery-linked and regular savings than with regular savings alone. Notably, respondents who saw an increase in total savings shifted away from lottery expenditures and consumption rather than from regular savings. Filiz-Ozbay et al. (2015) found respondents are more likely to delay payments with lottery-like returns compared to guaranteed interest of equivalent expected value. This finding suggests that lottery-linked schemes can be designed to be revenue neutral in expectation for account providers while still promoting savings. Filiz-Ozbay et al. (2015) additionally estimate structural parameters and argue that probability weights better explain the result than individual preferences for skewness. Outside the laboratory, evidence regarding LL-DAs is much more limited and diverges from experimental findings. Loibl et al. (2016) conducted a randomized evaluation of Individual Development Accounts (IDAs) in the U.S. that incorporated a lottery-based savings match. That study found no significant effect of the program relative to guaranteed matching, even when it was bundled with reminder calls and frequent deposit deadlines. They attribute the result to liquidity constraints among their sample, which potentially precluded the benefits of behavioral interventions. Research on the effect of other lottery-based incentive schemes relative to deterministic incentives remains inconclusive (Halpern et al. 2011; Gajic, Cameron, and Hurley 2011; Brune 2015).

The present study is a "lab-in-the-field" experiment testing the effects of lottery-linked savings on savings behavior. We provide a mobile savings product to 311 informal residents in Nairobi, Kenya and observe account activity over a 60-day period. We minimize barriers to saving by utilizing Safaricom's Sambaza mobile savings technology. This platform also allows us to collect detailed data on participant transactions to be able to examine savings behavior over time. One group is randomly assigned a savings account which provides a fixed 5% match to all deposits. A second group is assigned an account that yields stochastic returns, equal in expectation to the 5% match, through a lottery conducted on a daily basis. For each day a respondent makes a non-zero deposit, they receive a lottery ticket and an opportunity to win a prize instead of the fixed match. We compare the interest and lottery groups to determine how LLDAs impact savings behavior. A third group is given the same lottery-linked account with the additional feature that respondents receive a lottery ticket every day regardless of saving but cannot claim the prize untile after making a deposit. The key feature of this "regret" treatment is that respondents observe the lottery results and potential prize at the end of each day. We test this treatment against the lottery treatment to determine whether experienced regret from being unable to claim a prize affects decisions

We find that respondents with LLDAs made 30-40% more deposits on average over the project period than respondents receiving the matched incentive. Moreover, this increase in account activity is due to respondents making more deposits per day and depositing

over more days during the savingss period. Respondents saved 50% more per day and saved 5 days more compared to the control group. Effects on the number of deposits were more acute for respondents who are female, who are over 30 years old, with lower incomes, who are less prone to gambling, and who are risk-seeking, though we report no significant heterogeneous effects. There were no significant differences in effects on saving between the regular LLDA and the LLDA with regret framing. Interestingly, we find no effect of LLDAs on total amount saved or on the size of each deposit. Respondents made smaller, more frequent deposits compared to the control group. We find no evidence of the LLDA displacing savings from other sources. On gambling behavior, we find that 27% respondents in the regret framing self-report higher gambling activity compared to 12% in the control group.

This study contributes to the literature as one of the first randomized evaluations examining the impact of LLDAs on real-world savings behavior. Moreover, the study's unique experimental design allows us to identify dynamic effects – respondents make more frequent deposits to their accounts when given lottery-based returns. This result suggests that the appeal of gambling may be enough to induce a change in savings behavior. LLDAs may thus improve utilization among existing account holders and be able to attract new savers to open formal savings accounts. Frequent deposits may also have long-term benefits by encouraging the formation of a savings habit (Alessie and Teppa 2009). From a policy perspective, LLDAs are appealing because they are more effective at changing savings behavior than matched incentives without an increase in the incentive itself. However, LLDAs may not be revenue neutral if financial institutions incur greater transaction costs as a result of more frequent deposits.

Our study also shows that respondents with LLDAs with regret framing increased self-reported gambling activity relative to the control group. If LLDAs contribute to problem gambling, the program is potentially welfare-decreasing for poor households already susceptible to costly gambling behavior. Cookson (2016) reports a 50% reduction of casino gambling in Nebraska as a result of enrollment in an LLDA bundled with an anti-gambling advertising campaign. The difference from our results suggests that additional program components could diminish effects on outside gambling. Overall, we document several advantages of LLDAs over fixed-incentive schemes when it comes to promoting financial inclusion and show that product design is crucial in moderating adverse effects on gambling behavior.

The remainder of the paper is structured as follows. Section II describes our experimental design, Section III outlines our estimation strategy, and Section IV displays our main results.

II. Experimental Design

This study was conducted in conjunction with the Busara Center for Behavioral Economics in Nairobi, Kenya. We recruited 311 respondents through Busaras respondent pool. Participants were first invited to the lab at Busara where they completed a computerized questionnaire and behavioral tasks. The following outlines the schedule of tasks during the lab portion of the study:

- 1. Risk preference elicitation
- 2. Time preference elicitation
- 3. Willingness-to-pay to play a lottery
- 4. Internal locus of control
- 5. Baseline demographics questionnaire

Following the lab session, respondents were randomly assigned to one of the three treatment groups. Respondents then learned about their assigned savings program. Each respondent was given KSH 20 airtime credit and asked to practice saving using Sambaza. Respondents were then sent home with business-card sized handouts which described their savings program. We provided respondents simple instructions for saving and listed the number to our project phone. This was the number through which the savings program operated that also functioned as a help line for respondents.

Lab sessions took place over five weeks in May and June of 2014. Respondents were enrolled in the our savings program for two consecutive periods of 30 days starting from the day of a respondents lab session. On a respondents 30th day, a field officer called them and asked if they wished to withdraw any amount of their balance. Respondents who requested withdrawals were sent M-Pesa transfers equal to their request plus the M-Pesa withdrawal fee. These withdrawals were recorded in our systems ledger.

Following this, respondents moved on to their second 30-day savings period. Respondents were called and notified a few days before the end of their second 30-day period that the program would be ending soon. After receiving the end-of-day message on their 60th day, respondent were unenrolled from the program and were no longer allowed to save. Field officers called respondents to confirm final balances and sent M-Pesa transfers equal to total balance plus withdrawal fee shortly after. All respondents had completed the program by August 2014. In September 2014, we called respondents and conducted an endline survey. We obtained endline surveys for all but 27 of the 311 respondents.

A. Mobile savings program

We implemented our mobile-phone based savings program over Safaricom's *Sambaza* airtime sharing service. Using *Sambaza*, Safaricom users can send airtime to each other free of charge. Respondents saved into our program by sending airtime to a designated project phone that held the airtime in an account for each user.

Respondents received two SMS messages every morning after the first morning of the project period. The first message was an end-of-day message that reported how much the respondent saved the previous day, how much the respondent earned through interest or winnings, and their total balance. An hour later, respondents received a beginning-of-day message encouraging them to save that day. Respondents were allowed to send in savings at any time but any savings sent in after the end-of-day message would be counted towards the next days total. We used a custom-developed administrative system to manage the savings program. This system logged airtime sent to our project phone, maintained an internal ledger of balances, sent automated SMS confirmations after every transaction, and conducted the daily lottery game.

Respondents were enrolled in the savings program for a total of 60 days, split into consecutive 30-day periods. After the first 30 days, respondents were allowed to withdraw any amount of their savings up to the total balance. Outside of this opportunity, regular withdrawals were not allowed.

At the end of our experiment, we returned respondents savings and accumulated interest or winnings via an M-Pesa transfer. This M-Pesa transfer included the extra withdrawal fees needed to cash out an amount equal to the respondents full account balance. Therefore, respondents paid no explicit fees to participate in our program.

Respondents were randomized into one of three treatment groups and had the chance to earn either daily interest in the form of matching or play a daily lottery depending on assignment.

A.1 Interest-bearing savings

Respondents in the control group participated in a standard, interest-bearing savings program where they earned a 5% matching contribution on any amount that they saved in a particular day.

A.2 Lottery-linked savings

After saving a non-zero amount, respondents earned a lottery ticket - transmitted via SMS, which could win a cash prize in proportion to the amount they saved. A lottery ticket was a random sequence of four numbers between 1 and 9, inclusive. Each day, our administrative system randomly generated a winning sequence of four numbers. Prizes were awarded

according to how well a respondents lottery numbers matched the winning numbers. If the first or second numbers matched, a 10% match of savings was awarded. If both the first and second numbers matched, a 100% match of savings was awarded. Finally if all numbers matched, a prize of 200 times the daily savings was awarded. The earnings on this lottery ticket were equal in expectation to the 5% match earned in the control group. Our system processed the matching of lottery numbers and entered winnings into the internal ledger. Respondents could only earn one lottery ticket per day.

A.3 Lottery-linked savings with regret

This scheme is similar to the lottery treatment but respondents in this third group were sent lottery tickets in their beginning-of-day text message. These tickets only became redeemable, however, after respondents had saved a non-zero amount that day. Respondents with winning lottery tickets who did not save that day did not win money from their lottery ticket. However, they were informed whether they would have won in their end-of-day message the next morning.

III. Estimation Strategy

A. Treatment effect

We use the following econometric specification for basic identification of the treatment effect.

$$y_{i,t=1} = \beta_0 + \beta_1 \text{LOTTERY}_i + \beta_2 \text{REGRET}_i + \delta y_{i,t=0} + \varepsilon_i$$
 (1)

 $Y_{i,t=1}$ refers to the outcome variables for individual i at endline, LOTTERY $_i$ indicates assignment to the lottery group, and REGRET $_i$ indicates assignment to the lottery with regret framing group. The omitted group is the interest group. β_1 and β_2 respectively identify the treatment effects of the lottery and lottery with regret framing relative to the interest group. Following McKenzie (2012), we include the baseline level of the individual outcome $y_{i,t=0}$ in equation 1 where possible to improve statistical power. We will use an F-test to test the joint effect of both treatments to the comparison group and to compare the effects against one another. We also estimate a model which includes a vector of control variables measured at baseline.

We might expect that the errors for each of these regression are correlated. Instead of estimating these equations separately, we can estimate the system of seemingly unrelated regressions to improve the precision of the coefficient estimates (Zellner 1962). SUR estimation is equivalent to OLS when the error terms are in fact uncorrelated between regressions or when each equation contains the same set of regressors. Simultaneous estimation allows us to perform tests of joint significance on the treatment coefficients across equations.

We additionally control for the family-wise error rate (FWER) using the free step-down resampling method to compute adjusted p-values within each family of outcome variables (Westfall and Young 1993). This approach sets the size of the test to exactly the desired crticial value. For each variable, we report both unadjusted p-values as well as p-values corrected for multiple inference.

B. Heterogeneous treatment effects

We explore the extent to which the savings program produces heterogeneous treatment effects. We use the following specification for this analysis.

$$y_{i,t=1} = \beta_0 + \beta_1 \text{LOTTERY}_i + \beta_2 \text{REGRET}_i + \gamma_1 (\text{LOTTERY}_i \times x_{i,t=0}) + \gamma_2 (\text{REGRET}_i \times x_{i,t=0}) + \gamma_3 x_{i,t=0} + \delta y_{i,t=0} + \varepsilon_i$$
(2)

 $x_{i,t=0}$ is the dimension of heterogeneity measured at baseline. γ_1 and γ_2 respectively identify the heterogeneous treatment effects of the lottery and lotter with regret framing relative to when $x_{i,t=0} = 0$. We estimate this model with the following baseline variables:

- 1. Gender
- 2. Marriage status
- 3. Age
- 4. Education level
- 5. Use of a savings account
- 6. Monthly income
- 7. Employment status
- 8. Problem gambling
- 9. Risk aversion

IV. Results

Table 1: Treatment group by participation at endline

	Participation in endline							
	Attrited Completed Total							
Interest	11	94	105					
Lottery	8	95	103					
Regret	8	95	103					
Total	27	284	311					

Notes: This table reports a cross-tabulation between treatment assignment and selection into the endline survey.

Table 2: Attrition by treatment group

	Unobserved at endline
Lottery	-0.03
	(0.04)
Regret	-0.03
	(0.04)
Constant	0.10***
	(0.03)
Observations	311
Adjusted \mathbb{R}^2	-0.004
Difference p-value	1.00
Joint p-value	0.75

Notes: This table reports a regression of selection on each of the treatment arms. Standard errors are in parentheses. * denotes significance at 10 pct., ** at 5 pct., and *** at 1 pct. level.

Table 3: Summary statistics by treatment group

	Mean (SD, N)				Difference p-value			
	Control	Lottery	Regret	Lottery - Control	Regret - Control	Lottery - Regret		
Female	0.52	0.59	0.62	0.32	0.16	0.67		
	(0.50)	(0.49)	(0.49)					
	105	103	103					
Age	30.06	36.95	30.93	0.26	0.54	0.33		
	(10.52)	(61.21)	(9.96)					
	105	103	103					
Completed std. 8	0.99	0.97	0.97	0.31	0.31	1.00		
	(0.10)	(0.17)	(0.17)					
	105	103	103					
Married/co-habitating	0.42	0.52	0.51	0.15	0.21	0.83		
	(0.50)	(0.50)	(0.50)					
	104	101	102					
No. of children	1.75	1.98	1.99	0.34	0.33	0.97		
	(1.70)	(1.71)	(1.84)					
	105	103	103					
Constant relative risk aversion	1.16	1.25	1.13	0.64	0.85	0.52		
	(1.27)	(1.38)	(1.24)					
	105	103	103					
Locus of control	69.81	70.29	68.98	0.73	0.57	0.34		
	(10.78)	(9.41)	(10.30)					
	105	103	103					

Notes: The first three columns report means of each row variable for each treatment group. SD are in parentheses and N is displayed on the third line. The last three columns report the p-value for a difference of means t-test between each group. * denotes significance at 10 pct., ** at 5 pct., and *** at 1 pct. level.

Table 4: Summary statistics by treatment group

	N	Iean (SD, I	N)		Difference p-value	
	Control	Lottery	Regret	Lottery - Control	Regret - Control	Lottery - Regret
Monthly income	112.05 (137.13) 105	108.37 (117.43) 103	111.46 (104.85) 103	0.84	0.97	0.84
Receives regular income	0.06 (0.24)	0.11 (0.31)	0.17 (0.38)	0.36	0.08*	0.38
Employed	52 0.50 (0.50)	56 $ 0.54 $ $ (0.50)$	$ \begin{array}{c} 48 \\ 0.47 \\ (0.50) \end{array} $	0.49	0.68	0.27
Self-employed	$ \begin{array}{c} 105 \\ 0.24 \\ (0.43) \end{array} $	103 0.21 (0.41)	103 0.20 (0.40)	0.61	0.49	0.87
No. of dependents	78 3.18 (2.58)	72 3.49 (2.60)	81 3.27 (2.32)	0.40	0.79	0.53
Subject is a dependant	105 0.23 (0.42)	103 0.28 (0.45)	103 0.25 (0.44)	0.38	0.69	0.64
	105	103	103			

Notes: The first three columns report means of each row variable for each treatment group. SD are in parentheses and N is displayed on the third line. The last three columns report the p-value for a difference of means t-test between each group. * denotes significance at 10 pct., ** at 5 pct., and *** at 1 pct. level.

Table 5: Summary statistics by treatment group

	Mean (SD, N)				Difference p -value			
	Control	Lottery	Regret	Lottery - Control	Regret - Control	Lottery - Regret		
Currently saves	0.56 (0.50) 105	0.61 (0.49) 103	0.47 (0.50) 103	0.47	0.17	0.04**		
Total savings last mo.	58.82 (106.26)	41.01 (59.72)	51.79 (72.56)	0.14	0.58	0.25		
Currently saves with ROSCA	$ \begin{array}{c} 105 \\ 0.58 \\ (0.50) \end{array} $	$ \begin{array}{c} 103 \\ 0.57 \\ (0.50) \end{array} $	$ \begin{array}{c} 103 \\ 0.66 \\ (0.48) \end{array} $	0.91	0.24	0.20		
ROSCA savings last mo.	$ \begin{array}{c} 105 \\ 13.83 \\ (23.24) \end{array} $	$ \begin{array}{c} 103 \\ 15.46 \\ (28.42) \end{array} $	103 15.92 (23.41)	0.65	0.52	0.90		
M-Pesa savings last mo.	105 8.73 (30.53) 105	103 17.24 (87.04) 103	103 5.48 (20.51) 103	0.35	0.37	0.18		

Notes: The first three columns report means of each row variable for each treatment group. SD are in parentheses and N is displayed on the third line. The last three columns report the p-value for a difference of means t-test between each group. * denotes significance at 10 pct., ** at 5 pct., and *** at 1 pct. level.

Table 6: Treatment effects - Mobile savings by respondent

		No contro	ols	7	With cont	rols	Sample	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Lottery	Regret	Difference p -value	Lottery	Regret	$\begin{array}{c} \text{Difference} \\ p\text{-value} \end{array}$	Control Mean (SD)	N
Total no. of deposits	4.59*	5.71**	0.69	4.20*	5.55**	0.63	13.66	311
	(2.52)	(2.45)		(2.51)	(2.44)		(15.08)	
	[0.13]	[0.06]*		[0.17]	[0.06]*			
No. of days saved	3.93*	4.94**	0.66	3.49*	4.76**	0.58	11.78	311
	(2.05)	(2.08)		(2.02)	(2.09)		(12.93)	
	[0.13]	[0.04]**		[0.17]	$[0.06]^*$			
Avg. no. of deposits	0.08*	0.10**	0.69	0.07*	0.09**	0.63	0.23	311
	(0.04)	(0.04)		(0.04)	(0.04)		(0.25)	
	[0.13]	$[0.06]^*$		[0.17]	$[0.06]^*$			
Log total deposit amt.	-0.01	-0.01	0.95	-0.01	-0.00	0.89	0.32	311
	(0.06)	(0.05)		(0.06)	(0.05)		(0.43)	
	[0.81]	[0.84]		[0.82]	[0.94]			
Joint p-value	0.04**	0.00***	0.69	0.08*	0.00***	0.63		

Notes: Columns 1 - 2 report OLS estimates of the treatment effect. Columns 4 - 5 reports the estimates controlling for baseline covariates. Columns 3 and 6 report the p-values for tests of the equality of the two main treatment effects after estimation. Standard errors are in parentheses and FWER adjusted p-values are in brackets. * denotes significance at 10 pct., ** at 5 pct., and *** at 1 pct. level.

Table 7: Treatment effects - Self-reported savings behavior

		No controls			With cont	rols	Sample	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Lottery	Regret	$\begin{array}{c} \text{Difference} \\ p\text{-value} \end{array}$	Lottery	Regret	$\begin{array}{c} \text{Difference} \\ p\text{-value} \end{array}$	Control Mean (SD)	N
Log total savings last mo.	-0.15	-0.05	0.72	-0.14	0.01	0.60	3.80	284
	(0.32)	(0.29)		(0.32)	(0.29)		(2.11)	
	[0.86]	[0.89]		[0.90]	[0.98]			
Log M-Pesa savings last mo.	-0.22	-0.11	0.70	-0.18	-0.11	0.80	1.55	284
	(0.29)	(0.29)		(0.29)	(0.29)		(2.11)	
	[0.82]	[0.89]		[0.90]	[0.91]			
Log ROSCA savings last mo.	0.00	0.63**	0.04**	0.02	0.62**	0.04**	2.10	283
	(0.31)	(0.30)		(0.30)	(0.28)		(2.09)	
	[1.00]	[0.12]		[0.93]	[0.11]			
Currently saves with ROSCA	-0.02	0.14**	0.02**	-0.02	0.13**	0.02**	0.54	284
	(0.07)	(0.07)		(0.07)	(0.07)		(0.50)	
	[0.90]	[0.14]		[0.90]	[0.16]			
Joint p-value	0.85	0.20	0.72	0.79	0.18	0.60		

Notes: Columns 1 - 2 report OLS estimates of the treatment effect. Columns 4 - 5 reports the estimates controlling for baseline covariates. Columns 3 and 6 report the p-values for tests of the equality of the two main treatment effects after estimation. Standard errors are in parentheses and FWER adjusted p-values are in brackets. * denotes significance at 10 pct., ** at 5 pct., and *** at 1 pct. level.

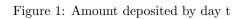
Table 8: Treatment effects - Gambling behavior

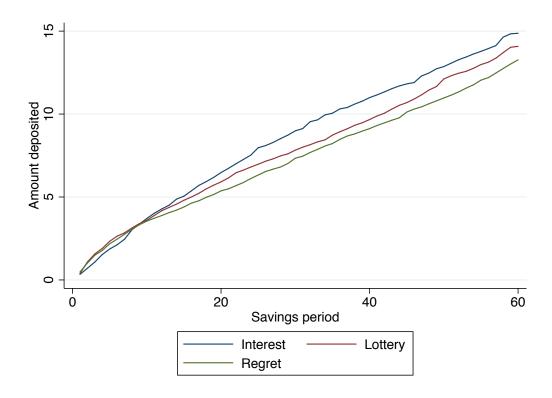
		No contro	ols	,	With cont	rols	Sample	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Lottery	Regret	Difference p -value	Lottery	Regret	Difference p -value	Control Mean (SD)	N
Gamble more	0.06	0.15***	0.16	0.08	0.16***	0.18	0.12	284
	(0.05)	(0.06)		(0.05)	(0.06)		(0.32)	
	[0.65]	$[0.04]^{**}$		[0.54]	$[0.02]^{**}$			
Gamble less	-0.02	0.04	0.24	-0.02	0.04	0.24	0.16	284
	(0.05)	(0.06)		(0.05)	(0.05)		(0.37)	
	[0.89]	[0.77]		[0.89]	[0.66]			
More tempted to gamble	0.09	0.05	0.56	0.08	0.04	0.57	0.47	284
	(0.07)	(0.07)		(0.07)	(0.07)		(0.50)	
	[0.65]	[0.77]		[0.60]	[0.66]			
Less tempted to gamble	-0.01	0.03	0.27	-0.00	0.04	0.30	0.06	284
	(0.03)	(0.04)		(0.03)	(0.04)		(0.25)	
	[0.89]	[0.77]		[0.94]	[0.64]		•	
Joint p-value	0.64	0.05**	0.16	0.59	0.02**	0.18		

Notes: Columns 1 - 2 report OLS estimates of the treatment effect. Columns 4 - 5 reports the estimates controlling for baseline covariates. Columns 3 and 6 report the p-values for tests of the equality of the two main treatment effects after estimation. Standard errors are in parentheses and FWER adjusted p-values are in brackets. * denotes significance at 10 pct., ** at 5 pct., and *** at 1 pct. level.

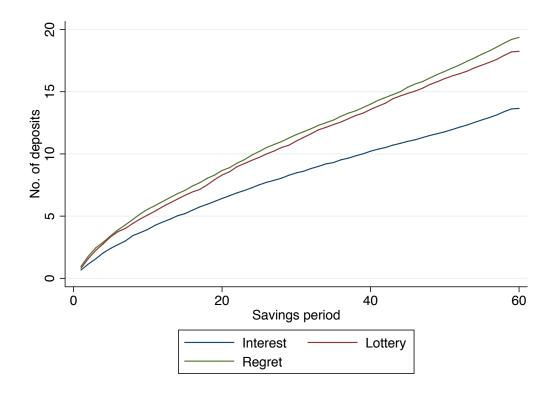
		Dependent var	iables	
	Total no. of deposits	Avg. no. of deposits	No. of days saved	Gamble mor
Cemale .	-			
$\hat{\beta} x_i = 1$	9.17***	0.15***	7.63***	0.11
	(0.00)	(0.00)	(0.00)	(0.00)
$\hat{\beta} x_i = 0$	0.33	0.01	0.67	0.19**
Below 30 y.o.	(3.57)	(0.06)	(3.06)	(0.09)
$\beta x_i = 1$	5.57*	0.09*	4.53*	0.15**
	(0.00)	(0.00)	(0.00)	(0.00)
$\hat{\beta} x_i = 0$	5.52	0.09	4.97	0.13
	(3.79)	(0.06)	(3.32)	(0.09)
Completed std. 8	5 94**	0.10**	5 11**	0.15**
$\hat{\beta} x_i = 1$	(0.00)	(0.00)	(0.00)	(0.00)
$\hat{\beta} x_i = 0$	4.67	0.00)	4.33	-0.00
P 21 - 0	(7.15)	(0.12)	(6.87)	(0.00)
Completed formal 4				
$\hat{\beta} x_i = 1$	4.10	0.07	4.53	0.16**
As a	(0.00)	(0.00)	(0.00)	(0.00)
$\hat{\beta} x_i = 0$	8.30** (3.78)	0.14** (0.06)	6.19* (3.24)	(0.09)
farried/co-habitating	(0.10)	(0.00)	(3.24)	(0.03)
$\hat{\beta} x_i = 1$	3.17	0.05	2.06	0.24***
/- /	(0.00)	(0.00)	(0.00)	(0.00)
$\hat{\beta} x_i = 0$	7.78**	0.13**	7.36**	0.06
7 141	(3.40)	(0.06)	(2.94)	(0.08)
Ias children				0.407
$\hat{\beta} x_i = 1$	6.34** (0.00)	0.11** (0.00)	4.99** (0.00)	(0.00)
$\hat{\beta} x_i = 0$	(0.00)	(0.00)	(0.00)	(0.00)
$\rho \mu_1 = 0$	(4.49)	(0.07)	(3.92)	(0.07)
Currently saves	()	(0.01)	(0-0-2)	(0.01)
$\hat{\beta} x_i = 1$	3.94	0.07	3.61	0.12
	(0.00)	(0.00)	(0.00)	(0.00)
$\hat{\beta} x_i = 0$	8.26**	0.14**	6.98**	0.18**
Above median monthly inc.	(3.23)	(0.05)	(2.71)	(0.07)
$\beta x_i = 1$	5.02	0.08	3.92	0.18**
$\rho x_i = 1$	(0.00)	(0.00)	(0.00)	(0.00)
$\hat{\beta} x_i = 0$	5.99*	0.10*	5.54"	0.09
	(3.43)	(0.06)	(2.88)	(0.07)
Receives regular income				
$\hat{\beta} x_i = 1$	-2.21 (0.00)	-0.04	-0.50 (0.00)	-0.05
$\hat{\beta} x_i = 0$	(0.00)	(0.00)	(0.00)	(0.00)
$\rho x_i = 0$	(4.06)	(0.07)	(3.44)	(0.10)
Employed	(4.00)	(0.01)	(0.44)	(0.10)
$\hat{\beta} x_i = 1$	2.20	0.04	1.74	0.13
	(0.00)	(0.00)	(0.00)	(0.00)
$\hat{\beta} x_i = 0$	9.02***	0.15***	7.96***	0.17***
elf-employed	(3.28)	(0.05)	(2.78)	(0.07)
$\hat{\beta} x_i = 1$	15.19**	0.25**	13.06**	0.19
$\rho x_i = 1$	(0.00)	(0.00)	(0.00)	(0.00)
$\hat{\beta} x_i = 0$	6.95**	0.12**	6.30**	0.14**
	(3.07)	(0.05)	(2.59)	(0.07)
Has dependant				
$\hat{\beta} x_i = 1$	6.51**	0.11**	5.37**	0.17**
4	(0.00)	(0.00)	(0.00)	(0.00)
$\hat{\beta} x_i = 0$	1.21	0.02	2.31	0.06
Subject is a dependant	(4.65)	(0.08)	(4.25)	(0.06)
$\hat{\beta} x_i = 1$	11.38**	0.19**	10.11**	0.22**
7-7-1	(0.00)	(0.00)	(0.00)	(0.00)
$\hat{\beta} x_i = 0$	3.84	0.06	3.24	0.12*
	(2.83)	(0.05)	(2.41)	(0.07)
lisk averse				
$\bar{\beta} x_i = 1$	3.21 (0.00)	(0.05	(0.00)	(0.00)
$\hat{\beta} x_i = 0$	7.83**	0.13**	7.01**	0.15*
$\rho x_i = 0$	(3.50)	(0.06)	(2.92)	(0.08)
Above median LOC	(0.00)	(0.00)	(=-0=)	(0.00)
$\hat{\beta} x_i = 1$	5.03	0.08	4.14	0.19**
	(0.00)	(0.00)	(0.00)	(0.00)
$\hat{\beta} x_i = 0$	6.14*	0.10*	5.44**	0.12
0	(3.15)	(0.05)	(2.68)	(0.07)
Above median i. point	1.77	0.03	1.45	0.10
$\tilde{\beta} x_i = 1$	(0.00)	(0.00)	(0.00)	(0.00)
$\hat{\beta} x_i = 0$	9.75***	0.16***	(0.00) 8.51***	0.19**
	(3.47)	(0.06)	(2.91)	(0.08)
Above median CPGI				()
$\hat{\beta} x_i = 1$	4.38	0.07	4.54	0.18**
	(0.00)	(0.00)	(0.00)	(0.00)
$\hat{\beta} x_i = 0$	6.17*	0.10*	4.78	0.11
	(3.59)	(0.06)	(3.03)	(0.08)

 $\rho|x_i=v \qquad \qquad 6.17 \qquad \qquad 0.10' \qquad 4.78 \qquad 0.11 \\ (5.59) \qquad (0.06) \qquad (0.359) \qquad (0.08) \\ Notes. This table reports heterogeneous treatment effects of regret on each of the obsens variables where such proof represents a dimension of between ground; The finite one code only nearly the treatment coefficient when the backford commary variables <math>x_i=1$ and the second cost is the treatment coefficient when $x_i=0$. Standard errors are in parentheses. A denotes significance at 10 pct., ** at 5 pct., and *** at 1 pct.









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