

Assignment_2

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2026-02-04

1. Problem to demonstrate that the population regression line is fixed, but least square regression line varies

Suppose the population regression line is given by $Y = 2 + 3x$, while the data comes from the model $y = 2 + 3x + \epsilon$. Step 1: For x in the range [5,10] graph the population regression line. Step 2: Generate $x_i(i = 1, 2, \dots, n)$ from Uniform(5, 10) and $\epsilon_i(i = 1, 2, \dots, n)$ from $N(0, 4^2)$. Hence, compute y_1, y_2, \dots, y_n . Step 3: On the basis of the data $(x_i, y_i)(i = 1, 2, \dots, n)$ generated in Step 2, report the least squares regression line. Step 4: Repeat steps 2-3 five times. Graph the 5 least squares regression lines over the population regression line obtained in Step 1. Interpret the findings. Take $n = 50$. Set the seed as seed=123.

```
set.seed(123)
x=seq(5,10,0.025)
y=2+3*x
y

## [1] 17.000 17.075 17.150 17.225 17.300 17.375 17.450 17.525 17.600
17.675
## [11] 17.750 17.825 17.900 17.975 18.050 18.125 18.200 18.275 18.350
18.425
## [21] 18.500 18.575 18.650 18.725 18.800 18.875 18.950 19.025 19.100
19.175
## [31] 19.250 19.325 19.400 19.475 19.550 19.625 19.700 19.775 19.850
19.925
## [41] 20.000 20.075 20.150 20.225 20.300 20.375 20.450 20.525 20.600
20.675
## [51] 20.750 20.825 20.900 20.975 21.050 21.125 21.200 21.275 21.350
21.425
## [61] 21.500 21.575 21.650 21.725 21.800 21.875 21.950 22.025 22.100
22.175
## [71] 22.250 22.325 22.400 22.475 22.550 22.625 22.700 22.775 22.850
22.925
## [81] 23.000 23.075 23.150 23.225 23.300 23.375 23.450 23.525 23.600
23.675
## [91] 23.750 23.825 23.900 23.975 24.050 24.125 24.200 24.275 24.350
24.425
## [101] 24.500 24.575 24.650 24.725 24.800 24.875 24.950 25.025 25.100
25.175
## [111] 25.250 25.325 25.400 25.475 25.550 25.625 25.700 25.775 25.850
25.925
## [121] 26.000 26.075 26.150 26.225 26.300 26.375 26.450 26.525 26.600
26.675
```

```

## [131] 26.750 26.825 26.900 26.975 27.050 27.125 27.200 27.275 27.350
27.425
## [141] 27.500 27.575 27.650 27.725 27.800 27.875 27.950 28.025 28.100
28.175
## [151] 28.250 28.325 28.400 28.475 28.550 28.625 28.700 28.775 28.850
28.925
## [161] 29.000 29.075 29.150 29.225 29.300 29.375 29.450 29.525 29.600
29.675
## [171] 29.750 29.825 29.900 29.975 30.050 30.125 30.200 30.275 30.350
30.425
## [181] 30.500 30.575 30.650 30.725 30.800 30.875 30.950 31.025 31.100
31.175
## [191] 31.250 31.325 31.400 31.475 31.550 31.625 31.700 31.775 31.850
31.925
## [201] 32.000

#1
plot(x,y,type='l')
#2
x_i=runif(50,5,10)
x_i

## [1] 6.437888 8.941526 7.044885 9.415087 9.702336 5.227782 7.640527
9.462095
## [9] 7.757175 7.283074 9.784167 7.266671 8.387853 7.863167 5.514623
9.499125
## [17] 6.230439 5.210298 6.639604 9.772518 9.447697 8.464017 8.202534
9.971349
## [25] 8.278529 8.542652 7.720330 7.970710 6.445799 5.735568 9.815121
9.511495
## [33] 8.453526 8.977337 5.123068 7.388980 8.792298 6.082040 6.590905
6.158129
## [41] 5.714000 7.072732 7.068622 6.844227 5.762224 5.694030 6.165170
7.329812
## [49] 6.329863 9.289139

ei=rnorm(50,0,4)
ei

## [1] -6.7467732 3.3511482 0.6134925 -4.5525477 5.0152597 1.7058569
## [7] -1.1802859 3.5805026 3.5125340 3.2863243 2.7545610 2.2156706
## [13] -0.2476468 -1.2238507 -1.5218840 -2.7788279 -0.8316691 -5.0615854
## [19] 8.6758239 4.8318480 -4.4924343 -1.6115393 -1.8666214 3.1198605
## [25] -0.3334763 1.0132741 -0.1141870 -0.1714818 5.4744091 -0.9030839
## [31] 6.0658824 -6.1950112 2.3384550 0.4954170 0.8637663 1.5185579
## [37] -2.0092938 -1.3328295 -4.0743015 -4.2871649 1.2141146 1.7928391
## [43] 0.2120169 3.6890699 8.2003387 -1.9641247 -9.2366755 4.0229541
## [49] -2.8368031 -2.7520345

y_i=2+3*x_i+ei
y_i

```

```

## [1] 14.56689 32.17573 23.74815 25.69271 36.12227 19.38920 23.74130
33.96679
## [9] 28.78406 27.13555 34.10706 26.01568 26.91591 24.36565 17.02199
27.71855
## [17] 19.85965 12.56931 30.59463 36.14940 25.85066 25.78051 24.74098
35.03391
## [25] 26.50211 28.64123 25.04680 25.74065 26.81181 18.30362 37.51125
24.33947
## [33] 29.69903 29.42743 18.23297 25.68550 26.36760 18.91329 17.69841
16.18722
## [41] 20.35611 25.01103 23.41788 26.22175 27.48701 17.11797 11.25884
28.01239
## [49] 18.15279 27.11538

points(x_i,y_i,col='red')
#3

linreg=lm(y_i~x_i)
abline(linreg,col='blue')
coef(linreg)

## (Intercept)          x_i
## -0.09638929  3.30539569

summary(linreg)

##
## Call:
## lm(formula = y_i ~ x_i)
##
## Residuals:
##    Min     1Q   Median     3Q    Max 
## -9.0231 -2.2314 -0.2627  2.1970  8.7445 
## 
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)    
## (Intercept) -0.09639   2.82610  -0.034   0.973    
## x_i         3.30540   0.36519   9.051 5.96e-12 ***
## ---        
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## 
## Residual standard error: 3.761 on 48 degrees of freedom
## Multiple R-squared:  0.6306, Adjusted R-squared:  0.6229 
## F-statistic: 81.93 on 1 and 48 DF,  p-value: 5.962e-12

#4

## beta not is 2 however the intercept is -0.0963... which are not similar

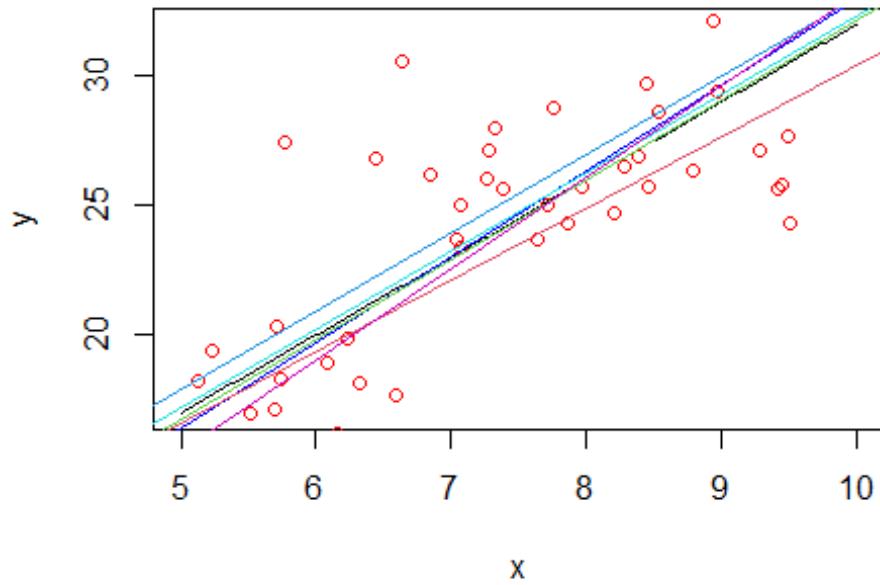
beta_hat=c()

```

```

for(i in 1:5){
  x_i=rnorm(50,5,10)
  ei=rnorm(50,0,4)
  y_i=2+3*x_i+ei
  linreg=lm(y_i~x_i)
  beta_hat=cbind(beta_hat,coef(linreg))
  abline(linreg,col=i+1)
  coef(linreg)
}

```



```

beta.hat=as.data.frame(beta_hat)
beta.hat

##           V1          V2          V3          V4          V5
## (Intercept) 2.792188 1.392997 2.823089 2.032506 -2.107763
## x_i         2.761042 3.073267 3.023608 3.028097 3.530691

```

2. Problem to demonstrate that $\hat{\beta}_0$ and $\hat{\beta}$ minimises RSS

Step 1: Generate x_i from Uniform(5, 10) and mean centre the values. Generate ϵ_i from $N(0, 1)$. Calculate $y_i = 2+3x_i+\epsilon_i$, $i = 1, 2, \dots, n$. Take $n=50$ and seed=123. Step 2: Now imagine that you only have the data on (x_i, y_i) , $i = 1, 2, \dots, n$, without knowing the mechanism that was used to generate the data in step 1. Assuming a linear regression of the type $y_i = \beta_0 + \beta x_i + \epsilon_i$, and based on these data (x_i, y_i) , $i = 1, 2, \dots, n$, obtain the least squares estimates of β_0 and β . Step 3: Take a large number of grid values of (β_0, β) that also include the least squares estimates obtained from step 2. Compute the RSS for each parametric choice of (β_0, β) , where $RSS = (y_1 - \beta_0 - \beta x_1)^2 + (y_2 - \beta_0 - \beta x_2)^2 + \dots + (y_n - \beta_0 - \beta x_n)^2$.

$\beta x_2)^2 + \dots (y_n - \beta_0 - \beta x_n)^2$. Find out for which combination of (β_0, β) , RSS is minimum.

```

set.seed(123)
x_ii=runif(50,5,10)
x_ii

## [1] 6.437888 8.941526 7.044885 9.415087 9.702336 5.227782 7.640527
9.462095
## [9] 7.757175 7.283074 9.784167 7.266671 8.387853 7.863167 5.514623
9.499125
## [17] 6.230439 5.210298 6.639604 9.772518 9.447697 8.464017 8.202534
9.971349
## [25] 8.278529 8.542652 7.720330 7.970710 6.445799 5.735568 9.815121
9.511495
## [33] 8.453526 8.977337 5.123068 7.388980 8.792298 6.082040 6.590905
6.158129
## [41] 5.714000 7.072732 7.068622 6.844227 5.762224 5.694030 6.165170
7.329812
## [49] 6.329863 9.289139

eii=rnorm(50,0,1)
eii

## [1] -1.68669331  0.83778704  0.15337312 -1.13813694  1.25381492
0.42646422
## [7] -0.29507148  0.89512566  0.87813349  0.82158108  0.68864025
0.55391765
## [13] -0.06191171 -0.30596266 -0.38047100 -0.69470698 -0.20791728 -
1.26539635
## [19]  2.16895597  1.20796200 -1.12310858 -0.40288484 -0.46665535
0.77996512
## [25] -0.08336907  0.25331851 -0.02854676 -0.04287046  1.36860228 -
0.22577099
## [31]  1.51647060 -1.54875280  0.58461375  0.12385424  0.21594157
0.37963948
## [37] -0.50232345 -0.33320738 -1.01857538 -1.07179123  0.30352864
0.44820978
## [43]  0.05300423  0.92226747  2.05008469 -0.49103117 -2.30916888
1.00573852
## [49] -0.70920076 -0.68800862

y_ii=2+3*x_ii+eii
y_ii

## [1] 19.62697 29.66236 23.28803 29.10712 32.36082 18.10981 24.62651
31.28141
## [9] 26.14966 24.67080 32.04114 24.35393 27.10165 25.28354 18.16340
29.80267
## [17] 20.48340 16.36550 24.08777 32.52552 29.21998 26.98917 26.14095
32.69401

```

```

## [25] 26.75222 27.88128 25.13244 25.86926 22.70600 18.98093 32.96183
28.98573
## [33] 27.94519 29.05587 17.58515 24.54658 27.87457 19.91291 20.75414
19.40260
## [41] 19.44553 23.66640 23.25887 23.45495 21.33676 18.59106 18.18634
24.99518
## [49] 20.28039 29.17941

model= lm(y_ii ~ x_ii)

# Least squares estimates
beta0_hat=coef(model) [1]
beta_hat=coef(model)[2]

beta0_hat

## (Intercept)
##      1.475903

beta_hat

##      x_ii
## 3.076349

beta0_grid= seq(beta0_hat-2, beta0_hat + 2, length.out = 100)
beta_grid = seq(beta_hat-2, beta_hat + 2, length.out = 100)

RSS <- matrix(NA, nrow = length(beta0_grid), ncol = length(beta_grid))

# Compute RSS for each (beta0, beta)
for (i in 1:length(beta0_grid)) {
  for (j in 1:length(beta_grid)) {
    RSS[i, j] <- sum((y_ii - beta0_grid[i] - beta_grid[j] * x_ii)^2)
  }
}
which(RSS == min(RSS), arr.ind=TRUE)

##      row col
## [1,] 47  51

# Find minimum RSS
min_RSS <- min(RSS)
index <- which(RSS == min_RSS, arr.ind = TRUE)

beta0_min <- beta0_grid[index[1]]
beta_min  <- beta_grid[index[2]]

beta0_min

## [1] 1.334489

```

```
beta_min
## [1] 3.096551
```

3. Problem to demonstrate that least square estimators are unbiased

Step 1: Generate $x_i (i = 1, 2, \dots, n)$ from Uniform(0, 1), $\epsilon_i (i = 1, 2, \dots, n)$ from $N(0, 1)$ and hence generate y using $y_i = \beta_0 + \beta x_i + \epsilon_i$. (Take $\beta_0 = 2$, $\beta = 3$).

Step 2: On the basis of the data $(x_i, y_i) (i = 1, 2, \dots, n)$ generated in Step 1, obtain the least square estimates of β_0 and β . Repeat Steps 1-2, $R = 1000$ times. In each simulation obtain $\hat{\beta}_0$ and $\hat{\beta}$. Finally, the least-square estimates will be given by the average of these estimated values. Compare these with the true β_0 and β and comment. Take $n = 50$ and seed=123.

```
#3
set.seed(123)
x_3=runif(50,0,1)
x_3

## [1] 0.28757752 0.78830514 0.40897692 0.88301740 0.94046728 0.04555650
## [7] 0.52810549 0.89241904 0.55143501 0.45661474 0.95683335 0.45333416
## [13] 0.67757064 0.57263340 0.10292468 0.89982497 0.24608773 0.04205953
## [19] 0.32792072 0.95450365 0.88953932 0.69280341 0.64050681 0.99426978
## [25] 0.65570580 0.70853047 0.54406602 0.59414202 0.28915974 0.14711365
## [31] 0.96302423 0.90229905 0.69070528 0.79546742 0.02461368 0.47779597
## [37] 0.75845954 0.21640794 0.31818101 0.23162579 0.14280002 0.41454634
## [43] 0.41372433 0.36884545 0.15244475 0.13880606 0.23303410 0.46596245
## [49] 0.26597264 0.85782772

e_3=rnorm(50,0,1)
e_3

## [1] -1.68669331 0.83778704 0.15337312 -1.13813694 1.25381492
0.42646422
## [7] -0.29507148 0.89512566 0.87813349 0.82158108 0.68864025
0.55391765
## [13] -0.06191171 -0.30596266 -0.38047100 -0.69470698 -0.20791728 -
1.26539635
## [19] 2.16895597 1.20796200 -1.12310858 -0.40288484 -0.46665535
0.77996512
## [25] -0.08336907 0.25331851 -0.02854676 -0.04287046 1.36860228 -
0.22577099
## [31] 1.51647060 -1.54875280 0.58461375 0.12385424 0.21594157
0.37963948
## [37] -0.50232345 -0.33320738 -1.01857538 -1.07179123 0.30352864
0.44820978
## [43] 0.05300423 0.92226747 2.05008469 -0.49103117 -2.30916888
1.00573852
## [49] -0.70920076 -0.68800862
```

```

y_3=2+3*x_3+e_3
model_3= lm(y_3 ~ x_3)
model_3

##
## Call:
## lm(formula = y_3 ~ x_3)
##
## Coefficients:
## (Intercept)          x_3
##           1.858        3.382

set.seed(123)

n= 50
R= 1000

beta0_hat <- numeric(R)
beta_hat <- numeric(R)

for (r in 1:R) {

  x_3 <- runif(n, 0, 1)
  e_3 <- rnorm(n, 0, 1)      # using our error term
  y_3 <- 2 + 3 * x_3 + e_3

  model_3 <- lm(y_3 ~ x_3)

  beta0_hat[r] <- coef(model_3)[1]
  beta_hat[r] <- coef(model_3)[2]
}
mean(beta0_hat)

## [1] 2.013053

mean(beta_hat)

## [1] 2.982112

var(beta0_hat)

## [1] 0.07969639

var(beta_hat)

## [1] 0.2360544

```

This demonstrate that the Least Squares Estimators are unbiased. The average of the 1,000 estimated coefficients ($\hat{\beta}_0$ and $\hat{\beta}_1$) is extremely close to the true parameter values ($\beta_0 = 2$ and $\beta_1 = 3$). The small differences observed in the table are due to sampling

variability and would approach zero as the number of simulations (R) increases towards infinity.

Question 4 : Comparing several simple linear regressions

Attach “Boston” data from MASS library in R. Select median value of owner- occupied homes, as the response and per capita crime rate, nitrogen oxides

concentration, proportion of blacks and percentage of lower status of the population as predictors.

- (a) Selecting the predictors one by one, run four separate linear regressions to the data. Present the output in a single table.
- (b) Which model gives the best fit?
- (c) Compare the coefficients of the predictors from each model and comment on the usefulness of the predictors.

```
library(MASS)
x1=Boston
y=Boston$medv#response
x=data.frame(x1$crim,x1$nox,x1$black,x1$lstat)#predictor
y

## [1] 24.0 21.6 34.7 33.4 36.2 28.7 22.9 27.1 16.5 18.9 15.0 18.9 21.7
20.4 18.2
## [16] 19.9 23.1 17.5 20.2 18.2 13.6 19.6 15.2 14.5 15.6 13.9 16.6 14.8
18.4 21.0
## [31] 12.7 14.5 13.2 13.1 13.5 18.9 20.0 21.0 24.7 30.8 34.9 26.6 25.3
24.7 21.2
## [46] 19.3 20.0 16.6 14.4 19.4 19.7 20.5 25.0 23.4 18.9 35.4 24.7 31.6
23.3 19.6
## [61] 18.7 16.0 22.2 25.0 33.0 23.5 19.4 22.0 17.4 20.9 24.2 21.7 22.8
23.4 24.1
## [76] 21.4 20.0 20.8 21.2 20.3 28.0 23.9 24.8 22.9 23.9 26.6 22.5 22.2
23.6 28.7
## [91] 22.6 22.0 22.9 25.0 20.6 28.4 21.4 38.7 43.8 33.2 27.5 26.5 18.6
19.3 20.1
## [106] 19.5 19.5 20.4 19.8 19.4 21.7 22.8 18.8 18.7 18.5 18.3 21.2 19.2
20.4 19.3
## [121] 22.0 20.3 20.5 17.3 18.8 21.4 15.7 16.2 18.0 14.3 19.2 19.6 23.0
18.4 15.6
## [136] 18.1 17.4 17.1 13.3 17.8 14.0 14.4 13.4 15.6 11.8 13.8 15.6 14.6
17.8 15.4
## [151] 21.5 19.6 15.3 19.4 17.0 15.6 13.1 41.3 24.3 23.3 27.0 50.0 50.0
50.0 22.7
## [166] 25.0 50.0 23.8 23.8 22.3 17.4 19.1 23.1 23.6 22.6 29.4 23.2 24.6
29.9 37.2
## [181] 39.8 36.2 37.9 32.5 26.4 29.6 50.0 32.0 29.8 34.9 37.0 30.5 36.4
31.1 29.1
## [196] 50.0 33.3 30.3 34.6 34.9 32.9 24.1 42.3 48.5 50.0 22.6 24.4 22.5
24.4 20.0
```

```

## [211] 21.7 19.3 22.4 28.1 23.7 25.0 23.3 28.7 21.5 23.0 26.7 21.7 27.5
30.1 44.8
## [226] 50.0 37.6 31.6 46.7 31.5 24.3 31.7 41.7 48.3 29.0 24.0 25.1 31.5
23.7 23.3
## [241] 22.0 20.1 22.2 23.7 17.6 18.5 24.3 20.5 24.5 26.2 24.4 24.8 29.6
42.8 21.9
## [256] 20.9 44.0 50.0 36.0 30.1 33.8 43.1 48.8 31.0 36.5 22.8 30.7 50.0
43.5 20.7
## [271] 21.1 25.2 24.4 35.2 32.4 32.0 33.2 33.1 29.1 35.1 45.4 35.4 46.0
50.0 32.2
## [286] 22.0 20.1 23.2 22.3 24.8 28.5 37.3 27.9 23.9 21.7 28.6 27.1 20.3
22.5 29.0
## [301] 24.8 22.0 26.4 33.1 36.1 28.4 33.4 28.2 22.8 20.3 16.1 22.1 19.4
21.6 23.8
## [316] 16.2 17.8 19.8 23.1 21.0 23.8 23.1 20.4 18.5 25.0 24.6 23.0 22.2
19.3 22.6
## [331] 19.8 17.1 19.4 22.2 20.7 21.1 19.5 18.5 20.6 19.0 18.7 32.7 16.5
23.9 31.2
## [346] 17.5 17.2 23.1 24.5 26.6 22.9 24.1 18.6 30.1 18.2 20.6 17.8 21.7
22.7 22.6
## [361] 25.0 19.9 20.8 16.8 21.9 27.5 21.9 23.1 50.0 50.0 50.0 50.0 50.0
13.8 13.8
## [376] 15.0 13.9 13.3 13.1 10.2 10.4 10.9 11.3 12.3 8.8 7.2 10.5 7.4
10.2 11.5
## [391] 15.1 23.2 9.7 13.8 12.7 13.1 12.5 8.5 5.0 6.3 5.6 7.2 12.1
8.3 8.5
## [406] 5.0 11.9 27.9 17.2 27.5 15.0 17.2 17.9 16.3 7.0 7.2 7.5 10.4
8.8 8.4
## [421] 16.7 14.2 20.8 13.4 11.7 8.3 10.2 10.9 11.0 9.5 14.5 14.1 16.1
14.3 11.7
## [436] 13.4 9.6 8.7 8.4 12.8 10.5 17.1 18.4 15.4 10.8 11.8 14.9 12.6
14.1 13.0
## [451] 13.4 15.2 16.1 17.8 14.9 14.1 12.7 13.5 14.9 20.0 16.4 17.7 19.5
20.2 21.4
## [466] 19.9 19.0 19.1 19.1 20.1 19.9 19.6 23.2 29.8 13.8 13.3 16.7 12.0
14.6 21.4
## [481] 23.0 23.7 25.0 21.8 20.6 21.2 19.1 20.6 15.2 7.0 8.1 13.6 20.1
21.8 24.5
## [496] 23.1 19.7 18.3 21.2 17.5 16.8 22.4 20.6 23.9 22.0 11.9

model1=lm(y~x1$crim)
model1

##
## Call:
## lm(formula = y ~ x1$crim)
##
## Coefficients:
## (Intercept)      x1$crim
##       24.0331     -0.4152

```

```

summary(model1)

##
## Call:
## lm(formula = y ~ x1$crim)
##
## Residuals:
##    Min     1Q Median     3Q    Max 
## -16.957 -5.449 -2.007  2.512 29.800 
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)    
## (Intercept) 24.03311   0.40914  58.74 <2e-16 ***
## x1$crim     -0.41519   0.04389  -9.46 <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 8.484 on 504 degrees of freedom
## Multiple R-squared:  0.1508, Adjusted R-squared:  0.1491 
## F-statistic: 89.49 on 1 and 504 DF,  p-value: < 2.2e-16

model2=lm(y~x1$nox)
model2

##
## Call:
## lm(formula = y ~ x1$nox)
##
## Coefficients:
## (Intercept)      x1$nox  
##        41.35       -33.92 

model3=lm(y~x1$black)
model3

##
## Call:
## lm(formula = y ~ x1$black)
##
## Coefficients:
## (Intercept)      x1$black  
##      10.55103     0.03359 

model4=lm(y~x1$lstat)
model4

##
## Call:
## lm(formula = y ~ x1$lstat)
##
## Coefficients:

```

```

## (Intercept)      x1$lstat
##            34.55          -0.95

models=c("model of y vs crim","model of y vs nox","model of y vs
black","model of y vs lstat")
beta_hat_0=c(coef(model1)[1],coef(model2)[1],coef(model3)[1],coef(model4)[1])
beta_hat=c(coef(model1)[2],coef(model2)[2],coef(model3)[2],coef(model4)[2])
Data=data.frame(models,beta_hat_0,beta_hat)
Data

##                         models beta_hat_0      beta_hat
## x1$crim    model of y vs crim   24.03311  -0.41519028
## x1$nox     model of y vs nox   41.34587 -33.91605501
## x1$black   model of y vs black  10.55103   0.03359306
## x1$lstat   model of y vs lstat  34.55384  -0.95004935

summary(model1)

##
## Call:
## lm(formula = y ~ x1$crim)
##
## Residuals:
##       Min     1Q   Median     3Q    Max
## -16.957 -5.449 -2.007  2.512 29.800
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 24.03311   0.40914  58.74 <2e-16 ***
## x1$crim     -0.41519   0.04389  -9.46 <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 8.484 on 504 degrees of freedom
## Multiple R-squared:  0.1508, Adjusted R-squared:  0.1491
## F-statistic: 89.49 on 1 and 504 DF,  p-value: < 2.2e-16

summary(model2)

##
## Call:
## lm(formula = y ~ x1$nox)
##
## Residuals:
##       Min     1Q   Median     3Q    Max
## -13.691 -5.121 -2.161  2.959 31.310
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 41.346     1.811   22.83 <2e-16 ***
## x1$nox     -33.916     3.196  -10.61 <2e-16 ***

```

```

## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 8.323 on 504 degrees of freedom
## Multiple R-squared:  0.1826, Adjusted R-squared:  0.181
## F-statistic: 112.6 on 1 and 504 DF,  p-value: < 2.2e-16

summary(model3)

##
## Call:
## lm(formula = y ~ x1$black)
##
## Residuals:
##     Min      1Q  Median      3Q      Max
## -18.884  -4.862  -1.684   2.932  27.763
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 10.551034  1.557463  6.775 3.49e-11 ***
## x1$black     0.033593  0.004231  7.941 1.32e-14 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 8.679 on 504 degrees of freedom
## Multiple R-squared:  0.1112, Adjusted R-squared:  0.1094
## F-statistic: 63.05 on 1 and 504 DF,  p-value: 1.318e-14

summary(model4)

##
## Call:
## lm(formula = y ~ x1$lstat)
##
## Residuals:
##     Min      1Q  Median      3Q      Max
## -15.168  -3.990  -1.318   2.034  24.500
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 34.55384   0.56263  61.41  <2e-16 ***
## x1$lstat    -0.95005   0.03873 -24.53  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 6.216 on 504 degrees of freedom
## Multiple R-squared:  0.5441, Adjusted R-squared:  0.5432
## F-statistic: 601.6 on 1 and 504 DF,  p-value: < 2.2e-16

extract <- function(m){
  c(

```

```

Intercept = coef(m)[1],
Slope = coef(m)[2],
R2 = summary(m)$r.squared,
p_value = summary(m)$coefficients[2,4]
)
}

results <- rbind(
  crim = extract(model1),
  nox = extract(model2),
  black = extract(model3),
  lstat = extract(model4)
)

results

##      Intercept.(Intercept) Slope.x1$crim        R2      p_value
##  crim           24.03311   -0.41519028 0.1507805 1.173987e-19
##  nox            41.34587   -33.91605501 0.1826030 7.065042e-24
##  black          10.55103    0.03359306 0.1111961 1.318113e-14
##  lstat          34.55384   -0.95004935 0.5441463 5.081103e-88

```

The model with lstat as predictor gives the best fit.

Highest R-squared (around 0.54, much larger than others), ,Very small p-value (≈ 0),Strong linear relationship with medv,Lowest residual standard error

Comparing coefficients and usefulness of predictors:

Crime rate (crim):

Negative coefficient → higher crime lowers home value

Statistically significant

Explains only a small portion of variability

Useful, but not a strong standalone predictor

Nitrogen oxides (nox)

Negative coefficient → more pollution lowers prices

Stronger effect than crime

Still limited explanatory power alone

Proportion of blacks (black)

Positive coefficient

Statistically significant, but low R^2

Weak predictor by itself

Lower status population (Istat)

Strong negative coefficient

Extremely significant

Largest magnitude effect

Best single predictor among the four