

PLSC 503 – Spring 2021

Generalized Linear Models

April 28, 2021

The Exponential Family

$$f(z|\psi) = \Pr(Z = z|\psi)$$

Exponential if:

$$f(z|\psi) = r(z)s(\psi) \exp[q(z)h(\psi)]$$

provided that $r(z) > 0$ and $s(\psi) > 0$.

$$f(z|\psi) = \exp[\underbrace{\ln r(z) + \ln s(\psi)}_{\text{"additive"}} + \underbrace{q(z)h(\psi)}_{\text{"interactive"}}]$$

$$y = q(z)$$

$$\theta = h(\psi)$$

$$f[y|\theta] = \exp[y\theta - b(\theta) + c(y)].$$

- $b(\theta)$ is a “normalizing constant”
- $c(y)$ is a function solely of y
- $y\theta$ is a multiplicative term

A Familiar Family Member: Poisson

$$f(y|\lambda) = \frac{\exp(-\lambda)\lambda^y}{y!}.$$

$$\begin{aligned} f(y|\lambda) &= \exp \{ \ln [\exp(-\lambda)\lambda^y / y!] \} \\ &= \exp \left[\underbrace{y \ln(\lambda)}_{y\theta} - \underbrace{\lambda}_{b(\theta)} - \underbrace{\ln(y!)}_{c(y)} \right] \end{aligned}$$

$$f(y|\theta, \phi) = \exp \left[\frac{y\theta - b(\theta)}{a(\phi)} + c(y, \phi) \right]$$

Familiar Family Member II: Normal

$$f(y|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left[\frac{(y - \mu)^2}{2\sigma^2} \right]$$

$$\begin{aligned} f(y|\mu, \sigma^2) &= \exp \left[-\frac{1}{2} \ln(2\pi\sigma^2) - \frac{1}{2\sigma^2}(y^2 - 2y\mu + \mu^2) \right] \\ &= \exp \left[-\frac{1}{2} \ln(2\pi\sigma^2) - \frac{1}{2\sigma^2}y^2 + \frac{1}{\sigma^2}y\mu - \frac{1}{2\sigma^2}\mu^2 \right] \\ &= \exp \left[\frac{y\mu}{\sigma^2} - \frac{\mu^2}{2\sigma^2} - \frac{y^2}{2\sigma^2} - \frac{1}{2} \ln(2\pi\sigma^2) \right] \\ &= \exp \left\{ \frac{y\mu - \frac{\mu^2}{2}}{\sigma^2} + \frac{-1}{2} \left[\frac{y^2}{\sigma^2} + \ln(2\pi\sigma^2) \right] \right\} \end{aligned}$$

$$f(y|\mu, \sigma^2) = \exp \left\{ \frac{y\mu - \frac{\mu^2}{2}}{\sigma^2} + \frac{-1}{2} \left[\frac{y^2}{\sigma^2} + \ln(2\pi\sigma^2) \right] \right\}$$

$\theta = \mu$, so:

- $y\theta = y\mu$
- $b(\theta) = \frac{\mu^2}{2}$
- $a(\phi) = \sigma^2$
- $c(y, \phi) = \frac{-1}{2} \left[\frac{y^2}{\sigma^2} + \ln(2\pi\sigma^2) \right]$

Other Family Members

- Binomial (\supset Bernoulli; also Multinomial)
- Exponential
- Gamma
- Logarithmic
- Inverse Gaussian
- Negative Binomial
- others...

$$\begin{aligned}\ln L(\theta, \phi|y) &= \ln f(y|\theta, \phi) \\ &= \ln \left\{ \exp \left[\frac{y\theta - b(\theta)}{a(\phi)} + c(y, \phi) \right] \right\} \\ &= \frac{y\theta - b(\theta)}{a(\phi)} + c(y, \phi)\end{aligned}$$

$$\begin{aligned}\frac{\partial \ln L(\theta, \phi|y)}{\partial \theta} &\equiv \mathbf{S} = \frac{\partial}{\partial \theta} \left[\frac{y\theta - b(\theta)}{a(\phi)} + c(y, \phi) \right] \\ &= \frac{y - \frac{\partial}{\partial \theta} b(\theta)}{a(\phi)}.\end{aligned}$$

Among family members:

- \mathbf{S} is a sufficient statistic for θ .
- $E(\mathbf{S}) = 0$.
- $\text{Var}(\mathbf{S}) \equiv \mathcal{I}(\theta) = E[(\mathbf{S})^2|\theta]$

$$E(Y) = \frac{\partial}{\partial \theta} b(\theta)$$

and

$$\text{Var}(Y) = a(\phi) \frac{\partial^2}{\partial \theta^2} b(\theta)$$

Example: Poisson Again

$$\begin{aligned} E(Y) &= \frac{\partial}{\partial \theta} \exp(\theta) \\ &= \exp(\theta)|_{\theta=\ln(\lambda)} \\ &= \lambda \end{aligned}$$

$$\begin{aligned} \text{Var}(Y) &= 1 \times \frac{\partial^2}{\partial \theta^2} \exp(\theta)|_{\theta=\ln(\lambda)} \\ &= \exp[\ln(\lambda)] \\ &= \lambda \end{aligned}$$

Example: Normal Again

$$\begin{aligned} E(Y) &= \frac{\partial}{\partial \theta} \left(\frac{\theta^2}{2} \right) \\ &= \theta|_{\theta=\mu} \\ &= \mu \end{aligned}$$

$$\begin{aligned} \text{Var}(Y) &= \sigma^2 \times \frac{\partial^2}{\partial \theta^2} \left(\frac{\theta^2}{2} \right) \\ &= \sigma^2 \times \frac{\partial}{\partial \theta} \theta \\ &= \sigma^2 \end{aligned}$$

$$Y_i = \mathbf{X}_i\boldsymbol{\beta} + u_i$$

$$E(Y_i) \equiv \boldsymbol{\mu}_i = \mathbf{X}_i\boldsymbol{\beta}$$

The “Generalized” Part

$$g(\mu_i) = \mathbf{X}_i\beta.$$

$$\begin{aligned}\eta_i &= \mathbf{X}_i\beta \\ &= g(\mu_i)\end{aligned}$$

$$\begin{aligned}\mu_i &= g^{-1}(\eta_i) \\ &= g^{-1}(\mathbf{X}_i\beta)\end{aligned}$$

Random component $\sim \text{EF}(\cdot)$ with

$$E(Y_i) = \mu_i.$$

Systematic component:

$$g(\mu_i) = \eta_i$$

“Link function”:

$$g(\mu_i) = \eta_i$$

or

$$g^{-1}(\eta_i) = \mu_i.$$

The Return of The Family

$$\begin{aligned}\theta_i &= g(\mu_i) \\ &= \eta_i \\ &= \mathbf{X}_i\beta\end{aligned}$$

$$g^{-1}(\theta_i) = \eta_i$$

GLM Example: Linear-Normal

$$f(y|\mu, \sigma^2) = \mathcal{N}(\mu, \sigma^2)$$

$$\mu_i = \eta_i$$

$$\begin{aligned} \mu_i \equiv \theta_i &= \eta_i \\ Y_i &\sim \mathcal{N}(\mu_i, \sigma^2) \end{aligned}$$

$$f(y|\pi) = \pi^y(1 - \pi)^{1-y}$$

$$\theta_i = \ln \left(\frac{\mu_i}{1 - \mu_i} \right)$$

$$\begin{aligned}\mu_i &= g^{-1}(\theta_i) \\ &= \frac{\exp(\eta_i)}{1 + \exp(\eta_i)} \\ Y_i &\sim \text{Bernoulli}(\mu_i)\end{aligned}$$

GLM Example: Counts (Independent Events)

$$f(y|\lambda) = \frac{\exp(-\lambda)\lambda^y}{y!}$$

$$\ln(\lambda_i) = \boldsymbol{\eta}_i$$

$$\boldsymbol{\mu}_i = \boldsymbol{g}^{-1}(\boldsymbol{\theta}_i)$$

$$= \exp(\boldsymbol{\eta}_i)$$

$$Y_i \sim \text{Poisson}(\lambda_i)$$

Common GLM Flavors

Distribution	Range of Y	Link(s) $g(\cdot)$	Inverse Link $g^{-1}(\cdot)$
Normal	$(-\infty, \infty)$	Identity: $\theta = \mu$ (Canonical)	θ
Binomial	$\{0, \dots, n\}$	Logit: $\theta = \ln\left(\frac{\mu}{1-\mu}\right)$ (Canonical) Probit: $\theta = \Phi^{-1}(\mu)$ C-Log-Log: $\theta = \ln[-\ln(1 - \mu)]$	$\frac{\exp(\theta)}{1 + \exp(\theta)}$ $\Phi(\theta)$ $1 - \exp[-\exp(\theta)]$
Bernoulli	$\{0, 1\}$	(same as Binomial)	(same as Binomial)
Multinomial	$\{0, \dots, J\}$	(same as Binomial)	(same as Binomial)
Poisson	$[0, \infty]$ (integers)	Log: $\theta = \ln(\mu)$ (Canonical)	$\exp(\theta)$
Gamma	$(0, \infty)$	Reciprocal: $\theta = -\frac{1}{\mu}$ (Canonical)	$-\frac{1}{\theta}$

Note: The Bernoulli is a special case of the Binomial with $n = 1$. The multinomial is the J -outcome variant of the Binomial, and is also related to the Poisson (see, e.g., Agresti 2002).

- Pick $f(Y)$
- Pick $g(\cdot)$
- Specify \mathbf{X}
- Estimate

- MLE
- IRLS (\approx MLE):

$$\hat{\beta}^{(t+1)} = [\mathbf{X}'\mathbf{W}^{(t)}\mathbf{X}]^{-1}\mathbf{X}'\mathbf{W}^{(t)}\mathbf{z}^{(t)}$$

with

$$\mathbf{W}_{N \times N}^{(t)} = \text{diag} \left[\frac{\left(\partial \mu_i^{(t)} / \partial \eta_i^{(t)} \right)^2}{\text{Var}(Y_i)} \right]$$

and

$$\mathbf{z}^{(t)} = \boldsymbol{\eta}^{(t)} + (Y - \boldsymbol{\mu}^{(t)}) \left(\frac{\partial \boldsymbol{\eta}^{(t)}}{\partial \boldsymbol{\mu}} \right).$$

At iteration t :

1. Calculate $\mathbf{z}^{(t)}$, $\mathbf{W}^{(t)}$
2. Regress $\mathbf{z}^{(t)}$ on \mathbf{X} , using $\mathbf{W}^{(t)}$ as weights, to obtain $\hat{\beta}^{(t+1)}$
3. Generate $\boldsymbol{\eta}^{(t+1)} = \mathbf{X}\hat{\beta}^{(t+1)}$
4. Generate $\boldsymbol{\mu}^{(t+1)} = \mathbf{g}^{-1}(\boldsymbol{\eta}^{(t+1)})$
5. Use $\boldsymbol{\eta}^{(t+1)}$ and $\boldsymbol{\mu}^{(t+1)}$ to calculate $\mathbf{z}^{(t+1)}$ and $\mathbf{W}^{(t+1)}$
6. Repeat until convergence.

“Response” Residuals:

$$\begin{aligned}\hat{u}_i &= Y_i - \hat{\mu}_i \\ &= Y_i - g^{-1}(\mathbf{x}_i\hat{\beta})\end{aligned}$$

“Pearson” Residuals:

$$\hat{P}_i = \frac{\hat{u}_i}{[\text{Var}(\hat{u}_i)]^{1/2}}$$

“Deviance” :

$$\begin{aligned}\hat{d}_i &= -2[\ln L_i(\hat{\theta}) - \ln L_i(\theta_S)] \\ &= 2 \left\{ \left[\frac{Y_i \theta_S - b(\theta_S)}{a(\phi)} + c(Y_i, \phi) \right] - \left[\frac{Y_i \hat{\theta} - b(\hat{\theta})}{a(\phi)} + c(Y_i, \phi) \right] \right\} \\ &= 2 \left[\frac{Y_i(\theta_S - \hat{\theta}) - b(\theta_S) + b(\hat{\theta})}{a(\phi)} \right]\end{aligned}$$

“Deviance” Residuals:

$$\hat{r}_{Di} = \left(\frac{\hat{u}_i}{|\hat{u}_i|} \right) \sqrt{\hat{d}_i^2}$$

Toy Example: Linear-Normal

$$X = \{1, 1, 2, 2, 3, 3, 4, 4, 5, 5\}$$

$$Y = \{0, 2, 1, 3, 2, 4, 3, 5, 4, 6\}$$

$$Y_i = 0 + 1X_i + u_i$$

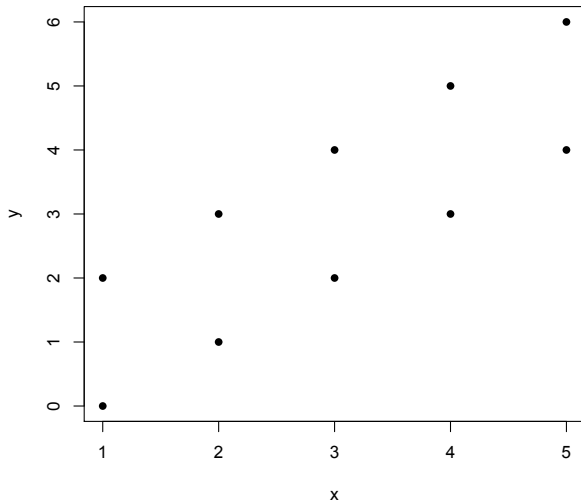
$$\hat{u}_i^2 = 1 \quad \forall i$$

$$\text{"TSS"} \equiv \sum (Y_i - \bar{Y})^2 = 30$$

$$\text{"RSS"} \equiv \sum \hat{u}_i^2 = 10$$

$$\text{"MSS"} / \text{"ESS"} = 20$$

Toy Example: Plot



Toy Example: OLS

```
> linmod<-lm(y~x)
> summary(linmod)
```

Call:

```
lm(formula = y ~ x)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-1.000e+00	-1.000e+00	1.110e-16	1.000e+00	1.000e+00

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-5.617e-16	8.292e-01	-6.77e-16	1.00000
x	1.000e+00	2.500e-01	4	0.00395 **

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.118 on 8 degrees of freedom

Multiple R-squared: 0.6667, Adjusted R-squared: 0.625

F-statistic: 16 on 1 and 8 DF, p-value: 0.00395

Toy Example: Linear-Normal GLM

```
> linalgm<-glm(y~x,family="gaussian")
> summary(linalgm)
```

Deviance Residuals:

	Min	1Q	Median	3Q	Max
	-1.000e+00	-1.000e+00	-5.551e-17	1.000e+00	1.000e+00

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-5.617e-16	8.292e-01	-6.77e-16	1.00000
x	1.000e+00	2.500e-01	4	0.00395 **

Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1 1

(Dispersion parameter for gaussian family taken to be 1.25)

Null deviance: 30 on 9 degrees of freedom
Residual deviance: 10 on 8 degrees of freedom
AIC: 34.379

Number of Fisher Scoring iterations: 2

Better GLM Example: Political Knowledge

- 2008 NES political knowledge
- Identify Speaker of the House, VP, British PM, and Chief Justice
- Y_i = number of correct answers (out of four)

$$f(Y_i, p_i) = \binom{4}{Y_i} p_i^{Y_i} (1 - p_i)^{4 - Y_i}$$

$$Y \sim \text{Binomial}(4, p)$$
$$E(Y_i) = \frac{\exp(\mathbf{X}_i \beta)}{1 + \exp(\mathbf{X}_i \beta)}$$

GLM Example Data (2008 NES)

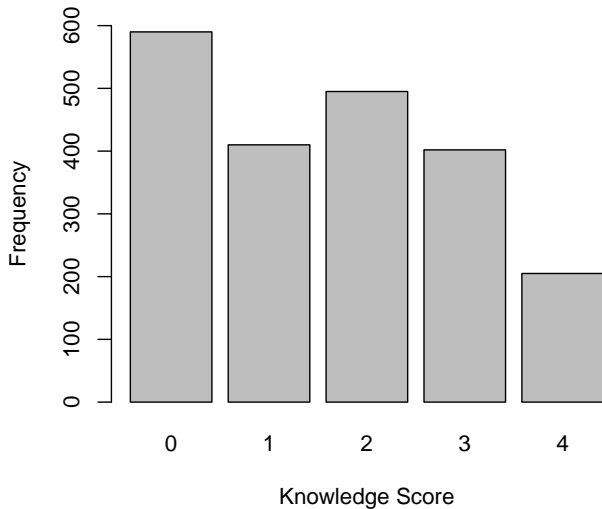
```
> summary(NES08[,4:16])
```

knowledge		sex		race	
Min.	:0.00	1. Male respondent selected	: 999	1. White	:1442
1st Qu.:	1.00	2. Female respondent selected:	1324	2. Black/African-American:	583
Median	:2.00			4. Other race	: 262
Mean	:2.37			5. White and another race:	16
3rd Qu.:	4.00			6. Black and another race:	6
Max.	:4.00			(Other)	: 2
NA's	:221			NA's	: 12

age		female		white		oftenvote		conservative	
Min.	:17	Min.	:0.00	Min.	:0.0000	Seldom	:621	Min.	:1.00
1st Qu.:	33	1st Qu.:	0.00	1st Qu.:	0.0000	Part of the Time:	287	1st Qu.:	3.00
Median	:46	Median	:1.00	Median	:1.0000	Nearly Always	:612	Median	:4.00
Mean	:47	Mean	:0.57	Mean	:0.6207	Always	:788	Mean	:4.14
3rd Qu.:	59	3rd Qu.:	1.00	3rd Qu.:	1.0000	NA's	: 15	3rd Qu.:	5.00
Max.	:90	Max.	:1.00	Max.	:1.0000			Max.	:7.00
NA's	:22							NA's	:697

prayfreq		heterosexual		married		yrssofschool		income	
Never	:235	Min.	:0.0000	Min.	:0.0000	Min.	: 0.00	Min.	: 1.00
Once/week	:321	1st Qu.:	1.0000	1st Qu.:	0.0000	1st Qu.:	12.00	1st Qu.:	5.00
Few times a week:	416	Median	:1.0000	Median	:0.0000	Median	:13.00	Median	:11.00
Daily	:525	Mean	:0.9591	Mean	:0.4224	Mean	:13.08	Mean	:10.52
Several/Day	:806	3rd Qu.:	1.0000	3rd Qu.:	1.0000	3rd Qu.:	15.00	3rd Qu.:	15.00
NA's	: 20	Max.	:1.0000	Max.	:1.0000	Max.	:17.00	Max.	:25.00
		NA's	:49	NA's	:15	NA's	:11	NA's	:151

Political Knowledge (2008 NES)



GLM Results

```
> nes08.binom<-glm(cbind(knowledge,4-knowledge)~age+female+white+oftenvote+conservative
+prayfreq+heterosexual+married+yrsofschool+income,data=nes2008,family=binomial)
> summary(nes08.binom)
```

Deviance Residuals:

	Min	1Q	Median	3Q	Max
	-3.59683	-1.01716	0.03124	1.34899	2.85336

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	2.097696	0.248976	8.425	< 2e-16 ***
age	-0.010789	0.001910	-5.650	1.60e-08 ***
female	0.213865	0.059534	3.592	0.000328 ***
white	-0.154109	0.064613	-2.385	0.017073 *
oftenvote	-0.097272	0.027511	-3.536	0.000407 ***
conservative	0.019421	0.019317	1.005	0.314704
prayfreq	0.048818	0.022248	2.194	0.028216 *
heterosexual	0.070894	0.138471	0.512	0.608665
married	-0.166501	0.058363	-2.853	0.004333 **
yrsofschool	-0.090790	0.013116	-6.922	4.45e-12 ***
income	-0.009015	0.005259	-1.714	0.086492 .

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 3181.4 on 1359 degrees of freedom
Residual deviance: 2952.9 on 1349 degrees of freedom
(963 observations deleted due to missingness)
AIC: 4563.1
Number of Fisher Scoring iterations: 4

GLMs: Other Topics + Extensions

Other Topics:

- Generalizations for Overdispersion (binomial)
- Diagnostics (leverage, etc.)
- Joint Mean-Dispersion Models

Extensions:

- Bias-reduced models (a la Firth 1993)
- “Generalized additive models” (GAMs)
- “Generalized estimating equations” (GEEs)
- “Vector” GLMs (Yee and Wild 1996; Yee and Hastie 2003)

McCullagh, P., and J. A. Nelder. 1989. *Generalized Linear Models*, 2nd Ed. London: Chapman & Hall.

Dobson, Annette J., and and Adrian G. Barnett. 2008. *An Introduction to Generalized Linear Models*, 3rd Ed. London: Chapman & Hall.

Faraway, Julian J. 2006. *Extending the Linear Model with R: Generalized Linear, Mixed Effects, and Nonparametric Regression Models*. London: Chapman & Hall / CRC.

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Hardin, James W., and Joseph W. Hilbe. 2012. *Generalized Linear Models and Extensions*, 3rd Ed. College Station, TX: Stata Press.