# HW4 - Lucas Fellmeth, Sven Bergmann

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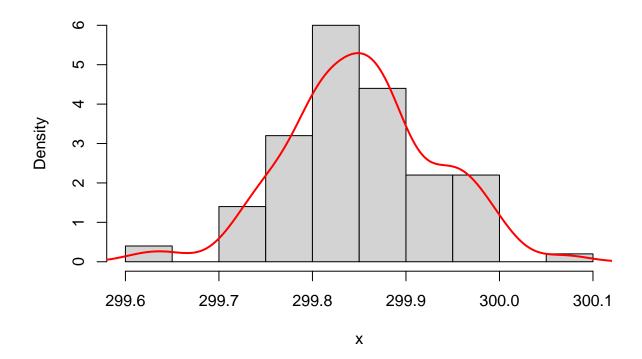
## Chapter 11

### Exercise 5

Recall Exercise 6.3 based on 100 measurements of the speed of light in air. In that chapter, we tested the data for normality. Use the same data to construct a density estimator that you feel gives the best visual display of the information provided by the data. What parameters did you choose? The data can be downloaded from http://www.itl.nist.gov/div898/strd/univ/data/Michelso.dat

```
x <- as.numeric(read.delim2("Michelso.dat.txt")[[1]])
hist(x, freq = F, nclass = 10)
kde <- density(x, bw = "SJ")
lines(kde$x, kde$y, col = "red", lwd = 2)</pre>
```

# Histogram of x



We chose 'nclass=10' and 'bw="SJ"' for the kernel density estimator. 'SJ' chooses the bandwidth automatically, we felt like this gives a good visual display of the data.

### Chapter 12

### Exercise 1

Using robust regression, find the intercept and slope  $\hat{\beta}_0$  and  $\hat{\beta}_1$  for each of the four data sets of Anscombe (1973) from p. 244. Plot the ordinary least-squares regression along with the rank regression estimator of slope. Contrast these with one of the other robust regression techniques. For which set does  $\tilde{\beta}_1$  differ the most from its LS counterpart  $\hat{\beta}_1 = 0.5$ ? Note that in the fourth set, 10 out of 11 Xs are equal, so one should use

$$S_{ij} = (Y_j - Y_i)/(X_j - X_i + \epsilon)$$

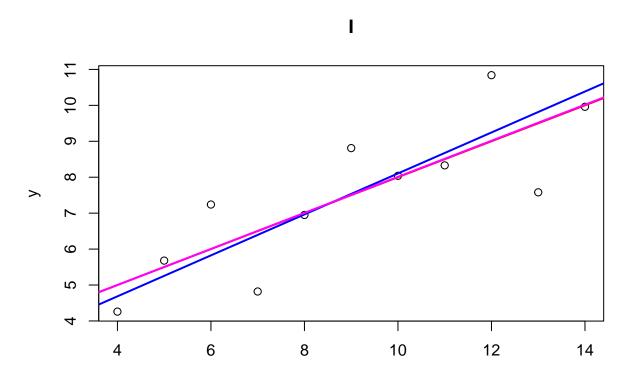
to avoid dividing by 0. After finding  $\tilde{\beta}_0$  and  $\tilde{\beta}_1$ , are they different than  $\hat{\beta}_0$  and  $\hat{\beta}_1$ ? Is the hypothesis  $H_0: \beta_1 = 1/2$  rejected in a robust test against the alternative  $H_1: \beta_1 < 1/2$ , for data set 3? Note here  $\beta_{10} = 1/2$ .

```
library(robustbase)
library(MASS)
library(L1pack)
```

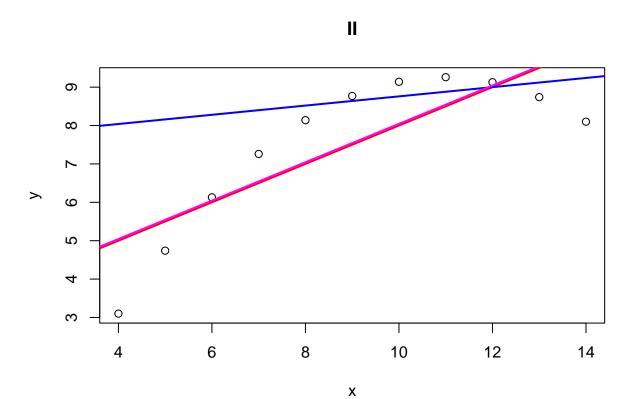
## Loading required package: fastmatrix

```
data <- read.csv("anscombe.csv")
data$dataset <- factor(data$dataset)
data1 <- subset(data, subset = data$dataset == "I")
data2 <- subset(data, subset = data$dataset == "II")
data3 <- subset(data, subset = data$dataset == "III")
data4 <- subset(data, subset = data$dataset == "IV")
datasets <- list(data1, data2, data3, data4)</pre>
```

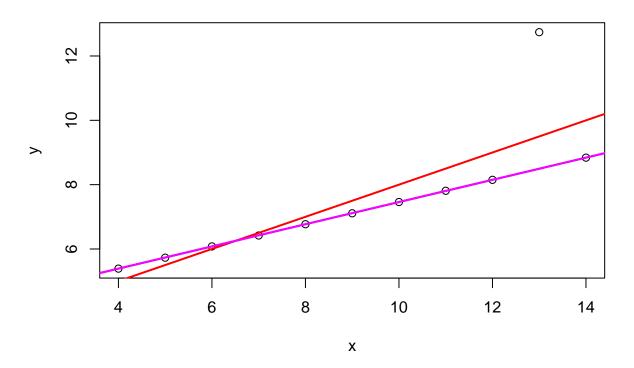
```
for (dataset in datasets) {
    plot(y ~ x, data = dataset, main = dataset[[1]])
    lmfit <- lm(y ~ x, data = dataset)
    abline(coef = coefficients(lmfit), col = "red", lwd = 2)
    lmedfit <- lmsreg(y ~ x, data = dataset)
    abline(coef = coefficients(lmedfit), col = "blue", lwd = 2)
    robfit <- rlm(y ~ x, data = dataset, method = "MM", psi = psi.bisquare, init = coefficients(lmedfit abline(coef = coefficients(robfit), col = "magenta", lwd = 2)
    cat("Coefficients for dataset", dataset[[1]][1], ":\n")
    print(coefficients(lmfit))
    print(coefficients(lmedfit))
    print(coefficients(robfit))
}</pre>
```



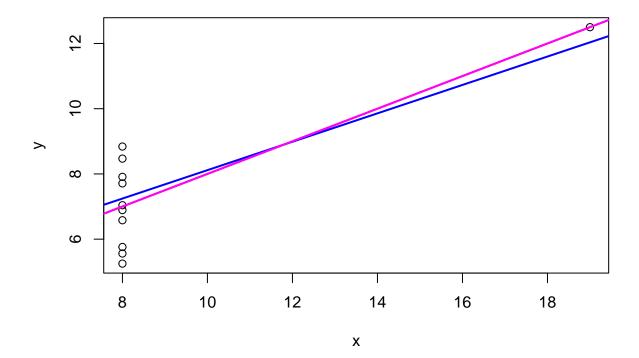
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For dataset 1  $\tilde{\beta}_1$  differs the at most 0.070 from the LS counterpart  $\hat{\beta}_1 = 0.5$ . This difference is attained with the least median of squares estimator.

For dataset 2  $\tilde{\beta}_1$  differs at most 0.38 from the LS counterpart  $\hat{\beta}_1 = 0.5$ . This difference is attained with the least median of squares estimator.

For dataset 3  $\tilde{\beta}_1$  differs at most 0.155 from the LS counterpart  $\hat{\beta}_1 = 0.5$ . This difference is attained with the least median of squares estimator & the M estimator using Tukey's bisquare function.

For dataset 4  $\tilde{\beta}_1$  differs at most 0.10 from the LS counterpart  $\hat{\beta}_1 = 0.5$ . This difference is attained with the least median of squares estimator.

Dataset 3 has the highest difference between  $\tilde{\beta}_1$  and  $\hat{\beta}_1 = 0.5$  with a difference of 0.155.

 $\tilde{\beta}_0$  and  $\tilde{\beta}_1$  are different than  $\hat{\beta}_0$  and  $\hat{\beta}_1$  for two of the three estimators in dataset 3.

Is the hypothesis  $H_0: \beta_1 = 1/2$  rejected in a robust test against the alternative  $H_1: \beta_1 < 1/2$ , for data set 3?

```
n <- length(data3$x)
x <- data3$x
y <- data3$y
beta_1 <- 0.5
beta_10 <- 0.5
U <- y - beta_10 * x
rho_hat <- beta_1 * sqrt(sum(rank(x)^2 - mean(rank(x))))/sqrt(sum(rank(U)^2 - mean(rank(U))))
p <- 1 - pnorm(rho_hat * sqrt(n - 1))
p</pre>
```

#### ## [1] 0.05692315

We accept the hypothesis  $H_0: \beta_1 = 1/2$  since p > 0.05