

Exercise 5

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1 Bias and variance of ridge regression

To show: $E[\hat{\beta}_\tau] = S_\tau^{-1} S \beta^*$

$$\text{Cov}[\hat{\beta}_\tau] = S_\tau^{-1} S S_\tau^{-1} \sigma^2$$

with $S := X^\top X$ and $S_\tau := X^\top X + \tau \mathbb{1}_D$

Proof: solution to ridge regression proven in the lecture

$$\begin{aligned}
 E[\hat{\beta}_\tau] &= E[(X^\top X + \tau \mathbb{1}_D)^{-1} X^\top y] \\
 &= E[(X^\top X + \tau \mathbb{1}_D)^{-1} X^\top (X \beta^* + \varepsilon)] \\
 &= E[(X^\top X + \tau \mathbb{1}_D)^{-1} X^\top X \beta^* + (X^\top X + \tau \mathbb{1}_D)^{-1} X^\top \varepsilon] \\
 &= E[(X^\top X + \tau \mathbb{1}_D)^{-1} X^\top X \beta^*] + E[(X^\top X + \tau \mathbb{1}_D)^{-1} X^\top \varepsilon] \\
 &= (X^\top X + \tau \mathbb{1}_D)^{-1} X^\top X \beta^* + E[(X^\top X + \tau \mathbb{1}_D)^{-1} X^\top \varepsilon]
 \end{aligned}$$

\downarrow

only ε is a random variable, thus

$$\begin{aligned}
 E[(X^\top X + \tau \mathbb{1}_D)^{-1} X^\top X \beta^*] &= (X^\top X + \tau \mathbb{1}_D)^{-1} X^\top X \beta^* \\
 &= (X^\top X + \tau \mathbb{1}_D)^{-1} X^\top X \beta^* + (X^\top X + \tau \mathbb{1}_D)^{-1} X^\top \underbrace{E[\varepsilon]}_{=0} \\
 &= (X^\top X + \tau \mathbb{1}_D)^{-1} X^\top X \beta^* \\
 &= S_\tau^{-1} S \beta^*
 \end{aligned}$$

$$\begin{aligned}
\cdot \text{Cov} [\hat{\beta}_\tau] &= \text{Cov} [(x^\top x + \tau \mathbb{1}_D)^{-1} x^\top (x \beta^* + \varepsilon)] \\
&= \text{Cov} [S_\tau^{-1} S \beta^* + S_\tau^{-1} x^\top \varepsilon] \\
&= \text{Cov} [S_\tau^{-1} x^\top \varepsilon] \quad (\text{Cov}[A y] = A \text{Cov}[y] \text{ for matrix } A \text{ and a random variable } y) \\
&= S_\tau^{-1} x^\top \underbrace{\text{Cov}[\varepsilon]}_{=\sigma^2 \mathbb{1}_N} (S_\tau^{-1} x^\top)^\top \\
&= S_\tau^{-1} x^\top \sigma^2 \mathbb{1}_N x (S_\tau^{-1})^\top \\
&= S_\tau^{-1} \underbrace{x^\top x}_{=S} (S_\tau^{-1})^\top \sigma^2
\end{aligned}$$

$S_\tau^\top = (x^\top x + \tau \mathbb{1}_D)^\top = (x^\top x)^\top + \tau \mathbb{1}_D^\top$
 $= x^\top x + \tau \mathbb{1}_D = S_\tau$

S_τ symmetric and nonsingular $\Rightarrow S_\tau^{-1}$ is also
 symmetric, i.e. $(S_\tau^{-1})^\top = S_\tau^{-1}$

Therefore : $\text{Cov} [\hat{\beta}_\tau] = S_\tau^{-1} S S_\tau^{-1} \sigma^2$

2. LDA - Derivation from the LSE

$$\hat{\beta}_{OLS} = \underset{\beta}{\operatorname{argmin}} \sum_{i=1}^N (\gamma_i^* - x_i \cdot \beta)^2$$

$$\frac{\partial}{\partial \beta} \sum_{i=1}^N (\gamma_i^* - x_i \cdot \beta)^2 = \sum_{i=1}^N \frac{\partial}{\partial \beta} (\gamma_i^* - x_i \cdot \beta)^2 =$$

$$\sum_{i=1}^N 2 (\gamma_i^* - x_i \cdot \beta) (-x_i^\top) = 0$$

$$\Leftrightarrow \sum_{i=1}^N (\gamma_i^* - x_i \cdot \beta) x_i^\top = 0$$

$$\Leftrightarrow \sum_{i=1}^N \gamma_i^* x_i^\top - \sum_{i=1}^N x_i \beta x_i^\top = 0$$

$$\Leftrightarrow \frac{1}{N} \sum_{i=1}^N \gamma_i^* x_i^\top - \frac{1}{N} \sum_{i=1}^N x_i \beta x_i^\top = 0 \quad (*)$$

$$\begin{aligned} \frac{1}{N} \sum_{i=1}^N \gamma_i^* x_i^\top &= \frac{1}{N} \sum_{i: \gamma_i^* = -1} -x_i^\top + \frac{1}{N} \sum_{i: \gamma_i^* = 1} x_i^\top \\ &= \frac{1}{2} (-\mu_{-1}^\top + \mu_1^\top) \\ &= \frac{1}{2} (\mu_1 - \mu_{-1})^\top \end{aligned}$$

$$\frac{1}{N} \sum_{i=1}^N x_i \beta x_i^\top = \frac{1}{N} \sum_{i=1}^N x_i^\top x_i \beta$$

$x_i \beta$ is a scalar

$$= \left[\frac{1}{N} \sum_{i: \gamma_i^* = -1} x_i^\top x_i + \frac{1}{N} \sum_{i: \gamma_i^* = 1} x_i^\top x_i \right] \beta$$

$$= \frac{1}{N} \left\{ \sum_{i: \gamma_i^* = -1} [(x_i - \mu_{-1})^\top (x_i - \mu_{-1}) + x_i^\top \mu_{-1} + \mu_{-1}^\top x_i - \mu_1^\top \mu_{-1}] \right\}$$

$$+ \sum_{i: \gamma_i^* = 1} \left\{ (x_i - \mu_1)^\top (x_i - \mu_1) + x_i^\top \mu_1 + \mu_1^\top x_i - \mu_1^\top \mu_1 \right\} \beta$$

$$\begin{aligned}
&= \frac{1}{N} \left[\sum_{i: y_i^* = -1} (x_i - \mu_{-1})^\top (x_i - \mu_{-1}) + \sum_{i: y_i^* = 1} (x_i - \mu_1)^\top (x_i - \mu_1) \right] \beta + \\
&\quad \left[\frac{1}{N} \sum_{i: y_i^* = -1} x_i^\top \mu_{-1} + \frac{1}{N} \sum_{i: y_i^* = 1} \mu_{-1}^\top x_i - \frac{1}{N} \cdot \frac{N}{2} \mu_{-1}^\top \mu_{-1} \right] \beta + \\
&\quad \left[\frac{1}{N} \sum_{i: y_i^* = 1} x_i^\top \mu_1 + \frac{1}{N} \sum_{i: y_i^* = 1} \mu_1^\top x_i - \frac{1}{N} \cdot \frac{N}{2} \mu_1^\top \mu_1 \right] \beta \\
&= \sum \beta + \left\{ \frac{1}{2} \mu_{-1}^\top \mu_{-1} + \cancel{\frac{1}{2} \mu_{-1}^\top \mu_{-1}} - \cancel{\frac{1}{2} \mu_1^\top \mu_{-1}} \right\} \beta + \\
&\quad \left\{ \cancel{\frac{1}{2} \mu_1^\top \mu_1} + \cancel{\frac{1}{2} \mu_1^\top \mu_1} - \cancel{\frac{1}{2} \mu_1^\top \mu_1} \right\} \beta \\
&= \sum \beta + \left(\frac{1}{2} \mu_{-1}^\top \mu_{-1} + \frac{1}{2} \mu_1^\top \mu_1 \right) \beta \\
&= \sum \beta + \frac{1}{4} (2 \mu_{-1}^\top \mu_{-1} + 2 \mu_1^\top \mu_1) \beta \\
&= \sum \beta + \frac{1}{4} (\mu_{-1}^\top \mu_{-1} + \mu_{-1}^\top \mu_{-1} - \mu_1^\top \mu_1 + \mu_1^\top \mu_1) \beta \\
&\stackrel{\mu_1 = -\mu_{-1}}{=} \sum \beta + \frac{1}{4} (\mu_{-1}^\top \mu_{-1} - \mu_{-1}^\top \mu_1 - \mu_1^\top \mu_{-1} + \mu_1^\top \mu_1) \beta \\
&= \boxed{\sum \beta + \frac{1}{4} (\mu_1 - \mu_{-1})^\top (\mu_1 - \mu_{-1}) \beta}
\end{aligned}$$

Hence, (*) can be rewritten as:

$$\sum \beta + \frac{1}{4} (\mu_1 - \mu_{-1})^\top (\mu_1 - \mu_{-1}) \beta = \frac{1}{2} (\mu_1 - \mu_{-1})^\top$$